

# **Implications of Recent Epidemiologic Studies for the Linear Nonthreshold Model and Radiation Protection**

*September 15, 2017*

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3 **Preface**

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5 The National Council on Radiation Protection and Measurements (NCRP) has a long history of  
6 issuing guidance on operational radiation safety including radiation exposure limits for radiation workers  
7 and the public. Effective dose limits are based on the linear nonthreshold (LNT) dose effects model, which  
8 is based almost entirely on the human epidemiology data. This Commentary provides a review of recent  
9 epidemiologic studies and an evaluation of whether the new observations are strong enough to support or  
10 modify the LNT model as used in radiation protection today. This Report represents an update of the  
11 guidance provided in NCRP Report No.136, *Evaluation of the Linear-Nonthreshold Dose-Response*  
12 *Model for Ionizing Radiation* (NCRP, 2001).

13

14 This Commentary was prepared by Scientific Committee 1-25 on Recent Epidemiologic Studies  
15 and Implications for the Linear Nonthreshold Model. Serving on Scientific Committee 1-25 were:

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240 **1. Executive Summary**

241

242

**1.1 Introduction**

243

244 Historically, epidemiologic studies have assessed the health effects of ionizing radiation exposure from multiple  
245 sources: occupational, accidental, environmental, military and medical. The several national and international  
246 reviews of the status of health risks associated with exposure to low levels of ionizing radiation that have been  
247 completed in the last few decades generally agreed that the expectations for health effects in humans, such as cancer  
248 induction or cardiac damage, observed at acute doses of 100 mGy and above are more reliable than those observed  
249 at <100 mGy, the low-dose region (NCRP, 2015). For the purpose of this Commentary, for low linear-energy  
250 transfer (LET) radiation, a low absorbed dose is <100 mGy delivered acutely, and a low absorbed-dose rate is  
251 <5 mGy h<sup>-1</sup> for any accumulated dose. See NCRP Commentary No. 24 (NCRP, 2015) Section 1.1 for additional  
252 discussion of low doses and low dose rates.

253

254 Our understanding of the shape of the dose-response relationship and the level of risk from low-LET types of  
255 radiation at low doses and low dose rates remains uncertain because of the intrinsic uncertainties in results from the  
256 epidemiologic and radiobiological studies of low doses of radiation. This uncertainty can impact actions taken  
257 regarding radiation protection guidance, medical practice, compensation programs, environmental contamination  
258 issues, technological advances, and communication with members of the public (NCRP, 2015). For over 40 y the  
259 linear nonthreshold (LNT) dose-response model has been commonly utilized for low-LET radiation when  
260 developing practical and prudent guidance on ways to protect workers and the public from the potential for harmful  
261 effects from radiation while balancing the beneficial, justified, and optimized uses of radiation in our society.  
262 Indeed, in developing its basic recommendations, as currently given in NCRP Report No. 116 (NCRP, 1993a), the  
263 Council reiterated its acceptance of the LNT model for the purposes of radiation protection.

264

265 The purpose of this Commentary is to provide a review of recent data from studies with low dose rates and from  
266 the Life Span Study of atomic-bomb survivors to determine whether these epidemiologic studies broadly support the  
267 LNT model of carcinogenic risk or, on the contrary, whether there is sufficient evidence that the LNT model is  
268 inappropriate for the purposes of radiation protection. The strength of epidemiologic support for the LNT model is  
269 evaluated for solid cancer incidence or mortality and secondarily for low dose-rate studies of leukemia. Briefer  
270 consideration is also given to low dose studies of thyroid cancer and breast cancer. The focus of this commentary is  
271 on low doses and low dose rates.

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## 1.2 Study Reviews

The primary approach was a critical review of each major study. The critique included an assessment of the quality of the epidemiology, dosimetry, and statistics of each study. The epidemiologic evaluation included a characterization of the study design and study population, quality of the data available, data collection methodology, the degree to which potential confounding variables or biases were assessed, and the quantitative results. Analysis of the dosimetry helped evaluate the robustness of the study in supporting the shape of the dose response curve at low doses and low dose rates. A statistical evaluation considered whether the analytic methods were appropriate, whether the study considered statistical alternatives to a linear dose-response trend, and whether sensitivity analyses or other clarifying analyses were undertaken. Based on those considerations, the contribution the study makes to the LNT model and to radiation protection is characterized. Several key studies about solid cancer mortality or incidence are briefly summarized below.

### 1.2.1 *Life Span Study*

Although this report focuses on studies with low doses and low dose rates, the Life Span Study (LSS) was included as a benchmark comparison study. The LSS cohort of atomic-bomb survivors (Section 4.1) has provided important data because it is a large cohort (~86,000 survivors of all ages) with relatively accurate dosimetry, a wide dose range (0 to 4 Gy colon dose, including ~68,000 with doses <100 mGy), over 50 y of high-quality follow-up for mortality and cancer incidence, and over 1,000 excess cancer cases associated with radiation exposure. These features provide relatively high statistical power and precision of risk estimates, resulting in a statistically significant dose response for all incident solid cancer over the dose range 0 to 100 mGy (Grant *et al.*, 2017). Formal dose-threshold analyses for both solid cancer incidence and mortality are compatible with no dose threshold, and a pure quadratic model provided a significantly poorer fit than a linear dose-response model (Grant *et al.*, 2017; Ozasa *et al.*, 2012).

A nonparametric analysis of the most recent mortality data indicated excess risk over the range of 0 to 200 mGy that was congruent with the LNT dose response model. Nevertheless, the most recent solid cancer mortality and incidence data provide some evidence for upward curvature in the dose response consistent with a linear-quadratic model. This implies a shallower, but still positive, dose-response slope at low doses than at higher ones, though this curvilinearity appeared to be confined primarily to males. In summary, the study provides strong support for the use of a LNT model, with consideration of a low-dose effectiveness factor (LDEF), for use in radiation protection. However, unlike most other studies reviewed in this report, the LSS assesses the effects of a single, brief exposure and the associated LDEF, but it does not assess protracted or highly-fractionated doses.

307

308 **1.2.2** *Worker Studies*

309

310 Radiation worker studies assess risks in worker groups exposed largely to low doses received at low dose-rates,  
311 addressing directly the validity of the LNT model for low dose-rate exposures. Further, cumulative doses can be  
312 several hundred mGy, especially for workers in early periods, so that some studies can offer reasonable statistical  
313 power.

314

315 **INWORKS Study:** Large studies that combine data from workers from numerous nuclear installations in a number  
316 of countries have now been conducted (Cardis *et al.*, 1995; 2007). An important study is the International Nuclear  
317 Workers Study (INWORKS), which included workers from nuclear facilities in the United States, United Kingdom,  
318 and France (Leuraud *et al.*, 2015; Hamra *et al.*, 2016; Richardson *et al.*, 2015) (reviewed in Section 4.2.2).  
319 INWORKS found an association between the cumulative external photon dose to the red bone marrow (RBM) and  
320 mortality from leukemia [excluding chronic lymphocytic leukemia (CLL) excess relative risk (ERR) Gy<sup>-1</sup> of 3.0,  
321 90 % confidence interval (CI) of 1.2 to 5.2]. External dose to the colon (used as the prototypic organ) was associated  
322 with mortality from all solid cancers combined (ERR Gy<sup>-1</sup> of 0.47, 90 % CI of 0.18 to 0.78). For solid cancer there  
323 was no evidence of nonlinearity ( $p = 0.44$ ). These risk estimates were similar to those in the LSS data. Even when  
324 the cumulative colon dose was restricted to 0 to 100 mGy, a statistically significant dose response was seen for all  
325 cancers excluding leukemia.

326

327 **Mayak Study:** The Russian Mayak workforce is of particular interest because of the high cumulative doses  
328 received (mainly at a low dose rate) by many workers during the early years of operations at this installation  
329 (Section 4.2.3). The investigators reported statistically significant associations between external dose and mortality  
330 from leukemia (excluding CLL) and from all solid cancers excluding lung, liver and bone (*i.e.*, excluding cancers at  
331 the major sites of plutonium deposition); and adjusting for plutonium exposure; ERR Gy<sup>-1</sup> of 0.12, 95 % CI: 0.03 to  
332 0.21 (Sokolnikov *et al.*, 2015). For solid cancer there was no indication of nonlinearity ( $p > 0.5$ ) based on external  
333 dose to the colon. For leukemia, excluding the chronic lymphocytic type, the linear ERR Gy<sup>-1</sup> estimate was 3.57  
334 (90 % CI 1.55, 8.22) for cumulative external radiation dose to the red bone marrow, adjusted for plutonium  
335 exposure. The linear-quadratic model fit marginally better than the linear model ( $p = 0.11$ ), and the pure linear and  
336 pure quadratic models fit about equally well.

337

338 **Summary of Worker Studies:** Overall the nuclear worker studies lend considerable support to the inference that an  
339 excess risk of cancer exists following protracted exposure to low doses received at a low dose rate, and the excess  
340 risk is compatible with a LNT model, perhaps modified by a DDREF. Although the accuracy of the risk estimates is

341 limited to some degree by uncertainties in dosimetry and epidemiology, the studies provide substantial support for  
342 the LNT model. The Million Worker Study is underway in the United States; when it is completed it is expected to  
343 augment appreciably the information available on worker radiation-related cancer risks and reduce the uncertainties  
344 in risk estimation after exposures at low dose rates (Boice, 2015a; 2017c; Boice *et al.*, 2014; Bouville *et al.*, 2015;  
345 Till *et al.*, 2014).

346

### 347 1.2.3 *Environmental Exposure Studies*

348

349 **Techa River Study:** Between 1949 and 1956 the Russian Mayak nuclear weapons facility released radioactive  
350 waste into the Techa River and exposed approximately 30,000 residents to relatively low doses at low dose rates  
351 from gamma rays (external) and  $^{137}\text{Cs}$  and  $^{90}\text{Sr}$  (internal). The recent studies of the Techa River Cohort have found  
352 associations between radiation dose and incidence and mortality rates for solid cancers and leukemia (other than  
353 CLL) that they report are linear in dose response (Davis *et al.*, 2015; Krestinina *et al.*, 2013a; Schonfeld *et al.*, 2013)  
354 (Section 4.3.1). However, inherent uncertainties in the dose reconstruction along with some limitation in the cancer  
355 ascertainment weaken inferences about the shape of the dose-response curves and the LNT model.

356

357 **Chernobyl Thyroid Cancer Studies:** New studies of cohorts of children in Ukraine and Belarus who had thyroid  
358 measurements of  $^{131}\text{I}$  shortly after the Chernobyl accident and systematic thyroid screening have added appreciably  
359 to our knowledge about thyroid cancer risk after protracted internal exposure (Brenner *et al.*, 2011; Zablotska *et al.*,  
360 2011) (Section 4.3.2). Both cohorts showed strong linear dose-response functions with no evidence of nonlinearity,  
361 though perhaps with a somewhat lower risk per unit dose than seen in studies of children exposed to external  
362 gamma radiation. The thyroid doses are believed to be sufficiently accurate to support a LNT interpretation.

363

### 364 1.2.4 *High Background Radiation Areas*

365

366 Studies of residents in areas of high natural background radiation have been conducted in Kerala, India and  
367 Yangjiang, China. However, it is exceedingly difficult to conduct a geographic study of background radiation, *e.g.*,  
368 it is difficult to find a suitable low exposure control group with highly similar lifestyles and natural disease rates to  
369 whom the highly exposed group may be compared. The better and larger of the two studies, the Kerala study of  
370 cancer incidence, included 70,000 individuals and over 1,300 cancers from high-background or low-background  
371 areas (Nair *et al.*, 2009) (Section 4.3.3). The dosimetry was based on measurements of ambient levels within and  
372 near homes, coupled with average house-occupancy factors by age and sex. They reported an ERR  $\text{Gy}^{-1}$  of  $-0.13$   
373 (95 % CI  $-0.58, 0.46$ ) for all cancer except leukemia, and there were too few leukemia cases to be informative. The  
374 Yangjiang study reported a positive, but nonsignificant, risk coefficient for all cancer except leukemia and liver

375 cancer [ERR Gy<sup>-1</sup> 0.19 (95 % CI -1.9, 3.0); Tao *et al.*, 2012]. These studies are nominally more supportive of little  
376 or no effect after low dose-rate exposures rather than the LNT model. However, the fact that much of the dose  
377 variation is attributable to geographic locations, which may be associated with risk factors other than radiation level,  
378 introduces ambiguity into the inference regarding radiation effects. Furthermore, the substantial uncertainties in  
379 dosimetry, the weaknesses in cancer ascertainment, and the wide confidence intervals on the risk estimates mean  
380 they need to be interpreted with caution.

381

### 382 1.2.5 *Childhood Radiation Studies*

383

384 Medical exposures are typically partial body and study results are subject to significant uncertainties including,  
385 but not limited to, historical exposure data, limited organ dosimetry for organs other than the target organ, and  
386 potential biases because radiologic procedures are often administered for an existing health condition. Recent  
387 epidemiologic studies have involved populations who had received computed tomography (CT) scans during  
388 childhood when risk might be greater because CT doses were relatively high and children may be more  
389 radiosensitive to cancer induction than adults (Section 4.4.2). However, information on organ doses from CT  
390 examinations in the 1980s and 1990s is sparse and individual doses have not been reconstructed. CT studies suffer  
391 from potential biases: confounding by indication (CT examinations more likely for those who have conditions that  
392 confer risk for cancer) and reverse causation (pre-existing but undetected malignancy). Because of the weak  
393 dosimetry and potential for bias, the results are considered unreliable for evaluating the LNT dose response model.

394

395 The data on postnatal diagnostic medical exposures and childhood leukemia risk are inconclusive (Wakeford,  
396 2008). Studies of juvenile irradiation and breast cancer generally support a linear dose response. A recent pooled  
397 analysis of external thyroid irradiation in childhood and subsequent thyroid cancer in nine studies showed a  
398 significant dose response from 0 to 100 mGy and no evidence of nonlinearity (Lubin *et al.*, 2017). An analysis of  
399 solid cancer incidence among the Japanese atomic-bomb survivors exposed prenatally or during childhood showed a  
400 clear dose response, but marginal upward curvature ( $p = 0.09$ ) suggested that the dose-response slope may be  
401 shallower in the low-dose range. In general, the low dose data on children are sparse, the number of specific types of  
402 cancer is small and uncertainties are large enough that such studies do not yield definitive information on the LNT  
403 model. In the case of thyroid cancer and breast cancer, the data broadly support the LNT model.

404

### 405 1.2.6 *Diseases Classified as Tissue Reactions*

406

407 Most of the available data on noncancer effects have large associated uncertainties and limitations that do not  
408 yet support a quantitative estimate of a specific threshold value for effects from either acute or protracted lens

409 exposures. However, the preponderance of evidence suggests the possibility that effects (such as lens opacities or  
410 cardiovascular disease) could occur at lower doses than previously thought.

411  
412 There is growing epidemiologic evidence to suggest a raised risk of cardiovascular disease (CVD) at lower  
413 levels of exposure to radiation than previously thought, implying that poorly understood radiobiological  
414 mechanisms associated with low-to-moderate doses and/or low dose rates may produce an increased risk of CVD  
415 (Section 5.1). Studies of nuclear workers and other exposed groups provide a mixed picture as to CVD risk, and  
416 most of them lack information on important confounding factors associated with lifestyle and concurrent conditions  
417 (*e.g.*, diabetes, obesity). Therefore, the evidence is too weak and inconsistent to support a LNT model for CVD at  
418 this time.

419  
420 Studies of cataracts in the atomic-bomb survivors and following Chernobyl exposures have revealed the  
421 development of minor lenticular opacities at doses lower than previously considered to be cataractogenic.  
422 Ophthalmologically detectable opacities are reported at doses of 0.5 to 2 Gy with large uncertainties below about  
423 0.5 Gy. So, at this time, the NCRP recommends use of a threshold model for cataracts (NCRP, 2016).

424

### 425 **1.3 Results of Study Evaluations**

426

427 Support by studies for any model requires adequacy of the study components, which for epidemiologic studies  
428 can be classified broadly as adequacy of epidemiologic methods, dosimetry, and statistics. For each component of  
429 the major studies, this commentary has critiqued both the methods used and the adequacy of the results of those  
430 methods (Section 7). The Committee evaluated these components for 26 principal studies or groups of studies of  
431 cancer risk. As a minimal criterion of study adequacy, 18 of the studies had no component on which they were  
432 scored as weak. Thirteen of the studies were scored moderate to strong on all three components of evaluation.

433

434 The Committee also rated each study or group of studies on their strength of support for the LNT model, as  
435 shown in Table 1.1. Twenty-one studies (80 %) provided some support for the LNT model, including five studies  
436 (19 %) providing strong support and seven providing moderate support.. Five of the studies (19 %) provided  
437 essentially no or inconclusive support for the LNT model. A rating of moderate versus strong support for LNT  
438 sometimes hinged upon the size of the study or other limitation and not on indications of nonlinearity. most of the  
439 larger,

Table 1.1—Ratings of the degree of support for the LNT model by the cancer studies reviewed.

Study (or groups of studies) a	Support for LNT Model
Life Span Study (LSS), Japan atomic bomb (Grant <i>et al.</i> , 2017) <sup>b</sup>	Strong
INWORKS (U.K., U.S., French combined cohorts) (Richardson <i>et al.</i> , 2015)	Strong
Tuberculosis fluoroscopic examinations and breast cancer (Little and Boice, 2003)	Strong
Childhood atomic-bomb exposure (Preston <i>et al.</i> , 2008)	Strong
Childhood thyroid cancer studies (Lubin <i>et al.</i> , 2017)	Strong
Mayak nuclear facility (Sokolnikov <i>et al.</i> , 2015)	Moderate
Techa River, nearby residents Davis <i>et al.</i> , 2015)	Moderate
Chernobyl fallout, Ukraine and Belarus thyroid cancer (Brenner <i>et al.</i> , 2011)	Moderate
Breast cancer studies, after childhood exposure (Eidemüller <i>et al.</i> , 2015)	Moderate
<i>In utero</i> atomic-bomb exposure (Preston <i>et al.</i> , 2008)	Moderate
<i>In utero</i> exposures, medical (Wakeford, 2008)	Moderate
Canadian worker study (Zablotska <i>et al.</i> , 2013b)	Moderate
Japanese worker study (Akiba and Mizuno, 2012)	Weak to moderate
Chernobyl cleanup workers, Russia (Kashcheev <i>et al.</i> , 2015)	Weak to moderate
U.S. radiologic technologists (Liu <i>et al.</i> , 2014; Preston <i>et al.</i> , 2016)	Weak-to-moderate
Mound facility (Boice <i>et al.</i> , 2014)	Weak-to-moderate
Rocketdyne facility (Boice <i>et al.</i> , 2011)	Weak-to-moderate
Medical x-ray workers, China (Sun <i>et al.</i> , 2016)	Weak-to-moderate
Background radiation levels and childhood leukemia (Kendall <i>et al.</i> , 2013)	Weak-to-moderate
Taiwan radiocontaminated buildings, residents (Hwang <i>et al.</i> , 2008)	Weak-to-moderate <sup>c</sup>
Pediatric CT examinations (Pearce <i>et al.</i> , 2012)	Weak-to-moderate <sup>c</sup>
Childhood leukemia studies (Wakeford and Little, 2003)	Weak-to-moderate
<i>In utero</i> exposures, Mayak and Techa (Akleyev <i>et al.</i> , 2016)	Weak-to-moderate
Hanford <sup>131</sup> I fallout study (Davis <i>et al.</i> , 2004)	None
Kerala, India, high natural background radiation area (Nair <i>et al.</i> , 2009)	None
Yangjiang, China, high natural background radiation area (Tao <i>et al.</i> , 2012)	Inconclusive <sup>c</sup>
U.S. atomic veterans (Beck <i>et al.</i> , 2017)	Inconclusive
Fallout studies (aggregate of eight studies) (Lyon <i>et al.</i> , 2006)	Inconclusive <sup>c</sup>

<sup>a</sup>A number of studies were excluded for various reasons described in the text, these include but are not limited to: ecological studies of residents around nuclear power plant facilities, studies of hereditary effects, studies of tissue reaction (or “deterministic”) effects, and the 15-country study

<sup>b</sup>A representative recent publication is listed for each study or study group.

<sup>c</sup>Considered “weak” support or “inconclusive” primarily because of epidemiologic method or dosimetric weaknesses. The other studies in these categories were reasonable methodologically but provided little or no support for the LNT model because their risk coefficients were essentially zero or negative.

440 stronger studies broadly supported a LNT model. The studies that provided no support for the LNT model either  
441 had a totally null dose response or had excessively unreliable data. It should be noted that all the studies being  
442 considered, except for the Life Span Study of atomic-bomb survivors, had exposures at low dose rates or multiple  
443 small exposures. Furthermore, the preponderance of study subjects had cumulative doses under 100 mGy. Thus  
444 these studies are very relevant for contemporary radiation protection concerns.

445

446

#### **1.4 Future Improvements**

447

448 To stimulate radiation epidemiology efforts to address the LNT model and low- dose risks, the Committee  
449 suggested a number of profitable areas of focus for future research (report Section 8), and a few are mentioned  
450 here.

451

452 **Atomic-Bomb Survivors:** The low-dose data need to be examined in more detail, using additional covariables,  
453 statistical methods and analytic strategies, not only for solid cancer and leukemia, but also to evaluate specific  
454 cancers or cancer groups, cardiovascular diseases, and various clinical health endpoints. An examination is needed  
455 of whether the dose-response LNT model applies to tumors of various organs or organ systems, insofar as  
456 statistical limitations permit, which will provide evidence regarding the generality of the LNT model across tumor  
457 sites. The large bank of blood and tissue samples should be studied more robustly by the biomedical community to  
458 identify bioindicators of drivers of adverse outcome pathways that mediate between radiation and disease  
459 development.

460

461 **Worker Studies:** Much of the statistical power of these studies derives from those workers who have  
462 accumulated moderate doses of several hundred milligray over many years, most of whom started work in earlier  
463 years. Continuing follow-up of worker cohorts is desirable, as much of the cancer incidence and mortality is yet to  
464 occur. Doses in the early years tended to be highest but also had the greatest uncertainties because most dose  
465 recording technologies and dose record keeping practices were less advanced. Therefore, scrutiny of dose records  
466 is necessary to identify any deficiencies in recorded doses. Issues of neutron exposures, internal exposures and  
467 missing doses need to be addressed further. Valid risk estimates depend, inter alia, upon reliable dose estimates, so  
468 this area should be pursued vigorously.

469

470 **Environmental Radiation Studies:** All the environmental study groups should consider measures to reduce  
471 individual dose uncertainties. The Kerala and Yangjiang studies should increase efforts to improve cancer  
472 ascertainment and diagnosis, and to closely examine sociodemographic and geographic factors that may affect the

473 adequacy of cancer ascertainment. Further validation of reconstructed doses by personal dosimetry measurements  
474 would also be valuable. The Techa River studies should continue to improve dosimetry, enhance their follow-up  
475 and outcome ascertainment and further address the medical exposures received.

476  
477 **Other Future Directions:** Uncertainties should be provided with the dose estimates and used to adjust risk  
478 coefficients and confidence intervals (Stram *et al.*, 2015; UNSCEAR, 2015). For radiation-induced adverse health  
479 outcomes, a clear need is to identify bioindicators that define the pathway from normal to malignant cells that can  
480 be used for developing biologically based dose-response models. Analyzing epidemiologic data in conjunction with  
481 relevant radiobiological concepts and data has the potential to provide insights about low-dose risk that augment  
482 knowledge gained from the empirical epidemiologic data in isolation (NCRP, 2015).

483

### 484 1.5 Summary

485

486 Quantitative solid cancer risk estimates, based on estimated individual doses, of cancer mortality or incidence  
487 have been reported for nearly one million individuals with low dose rates and mostly low doses from studies of  
488 radiation workers or those exposed to elevated environmental radiation levels (Shore *et al.*, 2017). The completion  
489 of the million person study will considerably augment the available information (Boice, 2012a). The more robust  
490 studies have many strengths, with relatively good quality dosimetry, high rates of cohort mortality/morbidity  
491 ascertainment, attention to potential confounding variables, and proper analysis. Nevertheless, it is recognized that  
492 all studies have limitations, ranging from minor to serious, in their contribution to the quantitative evaluation of  
493 the LNT model. The individual low- dose studies intrinsically have limited statistical power and precision in risk  
494 estimation. These studies complement the LSS study of atomic-bomb survivors with its high dose rate and high  
495 dose range.

496

497 Strengths of some of the large epidemiologic studies such as INWORKS and the LSS lie in the long follow-up  
498 and large numbers of cancers and person-years at risk. The length of follow-up of epidemiologic studies is  
499 particularly relevant since a large fraction of both spontaneous and radiation-related cancers occur at 60 y of age  
500 and beyond. Although an historic weakness of many worker and environmental radiation studies was inadequate  
501 dosimetry, in recent years investigators have been focusing more on improving the quality and accuracy of the  
502 dosimetry. However, most studies considered in this report did not consider the effects of shared vs unshared  
503 uncertainty and classical as opposed to Berkson error (UNSCEAR, 2015), and then adjust for the effects of dose  
504 uncertainty on the risk estimates. Nearly all studies have adjusted for potential confounding by sex, attained age  
505 and sometimes age at exposure. However, few studies have analyzed radiation risks with control for possible  
506 confounding by lifestyle (*e.g.*, smoking), other disease risk factors or other sources of radiation exposure; these

507 factors may diminish the consistency of findings. Nevertheless, it should be emphasized that lifestyle or other  
508 disease risk factors will cause confounding only if their frequency (or intensity) varies appreciably according to  
509 dose. The most prominent lifestyle risk factor is smoking. Adjustment for socioeconomic status is used in several  
510 studies as an indirect approach for controlling for smoking and other lifestyle factors. Indirect approaches to  
511 examine the impact of smoking in several major studies have not found that smoking introduced substantial bias  
512 (Akiba, 2013; Davis *et al.*, 2015; Hunter *et al.*, 2013; Richardson *et al.*, 2015). For a few studies, concomitant  
513 medical radiation exposures have been examined; for the Techa River study diagnostic medical exposures at the  
514 official clinic were included in the doses (Schonfeld *et al.*, 2013). Other factors, such as losses to follow-up or  
515 incomplete disease ascertainment, would cause bias in the risk estimates only if they occur differentially according  
516 to dose levels. Few studies currently have biological samples to evaluate genetic or phenotypic biological factors  
517 that might cause effect modification of radiation risks.

518  
519 Because individual studies with low doses (less than 100 mGy) almost inevitably have relatively low statistical  
520 power, the findings for radiation and solid cancer are often not statistically significant. Furthermore, studies may  
521 have sampling variation or confounding by other exposures (*e.g.*, smoking or other lifestyle factors) which can  
522 diminish the consistency or validity of findings. Nevertheless, most large and high quality low- dose studies show  
523 positive risk coefficients (Shore *et al.*, 2017), suggesting there may be cancer effects at low doses, which is  
524 consistent with, though not necessarily proving, the applicability of the LNT model for radiation protection.  
525 However, it should be recognized that, the risk of cancer at low doses is small.

526  
527 The data regarding noncancer effects at low doses—cardiovascular diseases, cataracts, thyroid dysfunction,  
528 central nervous system effects—are mixed or null, suggesting at this time that an LNT assumption for radiation  
529 protection purposes for noncancer effects is not appropriate.

530

### 531 **1.6 Overall Conclusions on the Use of the LNT Model**

532

533 While the ongoing development of science requires a constant reassessment of prior and emerging evidence to  
534 assure that the approach to radiation protection is optimal, though not necessarily perfect, NCRP concludes that,  
535 based on current epidemiologic data, the LNT model (perhaps modified by a DDREF) should continue to be  
536 utilized for radiation protection purposes. This is in accord with judgment by other national and international  
537 scientific committees, based on somewhat older data than in the present report (ICRP, 2007; NA/NRC, 2006;  
538 UNSCEAR, 2008), that no alternative dose-response relationship appears more pragmatic or prudent for radiation  
539 protection purposes than the LNT model.

540

541 **2. Introduction**

542

543

**2.1 Background**

544

545 For over 40 y the linear nonthreshold (LNT) dose-response model has been used to develop practical and  
546 prudent guidance on ways to protect workers and the general public from the potential harmful effects of radiation  
547 while, at the same time, balancing the beneficial, justified, and optimized uses of radiation in our society. Indeed,  
548 in developing its basic radiation protection recommendations, as currently given in NCRP Report No. 116 (NCRP,  
549 1993a), the Council reiterated its acceptance of the LNT for the dose-risk relationship. Specifically, “based on the  
550 hypothesis that genetic effects and some cancers may result from damage to a single cell, the Council assumes  
551 that, for radiation-protection purposes, the risk of stochastic effects is proportional to dose without threshold,  
552 throughout the range of dose and dose rates of importance in routine radiation protection. Furthermore, the  
553 probability of response (risk) is assumed, for radiation protection purposes, to increase linearly with dose. At  
554 higher doses, received acutely, such as in accidents, more complex (nonlinear) dose/risk relationships may apply”  
555 (NCRP, 1993a).

556

557 NCRP later reassessed the weight of scientific evidence for and against the LNT model without reference to  
558 associated policy implications in Report No. 136 (NCRP, 2001). As in previous reviews by the NCRP (1980;  
559 1993c; 1997) the Council concluded that there was no conclusive evidence on which to reject the assumption of a  
560 LNT dose-response relationship for many of the risks attributable to low-level ionizing radiation (although it was  
561 acknowledged that additional data were needed) (NCRP, 1993b). The NCRP then noted that while many, but not  
562 all, scientific data support this assumption (NCRP, 1995), the probability of effects at very low doses such as are  
563 received from ubiquitous low-LET background radiation (NCRP, 1987; 2015) is so small that it may never be  
564 possible to prove or disprove the validity of the LNT assumption at those dose levels.

565

566 The International Commission on Radiological Protection (ICRP) published a science evaluation report,  
567 Publication 99 (ICRP, 2005b), on low-dose extrapolation of radiation-related cancer risks and issued updated  
568 radiation protection recommendations based on the conclusion that “while existence of a low dose threshold does  
569 not seem unlikely for radiation-related cancers of certain tissues, and cannot be ruled out for all cancers as a group,  
570 the evidence as a whole does not favor the existence of a universal threshold, and there seems to be no particular  
571 reason to factor the possibility of a threshold into risk calculations for purposes of radiation protection (ICRP,  
572 2007).” ICRP concluded that a LNT theory, combined with an uncertain DDREF for extrapolation of risk from high  
573 doses received acutely remains a prudent basis for radiation protection at low doses and low dose rates (ICRP,  
574 2005b).

575  
576 The United Nations Scientific Committee on the Effects of Atomic Radiation (UNSCEAR) evaluates the  
577 evidence of radiation-induced health effects from studies of the health of survivors of the atomic bombings of  
578 Japan and of other exposed groups. It also reviews advances in understanding of the mechanisms by which  
579 radiation-induced health effects can occur. These assessments provide the scientific foundation used by the  
580 International Commission on Radiological Protection (ICRP) and other protection organizations in developing  
581 their recommendations on radiation protection. UNSCEAR has concluded that the simplest representation of  
582 tumorigenic response is a linear relationship, which is consistent with much of the available mechanistic and  
583 quantitative data and strongly supports the scientific rationale for the LNT model as used in radiation protection  
584 (UNSCEAR, 2000; 2006). A departure from linearity was noted for leukemia data, for which a linear-quadratic  
585 function was used. It was noted that linear or linear-quadratic functions are used for representational purposes only  
586 in evaluating possible radiation risks and that the actual response may involve multiple and competing processes  
587 that cannot be separately distinguished.

588  
589 Based on the available epidemiologic data, UNSCEAR derived risk estimates and noted, as a first  
590 approximation, linear extrapolation of the estimates at 1 Sv could be used for estimating solid cancer risks at lower  
591 doses. The rationale was re-evaluated in the UNSCEAR 2006 Report, Effects of Ionizing Radiation Vol. 1  
592 (UNSCEAR, 2008), that included several new cancer sites and used Bayesian methods for the incorporation of  
593 dose uncertainty in the atomic-bomb survivor cohort risk, and concluded that “the data reviewed for its 2006  
594 report do not necessitate changes in its current risk estimates for cancer and the hereditary effects of radiation.”  
595 However, this conclusion was based primarily on the LSS dose-response data following a high dose rate brief  
596 exposure and not on an evaluation of cancer excesses after exposure at low dose rates.

597  
598 The National Academies (NA) published the Biological Effects of Ionizing Radiation (BEIR) VII report  
599 (NA/NRC, 2006) that concluded that the available biological and biophysical data support a LNT risk model,  
600 whereby the risk of cancer proceeds in a linear fashion at lower doses without a threshold. The U.S. Environmental  
601 Protection Agency (EPA) in evaluating radiogenic risk models (EPA, 2011) noted that in general, results from  
602 epidemiologic and radiobiologic research are consistent with an LNT dose-response model in which the risk of  
603 inducing a cancer in an irradiated tissue by low doses of radiation is proportional to the dose to that tissue, while  
604 acknowledging that new research might conceivably lead to revisions in the future. In contrast, a report from the  
605 French Academy of Sciences (Tubiana *et al.*, 2005) that focused primarily on radiobiology raised doubts about the  
606 validity of using the LNT model for evaluating carcinogenic risks at low doses and suggested that since biological  
607 mechanisms and responses appear different at low doses and high doses, an empirical relationship of linearity  
608 validated at only doses >150 mSv may lead to an overestimation of risks at low doses.

609  
610 Box (1979) concluded that “all models are wrong but some are useful”. The LNT model is an assumption that  
611 has not been and likely cannot be scientifically validated in the low-dose range. Other dose-response relationships  
612 for the mutagenic and carcinogenic or detrimental effects of low-level radiation cannot be excluded, and there are  
613 notable exceptions to the LNT relationship seen in experimental and epidemiologic studies (Boice, 2015c; Dauer  
614 *et al.*, 2010). Nonetheless, on the basis of the scientific knowledge to date the current judgment by national and  
615 international scientific committees is that no alternative dose-response relationship currently appears more  
616 pragmatic or prudent for radiation protection purposes than the LNT model.

617

## 618 **2.2 LNT and the Estimation of Cancer Risk**

619

620 As part of the process for developing nominal dose limits for radiation protection purposes, it is the current  
621 practice of ICRP, for example, to calculate total health detriment values for exposure to low doses and low dose  
622 rates of radiation (ICRP, 2007). Detriment values are based largely on the risk estimates for fatal cancers, nonfatal  
623 cancers and heritable effects (so-called stochastic effects) and also factors such as quality of life and adjustment for  
624 DDREF. Until recently, little consideration has been placed on noncancer effects (harmful tissue reactions,  
625 previously called deterministic effects) for the calculation of nominal risk, largely because it has been assumed that  
626 noncancer effects have quite large threshold responses and that cancer is dominant at low doses and low dose rates.  
627 This assumption is being reassessed, most specifically for cataracts and cardiovascular diseases (ICRP, 2012;  
628 NCRP, 2007). How to use this information, if indeed it is to be used, for noncancer responses in a detriment  
629 calculation remains a matter of scientific debate and clearly requires additional human data on radiation-induced  
630 adverse health outcomes.

631

632 The process of estimating risks for adverse health outcomes (cancer and noncancer) has relied almost  
633 exclusively on the available human epidemiology data from exposed populations, in particular the survivors of  
634 the atomic bombs in Japan, but with additional support from other exposed populations, including those exposed  
635 occupationally, environmentally, or from medical diagnostic and treatment procedures. Rather little use of the  
636 extensive radiobiology data has been made in the risk assessment process, with the exception of calculations of  
637 the DDREF and radiation weighting factors.

638

639 The general approach used by ICRP, EPA, NA/NRC (2006) and NCRP for cancer risk estimation used for  
640 protection purposes has been to develop a dose-response curve for all solid cancers assessed in the LSS  
641 following acute exposures that are highly influenced by the mid to high dose ranges and to extrapolate from this  
642 range to estimate cancer frequencies at low doses assuming no threshold, a LNT extrapolation. A DDREF is

643 applied to the slope of the linear extrapolation to estimate the cancer risk at low dose rates and often also for  
644 low doses. It is important to note that the use of an LNT extrapolation model is really a default approach  
645 because of a lack of definitive evidence to the contrary (Preston, 2003). Considerations of nonlinear  
646 extrapolations for solid cancer risk from high-to-moderate doses are continually being investigated and received  
647 some support from the recent studies on solid cancer incidence and mortality for the LSS (Grant *et al.*, 2017;  
648 Ozasa *et al.*, 2012).

649  
650 A number of uncertainties are associated with the current approach to apply the scientific evidence for  
651 radiation protection, especially for DDREF and also with the model chosen for extrapolation (LNT) (NCRP,  
652 2012). It is difficult to conduct epidemiology studies that will allow for direct measurement of adverse health  
653 outcomes at low doses and dose rates, although the ongoing Million Person Study (Boice, 2012a; 2017) can  
654 enhance the assessment at low doses and low dose rates. The way forward is most likely to include an integration  
655 of epidemiology and radiobiology data (NCRP, 2015).

656

### 657 **2.3 Objective and Scope**

658

659 This Commentary is to provide a review of recent data from new epidemiologic studies and data of  
660 populations exposed to radiation at low dose rates and to review the new data from the Life Span Study of atomic-  
661 bomb survivors. The purpose is to determine whether these epidemiologic studies broadly support the LNT model  
662 of carcinogenic risk as used in radiation protection or, on the contrary, whether there is sufficient evidence that the  
663 LNT model is inappropriate? The strength of support for the LNT model is evaluated for solid cancer incidence or  
664 mortality and secondarily for low dose-rate studies of leukemia. The focus is on new human studies on low doses  
665 and low dose rates. In addition, the report will briefly review current evidence regarding certain noncancer  
666 outcomes, such as cardiovascular diseases, and risk from childhood exposure and heritable radiation risk.

667

668 This Commentary was written by a committee of multi-disciplinary experts based on a comprehensive review  
669 of recent (within approximately 10 y) relevant epidemiologic studies, especially those that have been extensively  
670 studied, with attention paid to epidemiologic methodology, dosimetry and statistical approaches. The Committee  
671 performed a critical but balanced evaluation of these epidemiologic studies, including a description of their  
672 strengths and limitations, similar to the approach utilized in recent related reviews by UNSCEAR (2008; 2013).  
673 The present evaluation includes a detailed assessment of the dosimetric and statistical approaches employed for the  
674 epidemiologic study. The aim was to develop a perspective for each study and evaluate its strength with regard to  
675 radiation protection implications. Future directions and ongoing research needs were identified.

676

677 **3. Important Considerations**

678

679

**3.1 Epidemiologic Considerations**

680

681 Epidemiologic studies of humans provide evidence that is critically important in establishing disease causation.  
682 Epidemiology is by nature primarily observational rather than experimental. Consequently, in virtually all  
683 epidemiologic studies there is always the possibility that biases or confounding factors of various sorts may give  
684 rise to spurious results. A well-designed study should attempt to minimize any potential biases and avoid or control  
685 for the effects of confounding factors.

686

687 Bias in a study can produce results or conclusions that differ systematically from the truth (Sackett, 1979).  
688 Therefore, high-quality studies will recognize and address various types of bias, including: selection and  
689 participation bias, dose-estimation bias, follow-up bias, disease ascertainment bias, recall bias, and others. Because  
690 a risk factor correlates with a disease (*e.g.*, smoking and lung cancer) does not necessarily mean it confounds the  
691 radiation association with that disease. It can confound the radiation-disease association only insofar as the risk  
692 factor is also correlated with the amount of radiation exposure. A few examples of mild (Cardis *et al.*, 2007) or  
693 even serious (Akiba and Mizuno, 2012) confounding by smoking or alcohol consumption have been seen in worker  
694 cohorts, but a few worker and environmental-exposure studies have found little evidence of such confounding (*e.g.*,  
695 the LSS, INWORKS, Mayak and Techa River studies) but individual data on specific confounding factors was  
696 often missing or only partially available and indirect ways to evaluate confounding were considered.. Confounding  
697 by indication or “reverse causation” has been seen in medical exposure studies (Berrington de Gonzalez *et al.*,  
698 2016; Boice *et al.*, 2015b; Journy *et al.*, 2015; UNSCEAR, 2013). In high-quality studies, confounding factors can  
699 usually be addressed at the analysis stage if information on the factors is available, either through regression  
700 procedures or stratification of the data according to the levels of the confounding factor (UNSCEAR, 2008), but the  
701 confounding factor data must be available for individuals. Confounding also can be indirectly assessed by  
702 examining whether target organs of the risk factors (*e.g.*, lung for smoking, liver for alcohol or hepatitis infection)  
703 show aberrantly high or low associations with radiation dose levels.

704

705 To convincingly establish causality, a number of criteria are relevant. However, it is not necessary to show  
706 that all the criteria are met to the same extent to be able to make a causal inference about disease risk. While  
707 exposure must, of course, precede outcome, another one of the most important criteria of causation is that  
708 there be a quantitative relationship between the factor of interest and the disease. The most convincing  
709 evidence is quantitative in nature, where the quantification is based on individuals’ exposure levels and health  
710 outcomes. In the case of ionizing radiation, personal data on exposure level and subsequent occurrence of

711 disease are needed to quantify the association. This is commonly referred to as a “dose response.” Almost as  
712 important as the existence of a dose gradient, it is very important that similar studies conducted by others  
713 around the world come up with similar results, *i.e.*, that there is consistency among multiple studies  
714 (UNSCEAR, 2008). That is, if one study shows a dose response but most other studies do not, then causal  
715 interpretations are tempered. In the current evaluation of the LNT model, final conclusions will be based on  
716 consistency of results in the nearly 20 epidemiologic studies evaluated and not on a single or very few  
717 investigations.

718  
719 In the context of radiation protection, comparable strength of the dose-response data across the spectrum of  
720 exposure, from the highest to the lowest would be ideal. Unfortunately, the data from epidemiologic studies do  
721 not provide a clear picture of the shape or magnitude of the dose response at all levels of exposure. Of particular  
722 interest is a better characterization of the dose response at low doses (100 mGy and below) and low dose rates  
723 because it is that part of the dose response curve where most human exposures to ionizing radiation occur today  
724 and are likely to remain so in the future, barring a major accident, terrorist event, intentional nuclear weapon  
725 detonation or natural disaster. Estimating risk in this dose range has not been possible with much certainty. A  
726 brief review of the primary epidemiologic research study designs illustrates some of the important strengths and  
727 limitations of epidemiologic methods.

728

729 There are three basic study designs that have been employed in radiation epidemiology are:

730

- 731 • cohort studies;
- 732 • case-control studies; and
- 733 • descriptive and/or ecologic studies.

734

735 Epidemiological investigations that quantify radiation effects are usually cohort studies and, to a much lesser  
736 extent, case-control studies. In a cohort study, a defined population (preferably with a wide range of exposures)  
737 is followed forward in time to examine the occurrence of effects. Such a study may be performed either  
738 prospectively (*i.e.*, by following a current cohort into the future) or retrospectively (*i.e.*, by constructing a cohort  
739 of persons alive at some time in the past and following it forward, possibly to the current time) (UNSCEAR,  
740 2000). Examples of informative cohort studies include the LLS study of Japanese atomic-bomb survivors (Ozasa  
741 *et al.*, 2012), the INWORKS study (Richardson *et al.*, 2015), and the Massachusetts tuberculosis study of patient  
742 monitored repeatedly with x-ray fluoroscopic examinations of the chest (Boice and Monson, 1977).

743

744 In case-control studies, people with and without a specified disease (the cases and controls, respectively) are  
745 compared and differences in exposures are examined (UNSCEAR, 2000). Some case-control studies are nested  
746 within a cohort study, in that the cases and controls are selected from the cohort. The nested case-control study  
747 design is used when it is difficult to obtain estimates of radiation dose or other exposures for all members of a  
748 cohort, but possible to collect them for a smaller number of individuals (*e.g.*, Cardis *et al.*, 2005; Kendall *et al.*,  
749 2013; Krille *et al.*, 2015; Schubauer-Berigan *et al.*, 2015; Zablotska *et al.*, 2013a). Of the approximately 25  
750 studies providing some quantitative information on radiation risk and thus relevant to LNT evaluation, there  
751 were only two referring to case-control studies which suggests a rather limited influence of this design in  
752 evaluating the LNT hypothesis for use in radiation protection.

753

754 The cohort design is less susceptible to biases than the case-control study design and has a number of  
755 important advantages over other observational designs:

756

- 757 • exposure is characterized without knowledge of disease status;
- 758 • disease free status is firmly established in a uniform way at the beginning of follow- up for the entire  
759 cohort;
- 760 • all cases of the disease under study are identified (with complete follow-up) incident cases and/or  
761 deaths;
- 762 • incidence/death rates can be calculated directly from the study and can be compared to assess risk  
763 across dose categories; and
- 764 • multiple outcomes can be evaluated in a single study.

765

766 A third type of observational epidemiologic study is often referred to as an ecologic study. These geographical  
767 (or temporal) correlation studies are those in which disease rates based on data aggregated over geographical areas  
768 (or time periods) are compared with aggregated data on levels of exposure. In such studies, groups rather than  
769 individuals are the unit of analysis and the correlation between disease rates and the groups' average levels of  
770 exposure is studied, with an intention to infer disease risk for individuals. These studies sometimes provide a good  
771 overview of the distribution of the disease of interest according to person (*e.g.*, age, sex and race), place and time.  
772 However, since the analyses are not based on individual-level data and are generally not able to control for  
773 possible confounding factors or effect modifiers. Such studies can be useful in generating hypotheses, but not in  
774 testing hypotheses. Because the exposure is not at the individual level, ecological studies cannot provide  
775 meaningful data regarding a dose response, and cannot be used to infer disease risk, and cannot be used to evaluate  
776 the LNT model.

777

778 In short, “When analyzing aggregated [ecological] data, we not only lose all ability to extend inferences  
779 reliably to less aggregated data but we even lose the ability to estimate the direction and magnitude of bias. We  
780 cannot rely on the addition of more grouped data to eliminate the bias” (Piantadosi, 1994). In general, “The  
781 investigator is never justified in interpreting the results of ecological analyses in terms of the individuals who give  
782 rise to the data” (Piantadosi *et al.*, 1988). Studies of this type therefore are not useful in determining the shape of  
783 the radiation dose response in humans (Brenner *et al.*, 1992; Greenland and Robins, 1994), and cannot substantiate  
784 or challenge the LNT model.

785

786 In addition, other types of epidemiologic studies include the randomized clinical trial (RCT), cross-sectional  
787 studies, case-cohort studies and variations of the case-control study using counter matching. An RCT, if the  
788 randomization is conducted properly, should not be subject to any biases (it is essentially a human experiment),  
789 and is generally regarded as the epidemiologic “gold standard.” There are few RCTs relevant to dose-response  
790 modelling, except perhaps the high-dose studies of breast cancer treatment and resulting deaths due to  
791 cardiovascular disease.

792

793 When assessing exposure-disease associations, cross-sectional studies (also called prevalence studies) are  
794 similar to ecological studies in providing limited evidence on the possible shapes of dose-response  
795 relationships, and are considered to be hypothesis-generating or exploratory. Cross-sectional studies measure  
796 exposure and health outcome simultaneously. They tend to assess the presence (prevalence) of the health  
797 outcome of interest at a single point of time without regard to duration. However, these should be distinguished  
798 from cross-sectional assessments of health outcomes of individuals who had previous quantified radiation  
799 exposures. Such studies can provide legitimate estimates of radiation risk (*e.g.*, Imaizumi *et al.*, 2006; 2015),  
800 albeit with potential qualifications regarding follow-up bias.

801

802

### 3.2 Dosimetry Considerations

803

804 A major objective of this Commentary was to provide detailed reviews of the dosimetry underpinning each  
805 epidemiologic study to gain a better understanding of the study’s strengths and weaknesses. Analysis of the  
806 dosimetry helped evaluate the robustness of each epidemiologic study in supporting the shape of the dose response  
807 curve at low doses and low dose rates. High quality dosimetry is essential in evaluating any dose response from  
808 radiation exposure especially at low doses (<100 mGy). because small numbers of excess cancers lead to  
809 uncertainty in the estimation of risk (*e.g.*, ERR Gy<sup>-1</sup>). Furthermore, it is now well established that shared dose  
810 uncertainty (*e.g.*, uncertainty in a source term) can result in further underestimating the uncertainty bounds in

811 ERR Gy<sup>-1</sup>, in addition to Berkson or random uncertainties (Kwon *et al.*, 2016; Land *et al.*, 2015; Stram *et al.*,  
812 2015). Further, unshared classical error (*i.e.*, random individual dosimetry error), if present, can bias the dose  
813 response toward the null (Stram *et al.*, 2015; UNSCEAR, 2014). However, adjustment for shared, Berkson and  
814 random measurement uncertainties is unlikely to change a significant dose response to a non-significant response,  
815 *i.e.*, if the confidence bound for a risk estimate does not include the null value, the uncertainty-adjusted bound  
816 usually will not include the null value either (Stram *et al.*, 2015).

817  
818 There does not seem to be a reasonable case that the positive dose-response associations that are consistent  
819 with a LNT model are due to dosimetry inaccuracies, especially for studies with measured doses. If one can assume  
820 that estimated individual doses in the various studies were well correlated with the true doses, and that the  
821 estimated doses were not based on individuals' disease status (which was generally true, except possibly for the CT  
822 studies), then dose errors would be unlikely to induce an apparent dose response where one does not exist. If  
823 anything, random ('classical') dose error will tend to diminish rather than heighten the slope of the dose-response  
824 model, while unbiased shared error will have little effect on the statistical significance. On the other hand, if  
825 individual doses were imputed based on a dose reconstructions from limited information, there may be unknown  
826 biases in the shared-dose estimates, but dosimetrists involved in the major studies have devoted much effort to  
827 providing reasonably accurate estimates of shared doses.

828  
829 Aspects of the dosimetry review have been reported by Till *et al.* (2017). It was carried out by a team  
830 working independently of the epidemiologic team and used evaluation criteria specifically focused on dose  
831 measurements and dose reconstructions. The criteria were divided into several categories:

- 832
- 833 • general study characteristics;
  - 834 • dose assignment;
  - 835 • uncertainty;
  - 836 • dose confounders;
  - 837 • dose validation; and
  - 838 • overall strengths and weaknesses of the dosimetry.
- 839

840 A template created for each study addressed these criteria and subcategories within them. The templates were  
841 used in the dosimetry review which became an integral part of the evaluation of the reported dose response  
842 relationships. A matrix summarizing the dosimetry characteristics of each epidemiologic study facilitated  
843 evaluations and comparisons of dosimetry quality. A summary of the dosimetry is included in the review of each

844 epidemiologic study in Section 4. In general, the most likely impacts of dose uncertainty are to reduce the  
845 statistical power of a study and (particularly in the case of unshared errors) to bias risk estimates toward the null. It  
846 is important, therefore, for studies to appropriately address and provide effective allowance for the effects of  
847 uncertainty in dose estimates; this may require more than a simple measure of the uncertainty of individual dose  
848 estimates themselves (Beck *et al.*, 2017; Bouville *et al.*, 2015; NCRP, in press; Stram *et al.*, 2015; Till *et al.*, 2014;  
849 2017). Additional comments on the importance of dosimetry in epidemiologic studies as well as recommendations  
850 for future considerations in planning and implementing studies ERR Gy<sup>-1</sup> are provided in the conclusions.

851

852

### 3.3 Statistical Considerations

853

854 Design and analysis features are relevant when assessing the evidence for and against the LNT model  
855 hypothesis from statistical details given in published studies. The type of study design (*e.g.*, cohort and case  
856 control) and the statistical analysis method applied (*e.g.*, Poisson regression with grouped data or Cox regression  
857 with individual data) are important features to consider. Such methods are often regarded to be the least open to  
858 bias.

859

860 The statistical precision of a study is a key determinant of the study's contribution toward evaluating the shape  
861 and slope of the dose response risk for detrimental health outcome. The precision is related to the statistical power  
862 of the study, and both depend on study features such as: the number of persons at risk and their years of  
863 observation (person-years at risk), the number of exposed and unexposed or minimally exposed cases with the  
864 health outcome of interest, the length of epidemiologic follow-up time, dose range and dose distribution within  
865 cohort members, the magnitude of the radiation effect, the availability of information on possibly explanatory  
866 covariables (that is, confounding factors). A consideration of estimated influences of any unavailable potentially  
867 explanatory covariables that might induce bias should be included. It is important to examine the choices made for  
868 the mathematical forms applied for the radiation dose effect, the modification of the radiation effect by other  
869 variables (*e.g.*, sex; time), and adequacy of modeling of the baseline rates of the outcome of interest. To assess the  
870 main dose-response model, functional forms such as linear, quadratic, linear-quadratic, nonparametric, categorical  
871 (the risk in each category of a set of predefined dose categories), and dose threshold ideally should be examined.  
872 Similarly, appropriate modeling of other sources of radiation exposure should be considered. For example, in  
873 assessing radiation related cancer risk per unit occupational gamma organ dose, adjustments may be needed to  
874 account for other types of radiation exposure, such as alpha particles or neutrons, or for other sources of external  
875 exposure, such as medical x-ray organ doses. Some investigators do not adjust but include the organ dose from  
876 other known exposures, *e.g.*, the LSS of atomic-bomb survivors incorporates the small neutron dose into the  
877 measure of “weighted Gy” organ dose to colon used for analyses. Interpretation, however, becomes difficult if the

878 organ dose from other exposures exceeds the gamma ray dose. Analyses excluding the organs with substantial  
879 dose from other occupational exposures, *e.g.*, internal intakes of alpha particle emitting radionuclides, are also  
880 conducted.

881  
882 In the baseline model it is usually appropriate to adjust for sex, age at exposure, attained age and sometimes  
883 calendar period or birth cohort to avoid confounding, as well as to explore whether those variables may be effect  
884 modifiers of the radiation dose response. When information on smoking, alcohol-intake or other lifestyle or  
885 sociodemographic factors is available, it is important to examine whether it may be a confounder and/or an effect  
886 modifier; often a measure of socioeconomic status is used in the analysis as an adjustment factor to reflect lifestyle  
887 variables, especially for the lung and other smoking-related sites.. Sometimes it may be appropriate to adjust for  
888 factors such as duration of employment in worker studies, and medical risk factors (*e.g.*, obesity or diabetes) for  
889 some types of outcomes such as cardiovascular disease.

890  
891 An important activity in assessing the evidence for and against the LNT model is to look for consistency of  
892 patterns in risks between a simple parametric model and a nonparametric model (*e.g.*, Pierce *et al.*, 2000) or dose-  
893 category specific risk model. The degree of consistency can either be assessed qualitatively or by applying  
894 statistical methods to provide quantitative evaluations. However, the limited statistical power of a particular study  
895 may restrict the number of parameters that can provide useful estimates, with informative confidence intervals, in  
896 both parametric and nonparametric types of models. Consequently it may be important to consider nonlinear one-  
897 parameter dose-response models, such as one purely quadratic in dose, when there is not enough statistical power to  
898 support a linear parameter and a nonlinear parameter in the same model. Epidemiologic studies are prone to many  
899 potential sources of bias and the results may be influenced by one or several of the many different possible types of  
900 bias, so thorough consideration of all sources of bias is also critically important in assessing evidence for and  
901 against the LNT model hypothesis.

902  
903 The studies reviewed in Section 4 have mostly used Poisson regression analyses and have adjusted for sex,  
904 age at exposure and attained ages at risk. Cox proportional hazards analyses are also used for estimating risk, *e.g.*,  
905 ERR Gy<sup>-1</sup>. Different statistical methods or notable changes in adjustment variables will be mentioned in the  
906 discussion of the individual studies when relevant.

907

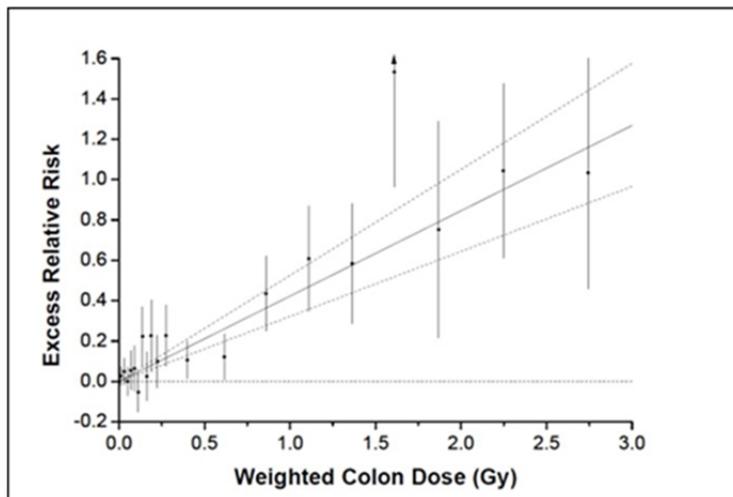
### 3.4 Dose-Response Considerations

#### 3.4.1 Dose Response and Solid Cancers: Cancer Type and Sensitivity

As an example to demonstrate dose-response considerations, we use the dose response for all solid cancers combined as obtained from the Life Span Study (LSS) data for mortality during 1950 to 2003 among the cohort of Japanese atomic-bomb survivors is shown in Figure 3.1. The proportionality of ERR of all solid cancers to the weighted absorbed colon dose (which includes a small neutron component) received during the bombings is apparent from the figure, although this is most evident for moderate and high doses ( $>0.1$  Gy), and the nature of the dose response for low doses ( $<0.1$  Gy) is not clear.

However, it must be borne in mind that by considering all solid cancers combined, any individual structure that may be present in the dose responses of the component solid cancer types may have been lost. This may be appreciated from Figure 3.2 showing the dose-averaged ERR for individual sites of solid cancer. While there is clear evidence of excess risks for certain cancer types, such as bladder, breast, colon and stomach, this is not the case for certain cancer types such as uterus, pancreas, testes, prostate, renal cell, rectum, bone and soft tissues, in part due to the relatively small numbers affected. Data on non-melanoma skin cancer incidence indicate substantial curvilinearity, consistent with a possible dose threshold of about 1 Sv to the skin (Little and Charles, 1997; UNSCEAR, 2000). Under these circumstances, whether it is valid to assume that each site of solid cancer may be adequately represented by the dose response for all solid cancers combined needs careful consideration, and at the minimum, different slopes may pertain for different sites of solid cancer, (Preston *et al.*, 2007), though there is some ambiguity in interpreting variations in slope because of statistical considerations (Pawel *et al.*, 2008). Indeed, one could not safely assume, on the basis of these LSS data, that radiation can cause certain solid cancers (for example, cancer of the gallbladder in addition to those mentioned above). There is also heterogeneity in the risk estimates for different histological types of some cancers, *e.g.*, the radiation risks for non-melanoma skin cancer a strong association is seen for basal cell carcinoma but there is little evidence of an association for squamous cell carcinoma (Preston *et al.*, 2007). Although for the purposes of radiation protection it has been considered prudent to adopt a linear dose response without a threshold dose for all solid cancers combined, this dose response may be driven by those cancer sites having a relatively high incidence and comparatively pronounced linear dose responses (such as breast cancer and thyroid cancer).

939



940

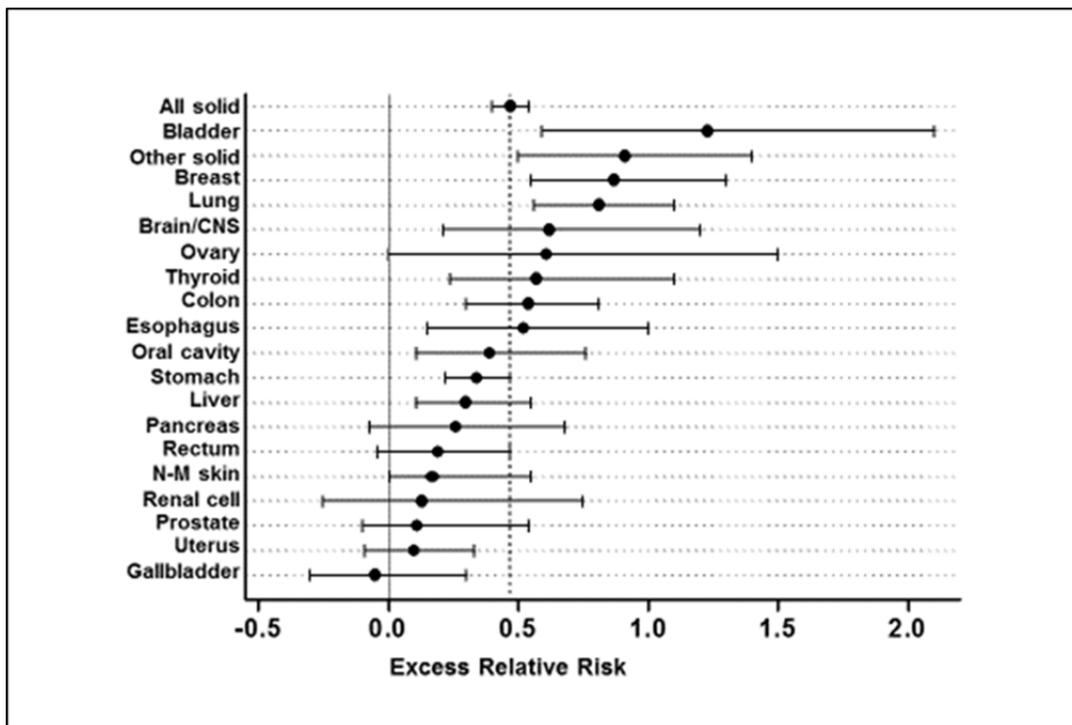
941

942

943 **Fig. 3.1.** Dose response of all solid cancer mortality among the LSS, 1950 to 2003. The solid line is the fitted  
944 linear, sex-averaged ERR dose response, and the dashed lines are its 95 % confidence range. The points are  
945 categorical estimates of the ERR in dose categories and the bars are their 95 % confidence intervals. The  
946 categorical estimates indicate a larger uncertainty in the risk estimate at low dose levels than that reflected in the  
947 linear fit because the confidence range of the linear fit is determined mostly by the association at higher dose levels  
948 (Kamiya *et al.*, 2015).

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**Fig. 3.2.** ERR Gy<sup>-1</sup> for the incidence of site-specific solid cancers in the LSS cohort. The risk is standardized as exposure at 30 y of age and estimated sex-averaged risk at age 70 y. The horizontal bars indicate 90 % confidence intervals (Kamiya *et al.*, 2015; Preston *et al.*, 2007).

974 **3.4.2** *Dose Response and Leukemia*

975

976 That the dose response may vary among different cancer types is illustrated by the difference between the dose  
977 response for all leukemias and that for all solid cancers. The dose response for leukemia is linear-quadratic in form  
978 whereas for solid cancers combined it is closer to linear, and the slope of the linear (low dose) segment of the  
979 leukemia dose response is steeper than the slope of the solid cancer dose response. Further, the dose responses for  
980 the main types of leukemia appear to differ, with acute myeloid leukemia exhibiting upward curvature while acute  
981 lymphoid leukemia and chronic myeloid leukemia show a linear dose response (Hsu *et al.*, 2013), with CLL having  
982 a low sensitivity to radiation induction (and technically speaking, CLL is actually a type of low-grade non-Hodgkin  
983 lymphoma).<sup>1</sup> This heterogeneity seen for leukemia may well extend to the various sites of solid cancer, and this  
984 possibility must be kept under consideration when assessing the nature of the dose response for solid cancer.

985

986 **3.4.3** *DDREF Considerations*

987

988 Cancer risk estimates obtained in the LSS for survivors of the atomic bombs are based on acute exposures,  
989 driven by the medium-to-high exposure levels. To predict effects at low doses and low dose rates, an  
990 extrapolation model (LNT) is used together with a DDREF that adjusts the slope of the linear curve for solid  
991 cancer. The value of this DDREF commonly used for radiation protection purposes ranges from about 1.5 (*e.g.*,  
992 NA/NRC, 2006) to 2 (*e.g.*, ICRP, 2007), although a value of one also has been proposed by the German  
993 Commission on Radiological Protection (SSK, 2014). The approaches for obtaining these values differ in a  
994 number of ways as do the data sets selected for use for the calculation. The data used are, broadly, cancer data  
995 from the LSS and studies with low dose rates, animal cancer data and cell and molecular radiobiology data.  
996 There is a high degree of uncertainty with the approaches used in all cases (NCRP, 2012).

997

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<sup>1</sup> See <https://www.lymphomas.org.uk/about-lymphoma/types/non-hodgkin-lymphoma/chronic-lymphocytic-leukaemia-ctl>

998 **4. Review of Epidemiologic Studies of Cancer and Genetic Effects from Low-Dose or Low Dose-Rate**  
999 **Irradiation**

1000  
1001 The primary question to be addressed is whether the new epidemiologic evidence sufficiently supports a LNT  
1002 model as a reasonable basis for radiation protection. This section provides brief reviews of a number of the  
1003 epidemiologic studies that contribute to an evaluation of quantitative information regarding the LNT model.  
1004 Because such evaluations are quantitative of necessity, most studies that rely on standardized mortality ratios  
1005 (SMRs) or on other simple group comparisons were not included, as they typically do not have quantitative  
1006 information regarding individual exposure levels, and SMRs of worker populations in particular have inherent  
1007 biases due to “healthy worker” selection effects that can be addressed, but interpretation of results uncertain.

1008  
1009 This review considers recent studies on occupational and environmental and a few medical radiation  
1010 exposures, although most medical exposures are typically limited to select organs. Brief summaries are provided  
1011 for *in utero* and childhood exposures (including pediatric CT examinations) and for heritable genetic outcomes.

1012  
1013 As an overview of a number of the important studies described in this Commentary, four summary tables are  
1014 given. The first table provides a description of the study designs, characteristics of the study populations and  
1015 underlying epidemiologic databases (Table 4.1). Table 4.2 provides a summary of the types of exposures  
1016 received, the dosimetry and the mean and range of doses for individual studies. Table 4.3 provides information  
1017 about the statistical methods and the epidemiologic and statistical results of each study. Table 4.4 gives a  
1018 summary of the strengths and weaknesses of each study as they pertain to the evaluation of the LNT model.

1019  
1020 In Section 4 a critical review is given for each major study. The epidemiologic evaluation includes a  
1021 characterization of the study design and study population, quality of the data available, data collection  
1022 methodology, the degree to which potential confounding variables or biases were assessed, and the quantitative  
1023 results. The critique includes an assessment of the comprehensiveness, quality and uncertainties in the dosimetry  
1024 used in each study, whether the analytic methods were appropriate and whether each study considered statistical  
1025 alternatives to a linear dose-response trend.. A brief statement characterizes the contribution the study makes on  
1026 providing information relevant to the LNT model and to radiation protection For selected studies, a more detailed  
1027 assessment of the dosimetry and the statistical and epidemiologic methodologies is hoped to be provided as  
1028 subsequent publications in the literature, as was the case recently for dosimetry (Till *et al.*, 2017).

1029  
1030 The more robust studies have a number of strengths, with relatively good quality dosimetry, good cohort  
1031 mortality/morbidity ascertainment, attention to potential confounding variables, and proper analysis. Nevertheless,

Table 4.1—Study population characteristics.

Study (Reference) [section of text]	Study Design Mortality/Incidence Country	Study population characteristics (N, PY, % female)	Mean Age at First Exposure or Study Entry (y)	Dates Follow-up (mean years of follow-up)	Information on Follow-up (F-U) and Cancer Ascertainment
1. Japanese atomic-bomb survivors (Grant <i>et al.</i> , 2017; Hsu <i>et al.</i> , 2013; Ozasa <i>et al.</i> , 2012; Preston <i>et al.</i> , 2007) [Section 4.1]	Cohort, Mortality and cancer incidence; Japan	N = 93,741 PY = 3,079,484 F = 59 %	~29 <sup>a</sup>	Mort: 1950–2003 (mean = 38.0 y) Solid cancer incid: 1958–2009 (mean = 29.2 y) Hematologic cancer incid: 1950–2001	Mort: >99.5 % follow-up (F-U) and ~98 % of deceased with death certificates; Incidence: Limited to cancer registries in 2 prefectures where ~85 % live – rates adjusted for out-migration. Good ascertainment and diagnosis.
2. 15-Country nuclear workers (Cardis <i>et al.</i> , 2007) [Section 4.2.1]	Pooled cohorts, Mortality; Europe, N. America, Asia, Australia	N = 407,391 PY = 5,192,710 F = 10 %	30.7	Maximum: 1943–2000 (mean = 12.7 y)	High F-U and death certificate ascertainment
3. INWORKS nuclear workers (Leuraud <i>et al.</i> , 2015; Richardson <i>et al.</i> , 2015) [Section 4.2.2]	Pooled cohorts, Mortality; UK, US, France	N = 308,297 PY = 8,222,000 F = 13 %	30.3	Maximum: 1944–2005 Mean = 26.7 y	High F-U and death certificate ascertainment

Study (Reference) [section of text]	Study Design Mortality/Incidence Country	Study population characteristics (N, PY, % female)	Mean Age at First Exposure or Study Entry (y)	Dates Follow-up (mean years of follow-up)	Information on Follow-up (F-U) and Cancer Ascertainment
4. Mayak nuclear workers (Hunter <i>et al.</i> , 2013; Sokolnikov <i>et al.</i> , 2015) [Section 4.2.3]	Cohort, Incidence and Mortality; Russia	N = 25,757 PY = 950,896 F = 25 %	24.7	Employed 1948–1982, follow-up 1948– 2008 (mean = 36.9 y)	23 % lost to F-U, mainly due to out-migration; Cause of death for >99 % of known deaths, but 9 % from relative reports; autopsy data for 21 % of deaths
5. Chernobyl Russian nuclear cleanup workers (Ivanov <i>et al.</i> , 2007; Kashcheev <i>et al.</i> , 2015; Kryuchkov <i>et al.</i> , 2009) [Section 4.2.4]	Cohort, Incidence and Mortality; Russia	N = 67,568 PY = mort 993,423; incid 972,660 F = 0 %	34.8 (at exposure; study entry ~6 y later)	1992–2009 (exposures in 1986–1987); (Mean = 14.7 y)	7 % lost to F-U Causes of death confirmed by multiple documents
6. Canadian nuclear workers (Zablotska <i>et al.</i> , 2013b)	Cohort, Mortality; Canada	N = 42,228 PY = 514,729 F = 16.8 %	30.6 y	1956–1994 (mean = 12.2)	2.4 % lost to follow-up; Cause of death for 99.9 % of deaths
7. Japanese nuclear workers (Akiba and Misuno, 2012; Hosoda <i>et al.</i> , 1997; Iwasaki <i>et al.</i> , 2003; Murata <i>et al.</i> , 2002) [Section 4.2.5]	Cohort, mortality; Japan	N = 200,583 PY ~1,373,000 F = 0	31.7 y <sup>c</sup>	1991–2002 (exposure in 1957–2000) (Mean = 6.8 y)	High rate of follow-up through Japan's <i>koseki</i> system; 99.4 % cause of death ascertainment <sup>d</sup>

Study (Reference) [section of text]	Study Design Mortality/Incidence Country	Study population characteristics (N, PY, % female)	Mean Age at First Exposure or Study Entry (y)	Dates Follow-up (mean years of follow-up)	Information on Follow-up (F-U) and Cancer Ascertainment
8. U.S. radiologic technologists (Liu <i>et al.</i> , 2014; Preston <i>et al.</i> , 2016) [Section 4.2.6]	Cohort, Incidence and Mortality (but no dose response performed for solid cancer mortality data); U.S.	N = 83,538 mort, 66,915 incid PY = 1,089,502 incid F = 100 % (breast cancer)	~22 y <sup>b</sup>	<1940–2008 (Mean = 20.6 y)	~97 % breast cancer detection rate. Based on self-report of breast cancer; 83 % with medical record verification, only 1 % of medical records disconfirmed reported diagnoses.
9. Rocketdyne nuclear workers (Boice <i>et al.</i> , 2011; 2006b) [Section 4.2.7.4]	Cohort, Mortality; U.S.	N = 5,801 radiation monitored + 41,169 non- radiation workers PY = 196,674 (+ 1,392,648 PY non-rad workers) F = 8 %	31 y	1948–2008 (33.9 y)	0.6 % lost to F-U; Death cause known, 98.1 %
10. Mound nuclear workers (Boice <i>et al.</i> , 2014) [Section 4.2.7.4]	Cohort, Mortality; U.S.	N = 7,269 (4,977 monitored for radiation) PY = 293,462 F = 24.8 %	n.a.	1944–2009 (40.4 y)	1.3 % lost to F-U Cause of death known for 98 % of deceased.

Study (Reference) [section of text]	Study Design Mortality/Incidence Country	Study population characteristics (N, PY, % female)	Mean Age at First Exposure or Study Entry (y)	Dates Follow-up (mean years of follow-up)	Information on Follow-up (F-U) and Cancer Ascertainment
11. Chinese x-ray workers (Sun <i>et al.</i> , 2016) [Section 4.2.8]	Cohort, Incidence; China	N = 27,011 x- ray workers; 25,872 unexposed PY = 1,446,347 F = 20 %	25.7	1950–1995 (27.4 y)	Medical x-ray workers in 1950 at major hospitals in 24 Chinese provinces. Question of socioeconomic comparability of exposed (“medical x-ray workers”) and unexposed (“physicians”). Diagnoses: 70 % histologic, remainder mainly radiologic exam.
12. Techa River, residents (Davis <i>et al.</i> , 2015; Schonfeld <i>et al.</i> , 2013) [Section 4.3.1]	Cohort, Incidence and Mortality; Russia	N = 29,730 mort; 17,435 incid PY = 927,743 mort; 472,788 incid F = 58 %	~28 (range 0 - >50)	Mort: 1950– 2007 (31.2 y) Incid: (27.1 y)	16 % lost to F-U through migration, but migrants censored so not a bias. 5.7 % of non-migrants lost to F-U Cause of death for 91 % of non- migrant deaths
13. Chernobyl: Ukrainian and Belarusian childhood <sup>131</sup> I exposure (Brenner <i>et al.</i> , 2011; Tronko <i>et al.</i> , 2006; Zablotska <i>et al.</i> , 2011; [Section 4.3.2]	Cohorts, Incidence Ukraine, Belarus	Ukraine: N = 13,127 F = 51 %. Belarus: N = 11,611 F =	~10 y	Screened 1998– 2000 (~13 y)	Ukraine: 67 % of subjects who could be traced were screened. Belarus: 74 % of eligible, traceable subjects were screened.

Study (Reference) [section of text]	Study Design Mortality/Incidence Country	Study population characteristics (N, PY, % female)	Mean Age at First Exposure or Study Entry (y)	Dates Follow-up (mean years of follow-up)	Information on Follow-up (F-U) and Cancer Ascertainment
14. Kerala, India, HBRA residents (Akiba, 2013; Nair <i>et al.</i> , 2009) [Section 4.3.3]	Cohort, Incidence India	N = 69,958 PY = 736,586 F = 54.1	47 (exposure began in utero)	1990–2005 (10.5 y)	0.7 % lost to follow-up; 6 % out-migration. Had histopathology/cytology on 73 % of cancers
15. Yangjiang, China, HBRA residents (Sun <i>et al.</i> , 2000; Tao <i>et al.</i> , 2012) [Section 4.3.4]	Cohort, Mortality China	N = 31,604 PY = 736,942 F = 49 %	31.7 (exposure began in utero)	1979–1998 (23.3 y)	Visited hospitals, reviewed medical records of deceased subjects every 3 to 4 y. When necessary, visited local village doctors, family members, or next of kin to collect further cause-of-death information. Had pathological information on 26 % of cancer deaths, and radiologic/ultrasonic on 62 %.
16. Taiwan, radiocontaminated dwellings (Hwang <i>et al.</i> , 2008) [Section 4.3.5]	Cohort, Incidence Taiwan	N = 6,242 PY = 118,000 F = 34.4 %	16.9	1983–2005 (18.9 y)	Good cancer ascertainment through Taiwan National Cancer Registry
17. UK pediatric CT patients (Berrington de González <i>et al.</i> , 2016; Pearce <i>et al.</i> , 2012) [Section 4.4.2]	Cohort, Incidence United Kingdom	N = 178,604 PY = 1,720,984 F = 45 %	~10.6	1985–2008 (9.6 y)	~97 % completeness of cancer registry.

Study (Reference) [section of text]	Study Design Mortality/Incidence Country	Study population characteristics (N, PY, % female)	Mean Age at First Exposure or Study Entry (y)	Dates Follow-up (mean years of follow-up)	Information on Follow-up (F-U) and Cancer Ascertainment
18. Australia pediatric CT patients (Mathews <i>et al.</i> , 2013) [Section 4.4.2]	Cohort, Incidence Australia	N = 680,211 exposed; 10,000,000 unexposed PY = ~6,460,000 for exposed F = 47 %	(ages 0– 19)	1985–2007 (9.5 y)	Cancers ascertainment nearly complete through national cancer database and death index.
19. Pooled analysis of external radiation and thyroid cancer (Lubin <i>et al.</i> , 2017; Veiga <i>et al.</i> , 2016) [Section 4.5.3.1]	Cohorts, Incidence Europe, Asia, U.S.	N = 61,155 exposed; 46,439 unexposed PY = 2,588,559 exp.; 1,865,957 unexp. F = 54 %	5.7 (ages 0 – 19)	1920 – 2009 (inclusive dates for various studies) (Mean n.a.)	Varies among studies, but follow- up and cancer ascertainment generally good-to-excellent.

CT = computed tomographic examinations

F = female

F-U = follow-up

HBRA = high natural background area

incid. = cancer incidence

mort. = cancer mortality

N = number of individuals studied

PY = person-years at risk

<sup>a</sup> Based on age distribution given in Ozasa *et al.* (2012) and assuming mean ages of age intervals and mean of 58 for those >50 y of age.

<sup>b</sup> Calculated from age at entry distribution for females with questionnaires as given in Mohan *et al.* (2003)

<sup>c</sup> Data for the subset described in Vrijheid *et al.* (2007b).

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Table 4.2—Summary of dosimetry.

Study <sup>a</sup>	Type of Dosimetry	Types of Exposure	Mean Dose (range) - mGy	Comment on Dosimetry
1. A-bomb survivors (Cullings <i>et al.</i> , 2006; 2017; Young and Kerr, 2005).	Individual dose reconstruction based on physical modeling of reported location, shielding and other factors.	Gamma and neutron dose mostly <1 % of gamma dose	~200 (0 – 4000) Uncertainties: estimated GSDs of 1.25 – 1.55.	Dosimetry validated by mock-up explosion measurements, measurements of physical samples and chromosome aberrations. Doses corrected for uncertainty by regression calibration method.
2. 15-Country (Cardis <i>et al.</i> , 2005b; Thierry-Chef <i>et al.</i> , 2007)	Based on film/TLD badges. Developed database of correction factors for calibration practices, various dosimeters, geometries etc. to adjust and harmonize for various cohort/year combinations. Colon, RBM and lung doses estimated.	X and $\gamma$ rays mainly between 100 and 3,000 Kev. Some tritium exposure. Workers with potential for substantial neutron or internal exposures excluded, as were those with high dose-rate (>250 mGy in a year) exposures.	19.4 (3.8 – 1500) Estimated uncertainty factors (K) of 1.07 – 1.99 and bias factors (B) of 1.01–2.31.	Excluding workers with neutron or internal exposures removed many of the higher-dose individuals; therefore loss of statistical power. Authors concluded that most of the uncertainty was Berkson error, which would have little effect on risk estimates.
3. INWORKS (Thierry-Chef <i>et al.</i> , 2015)	Based on film/TLD badges. Estimated colon, RBM and other organ doses. Evaluated dosimetry comparability across different nuclear facilities and time to identify bias and uncertainties in different dose estimates. Phantoms employed to reconstruct dose for 3 geometries (anterior-posterior, uniform rotational, and isotropic) and different energies.	X and $\gamma$ rays mainly between 100 and 3,000 Kev. Small percent with tritium exposure, but not usually included in dose estimates. Those indicated as monitored for neutron (13 %) or internal exposures (17 %) were flagged, but such exposure information may have been incomplete in the early days.	Colon: 17.4 mGy among all workers; 20.9 mGy among exposed workers (90 <sup>th</sup> percentile 53.4, max. 1332; ~4000 PY with >500 mGy) RBM: 16 mGy, all workers (10 <sup>th</sup> – 90 <sup>th</sup> percentile, 0.0 – 40.8 mGy)	Used flags for neutron exposure in main analysis and for internal exposures in sensitivity analysis. Too little systematic information available to evaluate impact of “missed” dose in the early years or to quantify neutron doses.

Study <sup>a</sup>	Type of Dosimetry	Types of Exposure	Mean Dose (range) - mGy	Comment on Dosimetry
4. Mayak (Sokolnikov <i>et al.</i> , 2015; Vasilenko <i>et al.</i> , 2007a; 2007b)	Film/TLD badges, urine Pu bioassays. Corrected film badge doses account for energy and angular variation for the various film badge dosimeters, and used information on the nature of the fields at various workplaces to improve dose estimates.	X and $\gamma$ rays, neutrons, plutonium. Neutron dose generally not measured but inferred based on estimated neutron to gamma ratios for various workplace environments (NCRP, 2012).	354 mGy (0 - >3 Gy; 17 % >1 Gy, 35 % <0.1 Gy)	External dose estimates by unshielded dosimeters before ~1955 have large uncertainties due to range of photon energies and angular responses, and high-energy beta exposures. Prior to 1960, neutron doses substantial, particularly in the reactor complex, but would have been unrecorded. Poor agreement between plutonium dose estimates from autopsy vs. urinalysis. Tritium and polonium exposures not well characterized.
5. Cleanup, Russia (Chumak <i>et al.</i> , 2008; Kryuchkov <i>et al.</i> , 2009)	Official dose records; 15 % reconstructed. Accuracy of doses differs because different methods of dose assessment were applied including use of individual dosimeters, group dosimeters, or dose-rate measurements at the work place.	$\gamma$ exposures during 1986–1987	132 mGy (0.1 – 1240); 20,992 with 50–100 mGy, 572 with >300 mGy Uncertainties estimated as 0.5 to 3 times estimated doses but may be larger.	Questions regarding official film badge data, conversion of badge reading to organ dose in highly variable directional radiation fields, and uncertainty due to recall.
6. Canadian nuclear workers (Zablotska <i>et al.</i> , 2013b)	Doses from National Dose Registry of Canada, supplemented by additional review of records. Details given in Zablotska <i>et al.</i> (2004).	Gamma- and x-ray, neutron, tritium; other internal exposures rare	21.6 mSv total dose (0 – 491 mSv); 3.02 mSv tritium dose (0 – 169 mSv)	Careful review of dosimetry problems with the earlier reports on this cohort. Found that dose information from Atomic Energy Canada Limited (AECL) was incomplete before 1965.
7. Japanese nuclear workers (Akiba and Misuno, 2012; Hosoda <i>et al.</i> , 1997)	Annual dose records in “Radiation Dose Registration Center for Workers”. Applied quality factors for various types of radiation.	External and internal radiation, 1957–1992 (but virtually all external); X-, gamma-, beta-rays, neutrons	12.2 mSv; (75.4 % <10 mSv, 2.6 % $\geq$ 100 mSv)	Harmonized dose records for technical differences/advances in dose measurements and metrics. Conducted review by facility visits and dose manuals.
8. US radiation technologists (Simon <i>et al.</i> , 2006b; 2014)	680,000 annual badge doses between 1960 and 1997, dose reconstruction based mainly on literature etc. before ~1970.	X ray	42 mGy (badge dose; 37 mGy breast dose)	Individual annual dose estimates were the arithmetic means of 1000 dose realizations to model uncertainties. ~70 % of workers had at least some annual doses estimated by dose reconstruction.

Study <sup>a</sup>	Type of Dosimetry	Types of Exposure	Mean Dose (range) - mGy	Comment on Dosimetry
9. Rocketdyne (Boice <i>et al.</i> , 2006a; 2006b; 2011)	Film/TLD badges 38 % monitored for internal exposures: >30,000 bioassays Calculated doses to 16 organs or tissues	X or $\gamma$ rays; 14 radionuclides	13.5 mSv (0 – 1 Sv)	Obtained dose information on other places worked for cohort members; Primary uncertainties: photon energy, exposure geometry, type of dosimeter
10. Mound (Boice <i>et al.</i> , 2014)	Film/TLD badges. >200,000 polonium urine bioassays, also plutonium and tritium bioassays.	Gamma, alpha, beta emitters, neutrons. Polonium, plutonium, tritium exposures.	26.1 mSv (0 – 939) 4.6 % with >500 mSv	Extensive dosimetry work-up. Obtained dose information on other places worked for cohort members.
11. China x-ray (Sun <i>et al.</i> , 2016)	Dose reconstruction. Simulated measurements for multiple X-ray machines, workplaces and working conditions, protective measures and work histories for 3805 (14.1 %) workers.	25–40 keV X rays	86 mGy – colon dose. 60 % had cumulative doses <50 mGy, and <1 % >500 mGy.	Used simulations to estimate average calendar year doses for x-ray workers and assigned average dose for the year to all who worked that year.
12. Techa River (Degteva <i>et al.</i> , 2009)	Individual dose reconstructions based on estimated village doses (based on distance down-river and distance from shoreline) with adjustments for age, sex etc. Estimated stomach dose	$\gamma$ , <sup>90</sup> Sr, <sup>137</sup> Cs and other radionuclides	35 mGy (0 – 960; <10 % of doses >100 mGy)	External dose peaked in 1951 Substantial revisions to dosimetry not yet reflected in published Epi studies.
13. Chernobyl childhood exposure (Drozdovitch <i>et al.</i> , 2013; 2015; Likhtarev <i>et al.</i> , 2003; 2006; 2014)	Individual measurements of thyroid activity soon after Chernobyl accident. Latest dosimetry takes into account uncertainties.	Internal <sup>131</sup> I was predominant exposure, and small external $\gamma$ exposure.	Ukraine: 670 mGy (0.35 mGy to 42 Gy; 19 % with >1 Gy) 96 % had dose uncertainty GSDs <2 (geometric mean of 1.47) Belarus: 560 mGy (0 to 32.8 Gy)	Ukraine: Evaluated shared and unshared uncertainties using Monte Carlo realizations. Adjusting for dose error resulting in changes in cancer risk estimates of only –11 % to +7 %, because errors modest and partly Berkson (Little <i>et al.</i> , 2014). Belarus dosimetry was similar to Ukraine dosimetry

Study <sup>a</sup>	Type of Dosimetry	Types of Exposure	Mean Dose (range) - mGy	Comment on Dosimetry
14. Kerala HBRA (Nair <i>et al.</i> , 2009)	Took indoor and outdoor ambient spot measurements for each house in the study using NaI scintillometers. Used model incorporating estimated age/sex specific occupancy factors to calculate yearly dose and cumulative doses for each individual. Validation sample of 800 houses with quarterly TLD readings in houses for year correlated 0.97 with scintillometers.	$\gamma$ and also had some radon and thoron exposures.	161 mGy, cumulative	Obtained survey of house occupancy for time spent indoors/outdoors for 2 % sample and used mean occupancies by age and sex to apply to ambient measurements to derive estimated cumulative doses. Had 160 individuals wear TLD badges for 2 months & correlated the badge readings with their modeled dose estimates. Found a correlation of 0.80, but only after discarding 15 % of the badge readings as “outliers”. No dose uncertainty analysis.
15. Yangjiang HBRA (Tao <i>et al.</i> , 2000)	Ambient dose rate survey: about 1/3 of houses and nearby areas in every hamlet;	$\gamma$ , but also had radon and thoron exposures positively correlated with $\gamma$ .	63.2 mGy – colon dose (excess above low-exposure area)	House occupancy information from 5,291 homes; developed aggregate age-sex occupancy factors.
16. Taiwan dwellings (Chen, 2002).	Exposure rates estimated from measurements of representative locations in each room; TLD measurements. Questionnaire used to reconstruct amount of time spent in each room of the contaminated apartment.	$\gamma$ from <sup>60</sup> Co-contaminated rebar used in building construction	Median 6.3 mGy (<1 – 2,363 mGy)	Primary uncertainties due to recall and individual locations in room; no uncertainty analysis included. Large number of exposure rate and TLD measurements taken.

Study <sup>a</sup>	Type of Dosimetry	Types of Exposure	Mean Dose (range) - mGy	Comment on Dosimetry
17. UK pediatric CT (Pearce <i>et al.</i> , 2012)	Doses to brain and RBM. “Obtained typical machine settings for CT in young people from UK-wide surveys undertaken in 1989 and 2003, combined those with a series of hybrid computational human phantoms and Monte Carlo radiation transport techniques to estimate absorbed doses to the RBM and brain for reference males and females for ages 0 – 22 years.” (Pearce <i>et al.</i> , 2012)	X-ray (CT)	2001 and after: For brain, dose 0 – 28 mGy per CT, depending on age and CT anatomical location; For RBM, dose of 0 – 9 mGy per CT depending on sex, age and CT anatomical location. Dose estimates before 2001 ~2–3 times higher because age-specific machine settings rarely used in earlier years.	No individual dosimetry. Uncertainty in reconstructed doses likely high for earlier exposure years. Recent validation work on doses vs. age, sex, size, time has been carried out (Kim <i>et al.</i> , 2012).
18. Australia pediatric CT (Mathews <i>et al.</i> , 2013)	Estimated “effective dose” based on literature information, and subsidiary brain and RBM doses. Effective dose modeled from anatomic scan site, year and age. Most analyses based on number of CT scans rather than dose.	X-ray (CT)	Estimated mean effective dose per CT scan: 4.5 mGy (but this would vary by year and scan site). Mean brain dose (brain CTs), 40 mGy; RBM dose, 4.6 mGy	No individual doses- estimated average doses for red bone marrow and brain and “effective dose”, based on literature information— taking into account the site of CT, age, and year of scan, but no details given. Organ dose estimates used only for subsidiary analysis.
19. Thyroid cancer, pooled analysis (Lubin <i>et al.</i> , 2004; 2017; Veiga <i>et al.</i> , 2016)	Variety of approaches based on modeling of doses from medical radiation exposures, plus atomic-bomb dose estimation.	Mostly X ray, but $\gamma$ for hemangioma radium needle study and $\gamma$ + neutron for atomic-bomb study	Used only study subjects with <200 mGy to thyroid.	For x ray procedures in early days, issues of distance from primary beam, degree of beam collimation, thyroid shielding etc. are primary sources of uncertainty.

<sup>a</sup> Study numbers refer to studies as referenced in Table 4.1.

3 CT = computed tomographic examinations

4 RBM = red bone marrow dose

5 TLD = thermoluminescent dosimeters

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Table 4.3—*Epidemiologic results and statistical features.*

Study <sup>a</sup>	Observed Cancers (excess cancers)	Estimated ERR Gy <sup>-1</sup> (95 % CI) for Solid Cancer (or all non-leukemia), and/or Leukemia	Statistical Methods	Statistical Models Evaluated	Covariate Adjustment Factors
1. A-bomb survivors	Incid: 22,538 (992 excess)	Mort: 0.47 (0.38, 0.56); Leukemia: 3.1(1.8, 4.3)  Incid: 0.64 (0.52, 0.77) –female 0.20 (0.12, 0.28) - male	Poisson regression. Sensitivity analyses for smoking, autopsy-only diagnoses, sex-related cancers	L, LQ, Q, threshold, semiparametric, nonparametric, dose categories <sup>B</sup>	City, sex, age at exposure, attained age, time since exposure, smoking (also, evaluated effect modification by these)
2. 15-Country	4,770	0.58 (90 % CI 0.10, 1.39) – (excluding the Canadian cohort because of it had missing dose data that created a marked bias in the results).	Poisson regression using time-dependent cumulative dose. Sensitivity analyses by sex, cohort, attained age, age at exposure, time since exposure, smoking.	L, polynomials in dose, dose categories	Sex, age, calendar period, facility, duration of employment, socioeconomic status
3. INWORKS	19,064 non-leukemias (209 excess);  531 non-CLL leukemias	Non-leuk: 0.48 (90 % CI 0.20, 0.79)  Non-CLL Leuk: 2.96 (90 % CI 1.17, 5.21)	Poisson regression using time-dependent cumulative dose. Sensitivity analyses for country, lag period, smoking (un)related cancers, neutron or internal radionuclide monitoring status, lower dose ranges; excluding lung, liver, bone	L, LQ, restricted dose ranges, nonparametric, dose categories	Country, sex, attained age, year of birth, socioeconomic status, duration of employment, neutron and internal-exposure monitoring status

Study <sup>a</sup>	Observed Cancers (excess cancers)	Estimated ERR Gy <sup>-1</sup> (95 % CI) for Solid Cancer (or all non-leukemia), and/or Leukemia	Statistical Methods	Statistical Models Evaluated	Covariate Adjustment Factors
4. Mayak	1,825 – solid cancer, excluding lung, liver, bone (97 excess)	Mort: 0.12 (0.03, 0.21) adjusted, or 0.16 (0.07, 0.26) unadjusted, for plutonium exposure; Threshold: 0.2 Gy (<0, 1.3)	Poisson regression using time-dependent cumulative dose. Sensitivity analyses by Pu exposure, attained age, time since exposure	L, LQ, Q, L+cell killing, threshold, dose categories	Sex, attained age, age at exposure, time since exposure, birth cohort, smoking, Pu exposure
5. Cleanup, Russia	Mort: 2,442 (excess 172)  Incid: 4,002 Leuk: 141	Mort: 0.58 (0.002, 1.25)  Incid: 0.47 (0.03, 0.96) Leuk: 0.44 (-1.68, 2.56)	Poisson regression; SMRs and SIRs; No sensitivity analyses	Linear, nonparametric, dose categories	Calendar year period, region, age at exposure, attained age
6. Canadian nuclear workers (Zablotska <i>et al.</i> , 2013b)	Mort: 324 solid cancers (excess n.a.); 12 non-CLL leukemias	-1.20 (<-1.47, 2.39), solid cancer; 9.79 (<-1.49, 107), non-CLL leukemia	Poisson regression; examined effect modification by facility, sex, attained age, time since first exposure	Linear, dose categories	Sex, attained age, calendar period, duration of monitoring, facility, monitoring status, socioeconomic status

Study <sup>a</sup>	Observed Cancers (excess cancers)	Estimated ERR Gy <sup>-1</sup> (95 % CI) for Solid Cancer (or all non-leukemia), and/or Leukemia	Statistical Methods	Statistical Models Evaluated	Covariate Adjustment Factors
7. Japanese nuclear workers (Akiba and Misuno, 2012; Hosoda <i>et al.</i> , 1997; Iwasaki <i>et al.</i> , 2003)	Mort: 2,636 non-leukemias; (Excess n.a.) 80 leukemias	1.26 (-0.27, 3.00) – cancers except leukemia; 0.20 (-1.42, 2.09) – cancers except alcohol-related and leukemia; -1.93 (-6.12, 8.57) – all leukemia	Poisson regression; sensitivity analyses for smoking and alcohol consumption	Linear, dose categories; no LQ model	Attained age, calendar year period, geographic region. Adjusted for smoking and alcohol consumption. Had data on but did not adjust for socioeconomic level, medical radiation exposures and other hazardous occupational exposures.
8. US rad techs	1,922 incident breast cancers (54 excess); 586 breast cancer deaths	0.7 (0.3, 3.9) – breast cancer incid. Effect confined to cohort that had worked before 1950; perhaps because low doses after 1950 limit statistical power.	Poisson regression using time-dependent cumulative dose.	L, LQ	Attained age, birth cohort, duration of employment, no. live births, menopausal status, age menarche, obesity, family breast cancer, alcohol intake, hormone replacement therapy, race, marital status, smoking
9. Rocketdyne	Mort: 651 non-leukemias; (excess n.a.)	-0.2 (-1.8, 1.7) <sup>c</sup> Non-CLL leukemia: 0.6 (<0, 12.3)	Cox regression using time-dependent cumulative dose. No sensitivity analyses.	Loglinear; Used 10y lag.	Years of birth and hire, sex, hourly/salary pay, duration of employment, rocket toxicant exposure

Study <sup>a</sup>	Observed Cancers (excess cancers)	Estimated ERR Gy <sup>-1</sup> (95 % CI) for Solid Cancer (or all non-leukemia), and/or Leukemia	Statistical Methods	Statistical Models Evaluated	Covariate Adjustment Factors
10. Mound	Mort: 26 non-CLL leukemias	Non-CLL leukemia: 0.4 (-3.7, 7.1) <sup>d</sup>	Cox regression using time-dependent cumulative dose. Cox analyses based on the radiation-monitored only, since the unmonitored had a different risk profile.	Loglinear 10y lag, except 2y lag for leukemia.	Year of birth, year of hire, sex, race, education level
11. China x-ray	1,643 non-leukemia (excess n.a.)	0.87 (0.48, 1.45) Male and female risk coefficients similar.	Poisson regression, but based on only 4 dose categories.	Linear, 5y lag; nonparametric, dose categories	Birth year, sex, year 1 <sup>st</sup> employment, age started work, attained age, calendar period
12. Techa River	Mort: 2,303 (50 excess); Incid: 1,933 (61 excess)	Mort: 0.61 (0.04, 1.27) Incid: 0.77 (0.13, 1.5)	Poisson regression using time-dependent cumulative dose. Sensitivity analyses: examined effect modifiers of sex, age entry, attained age, time since exposure; examined risk excluding bone and colon	L, LQ, Q, spline with knot at 0.1 Gy, threshold, dose categories; LQ not significantly ( $p = 0.2$ ) better than L, nor was spline better than L; Q fit as well as L ( $p > 0.5$ )	Gender, ethnicity, entry period, calendar time, attained age, age at entry, time since 1 <sup>st</sup> exposure, smoking. Effect modifiers: risk did not vary (incidence) or increased significantly with older age at exposure – not as expected.

Study <sup>a</sup>	Observed Cancers (excess cancers)	Estimated ERR Gy <sup>-1</sup> (95 % CI) for Solid Cancer (or all non-leukemia), and/or Leukemia	Statistical Methods	Statistical Models Evaluated	Covariate Adjustment Factors
13. Chernobyl childhood exposure	Ukraine: 45 prevalent (excess 34); 65 incident; Belarus: 87	Ukraine: 5.25 (1.70, 27.5) for prevalence, 1.91 (0.43, 6.34) for incidence; Belarus: 2.15 (0.81, 5.47) for dose range 0–5 Gy was linear, but negative exponential at higher doses.	Logistic regression, Poisson regression; Belarus: Binomial odds model (yielded excess odds ratio estimates)	L, LQ, L-exponential, nonparametric, dose categories	Age at screening; sex; place of screening & residence; urban/rural; marital status; personal history of cancer, thyroid diseases in self or relatives, iodine prophylaxis.
14. Kerala HBRA	1,349 (no excess)	–0.13 (–0.58, 0.46) (Dose group with >500 mGy cumulative dose “had no evident risk.”)	Poisson regression; 10 y lag; Exclusion of lung cancer (because of radon/thoron exposure) did not alter risk.	Linear	Sex, attained age, education, occupation, income, bidi smoking, tobacco chewing
15. Yangjiang HBRA	941 (excess n.a.)	0.19 (–1.87, 3.04) – excluding leukemia and liver cancer (liver disease common, and diagnosis as cancer or cirrhosis varied by region)	Poisson regression	Linear (used 10 y lag)	Sex, attained age, calendar year.
16. Taiwan dwellings	106 solid cancers (excess n.a.) 10 y lag	0.4 (90 % CI –0.3, 0.8) <sup>c</sup>	Cox regression; No sensitivity analyses; (CI seems narrower than expected, based on number of cancers.)	Linear (loglinear)	Sex, attained age, birth cohort

Study <sup>a</sup>	Observed Cancers (excess cancers)	Estimated ERR Gy <sup>-1</sup> (95 % CI) for Solid Cancer (or all non-leukemia), and/or Leukemia	Statistical Methods	Statistical Models Evaluated	Covariate Adjustment Factors
17. UK pediatric CT	74 leukemia; 135 brain tumors	Leukemia or myelodysplastic syndrome: 36 (5, 120) [Leukemia only: 19 (-12, 79)] Brain: 23 (10, 49)	Poisson regression with time-dependent cumulative dose. Sensitivity analyses: age at exposure, attained age, lag period, calendar year	L, LQ, L-exponential, nonparametric, dose categories ( $p > 0.4$ for LQ and L-exponential models for both leukemia and brain)	Sex, age at exposure, years since first and last CT scan
18. Australia pediatric CT	Leuk   Brain CT group: 211   283 Control group: numbers not reported, but ~20 times as many as in CT group.	Brain (after brain CT): 21 (14, 29) – for 5-y lag Leukemia (all CT exams): 39 (14, 70) – 1-y lag	Poisson regression with time.	L, nonparametric, time-dependent cumulative dose; main analysis used 1 y lag.	Age at exposure, sex, year of birth, year of exposure, time since exposure, socioeconomic status. No information on Down Syndrome or other markers of cancer susceptibility.
19. Thyroid cancer, pooled analysis	Exposed, 252; unexposed, 142	Thyroid: 11.1 (6.6, 19.7) – for 0–200 mGy range Dose threshold: 0 (95 % CI <0, 44 mGy)	Poisson regression. Sensitivity analyses: sex, no. of dose fractions, age at exposure, random effects model.	L, LQ, threshold, semiparametric, parametric, dose categories	Study, age at exposure, time since exposure, attained age, calendar year, number of radiation treatments, plus other variables specific to certain studies.

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11 CT = computed tomographic examinations  
12 Leuk = leukemia  
13 non-CLL = leukemia excluding chronic lymphocytic leukemia  
14 <sup>a</sup> Study numbers refer to studies as referenced in Table 4.1.

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- 15 <sup>b</sup> L = linear, LQ = linear-quadratic, Q = pure quadratic, Semiparametric = empirical dose-response without  
16 assuming a shape, Nonparametric = Risk estimates and CI for individual dose categories.
- 17 <sup>c</sup> Based on linear extrapolation of HR at 100 mGy of 0.98 (95 % CI 0.82, 1.17). For non-CLL leukemia, HR at  
18 100 mGy = 1.06 (0.50, 2.23).
- 19 <sup>d</sup> Based on HR at 100 mGy = 1.04 (0.63, 1.71).
- 20 <sup>e</sup> Based on HR at 100 mGy = 1.04 (90 % CI 0.97, 1.08)
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Table 4.4—Study strengths and limitations.

Study <sup>a</sup>	Strengths	Limitations
1. A-bomb survivors	<p>Large cohort, both sexes, exposed at all ages, with long F-U and wide dose range.</p> <p>Representative sample of general population.</p> <p>Low-dose data: 30,000 survivors with colon doses between 5 and 100 mGy.</p> <p>Evaluated shape of the dose-response curve, jointly and singly by sex</p> <p>Evaluated risk modifications by sex, age at exposure, attained age, smoking history</p> <p>Examined risk comparability for subsets of tumor types.</p> <p>Found statistically significant dose response over the range 0 – 100 mGy; dose threshold analysis consistent with no threshold.</p>	<p>Only one acute, high dose-rate exposure, not protracted exposures.</p> <p>Study started in October 1950, &gt;5 y after the bombings, so early data missing.</p> <p>Possible “healthy survivor effect”, particularly at high doses.</p> <p>Low proportion of men of military age.</p> <p>Malnourished Japanese population at time of bombing and for several years thereafter.</p> <p>Retrospective dosimetry and some doses uncertain.</p> <p>Incidence data for solid cancer available only beginning 13 y after exposure.</p> <p>Out-migration: could not ascertain tumor incidence outside of Hiroshima and Nagasaki prefectures, but mortality data available for all of Japan.</p> <p>Curvature is attributable to male data in the range of 0.2 to 0.75 Gy, not for 0 to 0.2 Gy; reasons for that unclear.</p>
2. 15-Country	<p>Extensive dosimetry effort.</p> <p>Large cohort of workers, some with long F-U (up to 62 y).</p>	<p>Dosimetry data provided were incorrect for a fraction of the Canadian cohort, which biased the risk estimates.</p> <p>Elimination of those with internal exposures or larger neutron exposures restricted the dose range of workers and reduced the statistical power.</p> <p>Exclusion of some workforces could have introduced bias.</p> <p>Suggestions that analyses confounded by worker socioeconomic status and duration of employment.</p> <p>Some confounding by smoking.</p>

Study <sup>a</sup>	Strengths	Limitations
3. INWORKS	<p>Large study sample with long, high-quality mortality F-U and personal dosimetry.</p> <p>Dosimetry based on an extensive effort. Overall, doses well characterized and uncertainties relatively small.</p> <p>Analyses restricted to those with &lt;200, &lt;150 or &lt;100 mGy were statistically significant.</p> <p>Showed good correspondence with a linear model and could reject curvilinearity.</p> <p>Results apparently not affected by smoking or asbestos exposure, since similar results after lung and pleural cancers removed from endpoint.</p>	<p>Dosimetry for the early time period includes relatively greater uncertainty, due to early technologies, and “missed” photon and neutron doses.</p> <p>Influence of excluded neutron and internal doses, and of “missed” doses, on reported dose response uncertain.</p> <p>Results regarding neutron exposure contrary to expectations; may represent confounding.</p>
4. Mayak	<p>Cohort with wide range of external doses from protracted exposures.</p> <p>Individual external dose measurements and detailed work histories.</p> <p>Long follow-up of high quality for Ozyorsk residents.</p> <p>Risk assessments adjusted for smoking; not adjusted for alcohol intake, but preliminary data show little correlation between alcohol intake and dose.</p> <p>Showed good correspondence with a linear model.</p>	<p>Autopsies more frequent among higher dose individuals, and cause of death obtained from a variety of sources, including 9 % from family members.</p> <p>Incidence data only available for Ozyorsk residents.</p> <p>Only ~38 % of workers with potential Pu exposure had Pu bioassay measurements.</p> <p>Issues of dose inaccuracies in early years due to dosimeter limits regarding photon energies, angular responses, high-energy betas and no/inaccurate neutron measurements.</p> <p>70 % of workers had one or more years with reconstructed doses.</p> <p>Possibility of surveillance bias – that higher dose workers were paid greater attention than others.</p>

Study <sup>a</sup>	Strengths	Limitations
5. Cleanup, Russia	<p>Majority of doses measured by personal dosimeters. 80 % of workers with recorded doses.</p> <p>Large cohort and fairly high cumulative doses.</p>	<p>Accuracy of “official” recorded doses unknown, and &gt;15 % based on dose reconstructions. Reliability sample: ~17 % of the official doses deviated 10-fold from RADRUE-estimated doses.</p> <p>Organ dose not used in the analyses of either solid cancers or leukemia.</p> <p>No information on smoking or alcohol consumption</p> <p>Nonlinear dose-response models not investigated.</p> <p>Potential difficulties of accurately following a large number of people over a wide geographical area, and of surveillance bias – greater attention paid to higher dose cleanup workers.</p>
6. Canadian nuclear workers	<p>Fairly large, well-defined cohort.</p> <p>Good individual dosimetry and elimination of AECL subgroup with missing doses.</p> <p>Excellent follow-up rate and cause of death determination.</p> <p>Had worker socioeconomic information.</p>	<p>Could not find the missing dose information for the AECL subgroup.</p> <p>Had no data on smoking, alcohol consumption or other lifestyle variables.</p> <p>Nonlinear models not investigated.</p>
7. Japanese nuclear workers	<p>Large cohort with good individual dosimetry.</p> <p>Had information on smoking and alcohol consumption for a substantial subset.</p> <p>High rate of follow-up and cause of death determination.</p>	<p>Apparent confounding by alcohol consumption and/or smoking but had analyses to consider these.</p> <p>Short follow-up.</p> <p>Study design weakness had potential survivor bias (follow-up began 30+ y after first exposure for some).</p> <p>Had data but did not adjust for socioeconomic level or medical radiation exposure.</p>

Study <sup>a</sup>	Strengths	Limitations
8. US rad techs	<p>Large nationwide cohort with individual dose estimates, long-term follow-up, detailed information on potential confounders, and high medical confirmation rate.</p> <p>Modeled both shared and unshared dose uncertainties for incorporation into analyses.</p>	<p>Only breast cancer and skin cancer analyzed with dose data to date (plus cataract with less accurate dosimetry).</p> <p>Substantial dose uncertainties for workers before ~1960 because had to rely on literature reports of doses for radiologic technologists.</p> <p>Potential intrinsic confounding between estimated cumulative breast dose and birth year (<math>r = -0.58</math>)</p>
9. Rocketdyne	<p>Doses well-characterized, including internal radionuclide exposures, and doses received at other places of employment.</p> <p>High rate of follow-up (99.4 %) and cause of death ascertainment (98 %).</p> <p>Lengthy follow-up.</p>	<p>Couldn't evaluate shape of dose response because risk estimate was negative for solid cancer.</p> <p>Relatively small study of radiation workers and thus relatively low statistical power.</p> <p>No lifestyle information.</p>
10. Mound	<p>Long F-U – up to 60 y</p> <p>Captured doses before and after Mound employment.</p> <p>High F-U rate and high percent with known cause of death</p> <p>Adjusted for education</p> <p>Had large amount of polonium bioassay data.</p>	<p>Relatively small number of workers.</p> <p>No lifestyle information.</p> <p>Uncertainties of polonium and other radionuclide measurements.</p>
11. China x-ray	<p>Substantial range of doses with long-term exposure and long F-U.</p>	<p>Dose response based on only 4 dose categories.</p> <p>Limited diagnostic accuracy (70 % with histology; others based on radiological exams)</p> <p>Assigned average estimated calendar-year doses to workers.</p> <p>Question about socioeconomic comparability of exposed and unexposed groups.</p>

Study <sup>a</sup>	Strengths	Limitations
12. Techa River	<p>Extensive dose reconstruction Unselected population; Large numbers with long follow-up; Many with <sup>90</sup>Sr measurements to help estimate individual exposures Could examine data for confounding by ethnicity, smoking. Evaluated L, LQ and Q models.</p>	<p>Intrinsic uncertainties in dosimetry; Recent revisions in dosimetry not reflected in published studies No personal gamma measurements Unusual age-at-exposure pattern of risk. Too little statistical power to discriminate between L and Q models. 16 % of cohort migrated away from the catchment area, thus reducing the effective sample size and power of the study, but there is no reason to believe that migration was a confounding factor.</p>
13. Chernobyl childhood exposure	<p>Fairly low thyroid dose uncertainties because of individual thyroid radioactivity measurements. Had detailed uncertainty analysis that took shared/unshared and other errors into account. Ultrasound and palpation screening with a standard protocol provided consistent, blinded assessment, and had cytologic indication for cancer before surgery. Non-participation rates of 26–33 % did not vary significantly by dose, so unlikely to bias the results.</p>	<p>Whereabouts, consumption details, etc. based on questionnaires with possible recall error. Direct thyroid measurements were conducted under difficult conditions within a few weeks after the accident.</p>
14. Kerala HBRA	<p>Had ambient measurements of external exposure for ~94 % of homes. Fairly high cumulative dose at a low dose rate. Used various resources to ascertain cancer. Have data on several potential risk factors, including smoking. Medical exposure infrequent, so little potential of dose-response bias from this source of exposure.</p>	<p>Dosimetric uncertainties because of having to use aggregate house-occupancy factors; had personal dosimetry on only a few individuals. No dose uncertainty analysis. Possible diagnostic bias: 72.4 % of cancers in 0–49 mGy group, and 64.8 % in &gt;200 mGy group, had histologic diagnosis. Cancer incidence rates increased more slowly than expected with age (2.4 power), suggesting under-diagnosis in older individuals. Access to adequate medical care potentially limited and unequal. No account for migration. Potential confounding in comparing largely coastal residents with largely inland residents.</p>

Study <sup>a</sup>	Strengths	Limitations
15. Yangjiang HBRA	<p>Dose groups did not differ regarding diet, drinking water, pesticide residue, aflatoxin in food, medical usage, smoking, alcohol intake.</p> <p>Indoor ambient measurements for ~1/3 of dwellings in each hamlet, and outdoor measurements in hamlet.</p> <p>Age/sex specific occupancy factors estimated from ~5,300 interviews.</p> <p>Stable population.</p>	<p>Dose was inversely related to mortality from external causes, TB and liver cancer – suggests potential bias.</p> <p>May have been geographic differences in quality of cancer ascertainment.</p> <p>Diagnosis weak: 26 % of cancer deaths based on pathological information, 62 % on radiography/ultrasound, remainder clinical impression etc.</p>
16. Taiwan dwellings	<p>Extensive ambient measurements made in dwellings but no personal measurements.</p> <p>Good quality tumor registry to ascertain cancers.</p>	<p>Low dose distribution, contributes to low statistical power and precision.</p> <p>No information on lifestyle or socioeconomic factors.</p> <p>Small sample size, young ages, so relatively few cancers.</p>
17. UK pediatric CT	<p>Large study.</p> <p>Mainly well-designed.</p> <p>Good cancer ascertainment.</p>	<p>Possible missed doses from retakes due to patient movement were not considered.</p> <p>No individual dosimetry.</p> <p>Information not available as to reasons for CTs or other clinical variables – susceptible to biases due to confounding by indication and reverse causation.</p> <p>Likely missed CT exams performed on study subjects at health care facilities not in the study.</p> <p>Myelodysplastic syndromes (MDS) included with leukemias; leukemia effect not significant without the MDS cases.</p>

Study <sup>a</sup>	Strengths	Limitations
18. Australia pediatric CT	Very large study. Good cancer ascertainment.	Information not available as to reasons for CTs or other clinical variables – susceptible to biases due to confounding by indication and reverse causation. Main risk estimates based on lagging only 1 y after exposure; thus risk estimates likely exaggerated by reverse causation. Found implausible results, <i>e.g.</i> , elevated risks for melanoma and Hodgkin lymphoma, but not for breast cancer; elevated risk for brain tumors after CT to sites other than the head. Possible missed doses from retakes due to patient movement not considered. Missed CT exams, including: “nearly all CT scans in state based tertiary hospitals” (where the majority of Australian CT scans are performed), and CTs before 1985 or after age 20. Estimated organ doses used only in subsidiary analyses.
19. Thyroid cancer, pooled analysis	Large number of exposed individuals, thyroid cancer cases and comparable controls. Adjusted for a variety of risk factors. Evaluated low-dose part of dose-response curve. Based on a number of high-quality studies with good follow-up and histologic verification of cancers.	Could not adjust for dose measurement error, although the 2 most influential studies had data corrected for measurement error, and measurement error uncertainty probably fairly small. Could not directly assess the possibility of differential medical surveillance by dose, but analyses suggested it was not a biasing factor.

23 <sup>a</sup> Study numbers refer to studies as referenced in Table 4.1.  
 24 CT = computed tomographic examinations  
 25 F-U = follow-up  
 26 HBRA = high natural background radiation area  
 27

28 it is recognized that all observational studies have limitations, mostly ranging from minor to moderate, that  
29 contribute to the evaluation of the LNT model. Except for the study of Japanese atomic-bomb survivors, the  
30 studies reviewed here have low doses, low dose rates, or both. The individual low-dose studies intrinsically have  
31 limited statistical power and precision in risk estimation. Therefore, a synthesis of study results regarding the  
32 LNT model, with consideration of study quality, will be the most informative epidemiologic evidence that can be  
33 provided for radiation protection purposes.

34

## 35 4.1 Japanese Atomic-Bomb Survivors

36

### Highlights

The Life Span Study (LSS) cohort of atomic-bomb survivors has provided important data because it is a large cohort (~94,000 survivors of all ages) with relatively accurate dosimetry, a wide dose range (0 to 4 Gy, including ~68,000 with doses <100 mGy), over 60 y of high-quality follow-up for mortality and cancer incidence, and nearly 1,000 excess solid cancer cases, besides leukemias. These features provide relatively high statistical power and precision of risk estimates, including a statistically significant dose response for all incident solid cancer over the dose range 0 to 100 mGy and no evidence of a dose threshold. Data are available on sociodemographic, lifestyle and other disease risk factors to assess confounding, and limited data suggest that neither atomic-bomb fallout exposures nor medical radiation exposures confound the results. In the latest analysis, cigarette smoking did not confound the association of radiation with solid cancer incidence.

Unlike most other studies reviewed in this report, the LSS assesses the effects of a single, brief dose and the associated low dose effectiveness factor (LDEF), but not protracted doses so as to examine a dose rate effectiveness factor (DREF). Though a pure quadratic dose-response curve provided a poor fit to the data, the most recent solid cancer mortality and incidence data provide evidence of linear-quadratic upward curvature at lower doses, implying a LDEF > 1 and suggesting a shallower dose-response slope at low doses than at higher doses. However, the fact that curvature was seen only in males, and that it was mostly because of a flat response between about 200 and 750 mGy, makes the curvature difficult to interpret.

37

### 38 4.1.1 *Dosimetry Considerations*

39

40 The cohort of survivors received a wide range of doses essentially instantaneously. Individual  
41 doses are estimated based on reported location, shielding, and other factors at the time of the bombing  
42 using complex radiation transport codes and the estimated height and yield of the two devices (Young  
43 and Kerr, 2005). Radiation doses have been estimated for about 87,000 of the 94,000 atomic-bomb

44 survivors for 15 different organs or tissues. Doses could not be estimated for the remaining 7,000  
45 because of complex shielding situations (Cullings *et al.*, 2006).

46  
47 The dosimetry has been refined periodically over a 50 y period with Cullings *et al.* (2017) representing the  
48 most recent update; this included a review of the original paper records to improve the accuracy of the location and  
49 shielding information that was used in the Dosimetry System 2002 (DS02) calculations (Young and Kerr, 2005).  
50 An electronic geographic information system was used to locate study subjects with greater accuracy on new  
51 orthophotographic maps which corrected for biases present in the original army maps dating from circa 1945, and a  
52 vastly improved method to estimate terrain shielding was applied (Cullings *et al.*, 2006; 2017).

53  
54 Extensive validation efforts of external gamma doses and neutron fluence have included thermoluminescent  
55 dosimetry (TLD) of roof tiles, sulfur activation of power line insulators, and activation of building materials and  
56 copper. Electron spin resonance of tooth enamel and fluorescence in situ hybridization (FISH) have been used to  
57 validate dose estimates for internal and external exposures (Nakamura *et al.*, 2012; Young and Kerr, 2005).

58  
59 Young and Kerr (2005) summarized the DS02 uncertainty range in individual doses as a coefficient of variation  
60 (CV) = 0.24 to 0.43 for Hiroshima and 0.28 to 0.47 for Nagasaki. Based on those data, 35 % was used as an  
61 uncertainty correction factor in estimating the doses for analyses. A revised method to estimate dose-response  
62 relationships that incorporates both individual (“classical”) and grouped (also called “Berkson”) dose measurement  
63 error has been proposed (Pierce *et al.*, 2008) but not yet routinely applied in papers. In test cases, with an individual  
64 CV of ~40 % and a grouped uncertainty CV of ~20 %, this method showed results only slightly different from the  
65 previous method that corrected for only individual measurement error.

66  
67 One issue with the dosimetry is the role of neutron exposures in the radiation effects observed. Neutron  
68 exposure levels were thoroughly studied as part of the DS02 (Young and Kerr, 2005) revision of the atomic-bomb  
69 dosimetry, but the relative biological effectiveness (RBE) of the neutrons is uncertain. In estimating doses, RERF  
70 applies a quality factor of 10 for the neutron component, whereas others have advocated 20 to 30 or a sliding  
71 RBE ranging up to 100 for neutron exposures at low total doses (Rühm and Walsh, 2007; Sasaki *et al.*, 2016).  
72 One recent analysis suggested an important role for neutron dose (Walsh, 2013) in atomic-bomb risk estimates,  
73 though another suggested intrinsic uncertainties because neutron and gamma doses are substantially correlated  
74 and neutrons are only a small fraction of the total dose (Cullings *et al.*, 2014).

75  
76 Other potential sources of uncertainty in radiation doses include possible additional exposure to some  
77 individuals from “rainouts” from the radioactive plumes of the bombs, or by neutron activation of soils. Available

78 exposure measurements suggest that radioactive fallout was not widespread (Okajima, 1987; Okajima *et al.*, 1987),  
79 except significant fallout did occur in the Nishiyama area of Nagasaki where one to two thousand resided, and  
80 those in the Koi-Takasu area of Hiroshima may have experienced a small amount of fallout. An analysis  
81 correlating reported fallout exposures with subsequent mortality and cancer incidence rates did not find significant  
82 associations (Sakata *et al.*, 2014). Significant exposure to radionuclides from neutron activation of soils would  
83 have occurred primarily to any who went quite near the hypocenter within a few days after the bombings (NCRP,  
84 2012), but individual information about “early entrance” to the proximal area is very limited.

85  
86 Although the dosimetry for the atomic-bomb survivors is considered to be fairly accurate, particularly in  
87 comparison to dose estimates for most other environmental dose reconstructions, the accuracy is still limited by the  
88 uncertainty in location and structural shielding for various individuals due to recall error and inability to precisely  
89 characterize shielding configurations. Furthermore, shielding configurations have greater uncertainty for the  
90 Nagasaki factory workers who comprise a large portion of the relatively high dose survivors in that city (Cullings  
91 *et al.*, 2006).

92  
93 Information on diagnostic medical radiation exposures was obtained for Adult Health Study (AHS; a dose-  
94 stratified random sample of LSS subjects in the catchment areas of Hiroshima and Nagasaki cities) participants  
95 from 1964 to 1982 (Yamamoto *et al.*, 1986). Those data did not show a correlation with atomic- bomb doses, so  
96 medical irradiation did not appear to be a confounding variable. There are relatively few socioeconomic or other  
97 barriers to utilization of medical care because it is free and available for essentially all atomic-bomb survivors.

98  
99 More information about dosimetry uncertainties for atomic-bomb survivors is presented in NCRP Report  
100 No. 171 (NCRP, 2012).

#### 101 102 **4.1.2** *Epidemiologic Methods and Uncertainties*

103  
104 The Life Span Study of ~94,000 atomic-bomb survivors includes ~54,000 who were within 2.5 km of the bomb  
105 hypocenters, a sampling of ~40,000 who were between 2.5 and 10 km away, matched on city, age and sex, as well  
106 as ~26,000 similarly matched individuals who were not in either city at the time of the bombing. The cohort was  
107 assembled based primarily on the 1950 Japanese census data which included a question about residence at the time  
108 of the bombing. The death certificate cause-of-death accuracy before and during the 1970s found some  
109 incompleteness of cancer death coding and still more inaccuracy of coding heart disease and certain other causes of  
110 death (Ron *et al.*, 1994), but cause-of-death accuracy has improved in more recent decades. The Hiroshima and  
111 Nagasaki city/prefecture (regional) tumor registries provide high-quality tumor incidence data. A limitation is that

112 such data are available for only the two prefectures, but AHS participation data provide a way to estimate the  
113 prefecture out-migration rates by age, sex and temporal period, so the incidence denominators are adjusted for  
114 population migration. Migration rates were not differential by dose.

115  
116 Information on sociodemographic, lifestyle and other disease risk factors is available for about two-thirds of  
117 the atomic-bomb survivors. Certain analyses have been conducted to account for possible biases associated with  
118 lifestyle activities such as smoking habits (Furukawa *et al.*, 2010; Grant *et al.*, 2017), but smoking habits did not  
119 alter the risk estimates materially (Grant *et al.*, 2017). Sociodemographic variations, such as urban/rural  
120 differences, have been examined to a limited extent. Data on a number of medical and lifestyle risk factors such as  
121 smoking, alcohol consumption, and dietary effects for heart disease, cataracts, and other noncancer effects are  
122 available for AHS subjects and have been used to adjust pertinent results in that subcohort. Broadly speaking, little  
123 confounding by lifestyle variables has been seen, because their distributions tend not to vary across the dose range.  
124 A more detailed summary of epidemiologic uncertainties is available in NCRP Report No. 171 (NCRP, 2012).

125  
126 Background disease rates in Japan have historically differed from those in western populations (*e.g.*, higher  
127 Japanese rates of stomach cancer, liver cancer and stroke; lower rates of breast cancer, colon cancer and ischemic  
128 heart disease), which creates uncertainties about how to extrapolate atomic-bomb survivor risk estimates to western  
129 populations. This has usually been approached as an across-the-dose-range generalization issue (*e.g.*, ERR vs. EAR  
130 extrapolation), and there is no information about how this might affect low-dose risk estimation.

131  
132 **4.1.3** *Statistical Results*

133  
134 The analyses of the LSS data have generally used fine gradations of dose (about 20 categories) and have  
135 incorporated adjustments for city, sex, age at exposure and attained age, plus various other factors in sensitivity  
136 analyses. Various past analyses of the LSS cohort had indicated that the dose response for all solid cancer  
137 incidence or mortality fits a linear model across the full dose range better than a linear-quadratic (upward  
138 curvature) or pure quadratic function. In the latest LSS report of mortality through 2003, a linear-quadratic model  
139 did not improve the fit ( $p = 0.36$ ) for all solid cancer across the full dose range, and a purely quadratic model  
140 provided a significantly poorer fit than the linear model (Ozasa *et al.*, 2012). The estimated lowest dose range with  
141 a significant ERR for all solid cancer was 0 to 0.20 Gy, and a formal dose-threshold analysis indicated no  
142 threshold; *i.e.*, zero dose was the best estimate of the threshold.

143  
144 However, the mortality report (Ozasa *et al.*, 2012) showed that when the data were analyzed over the range of  
145 0 to 2 Gy, there was statistically significant upward curvature, with a ratio of the dose-squared to linear dose

146 coefficients ( $\alpha/\beta$  ratio) of 0.81 (95 % CI 0.08, 8.6) (Ozasa *et al.*, 2012). Nevertheless, since curvature could be a  
147 result of nonlinear deviations in various parts of the dose range, an analysis was needed to clarify the risk and  
148 degree of uncertainty specifically in the low-dose range.

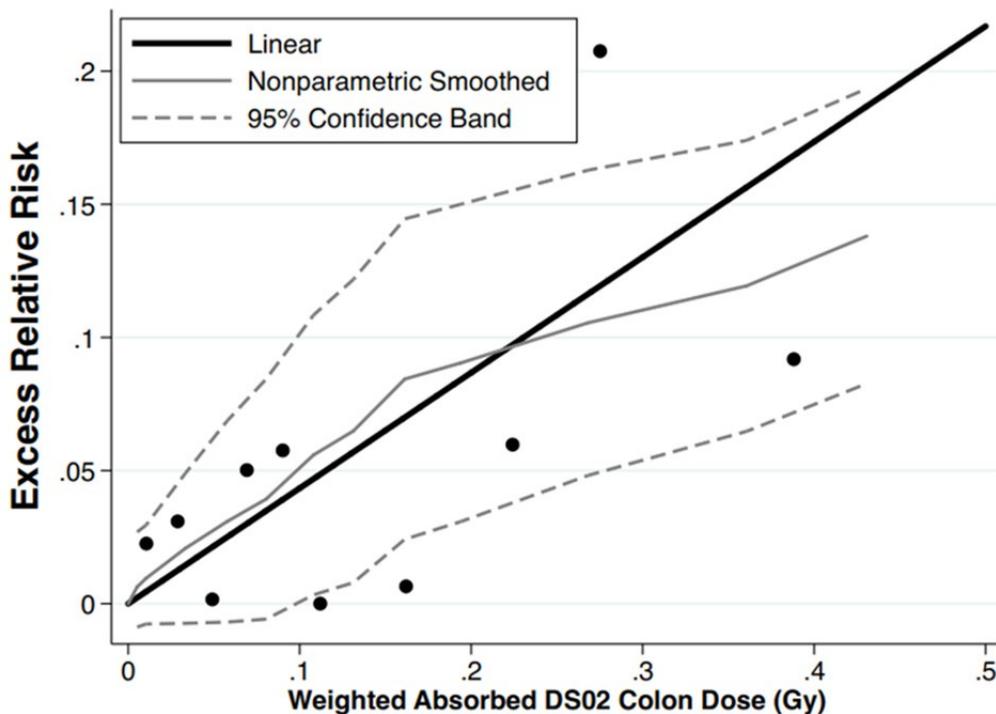
149  
150 An important issue regarding the LNT model is the nonparametric slope of the dose response in the low-dose  
151 range. About 68,000 LSS survivors had estimated colon doses less than 100 mGy. Evaluating the low dose  
152 uncertainties is especially problematic, because conventional parametric LNT model analyses implicitly assume the  
153 absolute width of the confidence intervals gets narrower as the dose goes lower (Figure 3.1), an assumption that is  
154 unlikely to be true. Figure 4.1 shows the confidence band at doses less than 500 mGy based on a nonparametric  
155 method that does not utilize the best-fitting LNT model for the most recent solid cancer mortality data. The 95 %  
156 confidence band is broad and compatible with no excess risk below about 150 mGy but is more compatible with the  
157 LNT model throughout the lower dose range. Another relatively nonparametric approach to fitting a model of the  
158 solid cancer incidence data for 1958 to 1998 (Preston *et al.*, 2007) found the nonparametric fit over the entire dose  
159 range was virtually identical to the best-fitting linear model (Furukawa *et al.*, 2015). But when doses of 200 mGy  
160 and under were examined, Furukawa *et al.* (2015) found that the confidence band in that dose range was  
161 considerably wider than the estimate from the best-fitting linear model. Their semiparametric analysis indicated  
162 clear excess risk above 100 mGy, but below 100 mGy the confidence bounds did not exclude either no risk or a  
163 linear dose response, though the slope was generally positive below 200 mGy.

164  
165 As seen in Figure 4.1 for solid cancer mortality, up to about 200 mGy the risk estimates for the LNT and  
166 nonparametric models were essentially identical, but above that level the nonparametric slope was flatter than the  
167 linear model. Examination of the dose response for the full dose range or the 0 to 2 Gy range suggests that excess  
168 risk was relatively depressed compared to the linear model over the range of roughly 0.2 to 0.7 Gy for unknown  
169 reasons, and this at least partly explains the significant upward curvature over the 0 to 2 Gy range (Ozasa *et al.*,  
170 2012).

171  
172 **New update of cancer incidence:** An 11 y update of solid cancer incidence in the LSS cohort was recently  
173 reported (Grant *et al.*, 2017), representing follow-up through 2009, 64 y after the atomic bombings. After  
174 reductions due to cancer or death before the incidence study began in 1958, or inability to estimate radiation doses,  
175 105,000 were included in the study cohort. As of 2009, 63 % of the cohort was deceased. The analysis utilized the  
176 improved individual dose estimates mentioned above and for the first time included smoking as a potential  
177 confounder or effect modifier. They estimated 992 excess solid cancers were attributable to radiation exposure.

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**Fig. 4.1.** Solid cancer mortality risk during 1950 to 2003 over the weighted absorbed colon dose range of 0 to 0.5 Gy, from the LSS cohort of atomic-bomb survivors (Ozasa *et al.*, 2012). The dark solid line represents the linear fit over the full dose range. The lighter solid line is a nonparametric, lowest- smoothed fit to the 0 to 0.5 Gy data with 95 % CI shown by the dashed lines. (Based on Report 14 mortality data available online at <http://www.rerf.jp>)

206 Some important patterns of risk have continued. The ERR Gy<sup>-1</sup> for total solid cancers is greater for those young  
207 at exposure, decreasing by 21 % per decade of age at exposure. The EAR estimates were likewise greater for those  
208 young at exposure. The ERR Gy<sup>-1</sup> also decreased with attained age, independent of age at exposure, and the  
209 decrease was significantly steeper in males than in females; the sex-averaged decrease in the ERR with attained age  
210 was proportional to age to the power of 1.66. They reported a decreasing ERR estimate with increasing attained age  
211 which occurred because, while the excess rates increase with increasing age, their rate of increase is slightly less  
212 than the increase in baseline rates (Grant *et al.*, 2017).

213  
214 The modeled sex-averaged ERR of 0.50 Gy<sup>-1</sup> (95 % CI 0.42, 0.59) for all solid cancer, with exposure at age 30  
215 and follow-up at age 70, was very similar to the prior report of 0.47 Gy<sup>-1</sup> (Preston *et al.*, 2007). Taking account of  
216 smoking made only a small difference in the radiation risk estimate (ERR Gy<sup>-1</sup> of 0.47, 95 % CI 0.39, 0.55 using a  
217 multiplicative smoking-radiation model). As in the prior report, females had a higher ERR Gy<sup>-1</sup> (0.64, 95 % CI  
218 0.52, 0.77 for females; 0.27, 95 % CI 0.19, 0.3.7 for males) based on linear models. Part, but not all, of the sex  
219 difference was attributable to the higher baseline cancer rates in males. The excess absolute risk (EAR per 10,000  
220 person-years) was 42.9 for males and 54.7 for females at 1 Gy, but the EAR for males at 100 mGy was only about  
221 55 % as large as for females because of the significant quadratic component of the male risk.

222  
223 The lowest dose range that showed a statistically significant dose response using the sex- averaged linear ERR  
224 model was 0 to 100 mGy with an excess relative risk estimate of 0.49 Gy<sup>-1</sup> (95 % CI: 0.026 to 1.01;  $P = 0.049$ ),  
225 virtually identical with the estimate of 0.50 over the full dose range. For the sexes combined, over the full dose  
226 range the linear-quadratic model fit the data better than the pure quadratic model ( $p < 0.001$ ), indicating there is a  
227 positive slope at low doses. Tests for a dose threshold did not indicate a statistically significant threshold.

228  
229 The main difference in risk estimation from the prior cancer incidence report (Preston *et al.*, 2007) was  
230 increased evidence of upward curvature in the dose-response curve ( $p = 0.03$ , sex- averaged). There was relatively  
231 strong evidence of curvature among males ( $\alpha/\beta$  ratio of 1.3,  $p = 0.002$ ) but no evidence of curvature among females  
232 ( $\alpha/\beta$  ratio of 0.08); the male/female difference in curvature was statistically significant ( $p = 0.02$ ). Similarly, when  
233 the excess absolute rate (EAR) model was examined, there was significant curvature in men but not in women.

234  
235 Closer examination of the dose-response curve for males indicated that the upward curvature occurred mainly  
236 because of a flat dose response over the range of about 0.2 to 0.75 Gy, whereas there was an upward slope below  
237 0.2 Gy. The dose-response slope for males over the low-dose range of 0 to 100 mGy, while quite uncertain, was  
238 0.33 Gy<sup>-1</sup> which was nominally higher than the male ERR Gy<sup>-1</sup> estimate over the full dose range of 0.27.

239

240 The interpretation of the gender-related difference in curvilinearity is complex. When sex- specific cancers  
241 were removed, the curvature increased among females, though not significantly, likely due to removing breast  
242 cancer which has linear dose-response characteristics, and also suggesting that different subsets of cancer types  
243 have varying degrees of curvilinearity.

244  
245 A major factor that drove the tendency for curvilinearity in these updated analyses was the newly revised  
246 dosimetry, which featured the incorporation of more extensive terrain shielding factors and improved accuracy of  
247 the dose data (Cullings *et al.*, 2017). Another indication of the role of changed dosimetry in the newfound  
248 curvature was the fact that, when the new dosimetry was applied to the data of the prior tumor incidence report  
249 through 1998, curvature also was found in those data, unlike with the previous dosimetry (Preston *et al.*, 2007).  
250 The upward curvature seen in males does not necessarily argue against LNT; it may rather suggest a LDEF > 1,  
251 *i.e.*, a lower slope at low doses than at high doses.

252  
253 It also is possible that shapes of the dose-response curves may differ for various individual cancer sites (Section  
254 3.4.1). However, the relatively small number of cancers for individual sites means that the statistical power to  
255 detect nonlinearity is limited. In the Preston *et al.* (2007) analysis of the incidence of solid cancers, no evidence for  
256 nonlinearity was found ( $p \geq 0.4$ ) for stomach, colon, liver, breast, bladder or brain/central nervous system cancers.  
257 There was strong evidence of radiation risk, along with weak suggestions of convex curvature for lung cancer  
258 ( $p = 0.2$ ) and thyroid cancer ( $p = 0.1$ ), and a clear indication of upward curvature for non-melanoma skin cancer  
259 ( $p = 0.005$ ). However, a new LSS report of lung cancer incidence, adjusted for smoking, found an ERR Gy<sup>-1</sup> for  
260 lung cancer of 0.81 (95 % CI 0.51, 1.18) but no indication of quadratic curvature ( $p > 0.5$ ) (Cahoon *et al.*, 2017a).  
261 As was noted in a previous NCRP report, “For analyses of various subtypes of cancer or other disease, the numbers  
262 are much smaller than for total solid cancer or broad categories of noncancer disease, so it is difficult to assess the  
263 specificity versus generality of particular shapes of the dose-response functions” (NCRP, 2012).

264

#### 265 4.1.4 *Study Strengths and Weaknesses*

266

267 The LSS cohort of atomic-bomb survivors has provided important data because it is a large cohort with  
268 accurate dosimetry, a wide dose range, all ages at exposure and over 60 y of high-quality follow-up, a relatively  
269 large number of excess cancer cases (992) and cancer deaths (527), and features that enable relatively high  
270 statistical power and precision of risk estimates, including a statistically significant dose response for all incident  
271 solid cancer over the dose range 0 to 100 mGy. Regarding the LNT model, the LSS is limited to assessing the  
272 effects of a single, brief dose (a low-dose effectiveness factor, LDEF) and not protracted doses (dose-rate  
273 effectiveness factor, DREF). Analyses of earlier LSS data had suggested a single linear dose-response function

274 from doses of 2 to 3 Gy down to doses of 200 mGy or below. There are certain limitations to the LSS data,  
275 including that it provides data only on acute exposures; there may be some residual sample selection effects; the  
276 (especially historical) coding of cause-of-death on death certificates has uncertainties; out-migration of study  
277 subjects, which affects the denominators of the tumor incidence data, is only estimated; and the dose-response  
278 patterns may differ by tumor type. However, those methodological limitations tend to be minor, and the study  
279 represents a benchmark for other radiation epidemiology studies.

280

#### 281 **4.1.5** *Implications for the LNT Model and Radiation Protection*

282

283 A pure quadratic model provided a significantly poorer fit to the dose-response data than a linear model for  
284 both solid cancer incidence and mortality, and there was no evidence of a significant dose-response threshold for  
285 either endpoint. An analysis of the most recent mortality data indicated excess risk over the range of 0 to 200 mGy  
286 that was congruent with the LNT slope, and the new tumor incidence data showed a statistically significant dose-  
287 response slope over the range of 0 to 100 mGy. Nevertheless, the most recent solid cancer mortality and incidence  
288 data provide some evidence for upward curvature at lower doses among males but not females. This suggests a  
289 shallower dose-response slope at low doses than at higher ones and would imply a LDEF > 1. However, little or no  
290 curvature is seen in the dose response for females. In summary, the study provides strong indirect support for the  
291 use of a LNT model, with consideration of a DDREF factor, for use in radiologic protection. More direct evidence  
292 will have to come from the studies with low dose rates or many small dose fractions.

293

294

## 4.2 Worker Exposure Studies

295

### Highlights

Radiation worker studies assess risks in worker groups exposed largely to many low doses received at a low dose rate, providing direct evidence of the validity of the LNT model. Further, cumulative doses, derived from dosimeter readings, can be moderate to high, especially for workers in early periods, so that studies can offer reasonable statistical power.

Large studies combining data from workers from more than one nuclear installation have now been conducted in a number of countries, and the largest study is the International Nuclear Workers Study (INWORKS), which has included workers from sites in the United States, United Kingdom, and France. INWORKS found associations between the cumulative dose from external sources of photons to the red bone marrow (RBM) and leukemia (excluding CLL) mortality,  $ERR\ Gy^{-1} = 2.96$  (90 % CI: 1.17, 5.21), and the external dose to the colon and mortality from all solid cancers combined,  $ERR\ Gy^{-1} = 0.47$  (90 % CI: 0.18, 0.79). For solid

cancer there was no evidence of nonlinearity ( $p = 0.44$ ). These risk estimates are compatible with predictions based upon LSS data.

The Russian Mayak workforce also is of particular interest because of the high cumulative doses received (mainly at a low dose rate) by many workers during the early years of operations at this installation. An association has been found between the external dose to the RBM and leukemia (excluding CLL) incidence, ERR Gy<sup>-1</sup> = 3.57 (90 % CI: 1.55, 8.22). For external dose to the colon and mortality from all solid cancers excluding lung, liver and bone (*i.e.*, excluding cancers at the major sites of plutonium deposition), ERR Gy<sup>-1</sup> = 0.12 (95 % CI: 0.03, 0.21) with no indication of nonlinearity ( $p > 0.5$ )

Overall, the nuclear worker studies lend support to the inference that an excess risk of cancer exists following protracted exposure to low doses received at a low dose rate, and the excess risk is compatible with a LNT model. Though there are dosimetry uncertainties, particularly for early periods when occupational doses tended to be highest, the best worker studies provide the strongest support currently for a LNT model based on low doses and dose rates. Several worker studies are reviewed, with greater detail accorded to the more important studies. The studies include the 15-Country study (Section 4.2.1), INWORKS study (Section 4.2.2), Mayak worker study (Section 4.2.3), Chernobyl cleanup worker study (Section 4.2.4), U.S. Radiologic Technologists study (Section 4.2.5), Million Worker study (Section 4.2.6), and Chinese medical x-ray worker study (Section 4.2.7). In addition to the text below, further systematic information is provided for most of these studies in Tables 4.1 to 4.4.

296

#### 297 4.2.1 15-Country Study

298

299 As a context for recent radiation worker studies, the earlier Three-Country and 15-Country Worker Studies are  
300 summarized here. In 1988 the International Agency for Research on Cancer (IARC) agreed to coordinate an  
301 international collaboration aimed at increasing the statistical power of radiation workers studies (Wakeford, 2014).  
302 The first outcome of this work appeared in 1995 and involved seven cohorts of workers from three countries, three  
303 from the United States (including Hanford), three from the United Kingdom (including Sellafield) and one from  
304 Canada (the workers of Atomic Energy of Canada Limited, AECL) (Cardis *et al.*, 1995). The Three- Country Study  
305 found a significantly elevated dose response for the risk of mortality from leukemia (excluding CLL) and the  
306 cumulative dose from external sources of radiation (ERR Sv<sup>-1</sup> = 2.18, 90 % CI: 0.13, 5.7), but the equivalent  
307 estimate for all cancers except leukemia was slightly, and nonsignificantly, negative (ERR Sv<sup>-1</sup> = -0.07, 90 % CI: -  
308 0.39, 0.30). The confidence intervals for both leukemia (except CLL) and other cancers were quite wide and  
309 compatible with a range of possibilities at low doses, including for other cancers either no risk or risks larger than  
310 the estimates from the atomic-bomb study (Cardis *et al.*, 1995).

311

312 The study was later extended to 15 countries, and reported in several publications (Cardis *et al.*, 2005b;  
313 2007; Thierry-Chef *et al.*, 2007; Vrijheid *et al.*, 2007a; 2007b). While the number of countries that contributed  
314 data on workers was considerably expanded in the 2007 reports, updating of the cohorts that contributed to the  
315 1995 analysis was incomplete. In particular, the U.S. cohorts contributed the same underlying data to the 15-  
316 Country Study as they had to the Three-Country Study, but a number of exclusions were made in the 15-  
317 Country Study of data that had been used in the Three-Country Study. Workers with substantial exposures to  
318 internal emitters and/or neutrons were excluded (Vrijheid *et al.*, 2007b), which was important because these  
319 workers tended to have the highest external doses.

320  
321 **4.2.1.1 Dosimetry.** The dosimetry for the Three-Country Study (Cardis *et al.*, 1995; Fix *et al.*, 1997) evolved from  
322 three separate investigations of dose response relationships for worker cohorts in the United States (Gilbert *et al.*,  
323 1993), the United Kingdom (Carpenter *et al.*, 1994) and Canada (Gribbin *et al.*, 1993). The dosimetry for the 15-  
324 Country Study (Cardis *et al.*, 2005b) that evolved from the Three-Country Study incorporated an analysis of  
325 exposure conditions in the workplace (predominant energy and geometry of exposure), dosimetry technology,  
326 calibration practices, and administrative procedures and is described by Thierry-Chef *et al.* (2007). Accounting for  
327 these factors in the dosimetry was important because of major changes in the evolution of radiation protection  
328 methods and policies over the years 1943 through the 1980s. This work by Thierry-Chef *et al.* (2007) was a  
329 significant improvement in the overall dose estimates available for epidemiologic study.

330  
331 The 15-Country investigators developed bias factors (B, systematic error) and uncertainties (K, characterizing  
332 the uncertainty/spread of the values) following principles described by the 1.05 to 1.2 (among lung, red bone  
333 marrow, and colon). Each source of uncertainty applied to the dosimetry record was considered and identified as  
334 shared or unshared, as was the uncertainty resulting from variation in the correct bias factor among workers  
335 (Thierry-Chef *et al.*, 2007). Despite these improvements, there are concerns about doses recorded during early time  
336 periods of the study, especially between 1944 and 1957 when annual recorded doses tended to be higher than in later  
337 years and major changes were occurring in dosimetry measurement technology and administrative practices.  
338 Furthermore, the impact of neutron dose and internal dose on the dose response is not clear, and there is a distinct  
339 possibility that better accounting of these doses could affect the estimates of risk. The methods of accounting for  
340 doses that were below the limits of detection were another source of uncertainty. In summary, although the 15-  
341 Country Study dosimetry effected a significant improvement in the overall dose estimates, questions of  
342 underestimation of dose due to missed dose, neutron dose, and internal dose remain, as acknowledged by the  
343 dosimetry investigators (Thierry-Chef *et al.*, 2015).

344

345 **4.2.1.2 Epidemiologic Methods, Findings and Issues.** The 15-Country Study included radiation workers from 154  
346 different facilities. For inclusion, each facility cohort had to meet certain defined criteria regarding completeness  
347 of the cohort, routine external radiation monitoring, socioeconomic status (*e.g.*, blue/white collar), and adequate  
348 vital status and cause of death information (Vrijheid *et al.*, 2007b). The epidemiologic analyses were adjusted to  
349 take into account the main sources of potential confounding: sex, attained age, calendar period, facility, and in  
350 some analyses, duration of employment and socioeconomic status. Although the trend of risk with cumulative  
351 external dose was positive for mortality from leukemia excluding CLL, somewhat surprisingly, given the findings  
352 of the Three-Country Study, it was not statistically significant ( $ERR\ Sv^{-1} = 1.93$ , 90 % CI < 0, 7.1). However, the  
353 association with external dose for all other cancers was both positive and significant ( $ERR\ Sv^{-1} = 0.97$ , 95 % CI:  
354 0.14, 1.97). Tests for nonlinearity were nonsignificant. The high risk estimate for cancers other than leukemia was  
355 barely compatible with the prediction of standard risk models (*e.g.*, Preston *et al.*, 2007); however, including the  
356 cohorts without socioeconomic-status data or removing the statistical adjustment for duration of employment  
357 decreased the risk estimate by 40 to 70 % (Cardis *et al.*, 2007), indicating potential confounding.

358  
359 The interpretation of the 15-Country Study was not straightforward for several reasons (Boice, 2010; Dauer *et al.*,  
360 *et al.*, 2010; Shigematsu, 2005; UNSCEAR, 2008; Wakeford, 2005; 2009), especially the surprisingly large influence  
361 of the Canadian workers on the risk estimate for all cancers except leukemia. Although the Canadian workers  
362 contributed around 4 % of the deaths, the exclusion of these workers caused a ~40 % reduction in the risk estimate  
363 (Wakeford, 2005) because of their anomalously high risk coefficient (Ashmore *et al.*, 2007; UNSCEAR, 2008).  
364 When cancers of the lung and pleura are removed from the 15-Country analysis of cancers other than leukemia, the  
365 statistical significance disappears ( $ERR\ Sv^{-1} = 0.59$ , 95 % CI: -0.29, 1.70). Among other issues, some of the  
366 concerns included bias, confounding (mixed evidence regarding smoking), the selection of workers to include (or  
367 exclude) from the study, analytical issues and low statistical power (Boice, 2010). Scrutiny of the previous findings,  
368 including the Three-Country Study (Cardis *et al.*, 1995), revealed an apparent upward change in risk estimates for  
369 the group of all cancers except leukemia that coincided with the start of the use of the Canadian National Dose  
370 Registry (NDR) data for analyses (Wakeford, 2009). Ashmore *et al.* (2010) examined the NDR data for the Atomic  
371 Energy of Canada, Limited (AECL) workers and identified a number of possible deficiencies in the AECL worker  
372 data used in the 15-Country Study, particularly those relating to the data before 1971.

373  
374 Zablotska *et al.* (2013b) recently reported the findings of an updated study of Canadian nuclear industry  
375 workers following a detailed check of dosimetry and employment records, which resulted in several changes in  
376 the AECL data in the NDR. It was suggested that findings of high risk for the early AECL workers was probably  
377 due to missing dose information, rather than a real effect of radiation exposure, and they believed that use of the  
378 pre-1965 AECL worker data could not be justified until further investigation is undertaken. When the early

379 AECL workers were excluded, there was a notable reduction in the ERR Sv<sup>-1</sup> for mortality from all solid cancers  
380 in the Canadian cohort. The Zablotska *et al.* (2013b) findings confirmed that both the original Canadian and the  
381 15-country solid cancer risk coefficients appear to be anomalously high. When the Canadian data were removed  
382 from the 15-Country analysis, the dose response for cancers other than leukemia was no longer statistically  
383 significant (ERR Sv<sup>-1</sup> = 0.58, 95 % CI: -0.22, 1.55), though the risk estimate was still as high or higher than for  
384 the matched worker subset of the LSS.

385  
386 **4.2.1.3 Study Strengths and Weaknesses.** The Three-Country and 15-Country Studies were the first major attempts  
387 at examining international pooled data on protracted radiation exposures and cancer risk and were successful in  
388 eliciting extensive international collaboration. The 15-Country Study investigators realized the importance of  
389 improving and harmonizing the dosimetry data and made a first attempt to do so. However, questions remain  
390 regarding underestimates of dose, due to missed dose and unincorporated neutron exposures, as acknowledged by  
391 the dosimetry investigators (Thierry-Chef *et al.*, 2015). An undetected problem with the dosimetry in the Canadian  
392 cohort led to issues of bias in the derived risk estimates. There were also questions about subject selection  
393 procedures: the exclusion from the analyses of certain cohorts (because occupational status information considered  
394 inadequate) or workers who had received substantial internal or neutron exposures (which eliminated many of the  
395 higher dose workers).

396  
397 **4.2.1.4 Implications for the LNT Model and Radiation Protection.** The recent re-evaluations illustrate the care that  
398 must be exercised in collating worker data and the problems that can arise, especially when using data that may  
399 have been collected for purposes other than epidemiology (Wakeford, 2014). The pooled 15-Country Study does  
400 not at this time provide sufficiently reliable information upon which to evaluate the LNT model for radiation  
401 protection purposes. However, the 15-Country Study has been superseded by the INWORKS study.

402  
403 **4.2.2 INWORKS Study**

404  
405 The International Nuclear Workers Study (INWORKS) is the latest international collaboration to be  
406 coordinated by the International Agency for Research on Cancer (IARC) for examining the health of workers in  
407 more than one country who were exposed occupationally to ionizing radiation (Daniels *et al.*, 2017; Gillies *et al.*,  
408 2017; Hamra *et al.*, 2016; Laurier *et al.*, 2017; Leuraud *et al.*, 2015; Richardson *et al.*, 2015; Thierry-Chef *et al.*,  
409 2015). The following summary considers the study's contribution to the evaluation of the LNT model. INWORKS  
410 combines three large cohorts of radiation workers from five nuclear facilities in the United States, nuclear industry  
411 workers in France, and workers included in the U.K. National Registry for Radiation Workers (NRRW), the results

412 of which have already been individually reported (Metz-Flamant *et al.*, 2013; Muirhead *et al.*, 2009; Schubauer-  
413 Berigan *et al.*, 2015). These represent updated data from those available for the 15-Country Study.

414  
415 The INWORKS findings published to date address mortality from leukemia and other hematopoietic and  
416 lymphatic cancers or all cancers except leukemia (or all solid cancers in the French study) and from  
417 cardiovascular diseases in relation to doses received in the workplace from external sources of photons (mainly  
418 gamma rays). The INWORKS study includes a total of 308,297 workers monitored for exposure to external  
419 sources of ionizing radiation, with 8.2 million person-years of follow-up. Follow-up ran from 1944 (or the start-  
420 up of the facility) to 2005 in the United States, from 1955 to 2001 in the United Kingdom, and from 1968 to 2004  
421 in France.

422  
423 **4.2.2.1 Dosimetry.** The INWORKS study included dosimetry for 20 different nuclear sites/organizations in three  
424 countries (Thierry-Chef *et al.*, 2015). Dosimetry was based on individual personal dosimeter readings dating from  
425 1944 to 2005, which provided an extraordinarily large dataset. Although the availability of individual dose  
426 information is a clear strength of INWORKS, the dosimetry is complex, with mixed radiation fields including  
427 varying gamma/x-ray energies, neutron exposure and internal exposure. The long time period presents a technical  
428 challenge due to changes in dosimetry technology, and the evolution of administrative exposure policies and  
429 recording practices. Furthermore, the dosimetry is difficult to verify from the open literature because it is builds  
430 upon dosimetric work performed for a series of independent epidemiologic studies that reach back almost four  
431 decades, and the original dosimetry is mostly inaccessible.

432  
433 The methods used to extract doses from the original records estimating systematic and random uncertainty and  
434 bias were described by Fix *et al.* (1997) and Cardis *et al.* (1995) for the Three-Country Study and by Thierry-Chef  
435 *et al.* (2007) and Cardis *et al.* (2005b; 2007) for the 15-Country Study. However, none of the published articles  
436 supporting INWORKS dosimetry provide sufficient details to completely understand all aspects of the  
437 methodology used, which requires a review of supporting literature back to its origins. Even this approach does not  
438 give a clear picture of the dosimetry history, but it allows us to make some observations about strengths and  
439 weaknesses of the methodology. For our review, we found it helpful to consider the individual facilities, when  
440 possible, which was critical to evaluate consistency in dosimetry among different sites and countries.

441  
442 The U.S. cohort of INWORKS consisted of 101,428 workers from five sites, each site having different  
443 administrative policies and methods for recording dose and exposure conditions. The Hanford site accounted for  
444 the largest number of workers (34,278) and highest cumulative collective gamma dose (~880 person-Sv) in the  
445 combined U.S. cohort (Schubauer-Berigan *et al.*, 2015). The characteristics and dosimetry of Hanford workers

446 were documented extensively in studies led by Gilbert (Gilbert and Marks, 1979; Gilbert, 1990, 1991; Gilbert *et*  
447 *al.*, 1989; 1993). Later, the importance of accounting for random and systematic errors and bias in the dosimetry  
448 was recognized (NA/NRC, 1989) which resulted in revised dose estimates (Fix *et al.*, 1994; 1997; Gilbert, 1998;  
449 2009; Gilbert and Fix, 1995; Gilbert *et al.*, 1996). These revisions significantly improved the quality of the  
450 dosimetry and accounted for a number of potential factors affecting individual exposures that were not specifically  
451 documented in the records. For the ORNL cohort (Frome *et al.*, 1997; Schubauer-Berigan *et al.*, 2015; Wing *et al.*,  
452 1991; 1993) a methodology was developed by Watkins *et al.* (1997) to adjust doses for uncertainties and bias  
453 similar to that at Hanford but applying site-specific parameters. Dosimetry for workers from Portsmouth Naval  
454 Shipyard (Daniels *et al.*, 2004), Idaho National Laboratory (Schubauer-Berigan *et al.*, 2005), and the Savannah  
455 River Plant (Richardson *et al.*, 2007) has also been reported but not as thoroughly investigated for uncertainties and  
456 bias as Hanford and Oak Ridge. Those sites included radiation workers beginning in 1952, 1952, and 1949,  
457 respectively.

458  
459 The U.K. cohort included some 175,000 workers from 11 organizations. It was based on the National  
460 Registry for Radiation Workers (NRRW) and is documented in a series of reports (Kendall *et al.*, 1992; Muirhead  
461 *et al.*, 1999; 2009). The largest number of workers (40,284) was employed by British Nuclear Fuels Ltd. (BNFL)  
462 and accumulated a collective dose of 2,159 person-Sv. Sellafield is the most important of the five BNFL sites,  
463 contributing approximately 1,700 person-Sv. The French study (Metz-Flamant *et al.*, 2013) included 59,021  
464 workers from four organizations with a cumulative collective dose of about 1,327 person-Sv. Among these cohort  
465 groups that form the INWORKS study, the dosimetry at Hanford and Sellafield are especially important to  
466 scrutinize because of their relative importance in contribution to cumulative dose and number of workers during  
467 early time periods compared to other facilities.

468  
469 As mentioned above, considerable improvements in dosimetry were made over time in both the U.S. and U.K.  
470 cohorts in addressing uncertainty and bias in recorded doses. Another adjustment to improve dose estimates was in  
471 the U.S. Oak Ridge and the U.K. cohort to account for missed dose. Missed dose in this context refers primarily to  
472 large numbers of workers whose dose was recorded as zero because readings were less than the level of detection.  
473 For example, dosimeters were exchanged weekly (or even daily) in many facilities until the late 1950s and, as in  
474 the case of Hanford workers, as many as 75 % of the workers were noted to have been assigned zero dose (s) in  
475 one or more years (Gilbert, 1990). In the early years the level of detection for film badge dosimeters was high at  
476 most facilities, ranging between 0.3 to 0.5 mGy, which with weekly badge readings could theoretically amount to  
477 even as much as 25 mGy of missed dose in a year. Recognition of the potential importance of missed dose among  
478 early nuclear workers in epidemiologic studies and methods for estimating missed dose have been reported (Inskip  
479 *et al.*, 1987; Kneale *et al.*, 1991; Maienschein and Peelle, 1992; Smith and Inskip, 1985; Strom, 1986; Strom *et al.*,

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480 1996; Tankersley *et al.*, 1996). Although the amount of total dose missed for many individuals may be small, if  
481 missed dose is not addressed appropriately the cumulative missed dose in the early time period could be significant  
482 for some individuals. The extent to which missed dose was accounted for in INWORKS, if at all, is not clear. In  
483 one of the INWORKS cohorts, US Oak Ridge workers, a detailed analysis led to the conclusion that the effect of  
484 missing doses was to introduce an upward bias in the dose-response risk estimates (Frome *et al.*, 1997), but  
485 comparable analyses are not available for other individual cohorts in the study.

486  
487 Considerable effort has been invested in the assessment of radiation exposures in the workplace by the 15-  
488 Country and INWORKS dosimetry teams (Thierry-Chef *et al.*, 2007; 2015). Various types of dosimeters were  
489 evaluated over the periods of exposure, and panels of experts were convened to characterize workplace conditions,  
490 monitoring routines, photon energies and exposure geometries over the periods of the study. A database of  
491 correction factors was developed to address geometries of exposure, energies of photons, and various other sources  
492 of bias and uncertainty (Thierry-Chef *et al.*, 2007). Administrative practices such as failure to document measured  
493 doses, missing (or not worn) badges, and the treatment of doses below the limit of detection were also evaluated,  
494 but were considered to be of minor importance by the INWORKS investigators. Additionally, some workers had  
495 recorded doses from neutrons (13 % of workers) or from internally deposited radioactive materials (“internal  
496 emitters”; 17 %), but these doses were not included in the dose estimates used in INWORKS. The dose records  
497 were used to derive absorbed doses to the RBM for the analyses of hematopoietic and lymphatic malignancies, and  
498 doses to the colon for analyses of all cancer other than leukemia (Thierry-Chef *et al.*, 2015). The average estimated  
499 cumulative dose to the colon among all workers was 17.4 mGy, or among the 257,166 exposed workers was 20.9  
500 mGy (90<sup>th</sup> percentile 53.4 mGy, maximum 1331.7 mGy) and the collective dose was 5,370 person-Gy. The mean  
501 RBM dose for all workers was 16 mGy (10<sup>th</sup> percentile 0.0, 90<sup>th</sup> percentile 40.8 mGy) (Leuraud *et al.*, 2015).

502  
503 The estimated uncertainties in the INWORKS dosimetry were generally quite low (K values for individual  
504  $H_p(10)$  doses ranging from 1.2 to 1.7, corresponding to geometric standard deviations (GSDs) of 1.1 to 1.3 or  
505 CV from ~0.2 to ~0.3) compared to other large scale dosimetry studies, especially environmental studies. These  
506 low values are to be expected because of the widespread availability of personal dosimetry data. An analysis of  
507 the various sources of shared and unshared uncertainty (Thierry-Chef *et al.*, 2007) concluded that most of the  
508 uncertainty was Berkson error and thus would probably introduce little attenuation into the risk estimates.  
509 Although no specific uncertainty was assigned to the bias factors used to correct film badge readings, the  
510 uncertainty in bias (B) is included in the estimates of uncertainty (K) for individual doses based on badge  
511 readings, as is explained by the National Research Council (NA/NRC, 1989). However, the authors’ assessment  
512 of uncertainties sometimes did not take into account potential “missed doses” below the limits of detection, nor  
513 the neutron or internal doses.

514

515 Addressing neutron exposure and internal exposure is complex for any analysis of dosimetry, especially when  
516 these exposures occurred during the early years when measurement technology was far less advanced than today.  
517 One complication is differences in recording practices between different facilities and countries at the time. Another  
518 challenge is the effectiveness of neutron dosimeters at various energies. None of the dosimeters was able to measure  
519 the full energy range, and some workers with the potential for receiving doses from neutrons may not have been  
520 monitored for such exposure. This deficiency is noteworthy because the highest exposures likely occurred during  
521 these early years. These technical deficiencies could result in the potential for considerable underestimation of dose  
522 from neutrons and internal exposure. Although this underestimation is acknowledged by Thierry-Chef *et al.* (2015)  
523 and considerable effort was expended trying to resolve this question of dose from neutrons and internal dose, there  
524 remains a question of the underestimation of dose from these sources during the early years of the study. If  
525 excluded neutron doses and internal doses are positively correlated with included photon doses, the dose response  
526 reported in INWORKS for photon dose will be an overestimate because of the excluded doses.

527

528 To summarize, while the availability of dose information on an individual basis is a clear strength of this  
529 INWORKS, the dosimetry is complex and presents a technical challenge to researchers because of the long time  
530 period covered, changes in dosimetry technology, evolution of administrative exposure and recording policies,  
531 mixed radiation fields including varying gamma/x-ray energies, internal exposure, neutron exposure, and missed  
532 dose. Some important details of the dosimetry are very difficult to verify from the open literature because they  
533 involve several layers of dosimetry efforts that reach back almost four decades and because of the inaccessibility  
534 of the original dosimetry. Even after our review to date, it is still not clear whether the individual dosimetry  
535 records included adjustments by previous investigators and how bias and uncertainty were addressed among the  
536 various cohort groups in the study. The technical challenge of the dosimetry and the positive findings of the  
537 epidemiology stress the importance of making the dosimetry as state-of-the-art, complete, thorough and  
538 transparent as possible, since minor variations in the cumulative dose could impact the outcome of the study.

539

540 It is extremely important to pay particular attention to the doses and their uncertainties for the early periods of  
541 exposure (1940s and 1950s) when doses tended to be highest, since those with higher cumulative doses tend to  
542 drive the analytic results. But this is the period when the least information from the historical records is available,  
543 so uncertainties potentially would be the greatest. It is unclear how adequately the investigators surmounted this  
544 challenge.

545

546 **4.2.2.2 Epidemiologic Methods, Findings and Issues.** Characteristics of the INWORKS consortium (nuclear  
547 workers in France, United Kingdom, and United States, 1944 to 2005) are given in Table 4.5, and further details are

548 Table 4.5—*Characteristics of cohorts included in the INWORKS consortium (nuclear workers in*  
 549 *France, United Kingdom, and United States, 1944 to 2005).*  
 550

Characteristics	
Calendar years of follow-up	1944 – 2005
Workers (no.)	308,297
Exposed workers (no.) <sup>a</sup>	257,166
Person years (millions)	8.2
All causes of death	
Collective dose (person Gy)	5370.3
Average individual cumulative dose (mGy) <sup>b</sup>	20.9
No. workers with cumulative dose >100 mGy (percentage of all workers)	18,384 (6.0 %)
All causes	66,632
All cancer	19,748
All cancer other than leukemia	19,064
Solid cancer	17,957
Solid cancer other than lung cancer	12,155
Leukemia, excluding chronic lymphocytic type	531
Ischemic heart disease	17,463
Cerebrovascular disease	4,444

<sup>a</sup>Workers with cumulative dose greater than zero.

<sup>b</sup>Average estimated cumulative dose to the colon, among exposed workers

551  
 552

553 given in the original publications (Gillies *et al.*, 2017; Hamra *et al.*, 2016; Leuraud *et al.*, 2015; Richardson *et al.*,  
554 2015; Thierry-Chef *et al.*, 2015). The table shows that 257,000 out of the 308,000 workers in the study monitored  
555 for exposure to radiation had at least one recorded dose greater than zero. The epidemiologic methods of cohort  
556 composition, occupational record compilation, follow-up and mortality documentation are described in the original  
557 cohort publications or the pooled analysis reports for the individual countries (Metz-Flamant *et al.*, 2013; Muirhead  
558 *et al.*, 2009; Schubauer-Berigan *et al.*, 2015). The length of follow-up was up to 61 y, with a mean of about 27 y.

559

560 The analyses of the data were conducted using conventional Poisson regression methods, with adjustment  
561 for attained age, country, sex, year of birth, socioeconomic status, duration of employment or radiation work,  
562 and neutron monitoring status. Nonlinear dose-response models were compared to the linear model, and a  
563 variety of sensitivity analyses were conducted, including analyses of the low-dose part of the dose-response  
564 curve.

565

566 The principal results thus far from INWORKS are for all cancers excluding leukemia (or for French workers,  
567 only solid cancers) and for all leukemias excluding CLL, which are summarized below. Table 4.6 compares the  
568 risk estimates reported in the original studies that constitute this study with the results of the pooled INWORKS  
569 analysis.

570

571 One superficially puzzling feature of the INWORKS results shown in Table 4.6 is that the ERR Gy<sup>-1</sup> estimates  
572 seem to be higher than might be predicted from the estimates reported by the separate studies of individual  
573 countries for the ERR Sv<sup>-1</sup>. Though some of the differences in risk estimates may be caused by different selection  
574 factors for inclusion in the studies (INWORKS, but not the original studies, used a minimum of 1 y of work at a  
575 nuclear facility), the primary explanation for the differences in risk estimates is likely to be the difference in the  
576 doses used in INWORKS and in the original studies. The original studies used the recorded *H<sub>p</sub>* (10) doses based on  
577 personal dosimeters, with no adjustment to derive organ/tissue-specific doses. INWORKS derived absorbed doses  
578 to the colon for the analysis of all cancers other than leukemia, and absorbed doses to the red bone marrow for the  
579 analysis of leukemia (Thierry-Chef *et al.*, 2015). Because the INWORKS study used estimated doses for deeper  
580 tissues (colon or RBM) instead of the *H<sub>p</sub>* (10) estimates of the original studies, the doses would be lower and  
581 consequently the risk estimates per unit dose would be higher. Thierry-Chef *et al.* (2015) have tabulated summary  
582 data for the organ/tissue-specific doses used in INWORKS; overall the mean *H<sub>p</sub>* (10) dose was 25.2 mGy while the  
583 mean estimated colon dose was 17.4 mGy and that to the RBM was 14.9 mGy. For all cancer except leukemia (or  
584 solid cancers in the French study), the ERR estimate was 0.47 (90 % CI 0.18, 0.79) using estimated colon doses  
585 and 0.33 (90 % CI 0.12, 0.56) using the *H<sub>p</sub>* (10) doses (Richardson *et al.*, 2015).

Table 4.6—ERR/Sv (ERR/Gy in INWORKS) estimates as presented in the original studies of French (Metz-Flamant et al., 2013), United Kingdom (Muirhead et al., 2009), and United States (Schubauer-Berigan et al., 2015) workers, respectively; and in INWORKS (Richardson et al., 2015).

	France [Hp(10) dose]	United Kingdom [Hp(10) dose]	United States [Hp(10) dose]	INWORKS [Hp(10) dose]	INWORKS (colon dose)
All cancers	N/A <sup>a</sup>	0.28	0.14	0.35	0.48
excluding leukemia		(0.02, 0.56) <sup>b</sup>	(-0.17, 0.48)	(0.14, 0.57)	(0.20, 0.79)
Solid cancers	0.34	N/A	N/A	0.33	0.47
	(-0.56, 1.38)			(0.12, 0.56)	(0.18, 0.79)
All leukemias	3.96	1.71	1.7		2.96
excluding CLL	(<0, 16.82)	(0.06, 4.29)	(-0.22, 4.7)		(1.17, 5.21)

<sup>a</sup> N/A not available

<sup>b</sup> 90 % CI, except 95 % CI for the U.S. study.

1 A comparison was made in INWORKS (Richardson *et al.* (2015), supplement) of risk estimates for all  
2 cancers other than leukemia (or for France, solid cancers) using colon doses with those using doses from original  
3 records. The ERR Gy<sup>-1</sup> and ERR Sv<sup>-1</sup> estimates are shown in Table 4.6, where the first INWORKS column uses  
4 doses from original dose records and the second INWORKS column uses colon doses.

5  
6 For the U.K. and U.S. workers (the cohorts that drive the INWORKS findings) the ERR Sv<sup>-1</sup> result based on  
7 Hp (10) doses from INWORKS (ERR Gy<sup>-1</sup> of 0.33) is still notably above the two ERR Sv<sup>-1</sup> values from the  
8 original studies, namely, 0.275 (90 % CI 0.02, 0.56) and 0.14 (90 % CI -0.17, 0.48), respectively. The reasons  
9 for this are unclear, though the inclusion of more short- term workers and of doses from neutrons and (some)  
10 beta-particle doses in the original studies but not INWORKS may affect the ERR Sv<sup>-1</sup>, depending on the  
11 distribution of such doses relative to the included photon doses.

12  
13 Table 4.7 shows several reported variations in INWORKS risk estimates according to which covariables were  
14 included in the model. The alternate risk estimates are 20 % to nearly 60 % lower than the main reported ones,  
15 depending on the covariables included. Particularly notable is the inclusion/exclusion of workers flagged for  
16 recorded neutron exposure in the last row of the table. The overall INWORKS ERR Gy<sup>-1</sup> estimate of 0.48 for all  
17 cancer except leukemia includes an adjustment based upon their neutron monitoring flag (neutron exposure: no  
18 indication, >0 to <10 mGy, ≥ 10 mGy), but if no such adjustment is made (*i.e.*, all workers are included, but no  
19 account is taken of neutron monitoring status) then the ERR Gy<sup>-1</sup> estimate reduces to 0.20 (90 % CI: -0.03, 0.45).  
20 The reasons for this large shift in ERR estimates from 0.48 to 0.20 Gy<sup>-1</sup> are unclear, but may relate to some  
21 confounding, especially since one might expect missing neutron doses in the risk modeling to cause  
22 overestimation, rather than underestimation, of risk if there was a positive correlation among workers between the  
23 neutron doses and external gamma doses, as might be expected. The potential for confounding is highlighted in the  
24 updated BNFL data for Sellafield (Gillies and Haylock, 2014).

25  
26 Linear and categorical risk estimates are shown for the INWORKS study in Figure 4.2. The linear ERR  
27 coefficient in the INWORKS study was 0.47 (90 % CI 0.18, 0.79) Gy<sup>-1</sup> for all cancer except leukemia. An added  
28 quadratic term was not statistically significant ( $p = 0.44$ ). The INWORKS analysis of solid cancers other than lung,  
29 liver and bone gave an ERR Gy<sup>-1</sup> of 0.51 (90 % CI: 0.15, 0.91), very similar to the risk estimate for all solid  
30 cancers. The exclusion of cancers of the lung, liver and bone excludes those sites of cancer most affected by doses  
31 from the internal deposition of plutonium, suggesting that internal plutonium doses did not have a marked effect  
32 upon the ERR Gy<sup>-1</sup> estimates. A parallel finding was that an analysis adjusted for a flag for any known internal  
33 emitter exposure also showed only negligible changes in the ERR Gy<sup>-1</sup>. The authors indirectly examined the  
34

35  
36 Table 4.7—*INWORKS Study: Comparison of risk estimates showing the effects of adjusting for*  
37 *different choices of covariables.*

Cancer Grouping	Primary Estimate <sup>a</sup> ERR (90 % CI) Gy <sup>-1</sup>	Partially Adjusted Estimate ERR (90 % CI) Gy <sup>-1</sup>
Solid cancer	0.47 (0.18, 0.79)	0.37 (0.14, 0.62) <sup>b</sup>
Solid cancer other than lung	0.46 (0.11, 0.85)	0.35 (0.07, 0.65) <sup>b</sup>
Cancer except leukemia	0.48 (0.20, 0.79)	0.20 (−0.03, 0.45) <sup>c</sup>

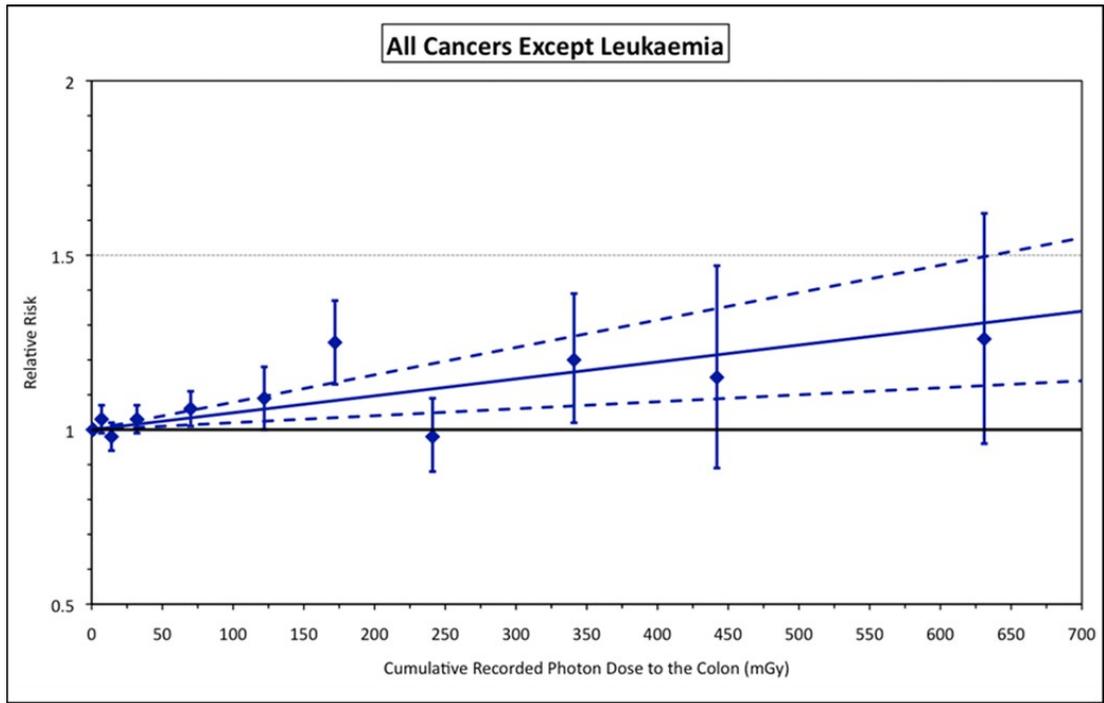
a The primary estimates reported by Richardson *et al.*(2015) were adjusted for country, age, sex, birth cohort, socioeconomic status, duration employed, and neutron monitoring status.

b These estimates were adjusted for country, age, sex and birth cohort, but not for socioeconomic status, duration employed, or neutron monitoring status.

c This estimate was adjusted for all the covariables listed in footnote A except for neutron monitoring status.

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**Fig 4.2.** Relative risk of all cancer mortality other than leukemia by cumulative colon dose, lagged 10 y. Points and vertical lines are estimates and 90 % CI for dose categories; dashed lines are = 90 % CI for the linear model (adapted from Richardson *et al.*, 2015)

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49 possible confounding effects of smoking and asbestos exposure by examining the risk when excluding lung cancer,  
50 or both lung and pleural cancer, and found that the risk estimates did not change materially.

51  
52 Richardson *et al.* (2015) examined the risk estimates for all cancers except leukemia (or all solid cancers for  
53 the French cohorts) for restricted ranges of worker dose (Table 4.8). Notably, even over the range of 0 to 100 mGy  
54 the risk was statistically significant (using their criterion of  $p < 0.05$  on a one-tailed test). Thus there appears to be  
55 risk consonant with LNT in the lower dose ranges for all cancer except leukemia. However, the full picture of  
56 solid cancer risks will not be apparent until risk estimates from INWORKS for individual sites of cancer based on  
57 appropriate organ/tissue-specific doses become available.

58  
59 Regarding the INWORKS leukemia findings, Leuraud *et al.* (2015) reported that for leukemia (excluding the  
60 chronic lymphocytic type; non-CLL) the linear ERR Gy<sup>-1</sup> was 2.96 (90 % CI 1.17, 5.21). This compares to the  
61 atomic-bomb LSS estimated ERR Gy<sup>-1</sup> of 2.60 at 60 y of age after exposure at age 30, for non-CLL leukemia  
62 incidence (Hsu *et al.*, 2013). Leuraud *et al.* (2015) indicated that linear-quadratic or pure quadratic models did not  
63 substantially improve the model fit for non-CLL leukemia, although the dose response had too little precision to  
64 discriminate between the model fit of a pure linear and a pure quadratic model. They also showed that when the  
65 analyses were restricted to data for <300 mGy or <100 mGy, the risk estimates for non-CLL leukemia were  
66 essentially the same as for the entire dose range.

67  
68 **4.2.2.3 Study Strengths and Weaknesses.** Owing to the size and nature of the workforces included in INWORKS,  
69 this study is the most statistically powerful of the radiation worker studies conducted to date; the major  
70 installations included in the study were largely established in the early years of the nuclear industry, when  
71 occupational doses were relatively high, and a substantial number of these workers employed in these early years,  
72 many with cumulative doses >100 mGy, have now died and thereby have contributed to the information and  
73 statistical power of the study. Attention was given to improving and harmonizing the dosimetric information. The  
74 high-quality follow-up of the worker cohorts was for up to 61 y, with good ascertainment of death due to  
75 malignancies. The statistics were strong, with examination of linear and linear-quadratic models and a number of  
76 sensitivity analyses.

77  
78 A principal strength of INWORKS is that cumulative individual photon dose estimates are provided for over  
79 300,000 workers using personal dosimeter data. This dosimetry is based on several decades of epidemiologic  
80 investigations and continued improvements in the dose assignments and associated uncertainties. The dosimetry is  
81 thorough, and technical methods applied seem to be advanced. Photon dosimetry was based on an extensive effort  
82 to harmonize a large database of dose measurements. However, the dosimetry was limited by uncertainties about

83 Table 4.8—*INWORKS* study, excess relative risk estimates for mortality from all cancer except leukemia  
84 over restricted colon dose ranges.

---

Dose Range (mGy)	ERR Gy <sup>-1</sup> (90 % CI)
Full range	0.47 (0.18, 0.79)
0 – 200	1.04 (0.55, 1.56)
0 – 150	0.69 (0.10, 1.30)
0 – 100	0.81 (0.01, 1.64)

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87 occupational doses received in the early years of the nuclear industry in general, which includes the photon doses  
88 that were the subject of the INWORKS analyses published to date.

89  
90 A special concern pertains to neutron doses which were measured poorly in the early years due to technical  
91 limitations, and there will have been neutron exposures that were not recorded. The limited neutron information  
92 was used only semi-quantitatively in the analysis, and there was a puzzling result associated with the analysis of a  
93 semi-quantitative surrogate of neutron exposure. Further, in addition to neutron doses, doses from intakes of  
94 radioactive materials were also excluded from the analyses conducted in INWORKS. The impact of these dose  
95 exclusions and uncertainties upon the published findings of INWORKS will depend upon the degree of correlation  
96 between the included and excluded doses. If more actual doses were missed/underestimated in the early years,  
97 which is possible given that dose uncertainties were the greatest for the years when the photon doses also tended to  
98 be highest, this would reduce the ERR Gy<sup>-1</sup> estimates that have been reported from INWORKS. The study also  
99 does not have any information on medical radiation exposures or on possible confounding lifestyle factors, but it is  
100 doubtful that these would change the radiation risk estimates much.

101  
102 **4.2.2.4 Implications for the LNT Model and Radiation Protection.** Because of its size and the availability of  
103 personal dosimetry information, INWORKS (Richardson *et al.*, 2015; Thierry-Chef *et al.*, 2015) has the potential to  
104 contribute significantly to our understanding of radiation effects at low doses and low dose rate. There are  
105 uncertainties about the magnitude of risk in the INWORKS study associated with dose uncertainties and some  
106 variant results in the sensitivity analyses. Nevertheless, all the risk estimates are positive and commensurate with  
107 risks from the LSS cohort of atomic-bomb survivors, even for restricted low dose ranges. There is no indication of  
108 upward curvature or a dose threshold in the dose-response curve. Further, the ERR Gy<sup>-1</sup> estimates from INWORKS  
109 are based on doses received at a low dose rate as many small doses over a number of years. Thus, this large study  
110 provides one of the strongest pieces of epidemiologic evidence that the LNT quantitative model is useful for  
111 radiation protection.

112  
113 **4.2.3 Mayak Worker Study**

114  
115 The Mayak Production Association (hereafter called “Mayak”) was the first and largest weapons-grade  
116 plutonium (<sup>239</sup>Pu) production plant in the U.S.S.R. It began operations in 1948 and includes radiochemical and  
117 plutonium production plants and nuclear reactor and auxiliary facilities. A large number of early workers were  
118 exposed to high levels of external gamma rays or internal alpha particles, or both. The studies have nearly 60 y of  
119 follow-up and individual dose estimates. Several recent reports have characterized radiation risks from low-LET  
120 exposures in the Mayak worker cohort, covering the incidence (Hunter *et al.*, 2013) or mortality (Sokolnikov

121 *et al.*, 2015; 2016) of solid cancers other than liver, lung and bone (to minimize the impact of plutonium  
122 exposure on risk estimates for external exposure), and hematopoietic cancer incidence (Kuznetsova *et al.*, 2016).  
123 The full cohort consisted of about 25,000 Mayak workers (mortality study, or 22,000 in the incidence studies)  
124 who were first employed during 1948 to 1982.

125  
126 **4.2.3.1 Dosimetry.** The 2008 Mayak Worker Dosimetry (MWDS-2008) provided annual external dose estimates for  
127 various organs for each cohort member. About 80 % of the external doses were based on personal dosimeter data  
128 and 20 % on reconstructed doses (Sokolnikov *et al.*, 2015). Although only crude estimates of uncertainties were  
129 made, the GSD was reported as only 1.26 (Vasilenko *et al.*, 2007a; 2007b). The raw film badge data were corrected  
130 for changes in badge types over time, photon energies and angular responses. For pre-1955 badges, fairly large and  
131 uncertain corrections were required to separate the high-energy beta response from the gamma response (Vasilenko  
132 *et al.*, 2007b). Doses from medical fluoroscopic and x-ray examinations were also estimated but have not been used  
133 in most cohort analyses; however, they are not highly correlated with worker cumulative doses, so have little  
134 potential to confound the data.

135  
136 About 15 % of the workers had indications of some exposure to neutrons. Neutron doses were very uncertain  
137 and mostly based on estimated ratios of neutron to gamma for a particular workplace (NCRP, 2012). Prior to  
138 1960, neutron doses could have been substantial, particularly in the reactor complex, but were largely unrecorded.

139  
140 Systematic urine monitoring for plutonium ( $^{239}\text{Pu}$ ) was not performed until the early 1970s. Internal doses from  
141 plutonium exposures were calculated using biokinetic models and urine and autopsy analyses. As a result,  
142 plutonium dose estimates are available for only 38 % of the 16,995 workers in the radiochemical and plutonium  
143 plants, the principal facilities with plutonium exposure. On average there were 7.51 (SD = 7.47) urinalysis  
144 measurements per worker among those who were measured (Azizova *et al.*, 2015). The estimation of time-  
145 dependent plutonium body-burdens and organ doses was based on elaborate models that were subject to substantial  
146 uncertainties and biases. Agreement between autopsy and urinalysis measurements of internal doses was poor,  
147 suggesting possible errors in the biokinetic models. Only a fraction of the cohort was monitored for plutonium via  
148 urinalysis, and monitoring selection may have been biased. For the 10,513 workers who were assessed as having  
149 nontrivial exposure to plutonium, but were unmonitored, six surrogate plutonium exposure categories were used,  
150 based on the worker's employment history (Shilnikova *et al.*, 2003). Tritium and polonium exposures were not well  
151 characterized.

152  
153 Only crude uncertainty estimates were provided for individual external exposures in the form of uncertainty  
154 bounds and no individual uncertainty estimates were made for internal exposures. Shared and unshared uncertainty

155 was not evaluated and dose uncertainty was not considered in any of the reported epidemiologic investigations.  
156 Considerable shared uncertainty is likely for internal doses. The overall quality of the external exposure data base  
157 is not well established because of limited access to the raw data.

158  
159 The mean cumulative external gamma  $H_p$  (10) dose was 0.51 Gy (0.54 Gy for men and 0.44 Gy for women  
160 (Hunter *et al.*, 2013). Over 3,800 workers had a cumulative dose of at least 1 Gy and 95 workers had >4 Gy).

161  
162 In summary, the overall quality of the MWDS-2008 data base remains to be assessed adequately. Other sources  
163 of dose uncertainty from medical fluoroscopic and x-ray examinations and from stack releases to those dwelling in  
164 Ozyorsk were not fully addressed (Koshurnikova *et al.*, 2012).

165  
166 For internal doses, the MWDS is being revised and more accurate internal doses are estimated using improved  
167 biokinetic models (Vostrotin *et al.*, 2016). MWDS-2013 will also account for doses from Mayak stack releases,  
168 particularly thyroid doses to residents of Ozyorsk. The MWDS-2013 dose system investigates uncertainty in great  
169 detail and uses multiple realizations to separate shared and unshared uncertainty. The new dose system  
170 appropriately models both shared and unshared errors, making it possible to explicitly account for the impact of  
171 shared uncertainties on risk estimates for external dose.

172  
173 A strength of the study is that most workers likely to have been exposed externally wore personal dosimeters  
174 (film badges or later TLDs). Overall, the weaknesses in the MWDS-2008 dosimetry, and in particular, the  
175 plutonium dosimetry, may have limited the accuracy of the risk estimates in the studies published to date more  
176 than the evaluation of LNT. But the newer MWDS-2013 as well as additional improvements in a projected  
177 MWDS-2017 should allow a better evaluation of LNT.

178  
179 **4.2.3.2 Epidemiologic Methods, Findings and Issues.** Follow-up was until 2004 for cancer incidence or 2008 for  
180 mortality (mean of ~37 y of follow-up). About 23 % were lost to mortality follow-up because of migration from  
181 the closed city of Ozyorsk after 2003 or other reasons. Hunter *et al.* (2013) investigated the incidence of other solid  
182 cancers (excluding lung, liver and bone because of plutonium exposure, and nonmelanoma skin cancers) for 1948  
183 to 2004 among 22,366 Mayak workers of both sexes, first employed at the site during 1948 to 1982. They  
184 identified 1,447 other solid cancer cases among workers who lived in Ozyorsk at the time of diagnosis, so data  
185 completeness was restricted by inability to identify cancer diagnoses after workers migrated from Ozyorsk, which  
186 neighbors Mayak, so a worker's follow-up was censored at that time. About 41 % had migrated from Ozyorsk by  
187 the end of 2004. About 91 % of known cases had morphological diagnosis (91.2 %), with autopsies for about 30 %  
188 of cancer deaths. In the first decades of Mayak operation, about 60 % of cancer deaths had postmortem or medico-

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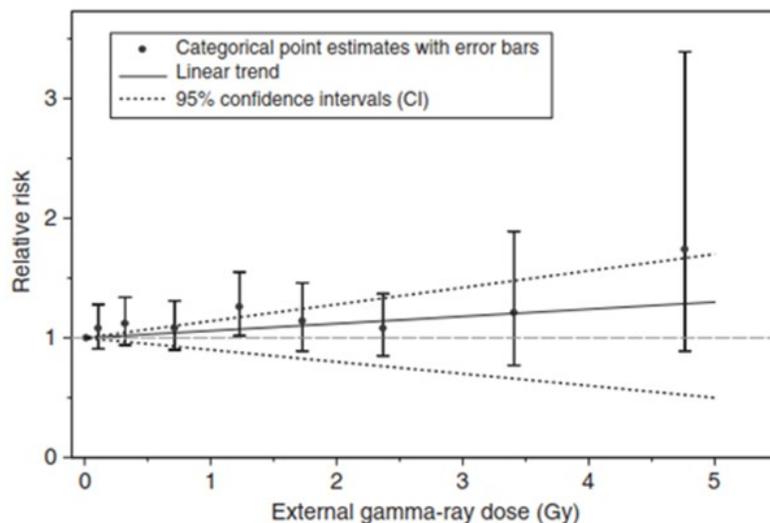
189 legal reports, but currently about 15 % are autopsied, with autopsies more likely to be performed on workers who  
190 had higher levels of plutonium exposure (Hunter *et al.*, 2013). It is unclear whether cancers discovered only at  
191 autopsy are included, and, if so, the degree to which that introduces a bias. Smoking information was available for  
192 89 % of cohort members and alcohol consumption for 78 %. Poisson regression analyses were conducted of “other  
193 solid cancers” (excluding lung, liver, bone and nonmelanoma cancers). External doses were obtained from  
194 individual film badges and expressed as Hp (10), while internal doses from plutonium were the estimated liver  
195 doses. A total of 1,447 incident cases of other solid cancers was registered among workers who were diagnosed  
196 during 1948 to 2004. Potential confounder covariates considered for the analyses included sex, particular Mayak  
197 plant, calendar year, attained age, age at first exposure, smoking, alcohol consumption and internal exposures. The  
198 final model for analyses of external dose needed adjustment only for attained age, sex and smoking status, and  
199 included estimated plutonium dose (measured dose or dose surrogate) as other additive sources of risk.

200  
201 There was no statistically significant association between the incidence of other solid cancers and internal liver  
202 dose for plutonium monitored workers ( $\text{ERR Gy}^{-1} = 0.10$ ; 95 % CI:  $-0.02, 0.26$ ) or using plutonium exposure  
203 surrogate categories for unmonitored plutonium workers ( $p > 0.5$ ) (Hunter *et al.*, 2013). A borderline statistically  
204 significant dose response for other solid cancers and cumulative external dose (0 y dose lag) was found ( $\text{ERR Gy}^{-1}$   
205  $= 0.07$ ; 95 % CI:  $0.01, 0.15$ ), and no substantial change was found when 5, 10, 15, or 20 y dose lags were used  
206 (Figure 4.3). When adjusting for monitored internal dose to the liver from plutonium, the external dose  $\text{ERR Gy}^{-1}$   
207 became  $0.06$  (95 % CI:  $-0.01, 0.14$ ;  $p = 0.12$ ), or  $0.07$  (95 % CI  $-0.005, 0.15$ ) when the surrogate exposure  
208 categories for unmonitored plutonium exposure also were adjusted for. There was no evidence of nonlinearity in  
209 the dose response for external exposure. Adding a quadratic term to the model did not improve the fit ( $p > 0.5$ )  
210 overall or when external dose was restricted to  $<3 \text{ Gy}$  ( $p > 0.5$ ).

211  
212 Sokolnikov *et al.* (2015) examined mortality from “other solid cancers”, excluding cancers of the lung, liver  
213 and bone, among 25,757 Mayak workers who were first employed during 1948 to 1982. A total of 1,825 deaths  
214 from these other solid cancers was recorded during 1948 to 2008; for those workers who had emigrated from  
215 Ozyorsk to other parts of the Russian Federation, follow-up stopped at the end of 2003 because of difficulties of  
216 determining vital status in the rest of Russia from 2004 onwards. The external dose used was the estimate of the  
217 dose to the colon. The internal dose from deposited plutonium was the dose to the liver. The dosimetry was based  
218 on the MWDS-2008, but whereas Hunter *et al.* (2013) used MWDS-2008 external doses expressed as Hp (10),  
219 Sokolnikov *et al.* (2015) used colon doses, which will be less than Hp (10) doses because of tissue shielding and  
220 may thereby yield a larger risk estimate.

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**Fig. 4.3.** Relative risks of the incidence of solid cancer (except lung, liver and bone) in relation to external exposure [ $H_p(10)$ ] categories and the linear trend (and 95 % CI), having adjusted for internal exposure (based on 0 y lag) (Hunter *et al*, 2013).

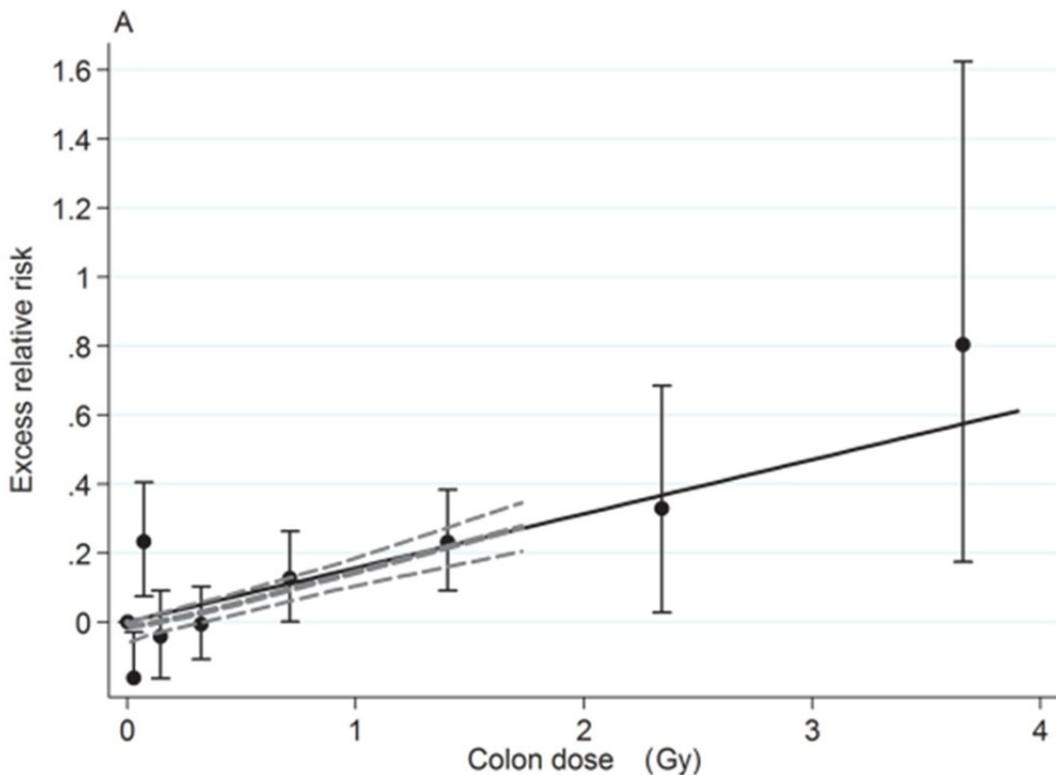
241  
242 The dose response for mortality from other solid cancers and cumulative external colon dose with a 5 y dose lag  
243 was statistically significant ( $ERR\ Gy^{-1} = 0.16$ ; 95 % CI: 0.07, 0.26;  $p < 0.001$ ) (Figure 4.4) (Sokolnikov *et al.*,  
244 2015). When the external dose response was adjusted for internal dose to the liver from plutonium and plutonium  
245 exposure surrogate categories, the  $ERR\ Gy^{-1}$  for external irradiation became 0.12 (95 % CI: 0.03, 0.21;  $p = 0.01$ ).  
246 Given that the age-adjusted rates of cancer mortality were 24 % higher among plutonium-exposed workers, the  
247 investigators examined subsets of workers to evaluate the effects of external doses, absent plutonium exposure  
248 (Sokolnikov *et al.*, 2016). For plutonium workers the  $ERR\ Gy^{-1}$  was 0.15 (95 % CI 0.06, 0.25), while for other  
249 workers it was 0.19 (95 % CI 0.02, 0.39). The risk estimates did not show statistical heterogeneity, so they  
250 concluded that concomitant plutonium exposure did not confound the external radiation estimates (although  
251 adjusting for it did reduce the risk estimate by about 25 %).

252  
253 There was no statistical indication of nonlinearity for external radiation, although most of the low-dose points  
254 suggested there might be less effect per unit dose in that range. Adding a quadratic term did not improve the fit of  
255 the model ( $p > 0.5$ ). The estimate of a dose threshold was compatible with no threshold; the maximum likelihood  
256 estimate of a threshold was 0.2 Gy (95 %  $< 0, 1.3$ ).

257  
258 Kuznetsova *et al.* (2016) have reported on radiation dose to the RBM and the incidence of leukemia,  
259 lymphoma and multiple myeloma during 1948 to 2004. Based on 31 cases of leukemia, excluding CLL, the  
260 linear  $ERR\ Gy^{-1}$  estimate was 3.57 (90 % CI 1.55, 8.22) for cumulative external radiation dose to the RBM,  
261 adjusted for the internal RBM dose from plutonium. The LQ model fit marginally better than the linear model  
262 ( $p = 0.11$ ), and the pure linear and pure quadratic models fit about equally well. The risk estimate for internal  
263 plutonium was not significant, nor were risk estimates for external irradiation and lymphoma or multiple  
264 myeloma.

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**Fig. 4.4.** External exposure dose response for solid cancer mortality, other than lung, liver and bone. The solid line is the fitted linear dose response, the points are ERR estimates in dose categories, with 95 % CI. The thick dashed line is a nonparametric smooth fit to the categorical estimates while the thin dashed lines indicate plus or minus one standard error from the smoothed curve. The models used in this analysis were not adjusted for plutonium exposure (Sokolnikov *et al.*, 2015).

293  
294 **4.2.3.3 Study Strengths and Weaknesses.** This fairly large cohort, with up to 60 y of follow-up, has a considerably  
295 wider dose distribution than any other worker study that reported a dose-response analysis. However, the estimated  
296 doses to the early workers have large uncertainties. A newly revised dosimetry, that has not yet been applied to the  
297 epidemiology, may change the risk estimates by some unknown degree. For the majority of individuals, who have  
298 remained in Ozyorsk, the follow-up for both cancer mortality and incidence is excellent. However, there is concern  
299 over possible bias in cancer detection due to dose-dependent rates of autopsy, the high levels of concomitant  
300 plutonium exposure for a subset, and the fact that about 60 % of those with potential plutonium exposure had no  
301 plutonium measurements. The investigators found evidence for a linear model of solid cancer (excluding lung,  
302 liver and bone), but at this time there is no obvious explanation as to why their risk estimates are lower than those  
303 in most other large worker studies, although dose misclassification ('measurement error') may be one factor. The  
304 discrepancy between the risk estimates of the INWORKS study and the Mayak worker study is explored further in  
305 Shore *et al.* (2017).

306  
307 **4.2.3.4 Implications for the LNT Model and Radiation Protection.** The studies of solid cancer risk supported a  
308 linear dose-response function with no clear evidence for a dose threshold. At face value, the solid cancer risk  
309 estimate in this study, by comparison to the atomic-bomb LSS risk estimate, does suggest that a DREFDREF >1  
310 would be appropriate, *i.e.*, a lesser dose-response slope for protracted exposure than for the LSS acute exposure.  
311 However, the uncertainties in dosimetry, concomitant plutonium exposures, and concerns about possible out-  
312 migration biases suggest caution in the interpretation of the dose response.

313  
314 **4.2.4 Japanese Worker Study**

315  
316 **4.2.4.1 Dosimetry.** Dosimetry for the Japanese nuclear worker studies was based on individual doses from  
317 personal dosimetry as recorded in the Radiation Dose Registration Center for Workers (RADREC) with records  
318 beginning in 1957 (Akiba and Minzuno, 2012; Hosoda *et al.*, 1997; Iwasaki *et al.*, 2003). As with other worker  
319 studies, dose measurement instrumentation, methods, units, facility types, and administrative policies changed  
320 over time and these changes were accounted for in the RADREC system. A radiation dosimetry committee was  
321 formed to evaluate the quality, consistency and procedures employed and concluded that the doses recorded were  
322 of sufficiently high quality and documentation that they could be used to support an epidemiologic study.

323  
324 Internal exposure to workers was very limited and not considered to be a confounder in the dosimetry.  
325 Neutron exposure was also taken into account but the number of workers affected was very small. Although  
326 doses below the limit of detection (~0.1 mSv) were assigned values of zero dose (Iwasaki *et al.*, 2003), it is

327 unlikely this resulted in any significant bias in reported cumulative doses. For dose records that were missing, the  
328 respective facilities reconstructed the dose based on the worker's job assignment and exposures received by  
329 workers in a similar position. The mean cumulative individual dose was approximately 12 mSv. The distribution  
330 of doses was heavily skewed towards lower doses with ~ 75 % of the workers having a cumulative dose less than  
331 10 mSv and only about 2.6 % receiving doses greater than 100 mSv or more (Akiba and Mizuno 2012). No  
332 uncertainties were estimated in the dosimetry nor was an evaluation of the importance of shared/unshared  
333 uncertainty. Investigators did a thorough job of considering inconsistencies that ordinarily increase uncertainty in  
334 other studies of workers over long time periods. In addition, the fact that exposures to the cohort began in 1957  
335 when methods and procedures for estimating dose were greatly improved over earlier time periods suggest that  
336 uncertainties would generally be smaller by comparison. Given that the exposures were almost exclusively  
337 external gamma radiation and that doses were based almost entirely on measured personal dosimetry with good  
338 quality assurance, the dosimetry was thus of relatively high quality.

339  
340 **4.2.4.2 Epidemiologic Methods, Findings and Issues.** Akiba and Mizuno (2012) have documented the mortality  
341 among 200,583 Japanese male nuclear workers from 1991 through 2002. The cohort was based on those  
342 included in a nationwide registry of nuclear workers as of 1989. Cancer mortality data for 1986 to 1990 had been  
343 obtained for an earlier publication (Iwasaki *et al.*, 2003) but was not included in the Akiba and Mizuno (2012)  
344 data. Mortality follow-up was through the Japanese *koseki* system which is very complete. Follow-up duration  
345 was only 6.8 y on average and did not begin until over 30 y after first exposure for some workers. During 1997  
346 to 1999 questionnaires were obtained from 48,000 workers pertaining to lifestyle factors, medical radiation  
347 exposures and occupational history (Murata *et al.*, 2002), with a subsequent survey in 2003 and 2004 of about  
348 45,000 workers. Murata *et al.* (2002) found that radiation dose was positively associated with both smoking and  
349 alcohol consumption, but those in higher dose groups received less medical radiation exposure than those in  
350 lower dose groups. These differences may reflect variations by socioeconomic variables (*e.g.*, job status) that are  
351 correlated with dose.

352  
353 The Poisson analysis of all cancer except leukemia, which stratified on attained age, calendar year period and  
354 geographic region, yielded an ERR Sv<sup>-1</sup> of 1.26 (95 % CI -0.27, 3.00, *n* = 2636). However, when alcohol-related  
355 cancers (upper digestive tract and liver) were excluded, the ERR Sv<sup>-1</sup> was 0.20 (95 % CI -1.42, 2.09, *n* = 1946).  
356 The difference between the risk estimates of these two analyses suggests there was confounding by levels of  
357 alcohol consumption. For all leukemia the ERR Sv<sup>-1</sup> was -1.93 (95 % CI -6.12, 8.57, *n* = 80). No analyses were  
358 reported to examine nonlinearity, a dose threshold or risk by time since exposure.

359

360 **4.2.4.3 Study Strengths and Weaknesses.** The study consists of a large cohort of workers who have had high-  
361 quality individual exposure monitoring and high rates of follow-up and death ascertainment. For a substantial  
362 fraction of workers, information was available on lifestyle factors, other occupational hazards and medical  
363 radiation exposures. Since there was some indication of confounding by alcohol consumption, they presented a  
364 risk estimate deleting alcohol related tumor sites. However, the cohort follow-up was short and began years after  
365 the inception of radiation exposures for some workers which introduces a potential for survival bias.

366  
367 **4.2.4.4 Implications for the LNT Model and Radiation Protection.** Although this is a large cohort with individual  
368 radiation exposure documentation and high-quality mortality follow-up, a short follow-up period, potential  
369 survivor bias, and apparent confounding by lifestyle variables reduce the study's informativeness regarding the  
370 LNT model and radiation protection.

371

#### 372 **4.2.5 Chernobyl Cleanup Worker Study**

373

374 Follow-up of the Russian cohort of Chernobyl emergency cleanup workers (1986 to 1987) was most recently  
375 updated from 1992 to 2009 by Kashcheev *et al.* (2015) for both cancer incidence and mortality. The study includes  
376 ~67,000 cleanup workers. Information on health status, both cancer incidence and mortality, was available for 1992  
377 to 2009. The mean age of the cohort was 34 y at the time of exposure.

378

379 **4.2.5.1 Dosimetry.** A variety of methods were used to estimate dosimetry for Chernobyl emergency cleanup  
380 workers, including individual dosimeters, group dosimeters, or dose-rate measurements at the work place.  
381 Individual dosimeter readings provided the most direct method and were used for 85 % of the doses. Group  
382 dosimeters and dose rate measurements in the work place were used to reconstruct doses for the remaining 15 %  
383 and were associated with more uncertainty (Pitkevitch *et al.*, 1997). The mean dose of the approximately 67,000  
384 workers was estimated as 132 mGy (median, 102 mGy), ranging up to 1240 mGy; 20,992 workers received 50 to  
385 100 mGy; 572 had >300 mGy. Estimated uncertainties in doses ranged from factors of 0.5 to 3, depending on the  
386 dosimetric method used, the working time, and the work location (Kashcheev *et al.*, 2015).

387

388 The reported doses are based on "official" film badge data. The quality of the official doses is suspect  
389 (Chumak *et al.*, 2008; Kryuchkov *et al.*, 2009). An attempt to validate the official doses found some large  
390 discrepancies, so there may be substantial uncertainty in the official doses that were used in the epidemiologic  
391 study. These concerns led to the development of RADRUE (Kryuchkov *et al.*, 2009) which estimated doses that on  
392 average are about 40 % lower than the official doses. The RADRUE technique is a reconstruction of external dose  
393 based on calculations of the product of the exposure rate and irradiation time, with shielding taken into account.

394 They reported GSD estimates measurement error of 1.1 to 5.8 (mean of 1.9) for individual dose estimates.  
395 Considering questions regarding the official film badge data, the conversion of badge reading to organ dose in  
396 directional highly variable radiation fields, and uncertainty due to recall, the uncertainties in dosimetry may be  
397 underestimated.

398  
399 **4.2.5.2 Epidemiologic Methods, Findings and Issues.** Follow-up was undertaken through the Russian National  
400 Medical and Dosimetric Registry (RNMDR) which was established in relation to the Chernobyl accident. The  
401 RNMDR obtained cancer mortality and incidence data from regional hospitals and clinics in Russia. About 7.2 %  
402 of the cohort was lost to follow-up. There were 2,442 solid cancer deaths and 4,002 incident solid cancers. The  
403 analyses were adjusted for calendar year, region, age at exposure and attained age. The linear dose response for  
404 solid cancer yielded ERRs of  $0.58 \text{ Gy}^{-1}$  (95 % CI 0.002, 1.25) for mortality and  $0.47 \text{ Gy}^{-1}$  (95 % CI 0.03, 0.96) for  
405 incidence (Kashcheev *et al.*, 2015). Only a linear model was fitted, but from the data presented it appears that a  
406 linear-quadratic model also should have been considered.

407  
408 Leukemia incidence also was examined in a nested case-control study within the Ukrainian cohort of 110,000  
409 Chernobyl cleanup workers (Zablotska *et al.*, 2013a). The individual cleanup worker doses were estimated using  
410 the RADRUE algorithm for cases diagnosed in 1986 to 2006 and for approximately five controls per case matched  
411 on place of residence and year of birth. The mean dose for cases was 132.3 mGy (range 0 to 3220), while that for  
412 controls was 81.8 mGy (range 0 to 2600). For 52 non-CLL leukemia cases and their controls the ERR  $\text{Gy}^{-1}$  was  
413 2.21 (95 % CI 0.05, 7.6), while for 65 CLL cases it was 2.58 (95 % CI 0.02, 8.4). For all leukemia, tests for a  
414 quadratic, exponential or power deviation from linearity were all nonsignificant ( $p = 0.93, 0.92$  and  $0.27$ ,  
415 respectively). In summary, within the limits of the dose uncertainties and the modest number of cases, the data tend  
416 to support the LNT model.

417  
418 **4.2.5.3 Study Strengths and Weaknesses.** Study strengths included a fairly large sample size and higher dose range  
419 than most other studies with protracted exposures; multiple sources of medical information to document cancer  
420 mortality and incidence, and recorded doses. Kashcheev *et al.* (2015) investigated whether the imputed doses may  
421 have created a bias and found that incidence rates were similar for those with measured and imputed doses.  
422 Limitations of the Kashcheev *et al.* (2015) study include the uncertainties in dosimetry; lack of data on smoking,  
423 alcohol use or sociodemographics; possible bias due to increased medical surveillance of the most heavily exposed;  
424 and failure to evaluate nonlinear dose-response models for solid cancer. The study of leukemia by Zablotska *et al.*  
425 (2013a), however, had data on smoking, alcohol use, sociodemographics, medical radiation exposures and  
426 evaluated several shapes for the dose-response association.

427

428 **4.2.5.4 Implications for the LNT Model and Radiation Protection.** Although the Kashcheev *et al.* results suggest  
429 that exposures at low dose rates confer radiation risk, the uncertainties regarding dosimetry and possible variations  
430 in health surveillance, plus failure to consider alternate dose-response models, lessen the contribution of the study  
431 for the LNT model and radiation protection. The Zablotska *et al.* leukemia study provides better support for the  
432 LNT model, though it is limited by having dosimetry uncertainties and only a moderate number of cases.

433

434 **4.2.6 U.S. Radiologic Technologists Study**

435

436 The x-ray technologist study was begun in 1983, with baseline questionnaires completed by ~90,000 radiologic  
437 technologists who had been registered by the American Registry of Radiologic Technologists during 1926 to  
438 1980 (Boice *et al.*, 1992). It inquired about lifetime work history as a radiologic technologist, with details about  
439 procedures performed, practices, and protective measures plus sociodemographic and lifestyle information, and  
440 employment, reproductive and medical histories. Subsequent questionnaires updated those data and provided  
441 information on health endpoints.

442

443 **4.2.6.1 Dosimetry.** Because radiation dose measurements were unavailable for many individual technologists,  
444 especially in the earlier years when exposures were often higher, self-reported work history data were used to  
445 construct categorical proxy measures of radiation exposure. Individualized cumulative dose estimates were  
446 developed (Simon *et al.*, 2006b; 2014) based on a combination of badges, when available, inference of exposure  
447 from self-reported work history and protective measures used, and literature estimates of exposures during the  
448 early years.

449

450 A dosimetry update (Simon *et al.*, 2014) includes many more badge readings than the earlier 2006 analysis and  
451 calculated doses to 12 organs and tissues. The 2014 study included numerous methodological improvements that  
452 reduced the uncertainty in doses substantially. For example, the GSD for individual cumulative occupational  
453 female breast doses ranged from 1.5–3.0, depending on particular information available, whereas Simon *et al.*  
454 (2006) had estimated the uncertainty in individual “badge-equivalent” doses as GSD = 2.4 to 3.9). The 2014  
455 analysis utilized additional information on apron shielding as well as more accurate dose conversion coefficients  
456 than the earlier dosimetry and included a more comprehensive assessment of uncertainties. As a result of the  
457 improved dosimetry, Simon *et al.* (2014) found higher badge doses than the earlier dosimetry (median cumulative  
458 badge dose 47 mGy vs. 29 mGy). A validation sub-study showed a positive correlation between imputed bone  
459 marrow doses and frequency of chromosome translocations, though the size of the correlation was not provided  
460 (Simon *et al.*, 2014).

461

462 **4.2.6.2 Epidemiologic Methods, Findings, and Issues.** Mortality follow-up was conducted using the technologist  
463 registry information and mortality sources, including the U.S. National Death Index. Data have also been  
464 published on the incidence of various types of cancer, based on self-reports with verification of a fraction of the  
465 reports with medical records. The most recent mortality analysis for the U.S. technologist cohort was published by  
466 Liu *et al.* (2014). There were 9,566 deaths (3,329 cancer and 3,020 circulatory). The work history from the  
467 baseline questionnaire was used as a categorical proxy for radiation exposure. The conclusion was that  
468 technologists working before 1950 had increased mortality from a few cancers (breast cancer, leukemia) and some  
469 cardiovascular diseases, based on observing SMRs >1. Other reports have examined cancer mortality or incidence  
470 in relation to work in interventional radiography (Linnet *et al.*, 2006; Rajaraman *et al.*, 2016) or nuclear medicine  
471 (Kitahara *et al.*, 2015), again based on categorical proxies for dose. To date, only a few U.S. radiologic  
472 technologist publications have used the individual dose estimates that were developed: for cataract (Chodick *et*  
473 *al.*, 2008), skin cancer (Lee *et al.*, 2015), and breast cancer (Preston *et al.*, 2016). The mean estimated dose to the  
474 breast was 47 mGy overall, but means ranged from 6 mGy among those who began work in the 1980s to  
475 1,168 mGy among those who began before 1930.

476  
477 **4.2.6.3 Study Strengths and Weaknesses.** Dosimetry weaknesses included the use of questionnaires to determine  
478 work history, shielding, use of protective devices, and badging, which can result in unreliable information. The  
479 high imputed doses during the early years were based mainly on early literature, so uncertainty estimates may be  
480 too low for some individuals. It also is unknown how faithfully dosimetry badges were worn in some settings. The  
481 most significant source of uncertainty was associated with dose estimates prior to 1977 when only a very limited  
482 amount of individual monitoring data for cohort members was available. Other significant sources of uncertainty  
483 involved radiological exam procedures, badge placement, spatial homogeneity of the radiation fields, the energies  
484 of radiations, and use of protective devices. On the other hand, a strength of the dosimetric work is the detailed  
485 investigation of uncertainty and the calculation of multiple realizations of individual doses that account for shared  
486 vs. unshared uncertainty.

487  
488 This large cohort has good mortality ascertainment and a wide range of doses delivered in a protracted fashion  
489 the investigators have attempted to reconstruct individual doses and model dose uncertainties. Weaknesses include  
490 inaccuracy of long-term recall of exposure-related information, inability to verify the dosimetry in the early period  
491 when doses were higher, and reliance on self-reports for much of the cancer incidence data.. The greatest  
492 limitation has been the use of self-reported work history as a surrogate for dose in nearly all publications to date.  
493 Only a few papers have actually used quantitative dosimetry, and it has not yet been applied to the study of total  
494 solid cancer or leukemia risks.

495

496 **4.2.6.4 Implications for the LNT Model and Radiation Protection.** Because nearly all the publications have relied  
497 on self-reported, categorical data from work histories rather than having individual dose estimates for quantitative  
498 dose analyses, the study currently does not address the LNT model [except for the one report on breast cancer  
499 (Preston *et al.*, 2016)], nor does it contribute quantitatively to radiation protection issues.

500

501 **4.2.7 Million Worker Study**

502

### Highlights

A critically important gap in knowledge surrounds the health consequences of exposure to radiation received gradually over time, which in the past was studied only piecemeal in the United States. The Million Worker Study (MWS) is an active study of U.S. workers and veterans to address the effects of radiation when doses are delivered gradually over years. The study is ongoing and will ultimately include historical cohorts of DOE workers, U.S. atomic veterans, nuclear power plant workers, industrial radiographers and medical radiation workers.

The relatively great statistical power and precision of the study will be related to the large numbers, a broad dose distribution, standard methods used across all exposed cohorts and comprehensive dose reconstruction. The dose distribution is broader than for most occupational studies, with more workers with cumulative doses greater than 50 mGy than reported among atomic-bomb survivors. The follow-up started as early as 1940 for Manhattan Project workers. Recent studies of leukemia among nuclear power plant workers and industrial radiographers are consistent with a linear dose-response below 100 mGy, thus providing support for use of the LNT model in radiation protection.

503

504 **4.2.7.1 Introduction.** The Study of One Million U.S. Radiation Workers and Veterans (MWS) (Boice, 2012a)  
505 was conceived over 25 y ago following an NCI request to the NRC to consider creating a radiation worker  
506 registry based on the NRC requirement that nuclear power plant licensees report radiation doses received by  
507 their employees (Boice, 2012a; Muirhead *et al.*, 1996). Subsequently, the MWS was expanded to include  
508 115,000 atomic veterans who participated in above-ground nuclear weapons testing (Boice, 2012b; 2014;  
509 Caldwell *et al.*, 2016), 360,000 workers during the Manhattan Project years (Boice, 2013c; Boice *et al.*, 2006b;  
510 2014), 250,000 early radiologists and medical workers, 150,000 nuclear power plant workers (Boice, 2013a;  
511 2016a; Boice *et al.*, 2017), and 130,000 industrial radiographers. The follow-up started as early as 1940 for  
512 Manhattan Project workers. The overriding goal of the MWS is to estimate the risk of organ-specific cancers  
513 from radiation doses received at low dose rates over the course of years. These risk estimates from healthy  
514 American workers and veterans are deemed more valid for American workers than those from the atomic-bomb  
515 survivor study which involved a brief, acute exposure to a Japanese 1945 population living in a war torn  
516 country. The MWS is providing risk estimates that are more valid for American workers than is the atomic-

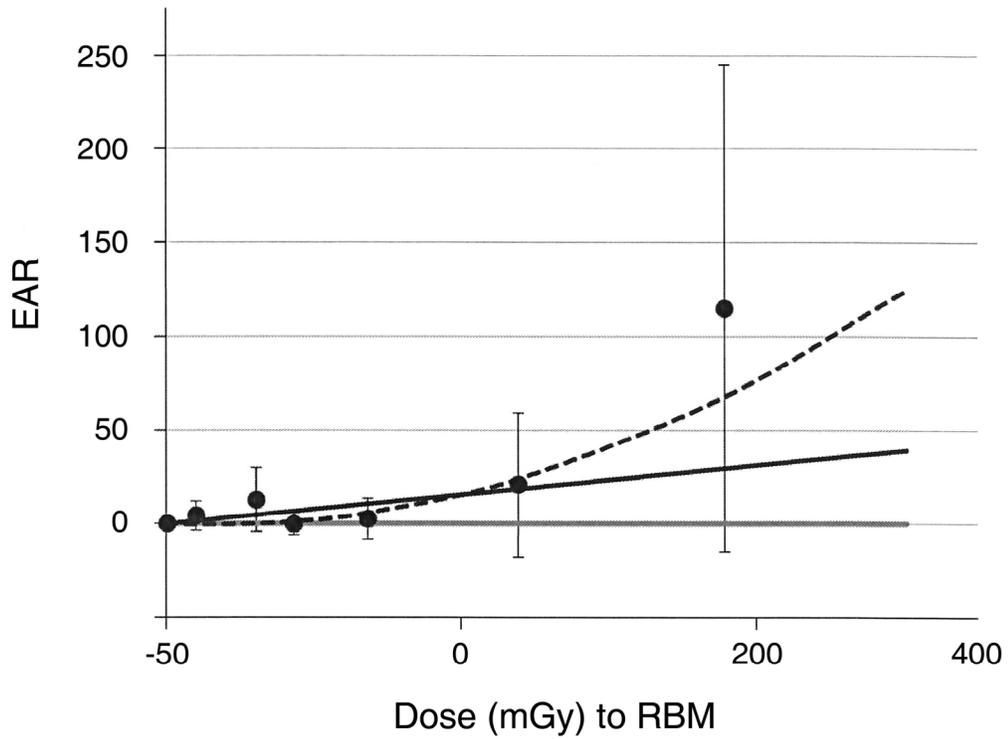
517 bomb survivor study which involved a brief, acute exposure to a Japanese population, who lived in a war-torn  
518 country with deprivation, malnourishment and infectious diseases, and who also have a profile of cancer types  
519 and other health conditions that differs from the U.S. population.

520  
521 Initial publications stressed the importance of dosimetry for both the DOE facility workers (Boice *et al.*,  
522 2006b; Leggett *et al.*, 2005), the atomic veterans (Beck *et al.*, 2017; Till *et al.*, 2014), and overviews of the  
523 entire dosimetry approaches for all cohorts studied (Boice, 2013b; 2014; 2016a; Bouville *et al.*, 2015).  
524 Statistical issues of uncertainty are being addressed (Stram *et al.*, 2015). Publications have included evaluations  
525 of the Rocketdyne (Atomics International) (Boice *et al.*, 2011) and Mound workers (Boice *et al.*, 2014). Papers  
526 have been submitted on the risk of male breast cancer among atomic veterans (Beck *et al.*, 2017) (Boice *et al.*,  
527 2017c), and on leukemia among nuclear power plant workers (Boice *et al.*, 2016b). The statistical power and  
528 precise risk estimates will come from the combination of the study cohorts. The earlier publications, while  
529 informative, are not sufficiently powerful to address radiation risks in the low-dose domain or the consistency  
530 or lack of consistency with an LNT model. However, as seen in Figures 4.5 and 4.6, more recent analyses of  
531 leukemia among the large populations of nuclear power plant workers and industrial radiographers have  
532 sufficient statistical ability to provide direct information on dose response relationships (Boice, 2017b) as the  
533 entire population epidemiology and dosimetry will eventually be combined into the large compiled effort. To  
534 date, subpopulations include: Rocketdyne Workers, Mound Workers, Mallinckrodt Workers, Atomic Veterans,  
535 Nuclear Power Plant Workers, Interventional Radiologists, and Medical Workers. Brief descriptions of the  
536 subpopulations that have been addressed to date and the results are provided below.

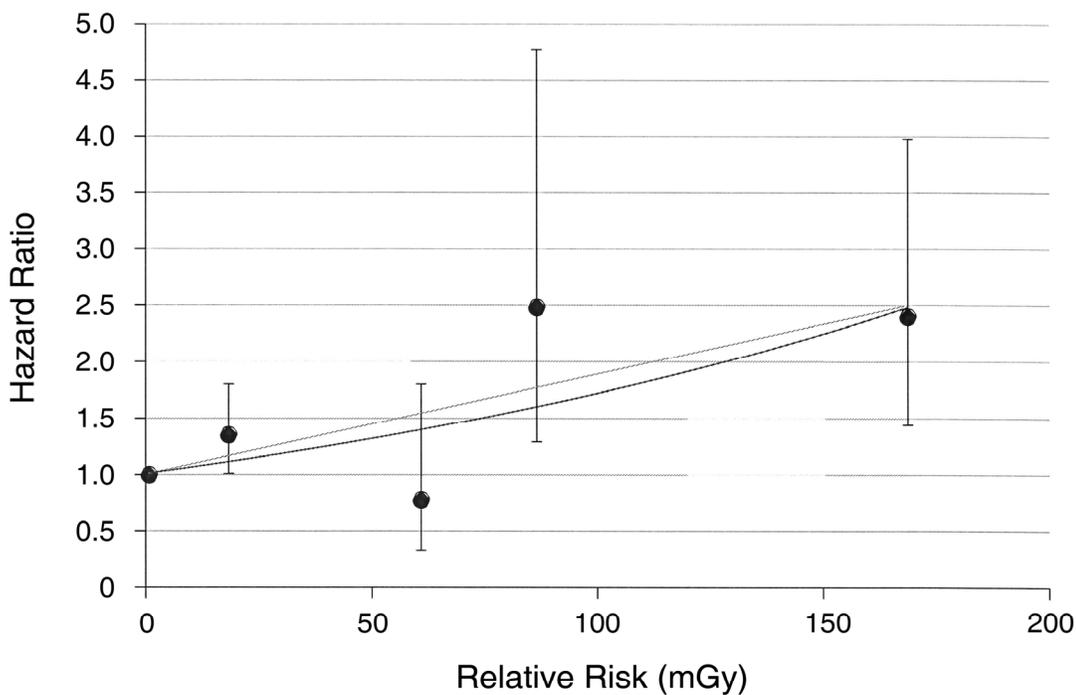
537  
538 **4.2.7.2 Dosimetry.** The approach to estimating occupational doses received by individuals included in the MWS  
539 follow the procedures outlined by Boice and colleagues (Boice, 2001; Boice *et al.*, 2006b; 2014). NCRP has  
540 undergone a review of both the dosimetry methods and the uncertainty associated with the reconstructions for the  
541 MWS and has provided an initial summary (Bouville *et al.*, 2015). A full report has been reviewed by the Council  
542 and is expected to be published in 2017. Specific details on the dosimetry aspects for each subpopulation are  
543 included in that report.

544  
545 **4.2.7.3 Epidemiologic Methods, Findings and Issues.** A brief summary of several of the MWS subpopulations,  
546 their exposures and results that have been or are being obtained are given below.

547  
548 **Rocketdyne Workers** (Boice *et al.*, 2006a; 2006b; 2011). Between 1948 and 1999, thousands of  
549 Rocketdyne/Atomics International workers were involved in a wide range of activities such as sodium-cooled



550  
551 **Fig. 4.5.** Excess absolute risk (EAR) dose response among nuclear power plant workers (preliminary  $n = 320$ )  
552 for leukemia (other than CLL). The dark solid line is the best linear slope, the dashed line is the linear-quadratic fit,  
553 and estimates (dots) and 95 % CI (vertical lines) are shown for individual dose categories (Boice, 2016b).  
554



555  
556 **Fig. 4.6.** Hazard ratio dose response among industrial radiographers (preliminary) for leukemia (other than  
557 CLL). The upper sloped line shows the best linear fit and the lower line the linear-quadratic fit. Hazard ratios for  
558 individual dose categories are shown by dots, with the 95 % CI shown by vertical lines.<sup>2</sup>  
559

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<sup>2</sup> John Boice, unpublished, 2017

560 breeder reactor technology, uranium fuel fabrication, spent fuel evaluation, radiography, hot lab chemistry,  
561 plutonium fuel fabrication and storage of nuclear material (Boice et al., 2006a; 2006b; 2011). During the 52 y  
562 covered by the study, 5,801 workers were monitored for external or internal radiation at Rocketdyne and  
563 1,833 workers were monitored at other facilities. Results were adjusted for year of birth, year of hire, sex, pay type  
564 (hourly/salary), duration of employment, and work as a test stand mechanic (which involved exposure to  
565 toxicants). Only 0.6 % were lost to follow-up, and cause of death was available for 98.1 % of deaths. There were  
566 651 deaths from all cancer except leukemia. The Cox regression hazard ratio (HR) at 100 mGy was 0.98 (95 % CI  
567 0.82, 1.17) which extrapolates to an ERR Sv<sup>-1</sup> of -0.2 (95 % CI -1.8, 1.7).<sup>3</sup> Given the null results, no analysis for  
568 nonlinearity was conducted. Similarly, for all leukemia other than CLL the HR<sub>100mSv</sub> was 1.06 (95 % CI 0.5, 2.23)  
569 which was extrapolated as an ERR Sv<sup>-1</sup> of 0.6 (95 % CI -5 to 12.3). Cox regression analyses revealed no  
570 significant dose-response trends for any cancer. Strong conclusions could not be drawn because of small numbers  
571 and relatively low career doses.

572  
573 **Mound Workers** (Boice *et al.*, 2014). Cancer mortality was examined among 7,270 workers at the Mound nuclear  
574 facility near Dayton, OH where <sup>210</sup>Po was used (1944 to 1979) in combination with beryllium as a source of  
575 neutrons for triggering nuclear weapons, including the Trinity and Nagasaki bombs (Boice *et al.*, 2014). Other  
576 exposures included external gamma radiation and to a lesser extent <sup>238</sup>Pu, tritium and neutrons. Radiation  
577 monitoring was conducted on 4,977 workers who also were followed up over the years 1944 to 2009. Lifetime  
578 occupational doses from all places of employment were sought and incorporated into the analysis. Over 200,000  
579 urine samples were analyzed to estimate radiation doses to body organs from polonium and other internally  
580 deposited radionuclides. The cohort was well defined from available records. Vital status of the cohort was  
581 achieved for 98.7 % of the cohort. There were 968 cancer deaths, including 31 leukemia cases (26 non-CLL).  
582 Analyses were adjusted for sex, education, year of birth and year of hire, and age at risk defined the follow-up time  
583 in the Cox regression analysis. External radiation dose-response analyses showed no significant association with  
584 death from any a priori cause. For leukemia (excluding CLL) the ERR was approximately 0.4 Gy<sup>-1</sup> (95 % CI -3.7,  
585 7.1). Combined internal and external radiation was not associated with lung cancer mortality ERR = 0.0 Gy<sup>-1</sup> (95 %  
586 CI -0.3, 0.4). Cox regression analysis revealed a significant positive dose-response trend for esophageal cancer  
587 [HR = 1.54 (95 % CI 1.15 to 2.07) at 100 mGy] and a negative dose-response trend for liver cancer HR = 0.55  
588 (95 % CI 0.23 to 1.32) at 100 mGy], but these are viewed as unexpected and possibly artifacts of chance or other  
589 factors.

590

---

<sup>3</sup> The ERR at 100 mSv was derived from the HR, *i.e.*,  $ERR_{100\text{ mSv}} \approx (HR_{100\text{ mSv}} - 1)$ . That ERR estimate was then extrapolated to an ERR Gy<sup>-1</sup> by multiplying by 10. The assumption of the approximate equivalence of the  $ERR \approx (HR - 1)$  at a low dose seems justified by the reasonably close correspondence shown in Cardis *et al.* (2007), Table 6.

591 **Mallinckrodt Workers** (Boice, 2017a; Golden *et al.*, 2016). An extended follow-up with comprehensive dose  
592 reconstruction was conducted of 2,514 white males employed at the Mallinckrodt Chemical Works in St. Louis,  
593 the earliest uranium processing facility in the United States, between 1942 through 1966 (Dupree- Ellis *et al.*,  
594 2000; Golden *et al.*, 2016). Workers processed pitchblende, a naturally occurring radioactive material containing  
595 uranium and silica. Over 75 % of the workers died during the up to 70 y of follow-up with cause of death known  
596 for 99 %. There was some evidence of a healthy worker effect, with the standardized mortality ratio (SMR) of 0.94  
597 for all causes, 0.97 for all cancer and 0.89 for all heart disease.

598  
599 The only significant radiation dose-response relationship was for kidney cancer [RR = 1.88 (95 % CI 1.13 to  
600 3.11) at 100 mGy] and was suggested for nonmalignant kidney diseases [RR = 1.36 (95 % CI 0.98 to 1.89) at  
601 100 mGy], such as nephritis. No notable observations were seen for any other cancer, leukemia or heart disease.  
602 Conceivably, dust could distort the association between kidney disease and radiation, related to nonradiogenic  
603 properties of uranium and silica (both kidney toxins), but the evidence was inconclusive.

604  
605 **Nuclear Power Plant Workers** (Boice, 2013a; 2016a; Boice *et al.*, 2017). In 1957 the United States became one  
606 of the first countries to produce electricity using nuclear power reactors. To date, 150,000 early workers in the  
607 nuclear industry, employed prior to 1985, have been selected for study from NRC files (the Radiation Exposure  
608 and Reporting System) and from dosimetry records from Landauer, Inc. Additional records, including early  
609 microfilm files, may be used to enhance the population size. Over the years, feasibility studies have been  
610 conducted utilizing NRC, Landauer, Inc., and utility records (Jablon and Boice, 1993; Muirhead *et al.*, 1996). A  
611 mortality study was conducted of 146,727 workers employed at nuclear power plants in the United States between  
612 1957 through 1984. Over 6 % of nuclear power plant workers had had cumulative exposure greater than 50 mGy.  
613 Follow-up through 2011 identified 30,993 deaths from all causes, including 68 from CLL and 320 from leukemia  
614 other than CLL. The SMR was not increased for CLL (0.90) or for non-CLL (1.02). Cox regression analyses  
615 revealed no evidence of a dose- response relationship for CLL. For non-CLL, preliminary results indicate a pattern  
616 of risk that was consistent with both a linear and linear-quadratic relationship with no elevation in the relative risk  
617 (hazards ratio) below 100 mGy and a nearly two-fold risk for exposures >250 mGy (Figure 4.6). A linear  
618 relationship between dose and leukemia could not be rejected, and the study is consistent with, though not strongly  
619 supportive of, the LNT model for radiation protection. Analyses are not yet available for solid cancers.

620  
621 **Industrial Radiography Workers** (Boice, ?) Industrial radiographic nondestructive testing typically utilizes <sup>192</sup>Ir  
622 and <sup>60</sup>Co sources. Industrial radiographers receive external irradiation, generally in an anterior- posterior (AP)  
623 geometry. Information on annual recorded dose has been collected by the MWS for 127,910 industrial  
624 radiographers. The average cumulative recorded dose for these industrial radiographers is ~20 mGy, with 10 % of

625 them receiving a cumulative recorded dose >50 mGy. Over 32,000 of the industrial radiographers are also known  
626 to have worked in naval shipyards. Approaches to follow-up, dose reconstruction, mortality ascertainment and  
627 analyses were similar to the study of nuclear power plant workers. Preliminary analyses for leukemia, other than  
628 CLL, have been conducted and the dose-response model is being assessed (Figure 4.6). The industrial  
629 radiographers include nearly 300 leukemia cases, in comparison with the atomic-bomb survivor data in adult  
630 males, *i.e.*, 94 cases, and the dose response curve is consistent with both a linear and a linear-quadratic model.  
631 Currently, the preliminary leukemia analyses for nuclear power plant workers and industrial radiographers are  
632 consistent with a dose response under 100 mGy and with the LNT model as used in radiation protection. For  
633 leukemia other than CLL, the MWS is consistent with a linear relationship for both the nuclear power plant  
634 workers and the industrial radiographers. A linear relationship would be expected when radiation is received  
635 gradually over time as these workers experienced. It might be considered an evaluation of the linear component of  
636 the linear quadratic relationship seen at higher dose rates. Thus, these data are consistent with the LNT model as  
637 used in radiation protection today.

638  
639 **Nuclear Weapons Test Participants (Atomic Veterans)** (Beck *et al.*, 2017; Caldwell *et al.*, 2016; Till *et al.*,  
640 2014). The United States conducted over 200 above-ground atmospheric nuclear weapons tests during the Cold  
641 War, many of which involved military maneuvers at the Nevada Test Site and the Pacific Proving Grounds (*e.g.*,  
642 Bikini Islands). A recent update of the SMOKY and PLUMBBOB veteran's studies included detailed dose  
643 reconstructions (Beck *et al.*, 2017; Caldwell *et al.*, 2016; Till *et al.*, 2014). There were 12,219 veterans at the  
644 PLUMBBOB test series, including 3,020 at the SMOKY nuclear test. Mortality follow-up was through 2010 and  
645 observed causes of death were compared with expected causes based on general population rates. Radiation dose to  
646 red bone marrow was based on individual dose reconstructions. Leukemia risk, initially reported to be significantly  
647 increased among SMOKY participants, remained elevated, but this risk diminished over time. Despite an intense  
648 dose reconstruction, the risk for leukemia was not found to increase with increasing levels of radiation dose to the  
649 red bone marrow. Based on a linear model, the estimated ERR mGy<sup>-1</sup> was -0.05 (95 % CI -0.14, 0.04). The  
650 observed null result for leukemia could be related to chance due to small numbers (27 cases) and low doses, or  
651 possibly to subtle biases in the data by lifestyle variations or other factors.

652  
653 For the Atomic Veterans, a full cohort study was also conducted of 114,270 military participants at over 100  
654 atmospheric nuclear weapons tests in Nevada, New Mexico (the Trinity test) and the Pacific from 1945 through  
655 1962 (Boice *et al.* 2017). Mortality follow-up was through 2010 and vital status for nearly 97 % of the veterans  
656 was determined. Radiation dose was based on comprehensive dose reconstructions included contributing causes of  
657 death (Beck *et al.*, 2017). No significant trends with dose were found, probably because of the dose uncertainties  
658 and the mostly small doses.

659  
660 **4.2.7.4 Study Strengths and Weakness.** For the MWS subpopulation studies, follow-up is generally long, with high  
661 success rates in follow-up and cause of death ascertainment. Dose reconstructions are detailed, they include  
662 individual badge measurements, individual internal dose evaluations from bioassays (where appropriate), and  
663 address specific uncertainties (Bouville *et al.*, 2015; NCRP, 2017). Organ doses are calculated for up to 16  
664 organs/tissues. Most studies can adjust for occupational status, sex, and other potential occupational radiation  
665 exposures (using information from other facilities before and after employment at the facility under study). The  
666 weaknesses for individual subpopulation studies are primarily the small numbers of cancers, appreciable dose  
667 uncertainties for some of the studies, and relatively low career doses that preclude strong conclusions. Data are  
668 often not available on smoking or other lifestyle sociodemographic factors. For the overall MWS, it is planned that  
669 subpopulations will be combined with other studies to generate the statistical power envisioned for dose response  
670 curves relevant to low dose risk estimation.

671  
672 **4.2.7.5 Implications for the LNT Model and Radiation Protection.** For each of the MWS subpopulation studies,  
673 because of the relatively small doses and number of deaths, and consequent wide confidence bounds, the individual  
674 studies make only a small contribution to the evaluation of the LNT model or radiation protection. Although  
675 several of the subpopulation studies include internal alpha emitters, and external photon exposures, and some study  
676 results have provided risk estimates for leukemia (with wide confidence intervals) most have not yet estimated  
677 quantitative risks for all solid cancer. At this time the individual studies provide little information relevant to  
678 considerations of LNT from low-LET irradiation. However, when all of the subpopulations are eventually  
679 combined, the statistical power will be much greater and it is expected that important information will likely be  
680 obtained along with a rigorous evaluation of associated uncertainties.

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682 **4.2.8 Chinese Medical X-Ray Worker Study**

683  
684 A total of 27,011 medical x-ray workers and 25,872 non-radiologist physicians in China between 1950  
685 and 1980 were followed up through 1995 for the incidence of all cancers excluding leukemia (Sun *et al.*,  
686 2016).

687  
688 **4.2.8.1 Dosimetry.** Personal dose monitoring did not begin until 1985. Therefore, earlier worker doses were  
689 estimated for 3,805 workers based on a simulation of multiple x-ray machines, workplaces, working conditions  
690 and protective measures used. Smoothed yearly averages, across all those conditions, without any individual  
691 information, were then applied to all active workers for each year of 1950 through 1995. For each work-year  
692 before 1949, the estimated 1949 average dose was applied. The mean cumulative  $H_p(10)$  dose was 0.25 Gy

693 (median, 0.12 Gy; ~60 % with cumulative dose <0.05 Gy, <1 % with >0.5 Gy); the corresponding mean colon  
694 dose was 0.086 Gy. However, for those who began work before 1950, the mean estimated cumulative colon dose  
695 was 0.583 Gy.

696

697 Since average yearly doses were applied to all workers, estimates of individual doses have substantial  
698 uncertainties, though they would be largely Berkson dose errors that would not be expected to bias the  
699 associations with risk. The use of colon dose for estimating risk to some organs may not be valid for low  
700 energy x rays, since the colon is one of the deepest organs.

701

702 **4.2.8.2 Epidemiologic Methods, Findings and Issues.** Investigators found 795 cancer cases in the exposed cohort  
703 and 848 in the unexposed cohort. Histologic information was available for about 70 % of cancers; the remaining  
704 were diagnosed mainly from radiographic evidence. Investigators calculated an ERR of  $0.87 \text{ Gy}^{-1}$  (95 % CI 0.48,  
705 1.45) based on estimated colon dose. They also showed risks by dose category but provided only linear ERR and  
706 EAR risk estimates.

707

708 **4.2.8.3 Study Strengths and Weaknesses.** Strengths include the fact that the exposed and unexposed groups  
709 came from the same hospitals, thereby controlling for regional differences. They observed differences between  
710 the exposed and unexposed in rates of lung, esophageal and liver cancers, which they surmised might be due to  
711 differences in smoking and hepatitis infection. However, an analysis with lung and liver cancer deleted gave a  
712 risk estimate very similar to the all solid cancer estimate.

713

714 Weaknesses of the study include the fact that the occupational radiation dose assigned to individual  
715 workers was the estimated average annual dose for all workers and that the medical diagnostic information  
716 was limited. They noted that the completeness of follow-up for cancer was uncertain. There is also concern  
717 that the exposed group of mixed radiologists and radiation technologists may have differed in socioeconomic  
718 status from the unexposed group consisting only of physicians, which might affect the risk estimate.

719

720 **4.2.8.4 Implications for the LNT Model and Radiation Protection.** Though the study shows an essentially  
721 linear dose-response association, the uncertainties in both dosimetry and cancer documentation mean it  
722 provides relatively weak information regarding the LNT model and radiation protection.

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725

### 4.3 Environmental Exposure Studies

#### Highlights

Several epidemiologic studies described below have been conducted to investigate whether exposure to ionizing radiation from a variety of environmental sources increases the risk of developing or dying from cancer. Between 1949 and 1956 the Russian Mayak nuclear weapons facility released radioactive waste into the Techa River and exposed approximately 30,000 residents to relatively low doses at low dose rates from gamma rays (external) and  $^{137}\text{Cs}$  and  $^{90}\text{Sr}$  (internal). A recent paper on cancer incidence among residents near Techa River reported an ERR  $\text{Gy}^{-1}$  estimate for all solid cancer of 0.77 (95 % CI 0.13, 1.5). Although the shape of the dose response was uncertain, especially at doses below about 100 mGy, the study provides fairly strong evidence that exposures at low dose rates confer risk for solid cancers.

New studies of cohorts of children in Ukraine and Belarus who had thyroid measurements of  $^{131}\text{I}$  activity shortly after the Chernobyl accident and systematic thyroid screening have added appreciably to our knowledge about thyroid cancer risk after protracted internal exposure. Both studies showed strong linear dose-response functions with no evidence of nonlinearity. The risks per unit dose were somewhat lower than, but statistically compatible with, those found for acute external radiation exposures.

Studies of residents of high natural background radiation have been conducted in Kerala, India and Yangjiang, China. However, it is difficult to conduct a study of background radiation, *e.g.*, to find a suitable low exposure control group with highly similar lifestyles and natural disease rates to whom the highly exposed group may be compared. The better and larger of the two studies, the Kerala study of cancer incidence, included 70,000 individuals and over 1,300 cancers from high-background or low-background radiation areas. The dosimetry was based on measurements of ambient levels within and near homes, coupled with aggregate house-occupancy factors by age and sex. They reported an ERR  $\text{Gy}^{-1}$  of  $-0.13$  (95 % CI  $-0.58, 0.46$ ) for all cancer except leukemia. The leukemia cases were too few to be informative. The Yangjiang study reported a positive, but nonsignificant, risk coefficient for cancer, excluding leukemia and liver cancer (ERR  $\text{Gy}^{-1}$  of 0.19, 95 % CI  $-1.87, 3.04$ ). These studies are nominally more supportive of little or no effect after low dose-rate exposures rather than the LNT model. However, the uncertainties in dosimetry, the weaknesses in cancer ascertainment, the low rates of histological verification of cancer cases, questions about the comparability of the low- and high-dose populations, and the wide confidence intervals on the risk estimates mean they need to be interpreted with caution.

A number of environmental studies are reviewed, with greater detail accorded to the more important or informative studies. The studies include Techa River residents (Section 4.3.1), Chernobyl residents (4.3.2), Kerala, India High Background Radiation Area (HBRA) residents (4.3.3), Yangjiang, China HBRA residents (4.3.4), Taiwan residents of radiocontaminated buildings (4.3.5), and radiation fallout studies (brief reviews)

(4.3.6) of Japan atomic-bomb fallout, Marshall Islands atomic testing fallout, Nevada Test Site (NTS) atomic fallout in Utah, Atomic testing fallout across the United States, Semipalatinsk fallout, Hanford <sup>131</sup>I fallout, Mayak fallout, Three Mile Island fallout and Fukushima Dai-Ichi fallout.

In addition to the text below, further systematic information is provided for the Techa, Chernobyl, Kerala, Yangjiang and Taiwan studies in Tables 4.1 to 4.4.

726

727 **4.3.1** *Techa River Resident Cohort*

728

729 Between 1949 and 1956 the Mayak Production Association, which is located in the southern Ural Mountains,  
730 released radioactive waste into the Techa River as part of the process of producing plutonium for the Soviet  
731 nuclear weapons program. Approximately 30,000 residents of 41 villages along the river were potentially exposed  
732 to radiation from these releases. External exposures were due to gamma rays from contaminated shorelines and  
733 flood plain soil. Internal exposures were from the consumption of water, milk and food contaminated with <sup>137</sup>Cs,  
734 <sup>90</sup>Sr, <sup>89</sup>Sr and other uranium fission products (Davis *et al.*, 2015; Degteva *et al.*, 2012; Kossenko *et al.*, 2005;  
735 Napier, 2014; Tolstykh *et al.*, 2011). This review updates the reviews by UNSCEAR (2008) and the BEIR VII  
736 committee (NA/NRC, 2006) of Techa River studies.

737

738 **4.3.1.1** *Dosimetry*. Analyses utilized dose estimates computed using the Techa River Dosimetry System 2009  
739 (TRDS-2009) (Degteva *et al.*, 2009), which is an improved version of TRDS-2000 (Degteva *et al.*, 2000). External  
740 exposure was estimated using exposure rate measurements available at various locations along the river bank, and  
741 a model of the river and flood plain contamination based on source term estimates (releases) derived from  
742 historical records of the facility. Doses from external exposure decreased with the distance along the Techa River  
743 and distance from the shoreline. Additional factors were characteristics of the riverbank, residence history, average  
744 gender- and age-dependent behavioral factors, and age-related coefficients to convert air to organ doses. (Degteva  
745 *et al.*, 2000; Schonfeld *et al.*, 2013). Exposure levels peaked in 1951 and since then have declined over time. The  
746 mean estimated cumulative stomach dose was 0.035 Gy (range 0 to 0.96 Gy).

747

748 Exposures to internal radionuclides (primarily <sup>90</sup>Sr and <sup>137</sup>Cs) also were estimated taking into account factors  
749 such as residence history and drinking water sources. Whole-body measurements of <sup>90</sup>Sr were available for some  
750 residents and were used to calibrate the model calculations.

751

752 While no formal uncertainty analysis was conducted and more precise individual doses and associated  
753 uncertainty are clearly preferred, using such individual dose estimates probably results in mostly Berkson

754 (grouped) errors, which introduces increased variance and reduced statistical power, but little bias provided the  
755 grouped estimates are unbiased. Less than 10 % of the person-years were at cumulative doses over 100 mGy,  
756 which indicates it is primarily a low-dose study.

757  
758 TRDS-2009 suffers from certain limitations. The  $^{90}\text{Sr}$  whole-body measurements, of necessity, were made  
759 decades after intake. Individual bone reabsorption rates vary by time, sex and other factors, and doses from the  
760 short-lived fission products had to be inferred with limited data. The source term for the activity in the river was  
761 not well established, and the river model itself has subsequently been revised substantially. Although diagnostic  
762 examinations were not accounted for in the dosimetry, evidence indicates they were not highly correlated with  
763 Techa River doses, so would not introduce substantial bias.

764  
765 External dose estimates were crudely validated on the village level using physical measurements of  
766 thermoluminescence in bricks. Tooth enamel electron spin resonance (ESR) was also used to validate the cumulative  
767 external exposure. Comparison with available but limited measurements was used to validate the river model. The  
768 consistency between the model calculations and the results of actual whole body counter measurements in humans  
769 was used to assess the reliability of the model used in TRDS-2009D for the calculation of internal doses due to  
770 strontium. FISH data in 2009 were considered preliminary but were roughly consistent with estimated doses.

771  
772 TRDS-2009 is being replaced by a more accurate and precise dose system (TRDS-2017) that among other  
773 improvements incorporates a more accurate river model and a more accurate source term. TRDS-2017 also  
774 includes estimates of confounding exposures from medical irradiations, exposure to the Kyshtym accident and  
775 Mayak stack releases, an improved internal dosimetry model for  $^{90}\text{Sr}$  bone dose, as well as other improvements.  
776 TRDS-2017 will provide more accurate individual doses along with shared and unshared uncertainty estimates  
777 based on multiple realizations of doses using a two-dimensional Monte-Carlo simulation.

778  
779 **4.3.1.2 Epidemiologic Methods, Findings and Issues.** The study cohort consisted of approximately 28,000  
780 individuals of either sex, 80 % of whom lived in villages along the Techa River between 1950 and 1953, the time  
781 of the greatest exposures, and met other eligibility requirements (Schonfeld *et al.*, 2013). About 40 % were under  
782 age 20 at the time. Vital status was determined from oblast (region) address bureaus, periodic contact letters to  
783 individuals, tax and passport records and personal interviews. About 22 % were lost to follow-up, mainly due to  
784 out-migration, but this was accounted for by censoring of the follow-up time. Cause of death information was  
785 obtained from oblast Civil Registrars Offices and was known for 91 % of deaths. Cancer incidence data were  
786 obtained from regional oncology dispensaries and a central cancer hospital and were thought to have nearly 100 %  
787 ascertainment, except for out-migrants from the region. Distant out-migrants tended to have higher doses, and

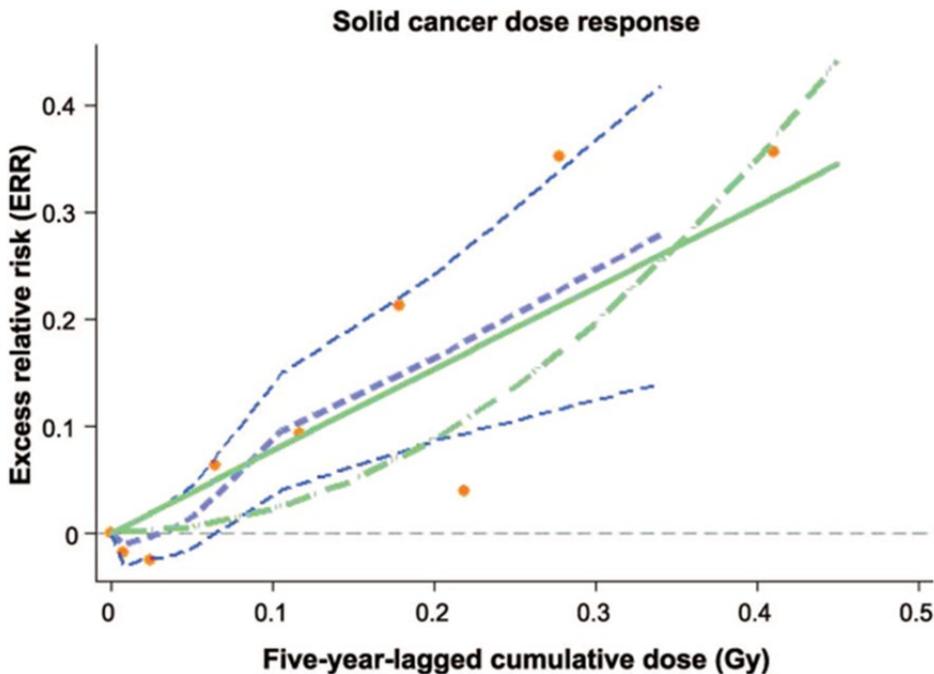
788 death and cancer incidence data were less known for them (Kossenko *et al.*, 2005), but the follow-up censoring in  
789 effect removed this as a biasing factor.

790  
791 The study of solid cancer mortality in the Techa River cohort during 1950 to 2007 included 2,303 deaths.  
792 Analyses adjusted for sex, ethnicity, attained age, birth cohort, and oblast (region). The linear ERR Gy<sup>-1</sup> was 0.61  
793 (95 % CI 0.04, 1.27) (Schonfeld *et al.*, 2013). Adding a quadratic term to the linear one did not improve the fit of  
794 the model ( $p > 0.5$ ). Nevertheless, a pure quadratic model fit the data as well as the linear model. To further focus  
795 on nonlinearity, the authors fit a spline model with a knot at 100 mGy; it suggested curvature but did not fit the  
796 data better than a pure linear model. The best estimate of a dose threshold was at 0.05 Gy, which had a lower  
797 confidence bound of  $<0$  and was therefore compatible with a nonthreshold model. It was reported that the highest  
798 risks were seen for the esophagus and uterus, findings at variance with other studies; however, given that there  
799 was a total of only 50 radiation-associated excess solid cancer deaths, variations by individual cancer type may  
800 not be meaningful. The investigators reported that those exposed at older ages had borderline larger risks (ERR)  
801 than those exposed or observed when young, which runs counter to other studies; it may simply represent small  
802 sample variation. In fact, the age effects were unstable, depending on how the background rates were modeled,  
803 but the age and tumor type findings temper the conclusions that can be drawn.

804  
805 The other major Techa River cohort report is of cancer incidence, 1956 to 2007 (Figure 4.7) (Davis *et al.*, 2015).  
806 The cohort consisted of about 17,000 rather than 28,000 because cancer incidence could be determined in only one  
807 oblast; in addition, 21 % had out-migrated while another 6 % was lost to follow- up. With adjustment for smoking,  
808 the linear ERR y<sup>-1</sup> for all solid cancer was 0.77 (95 % CI 0.13, 1.5). It was estimated that about 3 % (or 61) of the  
809 1,933 incident solid cancers were statistically associated with the radiation exposure. There was considerable  
810 uncertainty about the shape of the dose response, though at low doses there was no indication that a linear-quadratic  
811 model fit any better than a simple linear model ( $p = 0.2$ ). On the other hand, the pure quadratic model fit as well as  
812 the linear dose-response model ( $p > 0.5$ ), with an estimated ERR of 0.022 (95 % CI 0.005, 0.04) at 100 mGy, which  
813 was less than half as large as the linear model estimate (0.077) at that dose level. Using a two-stage clonal expansion  
814 (TSCE) biological model with follow-up data through 2003, Eidemüller *et al.* (2010) reported ERRs of 0.85 (95 %  
815 CI 0.36, 1.38) Gy<sup>-1</sup> for all cancer mortality and 0.91 (95 % CI 0.35, 1.52) Gy<sup>-1</sup> for cancer incidence. These values  
816 are somewhat higher than, but statistically compatible with, the empirical-model estimates by Schonfeld *et al.*  
817 (2013) and Davis *et al.* (2015).

818  
819 Krestinina *et al.* (2013) analyzed leukemia incidence in the Techa River cohort. The RBM doses ranged up  
820 to 2 Gy (0.3 Gy mean). The mean RBM dose rate decreased from 40 mGy y<sup>-1</sup> in 1950 to 1951 to 8 mGy y<sup>-1</sup> in  
821 1960, and about 92 % of the total RBM dose was from internal exposure, primarily <sup>90</sup>Sr. There were 70 cases of

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**Fig. 4.7.** Techa River dose response for solid cancer incidence. All results shown are based on models with adjustment for smoking in the baseline rates. The green solid lines are the fitted linear (solid) and quadratic (dash-dot-dot) dose-response curves. The orange points are ERR estimates in dose categories while the thick-blue-dashed curve is a nonparametric smooth fit to these points. The outer blue-dashed curves represent approximate (pointwise) +/- standard error limits on the nonparametric smooth (Davis *et al.*, 2015).

848  
849 non-CLL leukemia. Most leukemias (81 %) were diagnosed by qualified hematologists. The linear estimate of  
850 ERR Gy<sup>-1</sup> was 4.9 (95 % CI 1.6, 14). There was no indication of non-linearity ( $p > 0.5$ ); the estimated curvature  
851 was close to 0. About 59 % of the non-CLL cases were estimated to be associated with radiation exposure.

852  
853 **4.3.1.3 Study Strengths and Weaknesses.** The Techa River studies consist of a relatively large, unselected  
854 population of men, women and children with a long follow-up such that approximately 75 % have died. The  
855 range of estimated cumulative doses was fairly broad. It was possible to adjust for several covariables in the  
856 analysis, including smoking and ethnic differences. Though the mortality study had limitations in  
857 histopathologic verification of cancer deaths, it is notable that the risk estimates for solid cancer mortality and  
858 incidence, where the latter had better histopathologic verification, were similar. The study suffers from  
859 uncertainties in dose reconstruction, potential medical screening bias, concomitant radiation exposure from  
860 medical screening, losses to follow-up, and uncertain quality and completeness of outcome ascertainment  
861 (NCRP, 2012). While the dose reconstruction methodology is commendable, it has considered uncertainties  
862 because of the wide range of assumptions required. The measurements, of necessity, were made decades after  
863 intake. Another problem is that radiation exposures from medical fluoroscopic examinations were not accounted  
864 for, and those were sometimes very substantial and were correlated with Techa River exposure levels (Degteva  
865 *et al.*, 2007).

866  
867 The epidemiologic methods also had limitations: >20 % of the Techa River population was lost to follow-up,  
868 mainly due to out-migration, cause of death was missing for 9 % of deaths, and in the earlier years ~35 % of the  
869 solid cancer cases were morphologically or radiologically confirmed though since 1990 over 80 % have such  
870 confirmation (Davis *et al.*, 2015; Krestinina *et al.*, 2007). The small number of excess cancers limits the statistical  
871 precision and thus tempers interpretations when comparing results with larger studies. The dosimetric and  
872 epidemiologic limitations constrain the weight attached to this study in informing LNT and radiation protection  
873 (UNSCEAR, 2000; 2008).

874  
875 **4.3.1.4 Implications for the LNT Model and Radiation Protection.** Studies based on the TRDS-2009 dosimetry  
876 provide only modest support for the LNT model, particularly for leukemia where internal dose from <sup>90</sup>Sr is  
877 important. It is expected that new studies based on the upcoming TRDS-2017 will allow a better evaluation of the  
878 LNT model. The Techa River studies are quite important in showing that low dose-rate exposures over time may  
879 increase the risk of cancer in human populations. The recent studies of the Techa River Cohort have reported  
880 associations between radiation dose and incidence and mortality rates for solid cancers and non- CLL leukemia  
881 that appear to be linear in dose response (Davis *et al.*, 2015; Krestinina *et al.*, 2013a; Schonfeld *et al.*, 2013).

882 However, the small number of cancer cases, the challenging dosimetric issues, the possible confounding influence  
883 of medical irradiation, the peculiar inconsistency with other studies in age patterns and types of excess cancers,  
884 and other methodological limitations of follow-up and case ascertainment prevent strong inferences about the  
885 shape of the dose-response curves and DDREF. Thus the Techa River studies provide general support for the LNT  
886 model, but the results at doses under 100 mGy are uncertain, so implications for radiation protection are limited.

887

#### 888 4.3.2 *Chernobyl Resident Cohorts*

889

890 The Chernobyl accident in northern Ukraine introduced a new era of concern about the effects on humans of  
891 radiation at low doses and low dose rates. There was a clear and large excess of thyroid cancer cases among  
892 those heavily exposed to radioiodine at a young age in Ukraine, Belarus and parts of Russia following the  
893 Chernobyl accident (Brenner *et al.*, 2011; Jacob *et al.*, 2006; Tronko *et al.*, 2006; Zablotska *et al.*, 2011). The  
894 data on exposure to radioactive iodines have added considerable information relative to the dose-response  
895 relationship. In 2006, the BEIR VII committee published a report of Chernobyl research activities (NA/NRC,  
896 2006) and concluded with regard to thyroid cancer that “the linear-no-threshold model (LNT) provided the most  
897 reasonable description of the relation between low-dose exposure to ionizing radiation.” UNSCEAR followed  
898 with its own comprehensive report (UNSCEAR, 2011) on Chernobyl that was consistent with that conclusion as  
899 well. An update of the key studies of defined cohorts is summarized below.

900

901 4.3.2.1 *Dosimetry*. The thyroid dosimetry for both the Ukrainian and Belarusian cohorts was based on direct  
902 measurements of thyroid activity for all study subjects supplemented by estimates from a pathway model using  
903 ground deposition data [methods described in Likhtarev *et al.* (2003; 2006; 2014)] and reports obtained by personal  
904 interviews of behavioral factors *e.g.*, amount of local milk drunk at the time of the Chernobyl accident. The direct  
905 thyroid measurements were conducted under difficult conditions within a few weeks of the accident. The estimated  
906 uncertainties in thyroid doses are lower than most previous studies, mainly because of the availability of direct  
907 measurements of thyroid activity for all study subjects.

908

909 For the pathway model estimates, whereabouts of individuals, consumption details, and other variables were  
910 based on questionnaires with associated possible recall error. Although some inconsistencies have been observed,  
911 comparisons between the direct measurement-based doses and the pathway-based doses suggested that overall  
912 agreement was quite good. However, the earlier epidemiologic evaluations of the Ukrainian cohort used versions of  
913 the dosimetry that had some issues regarding thyroid mass, and the analysis did not take shared uncertainty into  
914 account. The most recent dosimetry analysis of the Ukrainian data (Likhtarov *et al.*, 2014) is considered more  
915 accurate, and among other things, reflects an increased understanding of how to account for thyroid mass.

916 Parameters of the model used in the previous dosimetry were also substantially improved. The GSDs obtained for  
917 the 13,204 cohort members in the newer dosimetry varied from 1.26 to 10.6, with a geometric mean of 1.47 and  
918 <4 % with GSDs greater than two.

919  
920 For the Belarusian studies, thyroid doses were based on methods identical to those used in Ukraine. Doses  
921 based on pathway modeling were used for validation and uncertainty estimation. Zablotska *et al.* (2011) used the  
922 deterministic doses reported by Drozdovitch *et al.* (2013). Their mean estimated dose was 0.56 Gy (range up to  
923 32.8 Gy), based on personal thyroid exposure measured shortly after the accident. The estimated individual dose  
924 uncertainties ranged from a GSD of 1.3 to 5.1 with an average GSD of about 1.7. A more recent dosimetry study  
925 investigated uncertainties in detail and estimated the shared vs, unshared components using two-dimensional  
926 Monte Carlo with multiple realizations (Drozdovitch *et al.*, 2015). They found that the GSD in individual thyroid  
927 doses varied among cohort members from 1.33 to 5.12, with a geometric mean of 1.73. The mean dose estimated  
928 by Drozdovitch *et al.* (2015), 0.68 Gy, was slightly higher than the earlier deterministic estimate of 0.56 Gy. For  
929 both the Ukraine and Belarus studies, the major sources of uncertainties were errors in the <sup>131</sup>I activities in the  
930 thyroids derived from the direct thyroid measurements and errors in assigning individual thyroid-mass values  
931 (Drozdovitch *et al.*, 2015). Little *et al.* (2014) concluded that dose-error adjustment has comparatively modest  
932 effects on regression parameters in both studies.

933  
934 **4.3.2.2 Epidemiologic Methods, Findings and Issues.** Among the earliest reports were ecological analyses of  
935 thyroid cancer risks. For example, a report by Jacob *et al.* (1998) showed a reasonably linear dose response  
936 relationship but at a slope lower than that for external radiation, whereas a later paper found a high linear slope with  
937 a significant negative quadratic term (Jacob *et al.*, 2006). However, these studies were based on ecologic, and not  
938 individual data, that were complicated by potential goitrogenic and screening effects, and limited <sup>131</sup>I thyroid  
939 measurements. A population-based case-control study by Cardis *et al.* (2005a) found that having low levels of  
940 dietary stable iodine did augment thyroid cancer risk, but reported a radiation effect independent of stable iodine  
941 intake levels.

942  
943 In the first decade or so after the Chernobyl accident, some thought that the sudden and sharp increase in  
944 thyroid cancer was mostly, if not entirely, due to aggressive screening programs instituted in the most contaminated  
945 areas, but an early case-control study showed that the increase was not merely a screening artifact (Astakhova  
946 *et al.*, 1998). Since then reports of cohorts of exposed children in Ukraine or Belarus with thyroid dosimetry based  
947 on thyroid <sup>131</sup>I measurements at the time of the Chernobyl accident and systematic thyroid screening have provided  
948 decisive information that the excess of thyroid cancer and thyroid nodules is radiation related. A systematic thyroid  
949 screening study of a cohort of children exposed to <sup>131</sup>I was initiated in Ukraine (Tronko *et al.*, 2006).

950 Approximately 13,000 individuals were exposed before age 18 and had thyroid measurements taken within two  
951 months after the accident (Likhtarev *et al.*, 2003; 2006). The principal radiation exposure to the thyroid was from  
952  $^{131}\text{I}$  (beta) and, to a much smaller extent, external gamma. The estimated mean thyroid dose in the screenees was  
953 0.79 Gy (median 0.26 Gy, maximum 47.6 Gy). About 53 % had estimated thyroid doses  $<0.3$  Gy, 27 % 0.3 to  
954  $<1$  Gy, and 19 %  $\geq 1$  Gy. The GSDs of 96 % of the dose estimates were less than 2.0.

955  
956 The participants were given thyroid screening by ultrasound and palpation to detect thyroid nodules and cancer  
957 between 1998 and 2000 (12 to 14 y after the accident). Forty-five pathologically confirmed cases of thyroid cancer  
958 were identified. Analyses were adjusted for sex and age at screening (or age at exposure). The dose response for  
959 thyroid cancer prevalence was essentially linear; an added quadratic term was nonsignificant ( $p > 0.99$ ). The linear  
960 ERR dose response was  $5.3 \text{ Gy}^{-1}$  (95 % CI 1.7, 28) overall and  $6.2 \text{ Gy}^{-1}$  over the dose range of 0 to 10 Gy, risk  
961 estimates similar to those for external radiation (Veiga *et al.*, 2016). The ERR estimates were nonsignificantly  
962 higher for females and for those youngest at exposure.

963  
964 The Ukrainian cohort was given three more thyroid screenings at approximately 2 to 3 y intervals to evaluate  
965 the incidence of thyroid tumors (Brenner *et al.*, 2011). An additional 65 thyroid cancers were diagnosed.  
966 Adjustment factors included sex, age at exposure, age at screening and calendar time, and sensitivity analyses  
967 examined oblast, urban/rural, smoking status, goiter, history of iodine prophylaxis and family history of thyroid  
968 disease. The dose response was approximately linear with no evident quadratic curvature ( $p = 0.31$ ). They found a  
969 somewhat lower, but still compatible ERR estimate of  $1.9 \text{ Gy}^{-1}$  (95 % CI 0.4, 6.3). There were no significant  
970 interactions of radiation risk with sex, age at exposure, screening age, urinary iodine status or history of iodine  
971 prophylaxis.

972  
973 A parallel thyroid cancer prevalence study was conducted in Belarus (Zablotska *et al.*, 2011). Mainly between  
974 1996 and 2001, approximately 12,000 young people who were  $\leq 18$  y old at the time of the Chernobyl accident  
975 were examined. The study participants were screened for thyroid cancer with ultrasound and palpation, and 85 cases  
976 were identified (excluding 53 additional cases diagnosed with thyroid cancer more than 3 y prior to the  
977 examination). The ERR was  $2.2 \text{ Gy}^{-1}$  (95 % CI 0.8, 5.5) among those with estimated doses up to 5 Gy, with a  
978 downturn in the dose response above 5 Gy. A number of covariables were examined as possible confounders and  
979 adjusted for as needed, including sex, age at screening, age at exposure, oblast of residence, urban/rural status,  
980 history of nodular goiter, past iodine deficiency and current urine iodine level.

981  
982 In summary, there was a small suggestion that risk may be lower with protracted irradiation from  $^{131}\text{I}$  than from  
983 acute external irradiation, but the risk estimates were statistically compatible. Evidence of an excess risk of thyroid

984 cancer in areas less affected by Chernobyl contamination was much less clear. In Finland, which was one of the  
985 countries most affected by Chernobyl contamination outside the former USSR, no increased risk of thyroid cancer  
986 after juvenile exposure was detected (But *et al.*, 2006), largely because the thyroid doses of a few milligray were far  
987 lower than those received around Chernobyl.

988  
989 Doses to other organs were much smaller than to the thyroid for residents downwind from Chernobyl.  
990 Because of the low and very uncertain whole-body exposure levels for individuals, mainly from  $^{137}\text{Cs}$ , the  
991 radiation results for other organs have been null or inconsistent [for instance, see a review of leukemia findings  
992 by Howe (2007) and the Ukraine-Belarus-Russia study by an international consortium (Davis *et al.*, 2006)].

993  
994 **4.3.2.3 Study Strengths and Weaknesses.** The principal Ukrainian and Belarusian cohort studies of thyroid cancer  
995 prevalence and incidence in young people after exposure to  $^{131}\text{I}$  from the Chernobyl accident included only  
996 children who had received radiation measurements of the thyroid gland shortly after the accident as the basis for  
997 individual dose reconstructions. They received several follow-up clinical and ultrasound examinations, performed  
998 blindly with respect to thyroid dose, and further fine needle aspiration cytology and surgery when indicated  
999 according to standard protocols. The cancers were verified by international pathology reviews. Limitations  
1000 include some potential sample selection bias due to prior thyroid cancers, voluntary participation and losses to  
1001 follow-up. In spite of the thyroid activity measurements, there were uncertainties in estimating doses due to  
1002 variations in thyroid mass, uptake and retention, timing of exposure, and unreliability of behavioral and milk-  
1003 intake reports.

1004  
1005 **4.3.2.4 Implications for the LNT Model and Radiation Protection.** The thyroid cancer experienced by children in  
1006 exposed areas of the Ukraine, Belarus and Russia conforms to the LNT model, though perhaps with a somewhat  
1007 lower risk per unit dose than seen in studies of children exposed to external gamma radiation (Little *et al.*, 2014).  
1008 The UNSCEAR (2011) report on Chernobyl concluded that the thyroid risk findings were consistent with the LNT  
1009 model. The dosimetry, although uncertain as described above, is believed sufficiently accurate to support that  
1010 conclusion.

1011

1012 **4.3.3 High Natural Background Radiation Area (HBRA) Studies: Kerala Study**

1013

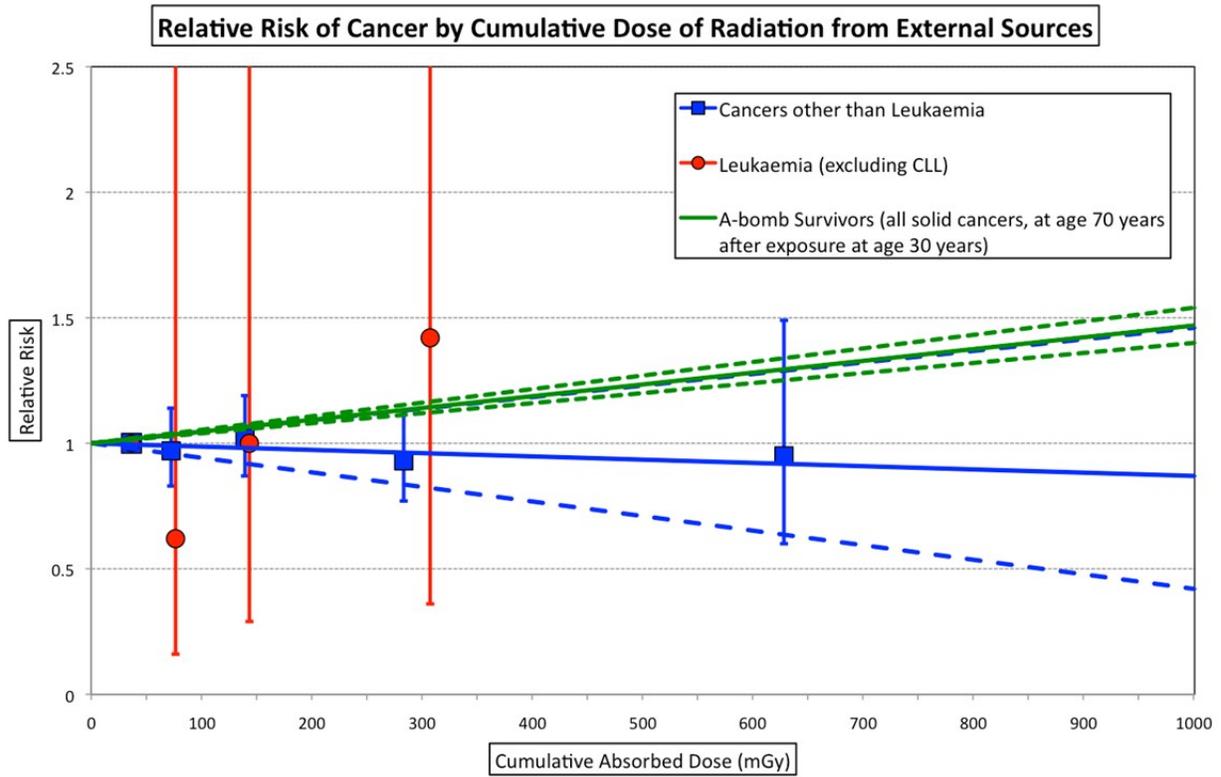
1014 Since exposure to ionizing radiation is ubiquitous, for epidemiologic studies to detect any effect of natural  
1015 background radiation upon health, it is necessary for different groups of people to receive materially different  
1016 doses from naturally occurring radiation sources. Differences in the doses received from naturally occurring low-  
1017 LET radiation do exist, although not to the extent of the variation of lung doses from radon and its progeny. There

1018 are areas of the world with high levels of natural background gamma radiation because of local geology, such as  
1019 thorium-bearing monazite sands. Such areas occur in Iran, Brazil, India and China (Hendry *et al.*, 2009). Studies  
1020 of high natural background gamma radiation areas are difficult to conduct because large numbers of exposed  
1021 people are required to achieve adequate statistical power to detect the predicted effects. Further, it may often be  
1022 difficult, especially with studies in developing countries, to find a suitable low exposure control group with which  
1023 the highly exposed group may be compared in terms of detailed lifestyles, completeness of disease ascertainment,  
1024 and baseline disease rates. However, there have been updates of the two major studies of high natural background  
1025 gamma radiation areas since the BEIR VII and UNSCEAR reports, described below, which might be of some use  
1026 in assessing the LNT model.

1027  
1028 For the Kerala epidemiologic study a cancer incidence registry was developed in Karunagappally, Kerala,  
1029 India that compiled data between 1990 and 2005 for about 70,000 persons of ages 30 to 84 y to compare low  
1030 and high background radiation areas (HBRA) (Akiba, 2013; Nair *et al.*, 2009).

1031  
1032 **4.3.3.1 Dosimetry.** The primary exposure was from external gamma radiation from thorium in the monazite sands.  
1033 Doses for the Kerala region were based on measured ambient exposure rates, factoring in estimates of average house  
1034 occupancy by age and sex determined for about 11 % of the cohort from interviews. Sodium iodide (NaI) detectors  
1035 were used to measure indoor and outdoor dose rates for 71,674 houses (~94 % of study residences) using the mean  
1036 of three dose measurements. The estimated dose rates were validated by TLDs and spot scintillometer measurements  
1037 made quarterly for 1 y in 800 homes selected at random. The annual dose, estimated from TLDs worn for two  
1038 months, and the personal doses, calculated by their algorithm for all subjects, showed a correlation of 0.80, but only  
1039 after 15 % of dosimeter measurements were discarded as “outliers” (Nair *et al.*, 2009). No uncertainties were  
1040 provided for the dose estimates.

1041  
1042 **4.3.3.2 Epidemiologic Methods, Findings and Issues.** The cancer incidence registry was based on records from a  
1043 cancer hospital, other hospitals and clinics, death certificates and visits to relatives to determine missing cause-of-  
1044 death information. Histopathology or cytology was available for 73 % of cases. There were 1,349 cancers other  
1045 than leukemia. The dose response analysis of all cancer except leukemia yielded an ERR estimate of  $-0.13$  (95 %  
1046 CI  $-0.58, 0.46, p > 0.5$ ) (Figure 4.8). The analysis was based on five dose categories and adjusted for sex, attained  
1047 age, follow-up interval, education, occupation and bidi (home-made cigarette) smoking, as needed. There was no  
1048 significant modification of the ERR by sex or age. When cancer cases were limited to those with pathological  
1049 verification, the estimated ERR  $\text{Gy}^{-1}$  became larger with wider 95 % CI (but data not shown). There was no  
1050 significant evidence of risk among those with 500 mGy or more. The risk estimate for leukemia was  
1051 noninformative, with large uncertainties.



1071 **Fig. 4.8.** Cancer risk by estimated cumulative external exposures in Kerala (adapted from Nair *et al.*, 2009), but  
1072 not shown there as a graph).

1074

1075 **4.3.3.3 Study Strengths and Weaknesses.** The HBRA study members had a fairly high cumulative dose (mean = 161  
1076 mGy), ranging to over 500 mGy. The investigators used various sources to try to ascertain cancers. The population  
1077 was more stable than in many regions of the world, so migration did not greatly alter the estimated doses or disease  
1078 ascertainment. Information was available on bidi smoking and other lifestyle and sociodemographic factors.

1079

1080 Over the course of the study, investigators made a considerable effort to characterize exposure levels.  
1081 Nevertheless, they had to rely primarily on ambient air sampling and aggregated reports on age- and sex-specific  
1082 house occupancy factors. They obtained personal dosimeters worn by 160 individuals for two months to compare  
1083 with estimated doses calculated by their usual methods, but the results are difficult to interpret.

1084

1085 The study period was only 15 y and the number of cancers therefore was fairly small. A possible problem is  
1086 that high cumulative doses occur in coastal areas where monazite sands are found, which raises the question of  
1087 whether background cancer rates in coastal (fishing) communities are the same as in inland (farming)  
1088 communities. Cancer incidence rates increased as the 2.4 power of attained age; this is somewhat lower than for  
1089 most studies, suggesting the possibility that there may have been cancer under-ascertainment among older people,  
1090 when many cancers occur. The investigators commented, "...the accuracy of information on cause of death is  
1091 questionable, particularly among poor people.... One may also suspect that the social structure and medical care  
1092 system in India might have made it difficult to obtain accurate information on cancer diagnosis" (Nair *et al.*,  
1093 2009). International Agency for Research on Cancer (Nair *et al.*, 2002) reported that the high percentages of  
1094 cancer cases indicated as an unspecified or ill-defined site, or detected through the death certificate only,  
1095 suggested the likelihood of cancer under-ascertainment. This raises the question of whether a non-negligible  
1096 fraction of the poor people may have received little or no medical attention for their cancers, so that cancer was  
1097 under-diagnosed in this study, and whether factors like individuals' distance from the sole cancer hospital might  
1098 affect the degree of cancer under-diagnosis, and might even bias the results if dose levels are correlated with  
1099 distance.

1100

1101 **4.3.3.4 Implications for the LNT Model and Radiological Protection.** This study is nominally more supportive of a  
1102 dose-effect threshold than of the LNT model. However, the uncertainties in dosimetry, the weaknesses in cancer  
1103 ascertainment and the wide confidence interval on the risk estimate mean it is inconclusive and needs to be  
1104 interpreted with caution.

1105

1106 **4.3.4** *Yangjiang, China HBRA Study*

1107

1108 Tao *et al.*(2012) reported an update on residents of a HBRA in the Yangjiang region of China for the period of  
1109 1979 to 1998. Most study families had resided there for six or more generations.

1110

1111 **4.3.4.1** *Dosimetry.* Tao *et al.* (2000) divided the region into three dose groups (high, medium, low) on the basis of  
1112 environmental dose rates per year. Sun *et al.* (2000) estimated individual annual doses and cumulative doses based  
1113 on location-specific exposure data for each hamlet and distinguished between outdoor environment in the hamlet or  
1114 on farmland. Arithmetic mean doses were calculated for public places and for farmland in each hamlet. Indoor  
1115 doses were based on sampling one-third of houses in each hamlet. Indoor doses differed somewhat from house to  
1116 house, being highly dependent on building materials and room size. A distinction was made between bedroom dose  
1117 and dose in other indoor places. Sex- and age-specific occupancy factors were used to represent the time spent in  
1118 bed, at other indoor places, and outside in public places or farmland. The occupancy factors were obtained from a  
1119 survey of 5,291 persons. Individual dose estimates calculated using the indirect method were compared to direct  
1120 measurements for 5,204 individuals that were obtained using electronic pocket dosimeters or TLDs and a good  
1121 correlation was reported (Morishima *et al.*, 2000). No uncertainties were calculated in the dosimetry. For the HBRA  
1122 areas and the control areas the estimated mean cumulative colon doses, lagged 10 y for solid cancer, were 84.8  
1123 mGy and 21.6 mGy, respectively, a difference of about 63 mGy.

1124

1125 **4.3.4.2** *Epidemiologic Methods, Findings and Issues.* The mortality experience of the approximately 31,000 cohort  
1126 members in the Yangjiang area of China was determined for ages 30 to 74 y based on periodic staff visits to area  
1127 hospitals, local village doctors, and family members (Tao *et al.*, 2012). Investigators ascertained 941 deaths due to  
1128 cancer excluding leukemia. Diagnoses were based on pathological determination for only 26 % of cancer deaths  
1129 and radiographic or ultrasound information for 62 %. Analyses used six dose categories and adjusted for sex,  
1130 attained age and follow-up interval. The estimated ERR Gy<sup>-1</sup> was -1.01 (95 % CI -2.53, 0.95). However, 29 % of  
1131 those deaths were coded as due to liver cancer, for which a strong negative risk estimate (ERR = -3.38 Gy<sup>-1</sup>) was  
1132 seen. Since that might be associated with area differences in other risk factors, *e.g.*, different prevalence of  
1133 hepatitis B infections, calculations were performed excluding both liver cancer and leukemia, which yielded an  
1134 ERR of 0.19 Gy<sup>-1</sup> (95 % CI -1.87, 3.04). Prior reports had compared socioeconomic and lifestyle factors between  
1135 the high- and low-background areas and found little difference. However, a survey of the frequency of having  
1136 received an x-ray examination showed about 30 % more in the low background area than in the HBRA (Tao *et al.*,  
1137 2000), suggesting a differential in medical care as a possible source of bias. Furthermore, the rates of infectious  
1138 disease, particularly tuberculosis, and external causes of death differed between the low- and high-background  
1139 areas, again raising a question of the comparability of the low- and high-background areas.

1140

1141 **4.3.4.3** *Limitations of the Kerala and Yangjiang HBRA Studies.* Dosimetry for the high natural background study  
1142 areas is limited by the lack of personal dose information among the study populations. Although doses were based  
1143 on measurement data within the regions, large uncertainties were likely within the study cohorts and little attempt  
1144 was made to determine these uncertainties. Based on the dosimetry, the use of these studies is limited in furthering  
1145 our understanding of LNT.

1146

1147 Since these studies compared persons residing in different geographical areas with relatively little migration,  
1148 the area basically defines the dose. This leads to questions about the comparability of the residents in different  
1149 areas regarding occupations, lifestyle and dietary habits, access to medical care, disease experience, and other  
1150 factors that might confound or otherwise modify cancer risk. These uncertainties may impact the risk estimates  
1151 and require caution in their interpretation.

1152

1153 **4.3.4.4** *Implications of the HBRA Studies for the LNT Model and Radiation Protection.* The failure of the HBRA  
1154 studies to find a risk for all cancer except leukemia is not supportive of the LNT model. However, the risk  
1155 estimates are largely statistically consistent with the LSS risk estimate owing to their broad confidence bounds. As  
1156 described above, the studies have large dose uncertainties and weaknesses in cancer ascertainment which may have  
1157 attenuated the risk estimates. Furthermore, the fact that much of the dose variation is attributable to geographic  
1158 locations, which may be associated with risk factors other than radiation level, introduces ambiguity into the  
1159 inference regarding radiation effects. Therefore, the studies are inconclusive regarding the LNT model, but at least  
1160 would tend to favor the use of a DREF greater than one in formulating radiologic protection guidelines.

1161

1162 **4.3.5** *Taiwan Residents of Radiation-Contaminated Buildings*

1163

1164 Between 1982 and 1984 Taiwanese manufacturers of building construction materials incorporated radioactive  
1165 <sup>60</sup>Co into rebar reinforcing rods. The contaminated rods were used in over 180 buildings in Taiwan, which  
1166 included more than 1670 apartments/houses and certain schools and industrial facilities and exposed about 10,000  
1167 individuals. The national government sponsored a program to conduct a detailed dosimetric survey of interior  
1168 exposure levels in those facilities (Chen, 2002).

1169

1170 **4.3.5.1** *Dosimetry.* A dosimetry survey was conducted from 1992 to about 2001 to estimate doses to residents of the  
1171 contaminated units (Chen, 2002). They surveyed 1,607 units in 181 contaminated buildings. They measured the  
1172 interior and exterior dose equivalent rate of each building. Based on the physical arrangement of contaminated  
1173 reinforcement bars they measured the dose equivalent rates at the surface of the contaminated locations and at

1174 likely person locations within the units. They measured the living room, bedroom, study room, dining room, bath  
1175 and kitchen. Interviews were conducted and records of the daily living activities of the residents were collected to  
1176 reconstruct the individual dose equivalent for each resident. Total cumulative doses were estimated for individuals  
1177 based on the ambient exposure measurements, the time and duration of residence in the contaminated apartments  
1178 and estimated occupancy-rate information. It was difficult for residents to provide explicit information on their  
1179 daily activities 10 to 20 y in the past. Survey information and the use of assumed room occupancy factors (50 %  
1180 living room; 33 % bedroom; 17 % others) contributed significantly to uncertainties in the individual dose  
1181 reconstructions.

1182  
1183 The estimated dose distribution was low. The mean cumulative dose was estimated to be 47.8 mGy (median  
1184 6.3 mGy, range <1 to 2,363 mGy), and the dose distribution was highly skewed, with only 15 % of the study  
1185 subjects exposed to more than 50 mGy (Hwang *et al.*, 2006). In general, the dosimetry effort was thorough and well  
1186 conducted, and the study authors appear to have done a rigorous job in reconstructing exposures. Nevertheless, the  
1187 individual estimates have considerable uncertainties due to faulty or nonspecific recall of daily activities 10 to 20 y  
1188 previously, and rough extrapolations of doses at various locations within their home. Based on those dosimetry  
1189 weaknesses, the study is limited in its use to evaluate dose response.

1190  
1191 **4.3.5.2 Epidemiologic Methods, Findings and Issues.** Interviews with study subjects also obtained information on  
1192 medical history, occupation, education, and detailed exposure history including dates moved in and out and  
1193 lifestyles in these buildings, and exposures to occupational and medical irradiation. The cumulative dose estimates  
1194 have been used to perform dose-response analyses of various health endpoints. Cancer incidence has been  
1195 ascertained from the national cancer registry (Hwang *et al.*, 2008), and various sub-studies of smaller groups have  
1196 evaluated lens opacities, chromosome abnormalities, hematological and immunological parameters, and  
1197 reproductive effects (Chang *et al.*, 1999a; 1999b; Hsieh *et al.*, 2002; 2010; Lin *et al.*, 2010). Various papers  
1198 presented categorical comparisons (*e.g.*, <50 mGy vs. ≥50 mGy) and most presented dose-response analyses using  
1199 Cox regression, or other statistics for continuous outcomes.

1200  
1201 In the latest report on cancer incidence the mean length of follow-up was 19 y (Hwang *et al.*, 2008). The Cox  
1202 regression analyses adjusted for sex, birth cohort and years since first exposure, using dose as a continuous variable.  
1203 Doses were not lagged other than allowing a minimum latency period after first exposure. The mean age was 17 y  
1204 (range was from intrauterine to 87 y) at first exposure and 36 y at follow-up. There was no indication of a significant  
1205 increase in risk for all solid cancers ( $p = 0.5$ ; hazard ratio (HR) at 100 mGy = 1.03 (90 % CI 0.96, 1.09,  $n = 106$ ;  
1206  $ERR\ Gy^{-1} \approx 0.3$ , 90 % CI -0.4, 0.9). They reported a significant dose response for non-CLL leukemia, ( $ERR\ Gy^{-1} \approx$

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1207 1.9, 90 % CI 1, 3.1,  $P = 0.08$ ,  $n = 6$ ) and a marginally significant dose response for breast cancer (ERR Gy<sup>-1</sup>  $\approx$  1.2,  
1208 90 % CI -0.1, 2.1), 1.21,  $P = 0.13$ ,  $n = 17$ ). There was no indication of risk for thyroid, lung or stomach cancers.

1209  
1210 **4.3.5.3 Study Strengths and Weaknesses.** The dosimetry effort to reconstruct exposures was thorough, albeit the  
1211 estimated doses were based on ambient exposure measurements and imputed occupancy factors. The ascertainment  
1212 of cancer incidence was of reasonably good quality. The investigators determined basic phenotypic data  
1213 (chromosome aberrations, blood cell counts) as well as overt health outcomes. However, the individual exposure  
1214 estimates have considerable uncertainties due to faulty or nonspecific recall of daily activities 10 to 20 y  
1215 previously, and necessary extrapolations of exposures at various locations within the home. The sample size was  
1216 small and the dose distribution was low, so one would not expect much statistical power, yet the authors have  
1217 reported a number of “positive” effects, some of them questionable. For example: they reported a narrow  
1218 confidence interval for leukemia (ERR Gy<sup>-1</sup>  $\approx$  1.9, 90 % CI 0.1, 3.1) based on only six cases (Hwang *et al.*, 2008).  
1219 Also, for eosinophils, they reported  $p = 0.03$  based on a mean ( $\pm$  SD) of 0.22 (0.17) in the exposed group and 0.21  
1220 (0.19) in the unexposed group.

1221  
1222 **4.3.5.4 Implications for the LNT Model and Radiation Protection.** The investigators did not compare a linear model  
1223 with quadratic or threshold models. The uncertainties associated with the dose reconstruction, small sample size  
1224 and other features mean the implications regarding dose response for radiation protection purposes cannot be  
1225 determined reliably .

1226  
1227 **4.3.6 Radiation Fallout Studies**  
1228

### Highlights

Fallout studies have been conducted in conjunction with the Japanese atomic-bomb survivors, nuclear bomb testing, nuclear production facilities and nuclear power plants (NPPs). Kim *et al.* (2016) reviewed, and conducted a meta-analysis of studies of thyroid cancer risk and residence near NPPs. Overall, they concluded that the evidence does not support an association, but noted some interpretive uncertainties. A weakness of most fallout studies has been the large uncertainties regarding fallout exposures: no individual dose information, or dose reconstructions with limited availability of individual measurements to validate the estimates, and small numbers of observed leukemias or thyroid cancers. Other issues with some radiation fallout studies have been the use of only geographic comparisons of populations (who may differ in other disease risk factors), exposure-related variations in intensity of disease surveillance, possible biases associated with self-reports of exposure-related behaviors (*e.g.*, fallout-contaminated milk consumption), and lack of information on lifestyle or other potential confounding variables. We conclude that, though a variety of fallout studies may have been needed to address public health concerns, the

studies have contributed little to our scientific understanding of the LNT model.

Epidemiologic studies of possible health effects from fallout episodes have included a brief summary of studies of health effects associated with several of the major fallout incidents follows.

1229  
1230 **4.3.6.1 Japanese Atomic-Bomb Fallout.** The Atomic Bomb Casualty Commission (ABCC) conducted interviews of  
1231 LSS members between 1951 and 1961 which included several questions about bomb fallout exposures from rain.  
1232 The results were published because of recent public concerns about “black rain” exposure (Sakata *et al.*, 2014). To  
1233 avoid outcome-dependent biases (*i.e.*, recall of exposure after a health outcome has already occurred), the primary  
1234 data analyses excluded health events prior to 1962. For 1962 to 2005 there was no significant association of  
1235 reported “black rain” exposure with deaths due to leukemia, all solid cancer or all causes for either Hiroshima or  
1236 Nagasaki. Similarly, there were no associations for the incidence of solid cancer or leukemia. On the other hand,  
1237 thyroid examinations by Nagataki *et al.* (1989) suggested an excess prevalence of thyroid nodules in the Nishiyama  
1238 district of Nagasaki, where there was a known radioactive rainout event, though the excess was based on small  
1239 numbers (9/184 exposed and 3/368 unexposed).

1240  
1241 Another group of investigators from Hiroshima universities analyzed data based on a cohort of atomic-bomb  
1242 survivors loosely defined in 1970, and reported positive associations of fallout with mortality through ad hoc  
1243 statistical modeling techniques (Tonda *et al.*, 2012) or use of reported fallout exposure information that was  
1244 obtained in 2008 (and therefore subject to outcome-dependent biases) (Otani *et al.*, 2012).

1245  
1246 Limitations of the atomic-bomb fallout data include: large uncertainties regarding fallout exposures and  
1247 considerable missing individual data; no indication of whether study members were protected from the rainout; no  
1248 data on where individual “black rain” exposures occurred (as much of the black rain, from carbonaceous fires,  
1249 contained little radioactivity); 6 to 16 y recall of a possible event that occurred when they were under heavy stress.  
1250 There were further limitations of the studies by the Hiroshima universities group because of the long lapse of time  
1251 before the cohort was formed (survival and migration bias) and the potential for outcome-dependent bias. Since  
1252 information on individual fallout exposures was very limited, the studies are not useful for evaluating the LNT  
1253 model.

1254  
1255 **4.3.6.2 Marshall Islands Atomic Testing Fallout.** Residents of the Marshall Islands were exposed to radioiodine as  
1256 a result of testing of nuclear weapons by the United States during 1946 to 1958, and particularly so by inadvertent  
1257 exposure from the Castle Bravo thermonuclear test explosion in 1954 that led to assessed thyroid doses of around  
1258 20 Gy for young children living on Rongelap Island (Simon *et al.*, 2010). Studies of small groups of highly

1259 exposed residents have been conducted over the years (Howard *et al.*, 1997; Robbins and Adams, 1989) but  
1260 provide little information on low-dose effects. As to lower dose studies, Hamilton *et al.* (1987) screened  
1261 approximately 7200 Marshall Islanders, including about 2300 who were residents of the northern Islands in 1954.  
1262 They found excess thyroid nodules among the residents of the northern islands and reported a linear distance-  
1263 response association based on distance from the BRAVO test. Takahashi *et al.* (1997) also found an excess of  
1264 thyroid nodules among residents of the northern islands in 1954. Limitations of both studies included no  
1265 individual doses and possible subject selection biases, so they provide little useful information for evaluating the  
1266 LNT model.

1267  
1268 **4.3.6.3 Nevada Test Site (NTS) Atomic Fallout in Utah.** The principal studies of potential effects of fallout from the  
1269 NTS in Nevada have centered on leukemia and thyroid cancer outcomes. A large case-control study (1,177 cases  
1270 and 5,330 controls) of leukemia mortality in Utah did not find a statistically significant overall association with  
1271 estimated bone marrow dose categories, though a significant excess was reported among children/adolescents  
1272 during 1952 to 1957 ( $p = 0.02$ ), especially for acute lymphocytic leukemia ( $p = 0.009$ ) (Stevens *et al.*, 1990). No  
1273 association was seen for those who were *in utero* during the fallout period, perhaps because of limited statistical  
1274 power. The median RBM dose was estimated as 3.2 mGy among all study subjects, and 19 mGy among residents of  
1275 the most highly exposed county.

1276  
1277 Two thyroid examinations in 1965 to 1966 and 1985 to 1986 were conducted of a cohort of school children in  
1278 parts of Utah, Nevada and Arizona (Rallison *et al.*, 1990). Some 2500 children received both examinations. The  
1279 maximum assessed individual thyroid dose was 1.4 Gy and the mean 0.12 Gy (Simon *et al.*, 2006a) in Utah based  
1280 on measurements of  $^{131}\text{I}$  deposition and interview recall of amount of milk and green vegetables consumed. A  
1281 detailed review of the study which revised the dosimetry and thyroid diagnoses (Lyon *et al.*, 2006) and the  
1282 application of a more sophisticated model for dose uncertainties found a high relative risk for total thyroid  
1283 neoplasms but the risk estimate was extremely wide (ERR =  $24.3 \text{ Gy}^{-1}$ , 95 % CI 3.9, 79;  $n = 20$ ) (Li *et al.*, 2007).  
1284 However, there was no clear indication of risk of thyroid cancer (ERR =  $0.8 \text{ Gy}^{-1}$ , CI < 0, 15;  $n = 8$ ) (Lyon *et al.*,  
1285 2006). Concerns about the study include possible outcome-dependent bias and having different examiners in the  
1286 various geographic areas that varied by dose which might affect the consistency of diagnostic sensitivity.

1287  
1288 The leukemia data were limited by reconstructed estimates of RBM doses and the small doses entailed. The  
1289 thyroid studies were limited by the potential for subject selection factors with a considerable loss to follow-up, large  
1290 thyroid dose uncertainties, limited blinding of screeners as to exposure, and the potential for outcome dependent  
1291 biases because some of the interviews that defined exposure were conducted after subjects' thyroid examination

1292 results were known. Because of wide uncertainties in individual doses and the potential for biases, the studies  
1293 provide little useful information for evaluating the LNT model.

1294  
1295 **4.3.6.4 Atomic Testing Fallout Across the United States.** Estimates of whole-body doses from external irradiation  
1296 by atomic fallout for the population of the United States have been reported as ~1.2 mGy. For U.S. children born in  
1297 1951 average thyroid fallout doses from <sup>131</sup>I were estimated to be ~32 mGy (Bouville *et al.*, 2002). Gilbert *et al.*  
1298 (2010) examined thyroid cancer rates in counties within the SEER cancer registries (SEER, 2001) in relation to  
1299 estimated <sup>131</sup>I exposure levels in those counties from NTS fallout. They found a suggestive but nonsignificant  
1300 exposure-thyroid cancer relationship for exposures before 1 y of age, but no indication of an association for  
1301 exposures at ages 1 to 15 y. Since this is an ecological study, the potential for ecologic biases means it should not  
1302 be used for evaluating the LNT model.

1303  
1304 **4.3.6.5 Semipalatinsk Fallout.** The more than 100 above-ground nuclear tests conducted by the Soviet Union in  
1305 1949 to 1962 at the Semipalatinsk nuclear test site (SNTS) in Kazakhstan delivered fallout to residents of  
1306 downwind villages. A nested case-control study of leukemia mortality ( $n = 22$  cases) and individual estimated  
1307 doses reported a significant relative risk of 1.9 for those with doses  $>2$  Gy (Abylkassimova *et al.*, 2000). A study  
1308 of solid cancers among ~19,000 residents around SNTS reported a high risk for all solid cancer (ERR Gy<sup>-1</sup> of  
1309 0.81) and for selected single cancer sites (Bauer *et al.*, 2005). However, the dose estimates used in the study are  
1310 now believed to be incorrect, so the dosimetry is being redone,<sup>4</sup> and the authors expressed concerns about  
1311 selection biases in the unexposed group. A recent report on radiation dose and cardiovascular diseases found that  
1312 all the differences observed were attributable to variation between the exposed and unexposed group and not to a  
1313 dose response within the exposed group, again suggesting selection biases (Grosche *et al.*, 2011).

1314  
1315 Land *et al.* (2015) investigated thyroid disease among almost 2400 people less than 21 y of age who were  
1316 resident downwind of the SNTS during 1949 to 1962. The prevalence of thyroid nodules was assessed by  
1317 ultrasound screening in 1998, and 35 cases of thyroid cancer were detected. Estimated thyroid doses ranged up to  
1318 several gray, with a mean of 100 to 200 mGy. Their methods for thyroid dose assessment were commendable  
1319 (Land *et al.*, 2015), but uncertainties were still substantial because of the limited data available. The dose  
1320 responses for thyroid nodules were difficult to interpret, with small numbers and estimated risks six times as  
1321 great for males as for females. The thyroid cancer results were positive, but not significantly so, for either males  
1322 or females.

1323

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<sup>4</sup> B Grosche, personal communication, 2015

1324 The incorrect dosimetry annuls the value of risk estimates of the cancer mortality studies regarding LNT,  
1325 while the sex disparity regarding thyroid nodularity and nonsignificant risks for thyroid cancer suggest caution  
1326 in the interpretation of thyroid results.

1327  
1328 **4.3.6.6 Hanford <sup>131</sup>I Fallout.** From 1944 to 1957, the Hanford Nuclear Site in Washington State reprocessed  
1329 nuclear fuel that had been stored for a comparatively short time, so that substantial quantities of <sup>131</sup>I (27 PBq)  
1330 were released to the atmosphere during this period; extensive efforts were made to reconstruct the magnitude and  
1331 geographic distribution of those releases (Napier, 2002). Concerns were raised about cancer risk in downwind  
1332 populations. A long-term ecological study gave no indication of excess risk of mortality from all cancer, thyroid  
1333 cancer, breast cancer, non-CLL leukemia or childhood leukemia during the years 1950 to 2000 (Boice *et al.*,  
1334 2006c).

1335  
1336 An extensive <sup>131</sup>I exposure reconstruction was conducted (Napier, 2002). During 1992 to 1997 Davis *et al.*  
1337 (2004) conducted a historical cohort study of thyroid disease among almost 3500 people who had potentially been  
1338 born during 1940 to 1946 in seven counties near the site and provided sufficient interview information on locations  
1339 and habitual behaviors to estimate individual doses (mean dose of 174 mGy; median 97 mGy, maximum 2823  
1340 mGy). The doses were around an order of magnitude greater than those received locally from the Windscale Fire,  
1341 but much less than the doses received after Chernobyl in heavily contaminated areas of the former USSR. The study  
1342 participants were examined for thyroid abnormalities by ultrasonography and palpation by two separate thyroid  
1343 specialists, blinded as to dose, plus laboratory functional thyroid tests. Based on 19 diagnosed thyroid cancer cases,  
1344 the ERR Gy<sup>-1</sup> was not statistically significant ( $p = 0.25$ ), nor was it significant for all detected thyroid nodules  
1345 ( $p = 0.65$ ) or any other thyroid disease.

1346  
1347 The authors noted that the study had sufficient statistical power to detect the magnitude of effects that had  
1348 been reported elsewhere following exposure to <sup>131</sup>I. However, another group re-examined the dose assessment and  
1349 other study aspects and concluded that the study results should be interpreted as inconclusive, rather than as  
1350 evidence for little or no disease risk, because the uncertainties were considerable (Hoffman *et al.*, 2007).

1351  
1352 The Hanford thyroid cancer screening study is methodologically one of the best in the literature, although the  
1353 ultrasound equipment was not as sensitive as that available today. The substantial uncertainties in reconstructing  
1354 individual thyroid doses plus the inability to examine a third of the potential study subjects make the null findings  
1355 less certain, so the study can make only a limited contribution to the question of LNT.

1356

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1357 **4.3.6.7 Mayak Fallout.** The Mayak nuclear complex in the Southern Urals of the Russian Federation commenced  
1358 reactor operations in 1948 and started reprocessing irradiated nuclear fuel in 1949. During 1948 to 1972 about 38  
1359 PBq of  $^{131}\text{I}$  was discharged to the atmosphere from Mayak (Eslinger *et al.*, 2014). A child born in 1947 in the  
1360 nearby closed city of Ozyorsk and living there until 1972 is estimated to have received a cumulative thyroid dose  
1361 during this period of 2.28 Gy. For young children 5 y of age living in Ozyorsk, the maximum annual thyroid dose  
1362 was approaching 1.0 Gy in 1949, with annual doses decreasing to around 10 mGy by the late-1950s.  
1363 Koshurnikova *et al.* (2012) studied thyroid cancer incidence during 1948 to 2009 in Ozyorsk and the neighboring  
1364 city of Kyshtym and compared rates based on registries in these cities with those derived from incidence data for  
1365 the regional center of Chelyabinsk during 1993 to 2006. They reported that thyroid cancer incidence rates in  
1366 Ozyorsk and Kyshtym were 50 % higher than the rate in Chelyabinsk, although details of the Chelyabinsk data  
1367 were not given. A thyroid screening study of 581 Ozyorsk residents born in 1952 to 1953 compared to 313 who  
1368 moved to Ozyorsk after 1967 showed a RR for thyroid nodularity of 1.4 (95 % CI 1.0, 1.9), but doses were not  
1369 estimated (Mushkacheva *et al.*, 2006).

1370  
1371 **4.3.6.8 Three Mile Island (TMI) Fallout.** About 550 GBq of  $^{131}\text{I}$  were released to the atmosphere during the  
1372 TMI reactor accident, leading to an estimated maximum individual thyroid dose of <0.2 mGy (Clarke, 1989).  
1373 The mean dose to the TMI population living within 5 miles of TMI was estimated to be ~0.1 mGy, with ~13 %  
1374 exposed to >0.2 mGy (Gur *et al.*, 1983). Levin *et al.* (2013) reported an elevated incidence of thyroid cancer in  
1375 the vicinity of the TMI nuclear plant (Levin *et al.*, 2013), but the ecological study is difficult to interpret,  
1376 particularly given that the natural incidence of thyroid cancer appears to be high in Pennsylvania (Bann *et al.*,  
1377 2014).

1378  
1379 Hatch and Susser (1990) examined rates of childhood leukemia and all childhood cancers near the TMI nuclear  
1380 plant for 6 y after the accident in 1979. Based on gamma radiation levels of areas within 10 miles of the plant, they  
1381 reported post-accident associations for both endpoints, but further analysis suggested that psychological stresses or  
1382 selection biases due to out-migration may have been alternate explanations (Hatch *et al.*, 1991). This was later  
1383 debated by Wing *et al.* (1997) and Hatch *et al.* (1997). After a 17 y follow-up of a defined cohort of 32,000, no clear  
1384 association with estimated radiation doses was seen, though there was a nonsignificant suggestion of associations  
1385 with breast and hematopoietic cancers (Talbot *et al.*, 2003). Because of small numbers, the ecological nature of the  
1386 earlier analysis, and large uncertainties, the TMI studies clearly do not contribute to understanding the LNT model.

1387  
1388 **4.3.6.9 Fukushima Dai-ichi Fallout.** Substantial quantities of  $^{131}\text{I}$  and other radionuclides were released during the  
1389 Fukushima Daiichi nuclear power plant accident in Japan in 2011, although the largest consequent thyroid doses  
1390 were considerably lower than those received locally after the Chernobyl accident (UNSCEAR, 2014). To date, no

1391 studies of health outcomes in relation to individual exposure levels from the Fukushima nuclear accident have  
1392 been published. However, about 100 histologically-diagnosed thyroid cancers were identified by the Fukushima  
1393 Health Management Survey (FHMS) in the first 3 y after the accident. Tsuda *et al.* (2015) advanced the hypothesis  
1394 that this represents a significant 30- to 50-fold excess of thyroid cancer among those exposed before age 18 in  
1395 Fukushima. This interpretation is based on their ecological study in which the primary comparison is between the  
1396 FHMS cohort with ultrasensitive ultrasound thyroid screening and corresponding all-Japan rates in which few  
1397 children have had thyroid screening. This represents a patent bias. However, they found no differences in thyroid  
1398 screening results between children in districts of Fukushima who had the highest <sup>131</sup>I exposures vs. little exposure  
1399 (odds ratio = 1.08, 95 % CI 0.60, 1.96). A recent report of the complete results of the first FHMS screening  
1400 showed suspected/confirmed (from fine needle aspiration cytology) thyroid cancer rates of 33 and 35 per 100,000  
1401 in the areas with the highest and lowest levels of exposure, respectively (Suzuki, 2016). Suzuki *et al.* (2016) also  
1402 noted there is a striking dissimilarity between the thyroid cancer cases around Chernobyl and in Fukushima  
1403 Prefecture, in that the youngest case in Fukushima was aged 6 y at the time of the accident whereas around  
1404 Chernobyl the excess cases occurred markedly among exposed infants and young children. Wakeford *et al.*  
1405 (2016) have summarized a variety of other reasons why the Tsuda *et al.* (2015) interpretation does not fit the  
1406 biology and epidemiology of radiogenic thyroid cancer. Hence, the Fukushima Dai-ichi studies to date do not  
1407 contain information relevant to the LNT model.

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1409

1410

#### **4.4 Medical Exposure Studies**

##### **Highlights**

Studies of patients treated with lung collapse for tuberculosis (TB) in the 1930s to 1960s are one of the few medically-exposed populations that provide consistent evidence for dose response relationships relevant to the LNT model used in radiation protection. Patients on average would receive on the order of 100 chest fluoroscopies over several years. When the studies were published in the 1970s and 1980s the data from both Massachusetts and Canada were consistent with a straight line relationship between breast tissue dose and breast cancer, adding substantial weight to the judgment on the use of the LNT model for radiation protection. Findings for lung cancer were negative and suggested that different tissues in the body respond differently to highly fractionated exposures received over a period of years. New studies have now been published from the Massachusetts and Canadian TB-Fluoroscopy cohort on cardiovascular disease. Though the dosimetry for breast was of good quality, to date that for heart and the circulatory system has been weaker. Overall, there is no convincing evidence for an association between fractionated doses to the heart and heart disease, which argues against including heart as a detriment for radiation protection.

Evaluation of epidemiologic studies of diagnostic medical exposures relative to assessment of the LNT model is particularly challenging. Medical exposures are typically limited to a relatively small portion of the body and the results of the studies are subject to significant uncertainties including, but not limited to, historical exposure data, uncertain organ dosimetry, uncertainty relative to the effectiveness of low energy photons compared to higher energy radiations, confounding by predisposing conditions, presence of disease and potential effects of associated therapy. The most recent epidemiologic studies have involved populations who had CT scanning during childhood when risk might be higher because CT doses are relatively high and children may be more radiosensitive than adults. Leukemia development has been of particular interest since there is a short latent period and non-CLL leukemia is quite radiosensitive. However, the CT studies from the U.K. and Australia, as well as other reviewed medical exposure studies, do not have the consistency, design, dosimetry, coherence or strength to provide strong evidence regarding the validity of the LNT model in the 10 to 100 mGy dose range. Hopefully, other studies currently in progress may add information over the next decade. Studies of multiple fluoroscopic examinations for tuberculosis (TB) and pediatric computed tomographic (CT) examinations are considered here.

1411

1412 **4.4.1** *TB Fluoroscopy Studies*

1413

1414 In the TB fluoroscopy cohorts, patients on average would receive on the order of 100 chest fluoroscopies to  
1415 monitor lung collapse (pneumothorax) and the treatment would last for several years with a fluoroscopic  
1416 examination every two to three weeks. When the studies were published in the 1970s and 1980s the data from  
1417 both Massachusetts and Canada were consistent with a straight line relationship between breast tissue dose and  
1418 breast cancer, adding substantial weight to the judgment on the use of the LNT model for radiation protection.  
1419 Findings for lung cancer were negative and suggested that different tissues in the body respond differently to  
1420 highly fractionated exposures received over a period of years.

1421

1422 **4.4.1.1** *Dosimetry*. Several investigators have estimated the cumulative radiation dose absorbed by breast, lung,  
1423 RBM, and several other organs for the Massachusetts TB Fluoroscopy Study (Boice and Hoover, 1981; Boice *et*  
1424 *al.*, 1978). Based upon interviews with both the physicians who conducted the examinations and the patients  
1425 themselves, probabilistic assumptions were made about duration of a fluoroscopic examination and patient  
1426 orientation, and cumulative dose was based on the number of fluoroscopies, calendar year of exposure (for  
1427 machine filtration), sex, age at treatment (for estimation of breast size and tissue), and phantom studies of organ-  
1428 specific doses using era-specific machine exposure settings to the extent possible (Boice *et al.*, 1978; Davis *et al.*,  
1429 1989). Uncertainty in organ dose/R (air) was estimated by Sherman *et al.* (1978) for the Canadian TB Fluoroscopy  
1430 Study and varied widely. Phantom doses were validated by Monte-Carlo simulations (Sherman *et al.*, 1978).

1431

1432 **4.4.1.2** *Epidemiologic Methods, Findings and Issues*. Since the 1970s studies of tuberculosis patients who  
1433 received repeated chest x-ray fluoroscopies to monitor lung collapse (pneumothorax) have provided important  
1434 information relevant to the LNT hypothesis for both breast cancer and lung cancer (Boice, 1978; Boice and  
1435 Monson, 1977; Boice *et al.*, 1979; Howe, 1995; Howe and McLaughlin, 1996; Miller *et al.*, 1989). Linear dose-  
1436 response relationships for the incidence of breast cancer were observed in both the Massachusetts and Canadian  
1437 studies with dosimetry that was state-of-the-art at that time (Boice *et al.*, 1978; 1981; Miller *et al.*, 1989; Sherman  
1438 *et al.*, 1978), and the risk coefficients were similar to those for the atomic-bomb survivors study and the acute  
1439 postpartum mastitis study (Boice *et al.*, 1979; Preston *et al.*, 2002). The Massachusetts study adjusted for age at  
1440 exposure, attained age or time since exposure, calendar time and TB subcohort (Boice *et al.*, 1981). The Canadian  
1441 study adjusted for age at first exposure, time since exposure and province (Nova Scotia versus other provinces)  
1442 (Miller *et al.*, 1989).

1443  
1444 The picture for lung cancer was different in that there was no evidence for an association between radiation  
1445 dose and lung cancer mortality in either the Massachusetts or Canadian studies, despite cumulative organ doses  
1446 larger than received by the atomic-bomb survivors and numbers of patients larger than the number of bomb  
1447 survivors (Davis *et al.*, 1989; Howe, 1995). Interpretation of the absence of a lung cancer effect is uncertain; one  
1448 possibility is that tissues differ in their response to radiation when the dose is delivered in fractions over time.  
1449 The studies took into account smoking histories, lungs at risk and other potentially confounding factors.

1450  
1451 **4.4.1.3** *Summary Studies of TB Patients Receiving Repeated Chest Fluoroscopies for Lung Collapse*. A strength of  
1452 the TB fluoroscopy studies is that multiple realizations of dose were used to account for uncertainty, and the doses  
1453 spanned a large range, for instance, in the Massachusetts study 0 to 11.6 Gy. The studies took into account smoking  
1454 histories and other potentially confounding factors. However, no adjustment was made to increase the dose for  
1455 patients who had special procedures that entailed lengthy fluoroscopic examinations.

1456  
1457 **4.4.1.4** *Implications for the LNT Model and Radiation Protection*. The TB fluoroscopy studies provide strong  
1458 support for the LNT model for breast cancer, but the lack of a radiation association for lung cancer indicates that  
1459 risk after multiple small doses may differ by cancer site.

1460  
1461 **4.4.2** *Computed Tomography Scanning Studies*

1462  
1463 Several recent publications have reported increased risk of leukemia and cancer following exposure of children  
1464 to CT scans (Huang *et al.*, 2014; Krille *et al.*, 2015; Journy *et al.*, 2014; Pearce *et al.*, 2012; Mathews *et al.*, 2013).  
1465 The concept is that children may be more sensitive than adults to radiogenic tumor induction, so risks might be seen

1466 at low dose levels. A typical CT scan delivers a tissue absorbed dose to a “slice” of the patient’s body in the range  
1467 of about 10 to 30 mGy. There is very little scattered radiation to body parts that are not in the field of interest. Any  
1468 potential cancer risk is concentrated in the directly irradiated tissues. For multiple CT scans the tissue doses can be  
1469 just below the range where studies of atomic-bomb survivors have shown a statistically significantly increased risk.

1470  
1471 Thus, large and well-designed studies are needed to discern whether there is any radiation- related effect.  
1472 Information on organ doses from CT examinations in the 1980s and 1990s is sparse and may have been skewed  
1473 toward “best practices”. In addition to the text below, further systematic information is provided for the two  
1474 major studies in tables 4.1 to 4.4.

1475  
1476 **4.4.2.1 Dosimetry.** No individual dosimetry was collected and used in any of the cohorts. In the Pearce *et al.* (2012)  
1477 study in the U.K., brain and RBM doses were estimated by reviewing typical CT settings for young people based on  
1478 U.K.-wide surveys and then combining the information with hybrid computational human phantoms and Monte-  
1479 Carlo techniques. Similarly, Journy *et al.* (2015) based the dose estimates on CT protocols from 916 different  
1480 hospitals which they combined with computational human phantoms and Monte-Carlo methodology. Cumulative x-  
1481 ray dose was estimated as effective dose, red marrow and brain dose. The uncertainties in doses were probably large,  
1482 but no formal dose uncertainty analyses were conducted.

1483  
1484 The Mathews *et al.* (2013) study in Australia used primarily estimated effective doses rather than organ doses,  
1485 and then based it on the literature according to age, calendar period of CT exposure and site of irradiation. There  
1486 was no correction for gender and size, but a different mean dose was assumed for exposures before 2001 as  
1487 compared to those after 2001. For the time period before 2001 effective doses were estimated from the adult CT  
1488 literature and a scaling factor (information not provided) applied to derive pediatric effective doses. The database  
1489 they used to identify CT exposures did not include CT scans received after age 19 and was based on fewer than  
1490 half the Australian hospitals, so individuals may have had additional CT scans that were not in the database. For  
1491 these and other reasons Mathews *et al.* (2013) does not provide reliable evidence regarding LNT in low-dose  
1492 radiation protection.

1493  
1494 **4.4.2.2 Epidemiologic Methods, Findings and Issues**

1495  
1496 **U.K. Pediatric CT Study.** Pearce *et al.* (2012) reported a retrospective cohort study of 180,000 patients in Great  
1497 Britain who had CT scans while 21 y old or less. Cases of leukemia and brain tumors were identified through  
1498 linkage to a national cancer registry. These two cancers were of particular interest because of the known radiation  
1499 sensitivity of bone marrow to leukemia induction and because about two-thirds of the scans were head scans. They

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1500 evaluated sex, age at exposure, number of CT scans, and lagged exposures as covariables. They analyzed leukemia  
1501 incidence commencing 2 y after the first CT scan and brain tumors after 5 y. Results showed an ERR of  $36 \text{ Gy}^{-1}$   
1502 (95 % CI 5, 120,  $n = 74$ ) for combined leukemia and myelodysplastic syndrome (MDS), which they reported was  
1503 compatible with estimates from the atomic-bomb LSS. For leukemia without MDS, the ERR  $\text{Gy}^{-1}$  was 19 (95 % CI  
1504  $-12, 79, n = 65$ ). The ERR for leukemias without MDS was positive but not significant. The ERR for brain tumors  
1505 was  $23 \text{ Gy}^{-1}$  (95 % CI 10, 46,  $n = 135$ ), appreciably higher than that reported for the LSS. From their results, the  
1506 authors calculated that children with doses of 50 to 60 mGy could triple their risk of leukemia and brain tumors.

1507  
1508 **Australian Pediatric CT study.** Mathews *et al.* (2013) reported results of a cohort study from Australia. From a  
1509 cohort of about 11 million people who were <20 y old at the start of 1985 or born during 1985 to 2005, about  
1510 60,500 cases of cancer were identified. Using a Medicare (Public Hospital) database it was determined that about  
1511 680,000 persons had a CT scan while less than 20 y of age. Cancer incidence in the cohort with CT scans was  
1512 compared to the cohort who did not have CT scans with adjustment for sex, attained age and year of birth. Risk of  
1513 all types of cancer was examined based on effective dose and beginning only 1 y after the CT scan. The calculated  
1514 risks were much higher than in the LSS. For example, with their preferred 1 y lag period, after a brain scan they  
1515 reported an ERR  $\text{Gy}^{-1}$  of 35 (95 % CI 26, 42) for all cancers except brain, which is highly implausible, and 29  
1516 (95 % CI 23, 37) for brain, compared to an ERR  $\text{G}^{-1}$  of about 1.7 for brain tumors after exposure at ages 0 to 5 y in  
1517 the LSS (Preston *et al.*, 2008). Information was not available on alcohol consumption, smoking, sun exposure or  
1518 markers of cancer susceptibility.

1519  
1520 **4.4.2.3 Strengths and Limitations.** The above CT studies (Mathews *et al.*, 2013; Pearce *et al.*, 2012) included a very  
1521 large number of individuals and were designed to provide a direct epidemiologic assessment of low-dose radiation-  
1522 associated excess cancer risk. However, the studies were subject to major biases which lead to doubt concerning the  
1523 causal nature, or at least the magnitude, of the reported associations. Walsh *et al.* (2013) and Boice (2015b) have  
1524 critiqued these studies and raised serious concerns about whether the relationship is causal. Among other concerns,  
1525 Mathews *et al.* (2013) reported a significant increase in brain cancer risk after abdominal/pelvic and extremity CT  
1526 scans, for which the brain dose was certainly negligible. They reported no increase in breast cancer, a highly  
1527 radiogenic organ, but increases in Hodgkin lymphoma and malignant melanoma, sites that have not been associated  
1528 with radiation exposure in other studies. In addition, they reported an increased risk for brain tumors and all other  
1529 cancers in years one to four after exposure which was somewhat larger than that in later years. This implied that  
1530 there was no latent period and increased the suspicion of confounding by indication (CT examination because of  
1531 having a predisposing factor for cancer) and reverse causation (pre-existing but undetected malignancy). There were  
1532 also major unanswered questions about missed doses.

1533

1534 In summary, the CT studies do not provide individual doses, and collection of the scan data and associated  
1535 exposure information for individual patients was not carried out. Information on organ doses from CT  
1536 examinations in the 1980s and 1990s is sparse; the estimates available may have largely reflected “best practices”  
1537 rather than typical practices at the time. Thus, assigned organ doses are of questionable validity. The uncertainties  
1538 in the CT doses are assumed to be large but no formal uncertainty analysis has been reported. Based on the  
1539 dosimetry review, these studies are of limited use in the evaluation of LNT.

1540  
1541 **4.4.2.4 Analyses to Evaluate Biases in CT Studies.** Krille *et al.* (2011; 2015) found that there may be reverse  
1542 causation bias (pre-existing but undetected malignancy) and Journy *et al.* (2015) found evidence of confounding by  
1543 indication (CT examination because of having a predisposing factor for cancer), although these confounding factors  
1544 accounted only partially for the effects observed.

1545  
1546 Berrington de González *et al.* (2016) analyzed predisposing factors or probable prior malignancy as biasing  
1547 factors in the previously published U.K. study (Pearce *et al.*, 2012) of leukemia and brain tumors after pediatric CT  
1548 examinations. The study included 74 leukemia cases and 135 brain tumor cases. For leukemia, deletion of study  
1549 subjects with predisposing conditions or suspicion of prior cancer derived from comparable medical sources  
1550 (radiology databases and death certificates) for the CT scan and control subcohorts revealed only a small decline in  
1551 the ERR Gy<sup>-1</sup>, from 36 to 33 (95 % CI 4, 114). However, when the full radiologist and pathology reports were  
1552 included for the leukemia cases, additional pre-existing cancers were identified and the association was no longer  
1553 statistically significant (ERR Gy<sup>-1</sup> was 20, 95 % CI -11, 86). A parallel pattern was seen for brain tumors: data  
1554 from comparable sources reduced the ERR Gy<sup>-1</sup> from 23 to 16 (95 % CI 6, 37), but additional exclusions based on  
1555 the full radiologist and pathology reports for the brain tumor cases reduced the ERR further to 10 (95 % CI 2, 26).  
1556 It should be noted that radiologist reports were available for only 40 % of patients and pathology reports for ~65 %  
1557 of cases. Furthermore, the reporting of pre-existing conditions and diagnoses in physician records was anecdotal  
1558 rather than systematic, so it is likely that some fraction of relevant information was missing. The reported  
1559 reductions of risk estimates in the study due to predisposing conditions or pre-existing malignancy were about 45 %  
1560 for leukemia and 55 % for brain tumors, but had the medical reports all been available and with systematic  
1561 information, the reductions would likely have been larger.

1562  
1563 **4.4.2.5 Implications for the LNT Model and Radiation Protection.** The large British and Australian CT studies and  
1564 risk of cancer in children have a number of caveats. The major methodological reason to doubt the causal nature of  
1565 the associations, or at least their magnitude, is bias from the absence of information on why CT scans were  
1566 performed, incompleteness of dosimetry and lack of individual dosimetry and unknown number of repeat  
1567 examinations or examinations performed at institutions not in the database. The published studies do not provide

1568 strong evidence that can be used to assess the validity of LNT at doses in the 10 to 50 mGy range, although the  
1569 most recent studies are beginning to address these issues (Berrington de González *et al.*, 2016; Journy *et al.*, 2015;  
1570 Krille *et al.*, 2015). New studies currently underway such as the EPI-CT study in Europe (Bosch de Basea *et al.*,  
1571 2015; Thierry-Chef *et al.*, 2013) will provide additional data and statistical power but may still have limited  
1572 information on predisposing conditions, pre-existing cancer and individual doses. The new studies will permit  
1573 credible insights into LNT only if the dosimetry is improved and all the potential sources of bias and missing data  
1574 are carefully assessed (Walsh and Nekolla, 2015).

1575

1576

1577

#### 4.5 Childhood Exposure Studies

##### Highlights

An analysis of solid cancer incidence among the Japanese atomic-bomb survivors exposed prenatally or during childhood showed a clear dose response, but upward curvature ( $p = 0.09$ ) suggested that the dose-response slope may be shallower in the low-dose range. The LSS of atomic-bomb survivors indicated a large leukemia response after childhood exposure, but the overall model showed clear upward curvature such that risk per unit dose at low doses was less than at high doses (Hsu *et al.*, 2013). The data on postnatal diagnostic medical exposures and childhood leukemia risk are inconclusive (Wakeford, 2008). Studies of juvenile irradiation and breast cancer generally support a linear dose response (UNSCEAR, 2013). A recent pooled analysis of nine studies of childhood external irradiation and thyroid cancer showed a significant dose response at 0 to 100 mGy and no evidence of nonlinearity (Lubin *et al.*, 2017).

For the most part, the low dose data on children are sparse, the number of specific types of cancer is small and uncertainties are large enough that such studies do not provide strong information regarding the LNT model. In the case of thyroid cancer and breast cancer, the data provide substantial support for a linear dose-response model.

In general, persons exposed at younger ages are at greater risk than those exposed as adults. One obvious reason for this is that those exposed at young ages are likely to live longer and have more time to express risk and detriment. In addition, tissue sensitivity for a given absorbed dose may vary with age. If children are indeed more radiosensitive, then review of the relevant epidemiology may shed some light on whether LNT or other dose response relationships are appropriate.

This is a complex issue and has been reviewed extensively in NCRP Report No. 136 (NCRP, 2001) and an UNSCEAR (2013) report. Studies of radiation effects on children come from a number of sources including atomic-bomb survivors, accidental exposures and post radiation therapy. Several of the key studies that report

dose-response data or risks at low-doses or low dose rates for childhood exposure will be mentioned briefly here. Studies considered include childhood atomic-bomb survivors (Section 4.5.1), childhood leukemia studies (Section 4.5.2), thyroid cancer studies (Section 4.5.3), and breast cancer studies (Section 4.5.4).

1578

1579 **4.5.1** *Childhood Atomic-Bomb Survivors*

1580

1581 Preston *et al.* (2008) reported that among the ~15,000 members of the LSS cohort who were ages 0 to 5 y at  
1582 exposure there were about ~650 incident cancers by age 55 y, of which 87 were estimated to be attributable to  
1583 radiation. They estimated an ERR Gy<sup>-1</sup> at age 50 y of 1.7 (95 % CI 1.1, 2.5), which is higher than the estimated  
1584 ERR of ~0.5 at all exposure ages, but may partly reflect the fact that the denominator baseline risks are lower at  
1585 younger ages. The EAR per 10,000 person-years per gray for the childhood exposure group was 56 (95 % CI 36,  
1586 79), with an estimated 87 radiation-associated excess cancers. However, there was no evidence of excess risk  
1587 among those exposed to <0.2 Gy. The combined *in utero* and childhood exposure groups yielded some evidence of  
1588 a linear-quadratic model [ $p = 0.09$  for curvature, with an  $\alpha/\beta$  ratio = 1.0 (95 % CI -0.1, 212)].

1589

1590 The atomic-bomb survivor data indicate that for some tumor types (leukemia, thyroid, breast, brain), but not for  
1591 all (bladder, lung), the risk is higher for those exposed in childhood than in adulthood, while for others there is too  
1592 little data to draw conclusions or little radiation effect. An additional complication is that depending upon the model  
1593 used (*e.g.*, time since exposure vs. attained age) differing conclusions are sometimes reached regarding the effect of  
1594 age at exposure. A recent UNSCEAR (2013) report considers age effects in more detail.

1595

1596 **4.5.2** *Childhood Leukemia Studies*

1597

1598 Lundell and Holm (1996) studied over 14,000 children treated with radiation therapy for skin hemangiomas  
1599 using <sup>226</sup>Ra applicators or x rays. For those with a marrow dose < 10 to 100 mGy the relative risk (RR) was not  
1600 elevated (0.9). For those with marrow dose >100 mGy there was a nonsignificant RR of 1.7 (95 % CI 0.7 to 3.4).  
1601 Another study (Lindberg *et al.*, 1995) showed a nonsignificant excess with a likely marrow dose of about 100 mGy.  
1602 One recent study of childhood cancer risk after conventional radiographic examinations found no evidence of risk  
1603 for leukemia and lymphoma (standardized incidence ratio = 1.05, 95 % CI 0.74, 1.45) (Hammer *et al.*, 2011), while  
1604 another of diagnostic radiography in early infancy reported a suggestive, but nonsignificant, increase in leukemia  
1605 (odds ratio = 1.39, 95 % CI 0.87, 2.23) (Rajaraman *et al.*, 2011). Neither study had estimates of RBM doses.  
1606 Wakeford (2008) characterized the literature on childhood leukemia after postnatal diagnostic medical radiation  
1607 exposure as equivocal and conflicting.

1608

1609 Studies also have been undertaken to determine if variation within the typical range of background external  
1610 gamma radiation affects childhood leukemia risk, which would indicate a risk at very low doses for a sensitive  
1611 cancer endpoint compatible with a nonthreshold dose response. However, since childhood cancer is rare and the  
1612 exposure levels are low, the numbers of children observed would need to be very large to achieve reasonable  
1613 statistical power (Little *et al.*, 2010b) and other leukemia risk factors could subtly bias the results. To avoid  
1614 participation bias and the impracticability of contacting many families, several recent studies have relied on using  
1615 existing databases of measured areal gamma radiation levels to estimate doses at the geographic locations of  
1616 childhood leukemia cases and controls (or cases and matched cohorts).

1617  
1618 A British study (Kendall *et al.*, 2013) of ~9,000 childhood leukemia cases and ~12,000 matched controls  
1619 found an association with assessed gamma radiation doses, with an ERR Gy<sup>-1</sup> of 120 (95 % CI 30, 220), which is  
1620 compatible with the atomic-bomb LSS risk estimate of about 50 for early childhood leukemia (Hsu *et al.*, 2013).  
1621 On the other hand, no association of ambient gamma radiation levels and childhood leukemia was found in a  
1622 French study of ~9,000 cases (Demoury *et al.*, 2016), although in smaller studies of ~1,800 cases in Switzerland  
1623 (Spycher *et al.*, 2015) and 1,000 cases in Finland (Nikkilä *et al.*, 2016) positive associations were reported. There  
1624 is broad compatibility of studies of the risk of childhood leukemia and natural background gamma irradiation with  
1625 a LNT model, although there are inconsistencies, owing in large part to the limited applicability of the areal  
1626 radiation measurements, potential for uncontrolled confounding and limited statistical power of such studies due  
1627 to dose uncertainties and low doses.

1628  
1629 **4.5.3** *Thyroid Cancer Studies*

1630  
1631 Ron *et al.* (1989) studied over 10,800 Israeli children treated with radiation therapy for ringworm of the  
1632 scalp. The average thyroid dose was about 90 mGy. In the most recent follow-up of this cohort, for up to 54 y,  
1633 Sadetzki *et al.* (2006) found 103 thyroid cancers in the irradiated group and reported an ERR Gy<sup>-1</sup> of 20 (95 % CI  
1634 12, 32). However, the high risk estimate in the Israeli scalp ringworm study may have been confounded by  
1635 surveillance or other differences between the exposed and unexposed groups; when (Ron *et al.*, 1995) fitted the  
1636 dose response, accounting for the exposed-unexposed difference, the ERR Gy<sup>-1</sup> was 6.6 (95 % CI < 0, 347). A  
1637 linear dose response was reported for thyroid cancer ( $p = 0.75$  for curvature) (Sadetzki *et al.*, 2006).

1638  
1639 **4.5.3.1** *Pooled Studies of Thyroid Cancer after X- or Gamma-Irradiation.* Veiga *et al.* (2016) conducted a pooled  
1640 analysis of 12 studies with primarily child/adolescent external irradiation of the thyroid gland and with a wide  
1641 range of doses which represented essentially all available studies of external irradiation and thyroid cancer risk that  
1642 had dose-response risk information. A subsequent analysis of radiation risk in the ranges of 0 to 200 mGy and 0 to

1643 100 mGy was conducted to further examine thyroid cancer risk from low-dose radiation exposures (Lubin *et al.*,  
1644 2017). In addition to the text below, further systematic information is provided for the Lubin *et al.* study in  
1645 Tables 4.1 to 4.4.

1646  
1647 **Dosimetry.** One criterion for acceptance in the pooled studies of Lubin *et al.* (2017) and Viega *et al.* (2016) was  
1648 that quantitative estimates of dose were required. Individual doses for the medical study cohort (treatment of  
1649 childhood cancer and treatment of various benign diseases such as tinea capitis and enlarged thymus) were  
1650 estimated based on the number of irradiations, irradiation equipment, and analyses of phantom exposures. For the  
1651 Life Span Study (LSS) dose estimation was discussed earlier (Section 4.1.1). The doses were due primarily to X  
1652 ray exposure, but  $\gamma$  for the hemangioma radium needle study and  $\gamma +$  neutron for the atomic-bomb study cohort.  
1653 For x-ray procedures in early days, issues of distance from primary beam, degree of beam collimation, thyroid  
1654 shielding, etc. are principal sources of uncertainty. Other dosimetric uncertainties are due to a variety of factors,  
1655 including dose uncertainty in the phantom studies, random differences in sizes of children of a given age, random  
1656 movements by children during treatment and missing data for ages at subsequent exposure for those with multiple  
1657 treatments.

1658  
1659 The two most influential studies in the analysis, the LSS atomic-bomb survivors (Furukawa *et al.*, 2013) and  
1660 children irradiated for scalp ringworm (Sadetzki *et al.*, 2006), accounted for 80 % of the exposed thyroid cancer  
1661 cases. The LSS results were adjusted for dose measurement uncertainty, while the evaluations of the scalp  
1662 ringworm study indicated that the effect on dose response due to measurement uncertainty was fairly small (Lubin  
1663 *et al.*, 2004; Schaefer *et al.*, 2001). Therefore, the effect of dose uncertainty on the meta-analysis dose response  
1664 was probably minor for the low-dose part of the study, though there were additional uncertainties for the high-dose  
1665 childhood cancer survivor participants.

1666  
1667 **Epidemiologic Methods, Findings and Issues.** The original pooled study (Veiga *et al.*, 2016) included 9 cohort  
1668 studies of children who received radiation treatments for various benign disorders, the LSS atomic-bomb  
1669 survivors, and two case-control studies of patients who received radiotherapy for childhood cancer. However, only  
1670 9 studies were contributive to the analysis of individuals with a <200 mGy thyroid dose (Lubin *et al.*, 2017). The  
1671 nine studies relied on tumor registries and/or individual reports with medical verification to identify thyroid cancer  
1672 cases, and the rates of follow-up were good to excellent. The analyses by Poisson regression adjusted for sex, age  
1673 at exposure, attained age, calendar year period, plus sensitivity analyses of other variables, such as number of  
1674 treatments, indicator for exposed/unexposed group, Jewish ethnicity, LSS participation in the clinical examination  
1675 program etc. Although the study as a whole did not have analyses corrected for dose uncertainties, uncertainties  
1676 had been evaluated in the two studies with the highest weights in the analysis, the LSS (Section 4.1.1) and the

1677 Israel tinea capitis study (Lubin *et al.*, 2004) both had assessments of dose uncertainties. After accounting for  
1678 downward curvature at doses > 10 Gy, Veiga *et al.* (2016) reported an ERR Gy<sup>-1</sup> of 5.5 (95 % CI 3.9, 7.5) based  
1679 on 1,070 thyroid cancer cases in 5.3 million person-years of observation. When the data were limited to the  
1680 subjects with ≤100 mGy, including 184 thyroid cancers, there was a statistically significant linear trend ( $p < 0.01$ )  
1681 with no evidence of a departure from linearity ( $p = 0.36$ ) and an ERR Gy<sup>-1</sup> of 11.2 (95 % CI 4.8, 20).

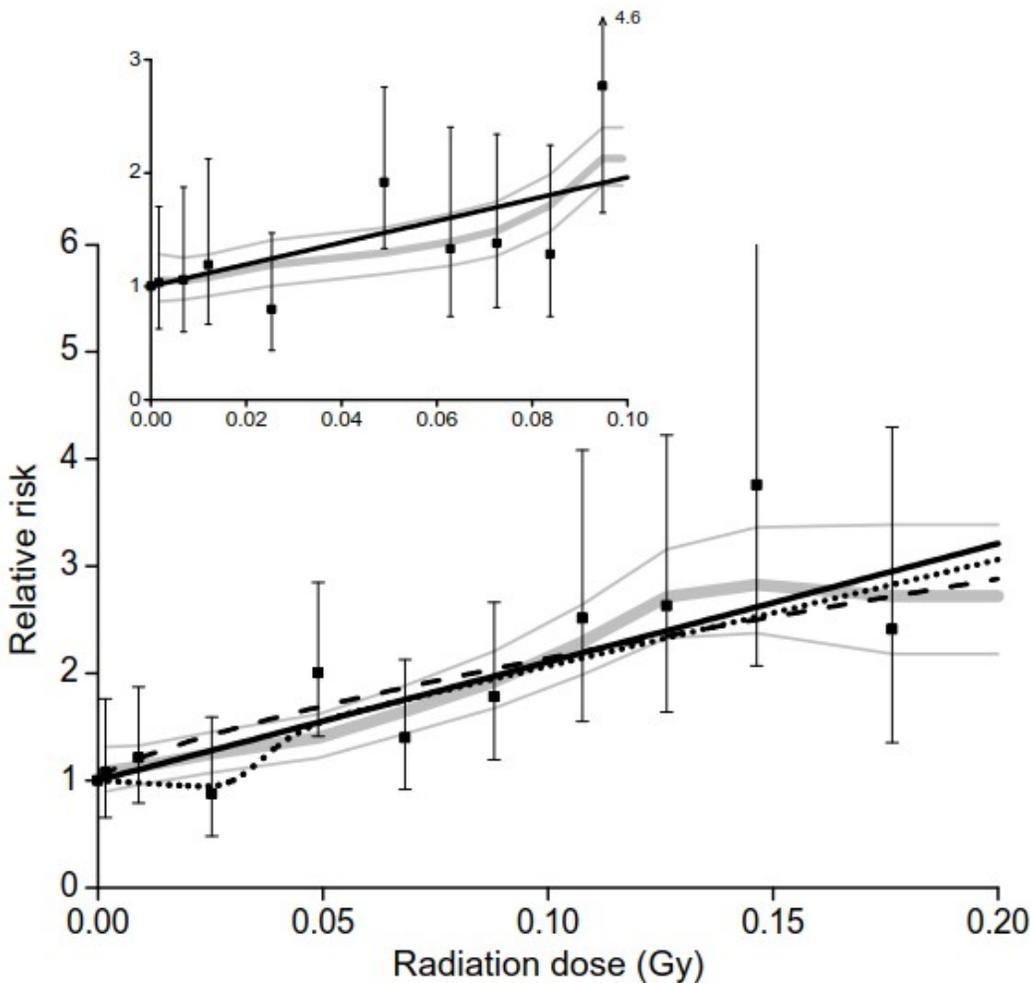
1682  
1683 The analysis of only the pooled data in the range of 0 to 200 mGy (Lubin *et al.*, 2017) included data from nine  
1684 studies. This analysis included 394 thyroid cancer cases among 107,594 individuals, with 252 cases among 61,155  
1685 individuals who received 1 to 200 mGy and 142 cases among 46,439 unirradiated individuals. For the range <200  
1686 mGy, RRs increased significantly with radiation dose ( $P < 0.01$ ), with an ERR Gy<sup>-1</sup> of 11.1 (95 % CI 6.6, 19.7). No  
1687 departure from linearity was evident ( $P = 0.77$ ) (Figure 4.9). A moving-average smoothing of the RRs (thick grey  
1688 line) showed consistency in the linear increase. A linear- quadratic model over the dose range 0 to 200 mGy  
1689 (dashed line) also showed near linearity. The results were similar for <100 mGy (inset panel), with no significant  
1690 departure from linearity ( $P = 0.66$ ) and a risk coefficient, ERR Gy<sup>-1</sup> of 9.6 (CI 3.7, 17.0). The best fitting model  
1691 was linear, compared to the quadratic or spline models. A dose-threshold analysis showed a maximum likelihood at  
1692 0 mGy and had a threshold upper bound of 40 mGy.

1693  
1694 ***Study strengths and weaknesses.*** The full dose-range (Veiga *et al.*, 2016) and low-dose (Lubin *et al.*, 2017)  
1695 meta-analyses included essentially all eligible dose-response studies of external irradiation and thyroid cancer.  
1696 At least for the two most statistically influential studies the impact of the dose measurement uncertainties was  
1697 small and the results were quite precise because of the large number of thyroid cancers. There might be some  
1698 possibility of cancer surveillance bias since head-and-neck irradiated populations tend to be screened more  
1699 often for thyroid cancer; however, sensitivity analyses reported in Lubin *et al.* (2017) suggested that there was  
1700 little or no surveillance bias in the data.

1701  
1702 ***Implications for the LNT Model and Radiation Protection.*** The analyses reported in the Lubin *et al.* (2017)  
1703 paper provide strong support for use of the LNT model. They indicate that, at least for the association of  
1704 radiation with thyroid cancer, there is a statistically significant dose-response over the restricted range of 0–  
1705 100 mGy that is compatible with linearity. This is strongly supportive of the use of the LNT model as prudent  
1706 for radiation protection.

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**Fig. 4.9.** Based on a pooled analysis of nine studies of external irradiation and thyroid cancer [Lubin *et al.*, 2017 (Fig 1)]. Shown are the models: dark solid line, linear; dashed line, linear-quadratic; thick gray solid line (with thin gray lines  $\pm 1$  standard deviation), smoothed nonparametric fit; dotted line, cubic spline.

1739 **4.5.3.2 Other Studies of Thyroid Cancer after Childhood Exposure.** Imaizumi *et al.* (2006; 2015) reported that a  
1740 long term study of the atomic-bomb survivors showed there was a significant linear dose-response relationship in  
1741 the prevalence of malignant thyroid tumors, benign thyroid nodules and cysts based on a clinical and ultrasound  
1742 examination. In addition there was significantly higher risk in those exposed at young ages. Three studies  
1743 included a total of more than 6,000 children who were administered known amounts of  $^{131}\text{I}$  for diagnostic  
1744 purposes (giving a mean thyroid dose of 1 Gy) but did not find any excess of thyroid cancer (Boice, 2005;  
1745 Dickman *et al.*, 2003; Hahn *et al.*, 2001; Hamilton *et al.*, 1989). However, the numbers of children exposed  
1746 while less than 10 y of age, who would be at most risk, were small.

1747

1748 In summary, children are more sensitive than adults to induction of thyroid cancer and nodules and  
1749 most studies show an approximately linear dose-response relationship.

1750

#### 1751 **4.5.4 Breast Cancer Studies**

1752

1753 The UNSCEAR 2013 report on risk from childhood radiation exposure presents a recent review of risk for  
1754 breast cancer according to age at exposure (UNSCEAR, 2013).

1755

1756 New dosimetric estimates for the breast have been calculated for the combined Swedish cohorts of young  
1757 children treated for hemangiomas and updated follow-up has been reported (Eidemüller *et al.*, 2015). An  
1758 analysis showed a linear dose response ( $p > 0.5$  for nonlinearity) below 5 Gy and an ERR risk estimate of 0.5  
1759  $\text{Gy}^{-1}$  (95 % CI 0.3, 0.7). A smaller French child hemangioma study (Haddy *et al.*, 2010) reported a nominal  
1760 increase in breast cancer risk with dose but it was not statistically significant.

1761

1762 An update of breast cancer risk over 57 y has been reported for 1,120 female infants irradiated for a presumed  
1763 enlarged thymus gland (Adams *et al.*, 2010). The median breast dose was 0.16 Gy (mean 0.72 Gy, range 0.02 to  
1764 6.2 Gy) and 96 incident breast cancers were found in the irradiated group. They found a better fit for a linear than  
1765 nonlinear dose-response function, with an ERR  $\text{Gy}^{-1}$  of 1.1 (95 % CI 0.6, 1.9).

1766

1767 A study of 3,010 females who received multiple fluoroscopic examinations for scoliosis (estimated mean  
1768 cumulative breast dose, 121 mGy) detected 78 breast cancers and reported an ERR  $\text{Gy}^{-1}$  of 2.9 (95 % CI -0.07,  
1769 8.6) (Ronckers *et al.*, 2008). No significant improvement in fit was seen for linear-quadratic, pure quadratic or  
1770 linear-exponential models. In the Taiwan study of individuals exposed in radiocontaminated dwellings at an  
1771 average age of 17 y, a marginal breast cancer risk was reported, ERR  $\text{Gy}^{-1}$  of 1.2 (90 % CI -0.1, 2.1,  $n = 17$ )

1772 (Hwang *et al.*, 2008). In summary, the studies of breast cancer after (primarily) juvenile irradiation broadly  
1773 support a linear model, although the small sample sizes of some studies limit their statistical power.

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1775

#### 4.6 *In Utero* Exposure Studies

1776

##### Highlights

In the atomic-bomb survivor studies, there were statistically significant dose-related increases in incidence rates of adult solid cancers among atomic-bomb survivors exposed to radiation *in utero*. Dose responses were estimated with a linear model; the data were too sparse to evaluate nonlinearity meaningfully. It appears that adult cancer risk from *in utero* exposure exists, but is not greater than that from early childhood exposure. There have been too few leukemia deaths (and data lacking on leukemia incidence during the first 4 y after the atomic bombing) to estimate the LSS radiation-related dose response for leukemia.

Although speculative, the absence of a radiation effect among the children of Mayak female workers exposed during pregnancy is also consistent with the possibility that chronic *in utero* exposures are less effective in causing solid cancers later in life than if the exposure had been received acutely, as was the case for the children of atomic-bomb survivors exposed *in utero*. However, there may be some indication of increased risks for leukemia by prenatal dose in the Mayak and Techa River cohorts.

In medicine, high doses to the embryo or fetus (*e.g.*, >0.5 Gy) increase the risk of cancer in offspring exposed *in utero*. However, there is disagreement over whether the reported risk of cancer after *in utero* exposure at a low dose such as <0.1 Gy is causal, and the available data have not provided definitive information with regard to LNT. Increased risks to the human embryo or fetus have not been observed at low doses for birth defects, growth retardation, or mental and neurobehavioral effects. NCRP Report No. 174 (NCRP, 2013), includes an in-depth review of radiation risks and potential outcomes. This section briefly reviews pregnancy risks from ionizing radiation, and effects from *in utero* exposures to occupational, environmental and medical sources.

The potential effects of ionizing radiation exposure *in utero* from medical procedures or other sources of radiation are of considerable importance and concern with regard to radiation protection. The NCRP previously evaluated prenatal effects in NCRP Report No. 54, Medical Radiation Exposure of Pregnant and Potentially Pregnant Women (NCRP, 1977). NCRP also published Commentary No. 9, Considerations Regarding the Unintended Radiation Exposure of the Embryo, Fetus or Nursing Child (NCRP, 1994), in which problems were discussed that could result from the medical administration of radioactive material to pregnant or nursing women. In 1997, NCRP held its Annual Meeting on The Effects of Pre- and Postconception Exposure to Radiation, and the proceedings were published in Teratology (Boice and Miller, 1999). These reports, commentaries, and previous

guidance documents were updated and expanded more recently in the comprehensive evaluation, NCRP Report No. 174, *Preconception and Prenatal Radiation Exposure: Health Effects and Protective Guidance* (NCRP, 2013) that included an in- depth review of radiation risks for potential outcomes.

1777

1778 **4.6.1** *Pregnancy Risks from Ionizing Radiation*

1779

1780 Any evaluations of prenatal radiation effects recognize the fact that pregnant women are faced with a baseline  
1781 risk to the embryo and fetus for reproductive and developmental problems. The normal background rate of adverse  
1782 pregnancy outcomes is considerable (NCRP, 2013). The background rate, absent radiation, for major congenital  
1783 malformations is ~3 %, with another 4 % of minor malformations. Pregnancy loss (spontaneous abortion,  
1784 miscarriage) in women who know they are pregnant occurs in 15 % of pregnancies with a wide standard deviation.

1785

1786 Increased human risks to the embryo or fetus have not been observed for birth defects, growth retardation,  
1787 neurobehavioral effects, mental retardation, decreased intelligence quotient (IQ), impaired school performance,  
1788 convulsive disorders, or embryonic or fetal death below a dose of 0.1 Gy (weighted uterine dose). Such effects are  
1789 generally attributed to the killing or functional disruption of critical cells during important stages of embryonic or  
1790 fetal development. All these effects are consistent with having a threshold dose below which there is no increased  
1791 risk. The data on IQ loss are somewhat difficult to interpret but even assuming that there might not be a true dose  
1792 threshold, any IQ effects at low doses would be so small as to be undetectable and therefore not of practical or  
1793 clinical significance.

1794

1795 NCRP Report No. 174 (NCRP, 2013) summarizes the health effects from ionizing radiation exposure of the  
1796 embryo or fetus during various gestational stages of pregnancy. That report also reviews cancer risk arising from  
1797 fetal exposure.

1798

1799 **4.6.2** *In Utero Exposures to Occupational or Environmental Sources*

1800

1801 The Japanese atomic-bomb study evaluated adult leukemia and cancer risks after *in utero* and early childhood  
1802 exposure (Preston *et al.*, 2008). They reported statistically significant dose- related increases in incidence rates of  
1803 solid cancers through 50 y of age after both *in utero* and childhood irradiation. Following *in utero* exposure the  
1804 ERR Gy<sup>-1</sup> was 1.0 (95 % CI 0.2 to 2.3) (weighted uterine dose), a risk not significantly less than that for survivors  
1805 exposed during the first 5 y of life, ERR Gy<sup>-1</sup> = 1.7 (95 % CI 1.1 to 2.5) (Preston *et al.*, 2008). The dose response  
1806 was estimated with a linear model; the data were too sparse to evaluate nonlinearity meaningfully. The temporal  
1807 patterns of the excess absolute rates which increased rapidly with age for early-childhood exposures did not appear

1808 to increase following *in utero* exposure, but the difference between the two was not statistically significant  
1809 ( $p = 0.14$ ). It can be concluded that adult cancer risk from *in utero* exposure exists but probably is not greater than  
1810 that from early childhood exposure. There have been too few leukemia deaths (and data lacking on leukemia  
1811 incidence during the first 4 y after the bombing) to estimate radiation-related dose response (Yoshimoto *et al.*,  
1812 1988).

1813  
1814 Because there are few prospective studies of sufficient numbers of children exposed *in utero* and followed into  
1815 adulthood, a recent cohort study of over 8,000 children of female workers at the Mayak Nuclear Facility is  
1816 noteworthy (Schonfeld *et al.*, 2012). The *in utero* gamma-ray doses were fairly small (mean 54.5 mGy; max ~800  
1817 mGy) and accumulated throughout the pregnancy, and thus not received briefly at one point in time. The  
1818 exposures were measured by film badges and not estimated by dose reconstruction; a notable strength and  
1819 uniqueness of the study. There was no evidence that prenatal gamma radiation received during pregnancy  
1820 increased the risk of solid cancer or leukemia mortality for up to 60 y of follow-up. The small number of cancers,  
1821 however, limited the precision of the study, and thus the negative findings were consistent in a statistical sense  
1822 with the positive observations of 2,452 children born to Japanese atomic-bomb survivors who received  
1823 comparable doses *in utero* (Preston *et al.*, 2008). Although speculative, the absence of a radiation effect among the  
1824 offspring of Mayak female workers is also consistent with the possibility that chronic exposures during pregnancy  
1825 are less effective in causing cancer later in life than if the exposure had been received acutely, as was the case for  
1826 the children of atomic-bomb survivors exposed *in utero*.

1827  
1828 Akleyev *et al.* (2016) conducted a follow-up through 2009 of the 5,331 offspring of female Mayak PA workers  
1829 and 16,821 offspring of women in the Techa River cohort who were born in 1950 to 1961. Mortality follow-up was  
1830 for 1950 to 2009 and cancer incidence for 1956 to 2009. Both prenatal and postnatal individual doses were estimated  
1831 for the combined cohort. The estimated mean *in utero* dose was 14.1 mGy (range 0 to 945 mGy) and the mean  
1832 postnatal dose, based on living near the Techa River or working at the Mayak facility, was 11.2 mGy (range 0 to  
1833 552 mGy). Based on 369 incident solid cancers, the linear prenatal dose response was null [RR per 10 mGy of  
1834 0.98 (95 % CI 0.96, 1.01); ERR Gy<sup>-1</sup>  $\approx -2$  (-4, 1)], while that for postnatal dose was statistically significant [RR per  
1835 10 mGy of 1.02 (95 % CI 1.00, 1.04) ; ERR Gy<sup>-1</sup>  $\approx 2$  (0, 4)], with mutual adjustment for the other exposure. For  
1836 solid cancer mortality, the risks were similar but not statistically significant, probably because of fewer cases. When  
1837 the authors further examined digestive, respiratory and breast cancer incidence, they found a suggestion of a linear  
1838 dose response for digestive cancer incidence [RR per 10 mGy of 1.04 (95 % CI 1.00, 1.07; ERR Gy<sup>-1</sup>  $\approx 4$  (0, 7)], but  
1839 not for respiratory or breast cancers. Analyses of all-solid, digestive or breast cancers restricted to those with  
1840  $\leq 10$  mGy of postnatal exposure found only null effects, but had very low statistical power.

1841

1842 Schüz *et al.* (2016), in a parallel study, examined hematologic malignancies by prenatal and postnatal exposure  
1843 in the Mayak and Techa sets of offspring. The ERR Gy<sup>-1</sup> for leukemia by prenatal dose was approximately 4 (95 %  
1844 CI 0.7, 24), based on 28 leukemias, but no risk was apparent for lymphoma. There were no significant effects for  
1845 postnatal exposure or for hematologic mortality endpoints

1846

#### 1847 4.6.3 *In Utero Diagnostic Radiology*

1848

1849 The risk of cancer in offspring that have been exposed to diagnostic x-ray procedures while *in utero* has been  
1850 debated for more than 50 y. While high doses to the embryo or fetus (*e.g.*, 0.5 Gy) increase the risk of cancer, the  
1851 risk of cancer in offspring exposed *in utero* at a low dose such as <0.1 Gy still has uncertainties, and the  
1852 available data have not provided definitive information with regard to LNT. NCRP Report No. 174 (NCRP,  
1853 2013) extensively reviewed the risk of specific and total childhood cancers and cancer mortality in offspring of  
1854 women who underwent diagnostic x-ray procedures during pregnancy. Tables 5.14 and 5.15 of NCRP Report No.  
1855 174 summarize the relative risks reported in the epidemiologic literature (NCRP, 2013).

1856

1857 Data from case-control studies (including two large studies that relied on medical records for exposure  
1858 determination) support a statistical association between childhood leukemia in offspring and the mother's exposure  
1859 to diagnostic x rays during pregnancy. The relative risk of childhood leukemia based on a meta-analysis of 32  
1860 case-control studies is estimated as 1.3 (95 % CI = 1.2 to 1.5) (Wakeford, 2008). Based on a review of the  
1861 available data through 2009, IARC recently noted a causal association between fetal exposure to diagnostic x rays  
1862 and the risk of childhood cancer (Ghissassi *et al.*, 2009; IARC, 2012; Wakeford, 2015), however, this  
1863 interpretation is not universally accepted (Brent, 2014). Although the statistical association from case-control  
1864 epidemiologic studies is not generally debated, investigators have disagreed about both the etiologic significance  
1865 (causality), and if the association is causal, the likely magnitude of the leukemogenic risk (ICRP, 2003; NCRP,  
1866 2013).

1867

1868 Meta-analyses of cohort studies (concerning exposure of mothers to diagnostic x rays during pregnancy) have  
1869 found small, not statistically-significant increases of total cancer, but confidence intervals (CI) were compatible  
1870 with a composite increase similar to that of the case-control studies of 30 % or a composite estimate compatible  
1871 with no increase in risk. Overall, the cohort studies are characterized by limited numbers of total childhood cancer  
1872 cases and the subset of childhood leukemia cases, and with insufficient statistical power, and potential  
1873 uncertainties and data unreliability (Doll *et al.*, 1994). These limit the ability to draw firm conclusions regarding a  
1874 causal LNT association.

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## 4.7 Genetic Studies (Heritable Effects in Human Populations)

### Highlights

Heritable genetic effects have not been demonstrated for either cancer or noncancer endpoints in human studies; most notably, no effects have been found in the F1 offspring of atomic-bomb survivors. Studies of children from parents exposed to radiation (atomic bombs, accidental, environmental or medical) have been conducted over many decades using a wide range of endpoints including recently developed molecular endpoints. There is no definitive evidence for an increase in mutation frequency in the first or subsequent generations of exposed parents. The reports of increases in repetitive DNA sequences that have been proposed as evidence for transgenerational effects have been broadly discussed, resulting in the general opinion that there is insufficient evidence to conclude that these constitute an increased risk to the F1 and subsequent generations (Bouffler *et al.*, 2006; Little *et al.*, 2013). For these reasons, the estimates of heritable risk are based on radiation-induced mutations from mouse studies where transgenerational effects are clearly observable.

For the purposes of incorporating heritable risk into the overall risk from ionizing radiation, heritable risk is calculated for continuous low dose-rate exposures over two generations. The present heritable risk estimate, developed by UNSCEAR (2001) and ICRP (2007) essentially using the same methods, is about 0.2 % per gray. This value is more than 20-fold less than that for cancer. For radiation protection purposes heritable risk is included with cancer in the overall risk for gonads (ICRP, 2007). In essence, the estimate of genetic risk is not based on a LNT model from epidemiologic data because there is no reliable evidence of radiation-induced heritable mutations in humans.

Measures of genetic disease in the children of exposed parents include cytogenetic syndromes, single gene disorders, malformations, stillbirths, neonatal deaths, cancer, common polygenic diseases, and cytogenetic and genetic markers (Grant *et al.*, 2015; ICRP, 2007; Nakamura, 2006; NA/NRC, 2006; NCRP, 2015; Tatsukawa *et al.*, 2013; Winther and Olsen, 2012; UNSCEAR, 2001).

This section briefly reviews studies of cancer in the offspring of Japanese atomic-bomb survivors (Section 4.7.1), offspring after parental preconception radiotherapy (Section 4.7.2) and after environmental preconception radiation (Section 4.7.3).

1878  
1879  
1880

### 4.7.1 *Studies of Atomic-Bomb Survivor Offspring*

1881 The atomic-bomb studies have not shown any indication of heritable genetic risks from radiation exposure to  
1882 either parent. No excess of birth defects was seen among about 77,000 F1 children conceived after the bombing

1883 (Neel and Schull, 1956; Otake *et al.*, 1990). Nor were dose-related excesses seen for either cancer or noncancer  
1884 mortality among 75,000 F1 offspring observed for up to 62 y in relation to maternal, paternal or combined  
1885 maternal + paternal doses (Grant *et al.*, 2015). In the clinical study of the prevalence of noncancer diseases or  
1886 conditions in the F1 offspring of atomic-bomb survivors, no associations were seen for any of the conditions—  
1887 stroke, myocardial infarction, angina pectoris, hypertension, hypercholesterolemia or diabetes— in relation to  
1888 either maternal or paternal radiation dose (Tatsukawa *et al.*, 2013).

1889  
1890 A number of other indicators of transgenerational effects have been evaluated over the years in the atomic-  
1891 bomb survivor studies but have not shown associations with parental radiation dose, including untoward  
1892 pregnancy outcomes (congenital malformations, stillbirths, neonatal deaths), cytogenetic abnormalities  
1893 [chromosome number (sex-aneuploidy or Down syndrome), chromosome structure (translocations)], sex of child,  
1894 growth and development, biochemical studies (electrophoretic variants and erythrocyte enzyme activity),  
1895 microarray-based comparative genomic hybridization, and mutations in minisatellite or microsatellite genetic loci  
1896 (Kodaira *et al.*, 1995; 2004; Neel, 1998; Neel and Schull, 1956; Satoh *et al.*, 1996).

1897  
1898 **4.7.2** *Studies of Offspring after Parental Preconception Radiotherapy*  
1899

1900 A variety of human studies in Japan, Denmark, Finland and the United States of birth defects occurring in  
1901 offspring of women with gonadal irradiation from radiotherapy treatment have produced null results and thus  
1902 provide no information on which to judge the use of the LNT model for radiation protection. There were no excess  
1903 heritable genetic changes in cytogenetic abnormalities, single gene disorders, birth defects, stillbirths, neonatal  
1904 deaths, and cancer in the children of men exposed to testicular irradiation or women exposed to ovarian irradiation  
1905 (Green *et al.*, 2009; NCRP, 2013; Signorello *et al.*, 2012; Winther and Olsen, 2012). The sex ratio among the live-  
1906 born children of cancer survivors treated with radiation therapy provided no indication of a possible  
1907 transgenerational or germline effect (Winther *et al.*, 2003). No radiation- related excess of infant mortality was  
1908 found in relation to preconception radiation exposure in Techa River residents (Ostroumova *et al.*, 2005).

1909  
1910 The molecular analyses provided no evidence of radiation-related excesses in offspring of genomic instability,  
1911 inherited mutations in minisatellite DNA or mitochondrial DNA, chromosome radiosensitivity and DNA  
1912 polymorphic variation, and the occurrence of cytogenetic abnormalities (Guo *et al.*, 2012; Kodaira *et al.*, 2004;  
1913 Tawn *et al.*, 2005; 2011; Wilding *et al.*, 2007).

1914

1915 **4.7.3** *Cancer in Offspring after Environmental Preconception Radiation Exposure*

1916  
1917 Subsequent to the report by Gardner (1984) of excess childhood leukemia and non-Hodgkin lymphoma in  
1918 Seascale, near the Windscale/Sellafield nuclear plant in Great Britain, a large number of studies have been  
1919 performed of children around various nuclear plants throughout Europe and the United States. Gardner (1984)  
1920 reported that their analyses tended to attribute the excess hematopoietic malignancies to preconception irradiation  
1921 of the father, though the potential for protracted exposures means that any association might be due to  
1922 preconception, *in utero*, or childhood exposure. Among alternate explanations for the initial Seascale findings was  
1923 a model, with some substantiation, of an infective etiology through new-population mixing (Kinlen *et al.*, 1991).  
1924 Overall, there have been a few studies imputing an association of projected radiation exposures around nuclear  
1925 plants (some mix of preconception, prenatal and postnatal exposure) and childhood hematopoietic malignancies  
1926 (*e.g.*, Kaatsch *et al.*, 2008). However, a large number of null studies (Laurier *et al.*, 2014) have been reported also,  
1927 as ably reviewed in various reports by the British Committee on Medical Aspects of Radiation in the Environment  
1928 (COMARE, 1986; 2005; 2011; 2016).

1929  
1930 **4.7.4** *Implications for Radiation Protection*

1931  
1932 Past and recent studies of heritable effects among the children of human populations exposed to ionizing  
1933 radiation have been mostly negative, with only weak positive evidence (NCRP, 2013). Extensive dosimetry for  
1934 atomic-bomb survivors and for the survivors of cancer has been conducted. There is no conclusive evidence for  
1935 either a radiation association or a dose response for any of the numerous measures evaluated. Thus there is little to  
1936 no evidence for heritable effects among the children of atomic-bomb survivors; children of cancer survivors;  
1937 children of residentially-exposed populations or children of radiation-exposed workers (NCRP, 2013).

1938  
1939 Mouse studies continue to be used to estimate genetic risks because of the lack of clear evidence in humans that  
1940 germline mutations caused by radiation result in demonstrable genetic effects in children (ICRP, 2007). For  
1941 radiation protection, it is assumed to be unlikely that the human is immune to heritable effects of radiation,  
1942 although the effect must be small since not detectable even at high doses. Gonadal exposure is included in the  
1943 “detriment” equation but the risk of heritable effects in the whole population associated with gonadal dose is now  
1944 estimated to be a factor of 5 lower than in the past, *i.e.*, about 20 cases per 10,000 people per Sv rather than about  
1945 100 cases per 10,000 per sievert (ICRP, 2007).

1946 **5. Review of Epidemiologic Studies for Tissue Reactions**

1947

**Highlights**

Since radiation protection is designed to prevent adverse tissue reactions (deterministic effects), assessment of the LNT model for use in radiation protection is predominantly directed toward cancer induction and, to a lesser extent, hereditary effects. In the last decade, however, for a number of effects which were previously thought to have clear thresholds, newer studies suggest that they may occur at low doses and may possibly be stochastic in nature. Therefore, the dose response, particularly in the low dose range, is briefly examined.

A number of studies with data on radiation and circulatory (cardiovascular) disease at doses under 1 Gy have been published in the last 10 to 15 y. Essentially all of these studies are complicated by potential confounding factors (especially tobacco), combining diverse types of cardiovascular diseases (likely of various etiologies), uncertain diagnostic criteria and poor pathological confirmation. There also is no clear understanding of the target cells or tissues and the underlying biological processes. The reported increase in cardiovascular diseases in atomic-bomb survivors is unclear statistically below 0.5 Gy and comes from a number of diagnoses which vary substantially over time and that may have explanations other than radiation exposure. Radiation therapy studies do not show a significant increase of cardiovascular disease below 1 Gy. Studies of radiologic technologists and radiologists usually are done by questionnaire, often lack individual dosimetry, are prone to recall bias and have not demonstrated a dose response. Studies of nuclear workers are often insignificantly positive and most lack information on important confounding factors.

Study of cataracts in the atomic-bomb survivors and particularly following Chernobyl exposures have revealed development of minor lenticular opacities at doses lower than previously considered to be cataractogenic. Ophthalmologically detectable opacities are reported at doses of about 0.5 to 2 Gy or more. Review of the mechanistic studies suggests a possible stochastic phenomenon while the epidemiologic studies suggest that a threshold model may be more appropriate. Presently, the link between the mechanistic and epidemiologic evidence is not clear. At this time, the NCRP recommends use of the threshold model. While this does not apply to the use of the LNT model for exposure of tissues other than the lens of the eye, it does require that the lower apparent threshold be taken into consideration for occupational eye exposures.

The studies of noncancer thyroid effects includes radiation therapy patients, atomic-bomb survivors, exposure from weapons fallout, Chernobyl populations, and Hanford public exposure. Essentially none of the most recent and well-designed studies show any increase in thyroiditis or hypothyroidism at doses of less than 1 Gy. Occasionally, the matter is complicated by authors inappropriately applying a LNT model to data that are clearly nonlinear. Overall, studies of adverse tissue reactions or of effects that may or may not be stochastic, do not clearly

show an effect below doses of 0.5 to 1.0 Gy and thus are not the limiting factor or helpful in establishing annual effective dose limits for workers or the public. The one exception is the equivalent dose to the lens of the eye during occupational exposure.

The appropriateness of a LNT model is briefly examined for cardiovascular diseases and cataract. The cardiovascular disease studies include studies of higher doses (Section 5.1.1), TB fluoroscopy studies (Section 5.1.2), nuclear worker studies (Section 5.1.3), and studies of environmental radiation exposures (Section 5.1.4). Recent cataract studies are reviewed in Section 5.2 and thyroid noncancer effects in Section 5.3.

1948

1949

### 5.1 Cardiovascular Effects Studies

1950

1951 It is well established that high acute doses of ionizing radiation increase the risk of circulatory system  
1952 (cardiovascular) diseases, such as heart attack and stroke (Darby *et al.*, 2013; HPA, 2010; ICRP, 2012; NCRP,  
1953 2011; Shimizu *et al.*, 2010; Stewart, 2012; Takahashi *et al.*, 2013; 2017; Travis *et al.*, 2014). The evidence for  
1954 this comes from the experience of the Japanese atomic-bomb survivors and of radiotherapy patients. The  
1955 underlying radiobiological mechanism is believed to be primarily cell killing by high absorbed doses (typically,  
1956 the doses received by tissues from whole-body doses >1 Gy) leading to tissue damage and a consequent raised  
1957 risk of cardiovascular disease that is clinically manifest some years later. However, there is growing evidence to  
1958 suggest a raised risk of cardiovascular disease at lower levels of exposure to radiation, although the mechanisms  
1959 for possible risk at those doses are poorly understood. Currently, the framework of radiation protection set out by  
1960 the ICRP (2007) does not include an increased risk of cardiovascular disease in the exposed individual from low-  
1961 level exposures. Hence, it is of some importance to understand and correctly interpret the evidence for  
1962 cardiovascular disease risk associated with low or moderate dose/dose-rate exposures to radiation to ensure that  
1963 the risks arising from these lower level exposures are appropriately assessed and (if necessary) incorporated into  
1964 the scheme of radiation protection. Given the high background rate of cardiovascular disease in economically  
1965 developed countries, a relatively small proportional increase in risk produced by low-level exposure to radiation  
1966 will imply a comparatively large number of excess cases. This matter is of direct relevance to occupational  
1967 exposure to radiation, for which there exists mixed evidence of increased cardiovascular disease risk. The  
1968 evidence for a risk of radiation-induced cardiovascular disease will be briefly summarized below.

1969

1970 The literature regarding cardiovascular disease (CVD) after radiation exposure is complicated by a number of  
1971 important factors. First, CVD is not a single entity and is a term used to describe a myriad of disparate conditions  
1972 with different causes. As an example, heart disease includes valvular abnormalities, capillary and blood vessel  
1973 lesions, aneurysms, effusions, muscle abnormalities, arrhythmia, endocarditis, malformations etc. A derivative

1974 issue is that of disease diagnosis and classification, *e.g.*, a diagnosis of hypertension varies widely by medical  
1975 practice, country, and over time, so that results of published studies about low dose radiation and CVD can be  
1976 misleading. There are many potential confounding causes of CVD, such as smoking, hereditary factors, diet-  
1977 related factors and concurrent conditions (*e.g.*, diabetes, obesity). Importantly, at present there is no clear  
1978 understanding of the biological mechanisms for cardiovascular diseases at low doses and there is little  
1979 understanding of the target cells or tissue. Without these, application of a linear dose response model at low doses  
1980 and assumption that this is a stochastic process remains debatable.

1981

### 1982 5.1.1 *Higher Doses and Cardiovascular Disease*

1983

1984 Increased risks of cardiovascular disease (including myocardial infarction, coronary artery disease and stroke)  
1985 are well documented effects after high radiation doses (>30 Gy) to the heart or neck that may occur with radiation  
1986 therapy [*e.g.*, for Hodgkin lymphoma and breast cancer (Darby *et al.*, 2013; Travis *et al.*, 2014)]. Overall,  
1987 however, the radiation therapy literature is of less interest for radiation protection standards than epidemiologic  
1988 studies of persons exposed to lower doses.

1989

1990 There are several reports of increased risk of cardiovascular disease in atomic-bomb survivors. In the LSS  
1991 cohort study of 8,400 heart disease deaths, researchers found an approximately linear dose response over the  
1992 dose range 0 to about 3 Gy (Shimizu *et al.*, 2010). The estimated linear dose response risk estimate was an  
1993 ERR Gy<sup>-1</sup> of 0.14 (95 % CI 0.06, 0.23). The dose response over the range 0 to 1 Gy was statistically significant  
1994 but over the range 0 to 0.5 Gy it was not. Excess risk was clear only above about 0.5 Gy. The increased risk of  
1995 14 % per gray seen in the LSS was driven primarily by the unexpected categories of heart failure, rheumatic  
1996 heart disease (after heart infections), hypertension and hemorrhagic stroke, and about 30 % of the nominal risk  
1997 was found to be potentially attributable to prior/concurrent diagnoses of cancer (Shimizu *et al.*, 2010). Further  
1998 detailed aspects of the atomic-bomb survivor data on cardiovascular diseases are given in (Ozasa *et al.*, 2017;  
1999 Takahashi *et al.*, 2017).

2000

2001 In the smaller atomic-bomb clinical morbidity study (Yamada *et al.*, 2004) the ERRs for hypertension,  
2002 hypertensive heart disease, ischemic heart disease, aortic aneurysm and stroke all were nonsignificant. There was a  
2003 linear quadratic dose response for hypertension and myocardial infarction but not a linear response.

2004

2005 There also are a number of issues regarding studies of atomic-bomb survivors that potentially limit the  
2006 application of findings to the United States population. These include the relatively high incidence of stroke to heart  
2007 disease in Japan (which is reversed in the United States), the high proportion of hemorrhagic stroke in Japan (which

2008 is much less common in the United States and may reflect major differences in diet and hereditary background).  
2009 The nature and magnitude of the risk (if any) at acute doses less than 0.5 Gy area unresolved (Takahashi *et al.*,  
2010 2013).

2011

### 2012 **5.1.2** *TB Fluoroscopy Studies*

2013

2014 A description of the Massachusetts (Boice and Monson, 1977; Boice *et al.*, 1978) and Canadian (Miller  
2015 *et al.*, 1989; Sherman *et al.*, 1978) TB fluoroscopy studies and their dosimetry is given in Section 4.4.1. The  
2016 patients on average received roughly 100 chest fluoroscopies to monitor lung collapse (pneumothorax), with a  
2017 fluoroscopic examination every two to three weeks, spread out over several years. The studies initially  
2018 concentrated on breast and other cancers, but recent publications have examined cardiovascular disease (Davis  
2019 *et al.*, 1989; Little *et al.*, 2016; Zablotska *et al.*, 2014).

2020

2021 **5.1.2.1** *Dosimetry*. The authors concluded that estimates of both the heart and carotid dose in the Massachusetts  
2022 study may be too low or high by a factor of two (Little *et al.*, 2016), but the factor could be much higher. Shared  
2023 uncertainty was not investigated but could be significant in that the same dose per treatment was used for many  
2024 individuals with the same characteristics: age, body size, calendar year, and beam orientation. These uncertainties  
2025 apply to both the Massachusetts and Canadian studies on cardiovascular disease, though the Canadian study  
2026 estimated lung doses using Monte-Carlo simulation techniques that should provide a reasonable estimate of  
2027 individual radiation doses to the lung and heart (Zablotska *et al.*, 2014).

2028

2029 A strength of the Massachusetts TB fluoroscopy study was that multiple realizations of dose were used to  
2030 account for uncertainty, and the doses spanned a large range, from 0 Gy to a maximum of 11.6 Gy. Dosimetry  
2031 weaknesses were that lung dose was used as a surrogate for cardiovascular diseases, actual estimates of  
2032 uncertainty in individual doses were not provided, and shared uncertainty was not considered. Thus, the risk  
2033 estimates could be biased by dose uncertainty. Further the inconsistencies in the Massachusetts and Canadian  
2034 studies, *e.g.*, a reported enhancement of risk for fewer fluoroscopies in the Canadian (Zablotska *et al.*, 2014) but  
2035 not the Massachusetts (Little *et al.*, 2016) studies point to potential biases or uncertainties in the assumptions for  
2036 derivative measures.

2037

2038 **5.1.2.2** *Study Strengths and Weaknesses*. The epidemiologic methods are described in Section 4.4.1.2. The  
2039 studies took into account smoking histories and other potentially confounding factors. These studies, overall,  
2040 provide no evidence that low-dose fractionated exposures accumulating to a moderate to high dose are associated  
2041 with heart or cardiovascular diseases in either the Canadian (Zablotska *et al.*, 2014) or Massachusetts (Davis

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2042 *et al.*, 1989; Little *et al.*, 2016) cohorts. There are dosimetry issues that hinder interpretation of the shape of any  
2043 dose-response relationship: average lung dose was used as a surrogate for heart dose and thyroid dose as a  
2044 surrogate for carotid artery dose. Further, no adjustment was made to increase the dose for patients who had  
2045 special procedures that entailed lengthy fluoroscopic examinations. Finally, there is no mention of how  
2046 corpulmonale (enlargement of the right side of the heart due to pulmonary hypertension caused by tuberculosis),  
2047 a potential confounder, was addressed.

2048  
2049 **5.1.2.3 Implications for the LNT Model and Radiation Protection.** The studies of heart and cardiovascular disease  
2050 at this time provide little evidence for an association between radiation and heart disease and much less evidence  
2051 for a dose response that might be relevant in the low-dose domain.

2052

### 2053 **5.1.3 Nuclear Worker Radiation Exposure and Cardiovascular Diseases**

2054

2055 A number of epidemiologic studies of the risk of cardiovascular diseases (CVDs) among groups of workers  
2056 exposed protractedly to radiation have been conducted and reported. These studies have been reviewed in the  
2057 literature, by Little and colleagues (Little, 2016; Little *et al.*, 2010a; 2012) and Kreuzer *et al.* (2015). Workforces  
2058 that have been studied include nuclear industry workers, Chernobyl emergency and recovery workers, and uranium  
2059 miners. The evidence from these occupationally exposed groups is summarised in Table I, which updates the  
2060 studies considered in the aforementioned reviews. Table 5.1 also includes selected medical and environmental  
2061 studies of relevance to low-dose and low-dose rate circulatory disease.

2062

2063 The studies considered are those with estimates of the slope of the CVD dose-response based on (primarily)  
2064 external photon exposure. It does not include other occupational studies that have reported only Standardized  
2065 Mortality Ratios (SMRs), because of the potential biases in SMRs and the fact that such studies are not based on  
2066 individual doses.

2067

2068 ERR/Gy estimates are shown for all CVD and for major groupings of these diseases: ischemic heart disease  
2069 (IHD) and cerebrovascular disease (CeVD). Other CVD groupings may be provided in some studies, but their  
2070 presentation is not consistent and so these results are not given in Table 5.1. Not all CVDs are related to radiation  
2071 at either high or lower doses, *e.g.*, rheumatic heart disease, so future studies should focus on clinically relevant  
2072 outcomes and not broad categories (Einstein *et al.*, 2017).

2073

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2074 Table 5.1—*Summary of estimates of ERR of various cardiovascular diseases per gray of cumulative whole-body*  
 2075 *external gamma-ray dose, and associated 95 % confidence intervals, obtained from studies of radiation workers*  
 2076 *and selected environmental and medical exposures. Results given for all cardiovascular diseases combined*  
 2077 *(CVD), IHD and CeVD separately, where available. Doses are lagged by 10 y, except where stated otherwise.*  
 2078 *Adaptation and update of tables presented by Little et al. (2010a; 2012) and Kreuzer et al. (2015).*

Study	Cohort Size	Mean Cumulative External Dose (Gy)	Cardiovascular Disease <sup>a</sup>	ERR Gy <sup>-1</sup> (95 % CI)			
Nuclear Workers							
INWORKS (Gillies <i>et al.</i> , 2017)	308, 297	0.025	CVD mortality	0.22 (0.08, 0.37) <sup>f</sup>			
			IHD mortality	0.18 (0.004, 0.36)			
			CeVD mortality	0.50 (0.12, 0.94)			
Mayak, Russia (Azizova <i>et al.</i> , 2014; 2015a)	18 856	0.593	CVD mortality	0.06 (0.01, 0.12)			
			IHD incidence	0.14 (0.08, 0.21)			
			IHD mortality	0.05 (−0.01, 0.13)			
			CeVD incidence	0.49 (0.39, 0.60)			
			CeVD mortality	0.05 (−0.03, 0.16)			
15 countries (Vrijheid <i>et al.</i> , 2007a)	275 312	0.021	CVD mortality	0.09 (−0.43, 0.70)			
			IHD mortality	−0.01 (−0.59, 0.69)			
			CeVD mortality	0.88 (−0.67, 3.16)			
NRRW-3, UK <sup>b</sup> (Muirhead <i>et al.</i> , 2009)	174 541	0.025	CVD mortality	0.25 (−0.01, 0.54)			
			IHD mortality	0.26 (−0.05, 0.61)			
			CeVD mortality	0.16 (−0.42, 0.91)			
BNFL, UK <sup>b,c</sup> (McGeoghegan <i>et al.</i> , 2008)	38 779 <sup>d</sup>	0.057	CVD mortality <sup>e</sup>	0.65 (0.36, 0.98) <sup>f</sup>			
			IHD mortality <sup>e</sup>	0.70 (0.33, 1.11) <sup>f</sup>			
			CeVD mortality <sup>e</sup>	0.43 (−0.10, 1.12) <sup>f</sup>			
Idaho National Engineering and Environmental Laboratory, USA <sup>j</sup> (Schubauer-Berigan <i>et al.</i> , 2005)	35,833	0.131	IHD mortality <sup>k</sup>	−0.30 (−1.24, 0.90)			
			France <sup>b</sup> (Metz-Flamant <i>et al.</i> , 2013)	59 021	0.023	CVD mortality	0.31 (−0.90, 1.74) <sup>f</sup>
						IHD mortality	0.71 (−1.20, 3.18) <sup>f</sup>
ORNL, USA <sup>b,g</sup> (Richardson and Wing, 1999)	14 095	NA	CeVD mortality	0.99 (<0, 5.05) <sup>f</sup>			
			IHD mortality	−2.86 (−6.90, 1.18)			

Russian Chernobyl liquidators (Ivanov, 2007; Ivanov <i>et al.</i> , 2006)	61 017	0.109	CVD incidence IHD incidence CeVD incidence	0.18 (−0.03, 0.39) 0.41 (0.05, 0.78) 0.45 (0.11, 0.80)
Eldorado uranium miners, Canada (Lane <i>et al.</i> , 2010)	16 236 <sup>d</sup>	0.052	IHD mortality <sup>h</sup> CeVD mortality <sup>h</sup>	0.15 (−0.14, 0.58) −0.29 (<−0.29, 0.27)
German uranium miners (Kreuzer <i>et al.</i> , 2013)	58 982	0.041	CVD mortality IHD mortality CeVD mortality	−0.13 (−0.38, 0.12) −0.03 (−0.38, 0.32) 0.44 (−0.16, 1.04)
French uranium miners (Drubay <i>et al.</i> , 2015)	1690 <sup>i</sup>	0.066	CVD mortality IHD mortality CeVD mortality	0.40 (−1.8, 3.0) <sup>j</sup> −1.1 (−4.0, 3.2) <sup>j</sup> 3.7 (−0.9, 10.6) <sup>j</sup>
Environmental and Medical Exposures				
Dwellers near Techa River (Krestinina <i>et al.</i> , 2013b)	29,735	0.035	CVD mortality IHD mortality CeVD mortality	0.24 (−0.08, 0.59) 0.40 (−0.11, 0.99) −0.14 (−0.64, 0.46)
Yangjiang, China high background radiation area (Tao <i>et al.</i> , 2012)	31,604	0.063	CVD mortality IHD mortality CeVD mortality	0.14 (−0.84, 1.29) 0.54 (−2.65, 6.13) 0.44 (−0.88, 2.08)
Canadian and Massachusetts patients with TB fluoroscopies (Tran <i>et al.</i> , 2017)	77,275	~0.20	CVD mortality IHD mortality CeVD mortality	0.27 (−0.02, 0.58) 0.22 (−0.11, 0.66) 0.54 (−0.20, 1.34)

2079 CVD = cardiovascular disease

2080 IHD = ischemic heart disease

2081 CeVD = cerebrovascular disease

2082 NA = not available

2083 CI = confidence interval

2084 <sup>a</sup> Mortality data are based on underlying cause of death

2085 <sup>b</sup> Substantial overlap with INWORKS and 15-country studies

2086 <sup>c</sup> some overlap with NRRW-3 study

2087 <sup>d</sup> men only included in analysis

2088 <sup>e</sup> 15 y dose lag

2089 <sup>f</sup> 90 % CI

2090 <sup>g</sup> analysis conducted in terms of the cumulative dose received after the age of 45 y

2091 <sup>h</sup> 2 y dose lag

2092 <sup>i</sup> restricted cohort from which cases and controls drawn

2093 <sup>j</sup> derived from hazard ratio obtained from nested case-control study. <sup>1</sup> Monitored study cohort (INEEL) adjusting for age, sex, calendar time, duration of employment, SES, migrant status, interaction between SES and migrant status, and internal dose (see Table 6-14).

2094 <sup>1</sup> Overlap with INWORKS and the 15-country study. The extent of overlap is uncertain because of exclusion criteria used in the studies [e.g., described in Vrijheid *et al.* (2007b) specifically for what is labelled as US-INL] and follow-up apparently started in 1960 and not 1942 when work at INEEL began or 1951 for the INEEL monitored workers.

2098 Table 5.1 and several reviews show elevated ERR Gy<sup>-1</sup> estimates for cardiovascular disease, ischaemic heart  
2099 disease and cerebrovascular disease associated with exposure to external sources of gamma radiation in the  
2100 workplace, and from environmental and low-dose rate medical exposures. However, a cautious interpretation is  
2101 required, as discussed below, that temper conclusions.

2102  
2103 Particular groups of workers are influential in generating the impression of a pattern of elevated ERR Gy<sup>-1</sup>  
2104 estimates. For example two of the estimates (Muirhead *et al.*, 2009; Vrijheid *et al.*, 2007a) are dependent upon  
2105 the significantly increased risk seen in the BNFL cohort, ERR = 0.65 Gy<sup>-1</sup> (90 % CI: 0.36, 0.98) (McGeoghegan  
2106 *et al.*, 2008). McGeoghegan *et al.* showed that the BNFL results have heterogeneity in risk related to  
2107 confounding by socioeconomic and lifestyle factors. The new report of cardiovascular disease risk in the  
2108 INWORKS study (Gillies *et al.*, 2017) also showed risk heterogeneity: significantly higher risk of IHD  
2109 among white collar (ERR Gy<sup>-1</sup> = 0.58, 90 % CI 0.22, 0.98) and a much lower risk among blue collar workers  
2110 (ERR Gy<sup>-1</sup> = 0.07, 90 % CI -0.11, 0.27), while there was a suggestion of an opposite difference for  
2111 cerebrovascular disease (ERR Gy<sup>-1</sup> values of 0.59 (90 % CI 0.18, 1.07) among white collar workers and a  
2112 significantly low risk of -0.08 (90 % CI <0, 0.77) among blue collar workers.

2113  
2114 Similar to BNFL, the Idaho National Engineering and Environmental Laboratory (INEEL) study  
2115 (Schubauer-Berigan *et al.*, 2005) is another example of overlap and heterogeneity. INEEL is included in both the  
2116 15-country study (Vrijheid *et al.*, 2007b) and in INWORKS (Gillies *et al.*, 2017; Hamra *et al.*, 2016). The  
2117 INEEL study cohort of monitored workers, followed from 1951 to 1998, showed a negative dose response for  
2118 IHD (ERR Gy<sup>-1</sup> = -0.30, 95 % -1.24, 0.902), and in contrast to INWORKS the blue collar workers (low SES)  
2119 had higher risks than the white collar workers (high SES), *i.e.*, just the reverse of INWORKS. Specifically in  
2120 INEEL, “Unskilled workers, skilled manual workers, and partially skilled workers showed greater risks (*for*  
2121 *IHD*) than the baseline risk of professional and skilled non-manual workers.” Other uncertainties included  
2122 different starts of follow-up (1951 for the INEEL defined study cohort and 1960 for the 15-country study),  
2123 different cohort numbers, *e.g.*, 64,000 for the 15-country study (Vrijheid *et al.*, 2007b). Comparisons with  
2124 INWORKS are problematic since few details are provided on the INEEL cohort, and a reference to the study is  
2125 not provided in any INWORKS publication. In supplementary tables, however, the INEEL ERR Gy<sup>-1</sup> estimates  
2126 for CVD, IHD and CeVD were negative (Gillies *et al.*, 2017, supplementary tables). Overlap between INEEL  
2127 and the ORNL and Hanford cohorts apparently was not removed (Vrijheid *et al.*, 2007b), suggesting that some  
2128 workers might be counted as many as three times in the analyses.

2129  
2130 In the case of the Mayak workers, information was available on smoking, alcohol consumption and obesity  
2131 (Moseeva *et al.*, 2014), but the results are difficult to interpret because of substantial differences between the

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2132 ERR Gy<sup>-1</sup> values for incidence and mortality from IHD and cerebrovascular disease; the risk estimates for  
2133 incidence are notably higher than for mortality, especially for cerebrovascular disease (Azizova *et al.*, 2014). The  
2134 Chernobyl liquidator data from Russia (Ivanov, 2007; Ivanov *et al.*, 2006) are also puzzling: the number of  
2135 incident cases of CeVD is surprisingly high, and the risk coefficient for incident CeVD is very high while that  
2136 for CeVD mortality is low.

2137  
2138 While cancer mortality was higher in U.K. radiologists who worked from 1897 to 1920, the mortality from  
2139 cardiovascular disease was lower compared to other medical practitioners (Berrington *et al.*, 2001), but there are  
2140 uncertainties in the dosimetry and statistics (Smith and Doll, 1981). The studies of about 90,000 U.S. radiologic  
2141 technologists have found increased risk of cardiovascular disease, particularly CeVD, in those starting work prior  
2142 to 1940 compared to those starting after 1960 (Hauptmann *et al.*, 2003), and other reports found weak evidence  
2143 of increased CVD risk related to work in interventional radiology (Linnet *et al.*, 2006; Rajaraman *et al.*, 2016) or  
2144 nuclear medicine (Kitahara *et al.*, 2015). The most recent report on cardiovascular mortality in the entire cohort  
2145 found that technologists working before 1950 had nonsignificantly increased mortality from some cardiovascular  
2146 diseases (IHD and stroke) (Liu *et al.*, 2014). However, none of these studies had individual dose estimates, so  
2147 their relevance to the LNT model is negligible.

2148  
2149 Table 5.2 presents estimates of ERR for all circulatory disease (CVD) and IHD per gray of cumulative  
2150 whole-body external gamma-ray dose and associated 95 % confidence intervals within the currently available  
2151 Million Worker Study cohorts (Boice 2017; Boice *et al.*, 2017; Bouville *et al.*, 2015).

2152  
2153 The four cohorts presented are nuclear power plant workers, industrial workers, aboveground nuclear  
2154 weapons test participants and workers at the Mound facility. They represent 387,532 or approximately 39 % of  
2155 the MWS population. The overall mean dose was 20 mGy. The dosimetry was conducted following the advice  
2156 from NCRP Scientific Committee (SC 6-9) as summarized in Bouville *et al.* (2015) and Boice *et al.* (2006b). The  
2157 follow-up for vital status and the determination of cause of death was also the same for all cohorts within the  
2158 MWS. Combining the cohorts within the MWS has begun for two cohort, Mound (Boice *et al.*, 2014) and  
2159 Rocketdyne (Atomics International) (Boice *et al.*, 2006a; 2011) for cardiovascular disease (Zhang *et al.*, 2014).

2160

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2161 Table 5.2—*Summary of estimates of ERR (and 95 % confidence intervals) of total heart disease and IHD per*  
 2162 *gray of cumulative whole-body external gamma-ray dose for cohorts within the Million Worker Study (Boice*  
 2163 *2017; Boice et al. 2017; Bouville et al. 2015). Results given for all heart diseases combined (CVD) and*  
 2164 *ischaemic heart disease separately, where available. Doses are lagged by 10 y.*  
 2165

Workforce	Cohort Size	Mean Cumulative Heart Dose (Gy)	Heart Disease <sup>a</sup>	ERR Gy <sup>-1</sup> (95 % CI)
Occupational Exposures				
Mound (Boice <i>et al.</i> , 2014; Golden <i>et al.</i> , 2017)	4,954	0.024	CVD mortality IHD mortality	-0.4 (-3.9, 3.2) -1.4 (-4.3, 1.4)
Nuclear Power Plant Workers (Boice and Bellamy, 2017; Golden <i>et al.</i> , 2017)	145,209	0.038	CVD mortality IHD mortality	0.06 (-0.5, 0.4) 0.1 (-0.5, 0.5)
Atomic Veterans (Boice 2017; Golden <i>et al.</i> , 2017)	113,813	0.006	CVD mortality IHD mortality	NA 0.06 (-0.05, 0.18)
Industrial Radiographers (Boice and Bellamy, 2017; Golden <i>et al.</i> , 2017)	123,556	0.011	CVD mortality IHD mortality	0.02 (-0.03, 0.08) 0.006 (-0.06, .05)

2166 CVD = cardiovascular disease

2167 IHD = ischemic heart disease

2168 NA = not available

2169 CI = confidence interval

2170 <sup>a</sup> Mortality data are based on underlying cause of death

2171

2172

2173 The ERR Gy<sup>-1</sup> estimates of risk for both CVD and IHD within the four MWS cohorts were all very low and  
2174 none were statistically significant, providing no evidence for an association between low dose and low-dose rate  
2175 exposures. The dose-response evaluation for the industrial radiographers provides no evidence for a dose  
2176 response relationship for IHD (Figure 5.1). Similar patterns are seen in the other three cohorts. The MWS studies  
2177 of medical workers and additional DOE workers are ongoing and will provide substantial statistical power when  
2178 combined.

2179

#### 2180 **5.1.4** Environmental Radiation Exposure

2181

2182 Few studies of environmental radiation exposure and cardiovascular diseases have individual dose information  
2183 that permits them to estimate the ERR Gy<sup>-1</sup>. Because of a recent indication of erroneous dosimetry,<sup>5</sup> the previous  
2184 report of residents downwind of the Russian Semipalatinsk nuclear test site in Kazakhstan (Grosche *et al.*, 2011) is  
2185 not included. In the Techa River study, positive but nonsignificant ERRs were reported for ischemic heart disease  
2186 mortality and all cardiovascular disease deaths (Krestinina *et al.*, 2013b). There were also positive but  
2187 nonsignificant risk coefficients for ischemic heart disease, cerebrovascular disease and all cardiovascular disease in  
2188 the Yangjiang, China study of high natural background irradiation (Tao *et al.*, 2012).

2189

#### 2190 **5.1.5** *Implications of Cardiovascular Disease for the LNT Model and Radiation Protection*

2191

2192 The study of atomic-bomb survivors had a complex pattern of types of cardiovascular disease associated with  
2193 radiation exposure and did not have clear evidence of excess risk below about half a gray. The occupational data  
2194 represent assessments at low doses and low dose rates. The majority of those studies were positive, but several had  
2195 inconsistencies or particular problems in interpretation. In addition, most of them did not have information on  
2196 smoking or other cardiovascular disease risk factors, so possible confounding could not be ruled out. The two  
2197 studies with environmental exposures had low statistical power to detect a radiation effect. So at this time, the data  
2198 on cardiovascular diseases at lower doses and dose rates do not present clear evidence regarding the LNT model or  
2199 applicability to radiation protection.

2200

---

<sup>5</sup> B Grosche, personal communication, 2015

2201



2226

2227 **Fig. 5.1.** Dose response relationship between IHD and cumulative dose to heart, industrial radiography workers.

2228 The data are fit with a linear ERR model.

2229

## 5.2 Cataract Studies

Studies of lens opacities and cataracts following radiation exposure have been reviewed extensively in NCRP Commentary No. 26 (NCRP, 2016). The commentary addressed radiation protection principles with respect to the lens of the eye, summarized the current understanding of eye biology and lens effects (including ionizing radiation effects), reviewed and evaluated the current epidemiology related to ionizing radiation and cataracts, and assessed exposed populations with the potential for significant radiation exposures to the lens. It pointed to a variety of strengths and limitations in the quality of the available epidemiologic evidence. NCRP determined that a threshold model should continue to be used for radiation protection purposes for lens of eye at this time. The value of the threshold for detectable opacities or vision-impairing cataracts (VICs) is less clear, with the epidemiologic evidence currently pointing to a threshold for VICs for doses in the region of 1 to 2 Gy. However, the NCRP concluded that it is not possible to make a specific quantitative estimate of lens effect thresholds.

NCRP also concluded that while the mechanisms underlying the transition of minor lens opacifications to clinically significant VICs are still not well understood, it is prudent to regard eye exposures and the potential for lens tissue effects in much the same way as whole-body exposures, using the ALARA principle, as was previously recommended by NCRP Report No.168 (NCRP, 2010). NCRP has determined that it is prudent to reduce the current recommended annual lens of the eye occupational dose limit from an equivalent dose of 150 mGy (NCRP, 1993a) to an absorbed dose of 50 mGy (NCRP, 2016). At this time the use of the LNT model for cataracts is not justified.

## 5.3 Thyroid Noncancer Effects Studies

This section reviews radiation effects on thyroid function and possible thyroiditis in an effort to evaluate use of LNT versus other potential dose-response models for radiation protection purposes. There are many large sources of human data on thyroid function and autoimmune issues including atomic-bomb survivors, fallout exposures, external radiation therapy, and radionuclide treatment for thyroid conditions. At high absorbed doses the main concern is reduced production of thyroid hormone (hypothyroidism) and at lower doses thyroiditis is more of a concern. The issue of thyroiditis is usually complicated by a number of issues including increase in thyroid antibodies with age, genetic predisposition, dietary factors, concurrent diseases and the criteria used to make the diagnosis. Several studies have reported statistically significant and relatively linear dose-related increases in the prevalence of thyroid nodules and cysts (Cahoon *et al.*, 2017b; Imaizumi *et al.*, 2006; 2015; Zablotska *et al.*, 2007; 2015).

2264 **5.3.1** *Atomic-Bomb Studies*

2265  
2266 An earlier study of atomic-bomb survivors showed a questionable increase in hypothyroidism in the 0.01 to  
2267 0.49 Gy group but not in the 0.50 to 0.99 Gy group (Nagataki *et al.*, 1994). However, a subsequent, larger study  
2268 among the survivors, which also included most of the individuals in the previous study, did not find any evidence  
2269 of dose-related hypothyroidism (Imaizumi *et al.*, 2006; 2017). Because the Nagataki *et al.* (1994) study had  
2270 applied a linear-quadratic model to their hypothyroid data, that model was also examined by Imaizumi *et al.*  
2271 (2006) but no quadratic curvature was found ( $p = 0.86$ ). Since testing positive for antithyroid antibodies is  
2272 strongly affected by increasing age, the subgroup of ages 0 to 9 y at the time of the bombing also was analyzed  
2273 separately. Antithyroid antibody-positive hypothyroidism among the younger cohort members was not associated  
2274 with radiation dose (excess odds ratio  $\text{Gy}^{-1} = -0.09$ ,  $p = 0.72$ ). Morimoto *et al.* (1987) also reported that in  
2275 survivors under the age of 20 y at exposure and with doses 1 Gy or more there was no increase in either  
2276 hypothyroidism or autoimmune thyroiditis.

2277  
2278 **5.3.2** *Chernobyl <sup>131</sup>I Fallout and Noncancer Thyroid Effects*

2279  
2280 There are quite a number of studies regarding Chernobyl with various methodologies, different dose sources  
2281 and often with conflicting results. The studies of thyroiditis are also complicated by the dietary deficiency of  
2282 nonradioactive iodine. Ostroumova *et al.* (2013) reported an increase in hypothyroidism [based on thyroid-  
2283 stimulating hormone (TSH), not thyroxine levels] but observed no evidence of autoimmune thyroiditis. They  
2284 estimated a risk of hypothyroidism at 1 Gy based on a linear model even though the best fit to the data was not  
2285 linear and there was no evidence of increased risk at doses <4 Gy. In addition, use of the linear model by the  
2286 authors implies a stochastic process even though it is well known that the relationship of thyroid tissue volume to  
2287 thyroid hormone or TSH is not linear.

2288  
2289 The largest study of the Chernobyl population was performed by the Sasakawa Foundation and reported by  
2290 Ito *et al.* (1995; Shibata *et al.*, 2002; Shigematsu, 2002). The study included 160,000 children and did not find  
2291 any increase in thyroid antibodies, hypothyroidism or hyperthyroidism that could be related to ionizing radiation.

2292  
2293 **5.3.3** *Other Studies of Radioactive Iodine Fallout and Thyroid Effects*

2294  
2295 Fallout from nuclear testing resulted in significant deposition of radioiodine in the Marshall Islands and  
2296 caused subclinical hypothyroidism in about 30 % of children who received thyroid doses of >2 Gy (Lessard *et al.*,  
2297 1985). Long-term follow-up of those with thyroid doses up to 4 Gy did not show an increase in autoimmune

2298 thyroiditis. A report on fallout of atomic tests in Nevada (Rallison *et al.*, 1990) indicated an increased risk of  
2299 autoimmune thyroiditis, which was supported by later analyses using corrected dosimetry (mean of 120 mGy,  
2300 with a standard deviation 167 mGy) and accounting for dose uncertainties (Li *et al.*, 2007), though this is in  
2301 contrast to negative findings at similar dose levels in most other studies.

2302  
2303 Davis *et al.* (2004) evaluated thyroid effects associated with low dose <sup>131</sup>I exposure from fallout from Hanford,  
2304 with individual estimates of absorbed dose the thyroid gland (median of 97 mGy, mean of 174 mGy, standard  
2305 deviation of 224 mGy, range of 0 to 2823 mGy) They found no evidence of dose response for either hypothyroidism  
2306 (ERR Gy<sup>-1</sup> of -0.006 (95 % CI -0.016, 0.047) or autoimmune thyroiditis (ERR Gy<sup>-1</sup> = -0.026, 95 % CI < -0.057,  
2307 0.044).

2308

#### 2309 **5.3.4** *Implications of Noncancer Thyroid Effects for Radiation Protection*

2310

2311 The available literature on hypothyroidism has negative results at low doses, which sometimes are  
2312 accompanied by upward curvature such that there is an indication of risk at doses well above 1 Gy. The data on  
2313 thyroiditis from the larger, well-conducted studies is also generally negative at low doses although there are  
2314 occasional scattered and conflicting small study reports. Given that hypothyroidism and immune thyroiditis are  
2315 considered tissue reactions, and not stochastic effects, findings of no effects at low doses are expected.

2316

2317

2318 **6. Study Quality**

2319

2320 The review of the current epidemiologic studies has identified various strengths and opportunities for  
2321 improvement, as briefly outlined below.

2322

2323 **6.1 Strengths**

2324

2325 Quantitative risk estimates, based on estimated individual doses, of cancer mortality or incidence have been  
2326 reported for nearly one million individuals with low dose rates and mostly low doses from studies of radiation  
2327 workers or those exposed to elevated environmental radiation levels (Shore *et al.*, 2017). The completion of the  
2328 million person study will considerably augment the available information (Boice, 2012a). The strongest studies  
2329 have high-quality follow-up, with a nearly complete ascertainment of cancer deaths or incidence. These studies  
2330 complement the LSS of atomic-bomb survivors with its high dose rate and high dose range. Although an historic  
2331 weakness of many worker and environmental radiation studies was inadequate dosimetry, in recent years  
2332 investigators have been focusing more on improving the quality and accuracy of the dosimetry. For example, the  
2333 million person study components (*e.g.*, atomic veterans, Rocketdyne, Mound), the INWORKS, Mayak, Techa River  
2334 and atomic-bomb studies have benefitted from devoting considerable resources to developing improved dosimetry,  
2335 and further improvements are underway. Several studies also have examined biodosimetry data (mainly  
2336 chromosome aberrations) for subsamples of the study subjects, and have broadly found validation of their  
2337 dosimetry. The dosimetry improvements are essential for accurate characterization of risks and assessment of the  
2338 LNT model at low doses and low dose rates.

2339

2340 Strengths of some of the large epidemiologic studies such as INWORKS and the LSS lie in the long follow-up  
2341 and large numbers of cancers and person-years at risk. The length of follow-up of epidemiologic studies is  
2342 particularly relevant since a large fraction of both spontaneous and radiation-related cancers occur after 60 y of  
2343 age. Nearly all studies have adjusted for potential confounding by gender, attained age and sometimes age at  
2344 exposure. There are also ongoing efforts to include other potentially explanatory variables such as life-style factors,  
2345 cancer predisposing factors and other sources of radiation exposure. Most of the studies reviewed here have  
2346 attempted to investigate potential sources of bias insofar as relevant data are available. A few studies examined  
2347 risks at restricted, lower-dose ranges to provide additional evidence regarding the applicability of a LNT model in  
2348 the dose range of greatest interest. Most notably, the LSS atomic-bomb survivors study of solid cancer (Grant  
2349 *et al.*, 2017) and the INWORKS study of all cancer except leukemia (Richardson *et al.*, 2015) found statistical  
2350 evidence for risk over the dose range of 0 to 100 mGy. A pooled analysis of studies of childhood external

2351 irradiation and thyroid cancer also showed a significant dose-response association over the dose range of 0 to  
2352 100 mGy (Lubin *et al.*, 2017).

2353

## 2354 **6.2 Opportunities for Improvement**

2355

2356 Although the worker studies and key environmental studies had individual dosimetry, a number of the other  
2357 studies briefly reviewed here did not (*e.g.*, most fallout studies), so they provided little or no information relevant to  
2358 the LNT model of risk. Most studies did not evaluate the effects of dose uncertainties on the risk estimates and their  
2359 confidence bounds. Several studies did not include concomitant exposures to neutrons or internal radionuclides in  
2360 their risk assessments.

2361

2362 For a few of the studies, a substantial fraction of the designated cancer cases were without histopathological  
2363 verification, and the completeness of enumeration of cancer or other diseases also was uncertain. Issues of possible  
2364 confounding effects of dose-dependent variations in health surveillance and access to medical care are also  
2365 considerations in several studies. Few studies have analyzed radiation risks with control for confounding by  
2366 lifestyle (*e.g.*, smoking) or other disease risk factors, and few studies currently have biological samples to evaluate  
2367 genetic or phenotypic biological factors in radiation risks. The studies of pediatric CT examinations are prone to  
2368 confounding in that the examination may have been performed because of a suspicion of cancer or a cancer-  
2369 predisposing condition; they also lacked dose estimates for individuals.

2370

2371 Important information for assessing the evidence for and against LNT is not currently available in some of the  
2372 studies reviewed here. This is partly because some study authors tend to ignore nonlinear models when there is not  
2373 enough statistical power to support both a linear parameter (which is usually applied first) and a second nonlinear  
2374 parameter in the same dose- response model. Consequently it is important to consider nonlinear one-parameter  
2375 dose-response models, such as pure quadratic in dose. Sometimes Excess Absolute Risk (EAR) models could  
2376 provide additional information relevant to assessing LNT, but not all studies have reported results with EAR  
2377 models. Attention should also be paid to temporal/age effects on the risk estimates, as the dose-response curve may  
2378 vary by age, time since exposure and birth cohort (Walsh and Schneider, 2016).

2379

2380

2381 **7. Strength of Support for LNT in Recent Epidemiologic Studies**

2382

2383 Support by studies for any model requires adequacy of the study components, which for epidemiologic studies  
2384 could be broadly characterized as adequacy of dosimetry, epidemiology and statistics. Dosimetric criteria include,  
2385 *e.g.*, personal vs. reconstructed dosimetric measurements; cumulative-dose range; accounting for neutron, internal  
2386 and medical exposures; validation of dosimetry; potential for bias in dosimetry; characterization and magnitude of  
2387 dose uncertainties; adequacy of dosimetry technology, calibration and administrative practices; and uncertainty in  
2388 dose reconstructions for individuals. Epidemiologic and statistical criteria include, *e.g.*, appropriateness of study  
2389 design and cohort definition; length of follow-up; completeness of follow-up and health outcome ascertainment,  
2390 equivalent across the dose range; sources and likelihood of bias; statistical power and precision; appropriate  
2391 statistical analysis method; sensitivity analyses; choices of mathematical models; types of adjustments for  
2392 explanatory covariables; and consideration of risk-effect modification. However, the listed criteria are not  
2393 exhaustive, as studies sometimes have unique characteristics that impinge on their ability to contribute information  
2394 regarding the LNT model.

2395

2396 For each component of the major studies, this report has critiqued both the methods used and the adequacy of  
2397 the results of those methods. To evaluate this, members of the committee made qualitative ratings of weak, medium  
2398 or strong for the adequacy of the dosimetry, epidemiology and statistics components of each of the major studies or  
2399 groups of studies of cancer risk after radiation exposure. Weak, medium and strong were scored as 1, 2, and 3, with  
2400 intermediate ratings (1.5 and 2.5) permitted; the scores were averaged across raters. The one exception to this was  
2401 that, for the dosimetry component only the ratings by the dosimetry team within the committee were considered  
2402 since they were more familiar with the details of the dosimetry than most of the remainder of the Committee. Many  
2403 of the implicit criteria that entered into the component ratings are given synoptically as study strengths or  
2404 weaknesses in Sections 6.1 and 6.2, although the items mentioned there are not exhaustive, as some unique features  
2405 may be important for a given study.

2406

2407 Table 7.1 summarizes the Committee's evaluation of the aforementioned components for 28 principal studies  
2408 or groups of studies of cancer risk. As a minimal criterion of study adequacy, 71 % of the studies had no  
2409 component on which they were scored as weak (score of 1). Fifty percent of the studies were scored moderate to  
2410 strong on all three components of evaluation.

2411

2412 The Committee further rated each study or group of studies on their strength of support for the LNT model  
2413 (Table 7.1). Eighty-two percent of the studies provided some support for the LNT model, including 18 % providing

2414 Table 7.1—Ratings of the quality of cancer studies reviewed and their degree of support for the LNT model.

Study (or groups of studies) <sup>a</sup>	Dosimetry <sup>b</sup>	Epidemiology <sup>b</sup>	Statistics <sup>b</sup>	Support for LNT Model <sup>c</sup>
Life Span Study (LSS), Japan atomic bomb (Grant <i>et al.</i> , 2017) <sup>d</sup>	3	3	3	4
INWORKS (UK, US, French combined cohorts) (Richardson <i>et al.</i> , 2015)	3	3	3	4
TB fluoroscopic examinations (Little and Boice, 2003)	2	3	2	4
Childhood A-bomb exposure (Preston <i>et al.</i> , 2008)	3	3	3	4
Childhood thyroid cancer studies (Lubin <i>et al.</i> , 2017)	3	3	3	4
Mayak nuclear facility (Sokolnikov <i>et al.</i> , 2015)	2	2	3	3
Techa River, nearby residents (Schonfeld <i>et al.</i> , 2013)	2	2	2.5	3
Chernobyl fallout, Ukraine and Belarus thyroid cancer (Brenner <i>et al.</i> , 2011; Zablotska <i>et al.</i> , 2011)	3	2	2	3
Childhood breast cancer studies (Eidemüller <i>et al.</i> , 2015)	2	3	3	3
In utero A-bomb exposure (Preston <i>et al.</i> , 2008)	2	3	3	3
In utero exposures, medical (Wakeford, 2008)	1	2	2	3
Canadian nuclear workers (Zablotska <i>et al.</i> , 2013b)	2.5	3	3	3
Japanese nuclear workers (Akiba and Misuno, 2012)	2.5	2	3	2
Chernobyl cleanup workers, Russia (Kashcheev <i>et al.</i> , 2015)	1	1.5	2	2
US radiologic technologists (Liu <i>et al.</i> , 2014; Preston <i>et al.</i> , 2016) <sup>e</sup>	1	2	2	2
Mound facility (Boice <i>et al.</i> , 2014)	2	1.5	1.5	2
Rocketdyne facility (Boice <i>et al.</i> , 2011)	2	2	2	2
Medical x-ray workers, China (Sun <i>et al.</i> , 2016)	1.5	1.5	2	2

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Background radiation levels and childhood leukemia (Kendall <i>et al.</i> , 2013)	1.5	2	2	2
Taiwan radiocontaminated buildings, residents (Hwang <i>et al.</i> , 2008)	2	1.5	1.5	2
Pediatric CT examinations (Pearce <i>et al.</i> , 2012)	1	1.5	1.5	2
Childhood leukemia studies (Wakeford and Little, 2003)	1	2	1.5	2
In utero exposures, Mayak and Techa (Akleyev <i>et al.</i> , 2016)	1	1.5	2	2
US atomic veterans (Beck <i>et al.</i> , 2017)	3	3	3	1
Kerala, India, high natural background radiation area (Nair <i>et al.</i> , 2009)	2	2	1.5	1
Yangjiang, China, high natural background radiation area (Tao <i>et al.</i> , 2012)	1.5	1	1	1
Fallout studies (aggregate of 8 studies) (Lyon <i>et al.</i> , 2006) <sup>f</sup>	1.5	1	1.5	1
Hanford <sup>131</sup> I fallout study (Davis <i>et al.</i> , 2004)	2	3	1.5	1

2415 <sup>a</sup> Studies excluded: studies of risks around nuclear sites (no dosimetry and extremely low exposures); the 15-country worker study  
 2416 (superseded by the INWORKS study); the Million Worker Study is not yet completed, but published components of it were evaluated  
 2417 separately; studies of genetic effects (since no human heritable risks have been shown); studies of tissue reactions (“deterministic effects”,  
 2418 because these are generally not believed to be LNT).

2419 <sup>b</sup> Judged quality of the Dosimetry, Epidemiology and Statistics scores: 1 = weak, 1.5 = weak-moderate, 2 = moderate, 2.5 = moderate-  
 2420 strong, 3 = strong.

2421 <sup>c</sup> Ratings of the Support for LNT: 1 = essentially no support (null, unreliable or inconclusive), 2 = weakly moderate support, 3 = moderate  
 2422 support, 4 = strong support.

2423 <sup>d</sup> A representative recent publication is listed for each study.

2424 <sup>e</sup> The dosimetry used in the Preston *et al.* (2016) study of breast cancer, based on the Simon *et al.* (2014) dosimetry, is significantly  
 2425 improved and would be rated “3” compared to the dosimetry that was used in other published epidemiologic analyses of this cohort.  
 2426 However, since little other epidemiology has been published using the new dosimetry, for the purposes of this report these studies are  
 2427 limited in their support for LNT.

2428 <sup>f</sup> Fallout studies were included as a group (they mostly had little or poor dosimetry and many were studies of aggregates rather than  
 2429 individuals; however, the Hanford <sup>131</sup>I fallout study was of better quality, so was identified separately).

2430

2431

2432 strong support and 25 % rated as providing moderate support. A rating of moderate versus strong support for LNT  
2433 sometimes hinged upon the size of the study or some other limitation, and not on indications of nonlinearity. The  
2434 studies that provided no support for the LNT model were either totally null studies, or ones with no dose-response  
2435 analysis or excessively unreliable data (“Inconclusive”). It should be noted that all the studies being considered,  
2436 except for the Life Span Study of atomic-bomb survivors, had either exposures at low dose rates or multiple small  
2437 exposures. Furthermore, the preponderance of study subjects had cumulative doses under 100 mGy. Thus these  
2438 studies are very relevant for contemporary radiation protection concerns.

2439  
2440 One question that could be raised is whether LNT underestimates the radiation risk of cancer. If the LNT  
2441 model had grossly underestimated cancer risk there would not have been a reasonably consistent picture of dose-  
2442 response slopes that were compatible with the LNT model, as was seen. Nonetheless, given the wide confidence  
2443 intervals in many studies, it is possible that LNT might underestimate risk, but not to any great extent. In fact, on  
2444 balance a number of the studies presented here suggest that a DREF >1 (*i.e.*, somewhat less risk at low dose  
2445 rates) may be appropriate. Similarly, for a linear-quadratic response, the curve at low doses will still be a linear  
2446 nonthreshold one but with a shallower slope than the overall curve at higher doses.

2447

2448 **8. Future Directions**

2449

2450 This section discusses several areas with regard to knowledge needs and opportunities for future  
2451 research.

2452

2453 **8.1 Epidemiology**

2454

2455 **8.1.1 Atomic-Bomb Studies**

2456

2457 A number of advances in the Japanese atomic-bomb studies potentially can enhance our understanding of the  
2458 magnitude of risk associated with acute radiation exposures and the shapes of dose-response curves. Some that are  
2459 briefly outlined below represent a continuation of present directions and others are suggestions for new advances.

2460

2461 Improved organ doses. Organ doses in the atomic-bomb studies have been calculated using mathematical  
2462 phantoms developed in the early 1980s and using only three prototypic ages (infant, child, adult) and a limited  
2463 set of organs (*e.g.*, not including lens of the eye, heart, oral cavity, esophagus, kidney, rectum, prostate). The  
2464 technology for organ dose estimation has improved significantly since then, and a more detailed set of ages,  
2465 organs and other features is needed (Cullings, 2012). Until now, there has not been adequate dosimetry for the  
2466 fetus, where the size and organ placement varies substantially by fetal age. This should be addressed as part of a  
2467 new organ dosimetry project. Adequate dosimetry for the heart and the lens of the eye (for neutron and gamma,  
2468 by posture and orientation vis a vis the bomb) is needed to improve the individual dose estimates.

2469

2470 Clarification of the dose-response shape and low-dose effects. Recent publications (Grant *et al.*, 2017; Ozasa  
2471 *et al.*, 2012) have indicated more curvature in the dose response for all solid cancer and differences between males  
2472 and females in the dose-response shape. Further evaluations may help clarify current discrepant views of the  
2473 magnitude of the neutron RBE (Cullings *et al.*, 2014; Sasaki *et al.*, 2016; Walsh, 2013). A detailed analysis of  
2474 possible factors involved in the curvature needs to be undertaken, *e.g.*, urban/rural differences, socioeconomic  
2475 factors, close examination of possible selection factors for males (*e.g.*, health related selection because of the  
2476 wartime male military draft), or quality of dosimetry for subsets of males in different shielding situations. Males  
2477 could have different background risks from females for the influential cancers, and the interaction (or lack of it)  
2478 between these background risk factors and radiation could influence the dose responses in different ways (as could  
2479 confounding). The low-dose data need to be examined in more detail, using existing and new statistical methods  
2480 and analytic strategies. This applies not only to all solid cancer or to leukemia, but to specific cancers or cancer  
2481 groups, cardiovascular diseases, and various clinical health endpoints.

2482

2483 Evaluation of LNT for various organs or organ systems. Evaluate whether dose-response LNT is similar for  
2484 tumors of various organs or organ systems, insofar as statistical limitations permit. This will provide evidence  
2485 regarding the generality of the LNT model and the need for a low-dose effectiveness factor (LDEF) across tumor  
2486 sites. Similarly, the new ophthalmological examination data currently being obtained need a careful evaluation of  
2487 dose-response shape, as well as analysis by age, opacity grade, etc.

2488

2489 Risk at young ages (or fetal) at exposure. The mortality data and cancer incidence data for those exposed in  
2490 childhood and adolescence are accruing extensively as they reach older ages, so there is increasing potential to  
2491 learn more about risks following early-life exposure to radiation. Also, the most recent publication of solid cancer  
2492 risk following *in utero* and early-childhood exposure yielded an intriguing result that risk following childhood  
2493 exposure continues to increase over time, while it does not after *in utero* exposure (Preston *et al.*, 2008). However,  
2494 the difference between the two did not attain statistical significance. An update of those results is needed to clarify  
2495 whether or not the difference in temporal trends continues and whether the LNT model applies to *in utero* exposure.

2496

2497 F1 generation risk. The risk in the F1 cohort with respect to parental irradiation has been purely null to date for  
2498 cancer and noncancer mortality, for cancer incidence, and for clinical indications of cardiovascular diseases and  
2499 diabetes. However, the cohort was not yet 60 y old on average, so 30 more years of observation are needed to see  
2500 the full expression of disease and determine if any risk is present. Even though the numbers would likely be small,  
2501 it also would be valuable to publish information on the occurrence of diseases with a strong genetic component  
2502 (*e.g.*, rheumatoid arthritis) in F1 clinical study subjects in relation to radiation exposure.

2503

2504 Bioindicators for cancer and CVD outcomes. The large bank of blood and tissue samples needs to be exploited  
2505 more robustly by the biomedical community to evaluate genotypic and phenotypic alterations associated with both  
2506 radiation exposure and disease risk, to thereby discover key events related to adverse outcome pathways that  
2507 mediate between radiation and disease development. It is also recommended that the studies of chromosome  
2508 translocations among those exposed *in utero* or at a young age be continued. If the existing *in utero* translocation  
2509 findings are an indication of a drop-off of risk at moderate (rather than high) doses (Ohtaki *et al.*, 2004), and if this  
2510 extends after birth, then the usual extrapolation of risk from moderate-to-high doses to low doses might be  
2511 modified for risk of leukemia, and possibly other cancers, after low-dose exposures at prenatal/early-childhood  
2512 ages.

2513

2514 **8.1.2** *Radiation Worker Studies*

2515

2516       Recent large-scale studies of radiation workers have started to make meaningful estimates of risks arising from  
2517 protracted exposures to many low doses. It is important to continue to have accurate linkage to high quality cancer  
2518 and mortality registries. Much of the statistical power of these studies emanates from those workers who have  
2519 accumulated moderate-to-high doses over many years, and most of these workers will have started work in the  
2520 early years of the nuclear industry. Therefore, the continuing follow-up of worker cohorts to obtain as much  
2521 information as possible from high cumulative-dose workers is desirable, and should be pursued. International  
2522 collaboration permitting the combining of worker cohorts, particularly those with reasonable numbers of high  
2523 cumulative dose workers, is another valuable way of increasing the power of worker studies. Unfortunately, the  
2524 earlier workers who will tend to have received the highest doses also are likely to have the greatest uncertainty  
2525 attached to their doses because dose recording technologies, possibly combined with dose record keeping practices,  
2526 were less advanced in the earlier years when doses tended to be highest. Therefore, careful scrutiny of dose records  
2527 is necessary to identify any deficiencies in recorded doses. For example, during the 1950s the design of shielding  
2528 of early reactors may have permitted neutron streaming into workspaces, but the inadequacy of personal neutron  
2529 dosimeters to accurately measure intermediate and fast neutrons during this period may have led to underestimated  
2530 neutron doses, or even the absence of recorded neutron exposures when neutron doses had been received. It is  
2531 desirable to make use of all available dose information in the worker studies so as to identify inadequacies in dose  
2532 records and to compensate for these wherever possible through, for example, dose reconstruction methods (see  
2533 further discussion in Section 7.2). Reliable risk estimates depend, inter alia, upon reliable dose estimates, and this  
2534 is an area that should be pursued. In addition, a further desirable refinement is the calculation of organ/tissue-  
2535 specific doses for workers. Epidemiologic studies of medical workers pose problems in estimating doses because  
2536 of factors such as partial-body irradiation, low photon energies, location of the dosimeter, use or non-use of the  
2537 apron or other protective devices, missed dose etc.

2538

2539 **8.1.3** *Environmental Radiation Studies*

2540

2541       Several improvements in the Techa River study are suggested: the new forthcoming dosimetry (TRDS-2017)  
2542 will improve several aspects of the dose estimation; the shape of the dose response for solid cancer can be  
2543 evaluated more fully in the low dose range; the potential for dose-related screening bias should be evaluated and,  
2544 if possible, adjusted for; efforts should be made to improve the completeness of follow-up and the rate of  
2545 histological verification of cancers. The Chernobyl <sup>131</sup>I studies of thyroid cancer should focus more fully on the  
2546 lower dose part of their dose-response curves. The Kerala and Yangjiang studies need to increase efforts to  
2547 improve their cancer ascertainment and diagnosis, and to closely examine the impact that low income/education

2548 and distance from the principal cancer facilities may have on cancer ascertainment rates. Further carefully  
2549 designed validation studies of reconstructed doses by personal dosimetry measurements would also be valuable  
2550 for these studies. All of the environmental study groups should seek to implement measures to reduce individual  
2551 dose uncertainties. (See further discussion in Section 6.2). Collection and storage of biospecimens from  
2552 strategically defined subgroups may also be useful for future biodosimetry, molecular epidemiology and  
2553 bioindicator studies.

2554

#### 2555 **8.1.4** *Computed Tomography Studies*

2556

2557 Within the next few years several major reports of the associations between pediatric CT examinations and  
2558 subsequent cancer outcomes are expected to be published, including reports on cohorts in Canada, Australia and the  
2559 EPI-CT study in Europe. Among other things, these studies will need improvements in the dosimetry compared to  
2560 existing reports, including use of individual organ/tissue doses, more accurate characterization of historical CT  
2561 doses, and accounting for all past CT examination doses. Appropriate incorporation of minimum latency periods for  
2562 cancer induction is also needed. Avoiding bias: there is a fundamental need to fully disentangle radiation effects  
2563 from the reasons why the CT examination was given, *i.e.*, confounding by indication (triggered by conditions that  
2564 predispose to cancer) and reverse causation (pre-existing, but as yet undetectable, malignancy as the reason for the  
2565 CT scan). However, owing to limitations in existing individual medical records, that may not be possible unless the  
2566 impact of all of the potential sources of biases and missing sources of data are carefully assessed (Bosch de Basea  
2567 *et al.*, 2015). The identification of children who were not scanned for some pre-existing medical reason, such as  
2568 children who are passengers in vehicles involved in collisions, is desirable to overcome confounding by indication  
2569 and reverse causation, but this may be difficult with available records, especially for historically early exposures  
2570 when the doses received from a scan tended to be highest.

2571

#### 2572 **8.1.5** *Childhood and In Utero Exposures*

2573

2574 The latest report on the cancer incidence of Japanese atomic-bomb survivors who were exposed prenatally  
2575 or during childhood (Preston *et al.*, 2008) considered cancer only through 1999 and needs an analysis of  
2576 updated data. They found some indication ( $p = 0.09$ ) of upward curvature in the dose response for all solid  
2577 cancer and marginally greater risk among those exposed during childhood than prenatally, which further data  
2578 may clarify. Further follow-up in the Techa River and Chernobyl studies also can yield greater statistical  
2579 power to evaluate dose- response associations for cancer endpoints.

2580

2581 It is possible that future studies could also evaluate available data from radiation therapy patients such as those  
2582 involved in the childhood cancer survivor studies (<https://www.cancer.gov/types/childhood-cancers/ccss>) if more  
2583 accurate doses outside of the treatment fields can be ascertained (AAPM, 2017; Xu *et al.*, 2008) and information  
2584 on secondary malignancies can be assessed, along with the obvious confounding factors. NCRP has previously  
2585 evaluated second primary cancers and cardiovascular disease after radiation therapy in NCRP Report No. 170  
2586 (NCRP, 2011) with such recommendations.

2587

#### 2588 **8.1.6** *Studies of Tissue Reactions*

2589

2590 In general, a greater understanding of the mechanisms for specific tissue reactions is needed, along with  
2591 additional, quality epidemiologic data. With regard to the effects of ionizing radiation on the lens, while the  
2592 currently available information has provided input on appropriate guidance with regard to radiation protection,  
2593 much more work is needed to develop a complete understanding of such detriments. For the lens of the eye, NCRP  
2594 (2016) recommends ongoing evaluation and additional research in the following areas: comprehensive evaluation  
2595 of the overall effects of ionizing radiation on the eye; dosimetry methodology and dose-sparing optimization  
2596 techniques; additional high-quality epidemiology studies; medical countermeasures; and a basic understanding of  
2597 the mechanisms of cataract development. Similarly, for circulatory effects, it will be important to develop an  
2598 understanding of the mechanisms for the plethora of associated diseases and to carefully address confounding  
2599 factors in epidemiologic studies (especially with regard to lifestyle factors).

2600

### 2601 **8.2 Dosimetry: Future Directions**

2602

2603 The dosimetry review highlighted several recommendations about future directions of studies focusing on dose  
2604 response. Clearly, individual doses based on personal monitoring technology, bioassay data, and measurements  
2605 characterizing exposure scenarios provide the highest quality dose estimates. Dosimetrists must use the information  
2606 available and address missing information through dose reconstruction and by taking into account uncertainties.  
2607 When opportunities exist to plan dosimetry for new epidemiologic studies, the use of measurement data should be  
2608 the first recourse, if available.

2609

2610 Ideally, uncertainties should be provided along with estimates of dose, and the dose uncertainties should be  
2611 used to adjust risk coefficients and confidence intervals (Kwon *et al.*, 2016; Stram *et al.*, 2015; UNSCEAR, 2015;  
2612 Zhang *et al.*, 2017). Both individual and shared uncertainty can result in the width of uncertainty bounds of the  
2613 risk estimates being significantly underestimated. Dosimetry upgrades to the Techa River and Mayak dosimetry  
2614 are expected to provide better estimates of doses and multiple realizations of the doses that can be used to estimate

2615 shared uncertainty and correct the error bounds of the estimated dose response. Future dose response analyses  
2616 should include dose uncertainties in the analysis of ERR Gy<sup>-1</sup>.

2617  
2618 Most studies reviewed focused on external exposures from gamma radiation. Some of these studies also had  
2619 components of external dose from neutrons and from internal exposures. The general practice followed in many  
2620 studies has been to dismiss the dose from neutrons and internal dose because of the complexity and difficulty of  
2621 estimating their contribution to dose. More effort should be given to exploring the impact of doses from neutrons  
2622 and internal exposure.

2623  
2624 In the future, better integration of the dosimetry and epidemiology both in the study design and implementation  
2625 of the study should be achieved. When dosimetrists work closely with the epidemiologists, as in the case of the  
2626 Atomic Veteran Study (Beck *et al.*, 2017; Till *et al.*, 2014) or the Techa River and Mayak studies (Fountos, 2016;  
2627 Napier, 2014), doses and related uncertainty provided will more likely be in a form that is most useful for the dose  
2628 analysis, and both groups will have a better understanding of the strengths and weaknesses of the dosimetry with  
2629 respect to estimating a dose response. Integration of the dosimetry team with the epidemiologic team from the  
2630 conceptual design through the implementation and documentation of any analysis of dose response is critical to the  
2631 success of the effort.

2632

### 2633 **8.3 DDREF: Future Directions**

2634

2635 As noted above (Section 8.1.1), cancer risk estimates obtained in the LSS for survivors of the Japan atomic  
2636 bombs are driven by medium-to-high, acute exposures. To predict effects at low doses and low dose rates, an  
2637 extrapolation model (LNT) has been used together with a DDREF that adjusts the slope of the linear curve. The  
2638 value of this DDREF used for radiation protection purposes ranges from about 1.5 (*e.g.*, NA/NRC, 2006) to 2 (*e.g.*,  
2639 ICRP, 2007). The use of a DDREF has been debated extensively and the issue remains to be resolved. Further, it  
2640 has been proposed that a separation of DDREF into a LDEF and a DREF more accurately reflects the extrapolation  
2641 approach used (Rühm *et al.*, 2015).

2642

2643 Without a LDEF the LNT model would be a simple linear extrapolation from medium-to-high dose to low  
2644 dose. If a linear-quadratic curve is deemed to better fit the LSS cancer data (as has been proposed), then a  
2645 shallower linear coefficient based on the low-dose cancer frequencies would be appropriate, which implies the  
2646 incorporation of an LDEF. A recent paper by Haley *et al.* (2015) described a comprehensive analysis of 15 animal  
2647 life span studies for acute and protracted exposures. The conclusions were that a linear-quadratic curve did not  
2648 provide the best fit to the data sets (though there were few low-dose data) and that a direct comparison of data from

2649 acute and protracted exposures was the preferred approach. It is considered that the question of applying an LDEF  
2650 remains an open one because of a lack of pertinent data from human studies or animal and cellular studies.

2651  
2652 Nevertheless, a DREF may be needed in the current risk estimation process, although the effect of dose rate on  
2653 cancer risk has not been firmly established. However, a recently propounded view is that there is no dose rate  
2654 effect per se for total cancers (*e.g.*, SSK, 2014). UNSCEAR handled this dilemma by proposing that the dose-  
2655 response curve for cancers followed a linear-quadratic (LDEF; or a linear-quadratic-exponential) model, such that  
2656 at low doses, where the linear component prevailed, there would be a DREF of 1 corresponding to the low-dose  
2657 slope (UNSCEAR, 2008). A recent meta-analysis of solid cancers in worker and environmental studies with low  
2658 dose rates found suggestive discrepancies in DREF estimates, depending on the inclusion/exclusion of two large  
2659 studies (Shore *et al.*, 2017); future improvements in dosimetry and longer follow-up may reduce uncertainty in the  
2660 value of DREF from epidemiologic studies.

2661  
2662 It appears to be appropriate to separate the DDREF into its component parts, LDEF and DREF, for subsequent  
2663 deliberations. It is recommended that a combined approach using low dose and low dose-rate epidemiology data  
2664 together with information from animal experiments and informative bioindicators collected with human and  
2665 animal models be employed to evaluate such factors (NCRP, 2015).

#### 2667 **8.4 Key Events, Bioindicators and Risk Assessment: Future Directions**

2668  
2669 For estimating the cancer risks from exposure to environmental chemicals, the U.S. Environmental Protection  
2670 Agency (EPA) has developed a mechanistic approach, in particular because human data are largely unavailable  
2671 (EPA, 2005). More recently, this process has been enhanced through the inclusion of a so-called Adverse Outcome  
2672 Pathway (AOP)/Key Event approach that incorporates data on key events as parameters in a Biologically-Based  
2673 Dose- Response (BBDR) Model (Edwards *et al.*, 2016). An AOP “is a conceptual construct that portrays existing  
2674 knowledge concerning the linkage between a direct molecular initiating event (*e.g.*, a molecular interaction between  
2675 a xenobiotic and a specific biomolecule) and an adverse outcome at a biological level of organization relevant to  
2676 risk assessment” (Ankley *et al.*, 2010). In this context, a Key Event is an empirically observable step, which is a  
2677 necessary element of the mode of action critical to the outcome (*i.e.*, necessary, but not necessarily sufficient in its  
2678 own right); key events are measurable and reproducible (Meek *et al.*, 2014). The characterization of AOPs and their  
2679 associated key events is likely to require a concerted effort. However, it is a feasible task as demonstrated by the  
2680 ongoing effort of the Organisation for Economic and Co-operation and Development (OECD, 2015). For radiation-  
2681 induced adverse health outcomes, a clear need is to design research programs that are targeted towards a risk  
2682 assessment framework that includes the identification of key events along the pathway to disease. Such key events

2683 have been described as specific bioindicators of effect, in contrast to biomarkers that are surrogates for exposure  
2684 (*e.g.*, chromosome alterations in peripheral blood) or that suggest an enhanced likelihood of disease outcome in a  
2685 qualitative sense (*e.g.*, molecular alterations in nontarget tissues). Ultimately, the set of bioindicators that define the  
2686 pathway from normal to malignant can be used for developing a specific BBDR model.

2687  
2688 **Recommendation:** It is recommended, that in planning for new radiation biomarker studies, further consideration  
2689 be given to using an adverse outcome pathway/key events approach to aid in the integration of epidemiology and  
2690 radiation biology (NCRP, 2015). The aim will be to reduce uncertainty in risk assessments for low doses and low  
2691 dose rates.

2692

### 2693 8.5 Other Future Direction Recommendations

2694

2695 Better characterization of doses to organs/tissues remote from sites of localized radiotherapy might permit a  
2696 variety of new studies of radiation effects in the low-to-moderate dose range (Mazonakis *et al.*, 2016; Xu *et al.*,  
2697 2008), though characterizing the doses may be difficult.

2698

2699 Another area is background radiation. Continuing with the Kerala, India and Yangjiang, China HBRA studies is  
2700 useful, but accurate dosimetry, complete and accurate diagnosis and cancer registration, and the elimination of  
2701 confounding require more attention. Large numbers of cases are required (hundreds to thousands) to obtain  
2702 reasonable power. Opportunities could be sought to link large historical registries of childhood cancer with  
2703 dosimetric databases in regions where there is substantial variation in the ambient radiation exposure levels.  
2704 Because the Kerala and Yangjiang studies are relatively recent and limited in numbers, large numbers of cancers are  
2705 unlikely (*e.g.*, only ~30 cases of leukemia in the Kerala study and 15 in the Yangjiang study), so these studies  
2706 would have to be conducted in other areas with well-established registries and sufficiently high doses. Such areas  
2707 have not been identified at this time.

2708

2709 Long term opportunities for improvement include encouraging more widespread bio-banking in connection  
2710 with epidemiologic studies. Information on new and as yet undiscovered biomarkers of radiation risk (rather than  
2711 of exposure) of cancer or cardiovascular endpoints need to be explored as potential mediators or modifiers of  
2712 radiation effects. Such studies could increase our understanding of the biology of radiation risk and have potential  
2713 implications for personalized medicine. These could eventually be built into the statistical analysis of cancer risk at  
2714 low doses. Analyzing epidemiologic data in conjunction with relevant radiobiological concepts and data also has  
2715 the potential to provide insights into LNT that go beyond those gained from merely analyzing the empirical  
2716 epidemiologic data in isolation (Eidemüller *et al.*, 2015; Kaiser *et al.*, 2016).

2717 **9. Conclusions**

2718

2719 **9.1 Overall Conclusion on Use of the LNT Model**

2720

2721 While the ongoing development of science requires a constant reassessment of prior and emerging evidence  
2722 to assure that the approach to radiation protection is optimal, though not necessarily perfect, NCRP concludes  
2723 that, based on current epidemiologic data, the LNT model (perhaps modified by a DDREF) should continue to be  
2724 utilized for radiation protection purposes. This is in accord with judgment by other national and international  
2725 scientific committees, based on somewhat older data than in the present report (ICRP, 2007; NA/NRC, 2006;  
2726 UNSCEAR, 2008), that no alternative dose-response relationship appears more pragmatic or prudent for  
2727 radiation protection purposes than the LNT model.

2728

2729 **9.2 Supporting Conclusions**

2730

2731 This report has examined the evidence for or against the appropriateness of using LNT as a practical approach  
2732 for managing radiation exposures to individuals. It is important to point out that there may be a DDREF involved  
2733 that is greater than one, so that the LNT does not imply a single straight-line proportionality of effects from high,  
2734 acute doses to low doses and/or low dose rates. Rather, low-dose or low dose-rate cumulative exposures may have  
2735 a shallower linear slope than seen for high, acute doses.

2736

2737 Because individual studies with low doses (less than 100 mGy) almost inevitably have relatively low statistical  
2738 power, the findings for radiation and solid cancer are often not statistically significant. Furthermore, studies may  
2739 have sampling variation or confounding by other exposures (*e.g.*, smoking or other lifestyle factors) which can  
2740 diminish the consistency or validity of findings. Nevertheless, most large and high quality low- dose studies show  
2741 positive risk coefficients (Shore *et al.*, 2017), suggesting there may be cancer effects at low doses, which is  
2742 consistent with, though not necessarily proving, the applicability of the LNT model for radiation protection.  
2743 However, it should be recognized that, the risk of cancer at low doses is small.

2744

2745 The most recent epidemiologic studies show that the assumption of a dose-threshold model is not a prudent  
2746 pragmatic choice for radiation protection purposes. The consistency of the better-designed and larger studies with  
2747 dose-response functions that are essentially linear or linear-quadratic, argues for some risk at low doses. Some  
2748 studies explicitly found risk in the dose range of 100 mGy or less, *e.g.*, the atomic-bomb survivor studies, the  
2749 INWORKS worker study, and the pooled radiation and thyroid cancer analysis. Several studies also performed

2750 explicit dose-threshold analyses and found the estimates of dose thresholds to be compatible with zero dose (*i.e.*,  
2751 no threshold).

2752  
2753 The data regarding noncancer effects at low doses—cardiovascular diseases, cataracts, thyroid dysfunction,  
2754 central nervous system effects—are mixed or null, suggesting at this time that an LNT assumption for radiation  
2755 protection purposes for noncancer effects is not appropriate.

2756

### 2757 **9.3 Radiation Protection Implications**

2758

2759 While the LNT model is an assumption that likely cannot be scientifically validated by radiobiologic or  
2760 epidemiologic evidence in the low-dose range, the preponderance of epidemiologic data is consistent with the  
2761 LNT assumption, although there are a few notable exceptions. The current data are not precise enough to  
2762 exclude other models, and there appears to be curvature in some datasets. The current judgment by national and  
2763 international scientific committees is that no alternative dose-response relationship appears more pragmatic or  
2764 prudent for radiation protection purposes than the LNT model on the basis of available data, recognizing that  
2765 the risk below 100 mGy is uncertain but small (ICRP, 2007; NA/NRC, 2006; UNSCEAR, 2008).

2766

2767 When discussing the LNT model for radiation protection purposes UNSCEAR (2015) notes that “[t]his and the  
2768 other dose-response relationships are plausible but currently none of them are definitively verifiable and therefore  
2769 cannot be deemed proven or disproven.” They further point out that while some have argued in support of a  
2770 practical threshold for management of the risk of radiation-induced cancer, epidemiology alone will not be able to  
2771 resolve the issue of whether there are dose thresholds for radiation risks, thus also supporting the need for further  
2772 integrated radiobiology and epidemiology research (NCRP, 2015). Indeed, the LNT model, “...has practical  
2773 advantages for radiation protection purposes in order to derive nominal radiation risk coefficients for a  
2774 ‘representative population’ ...[and are] coherent with radiobiological knowledge, epidemiologic information, and  
2775 incorporate ethical judgments on the relative harm of different health effects” (UNSCEAR, 2015).

2776

2777 NCRP has carefully assessed the most currently available epidemiologic evidence and concludes that the LNT  
2778 model (perhaps modified by a DDREF) should continue to be used for radiation protection purposes.

2779

2780 **Glossary**

2781

2782 **acute radiation exposure:** Radiation exposure received during a short time period (*e.g.*, hours).

2783 **angiography:** The radiographic visualization of blood vessels following introduction of contrast material.

2784 **as low as reasonably achievable (ALARA):** A principle of radiation protection philosophy that requires  
2785 that exposures to ionizing radiation be kept as low as reasonably achievable, economic and societal factors  
2786 being taken into account. The ALARA principle is satisfied when the expenditure of further resources would  
2787 be unwarranted by the reduction in exposure that would be achieved.

2788 **cataract:** A cataract is a clouding or opacification that occurs in the normally clear lens of the eye. Some  
2789 cataracts are clinically unimportant and do not impair vision in any way. But, without intervention, cataracts  
2790 remain the most common cause of blindness.

2791 **confidence interval (CI):** A measure of the extent to which an estimate of risk, dose or other parameter is expected  
2792 to lie within a specified interval (*e.g.*, a 95 % confidence interval of a risk estimate means that, based on  
2793 available information, the probability is 0.95 that the true but unknown risk lies within the specified interval).

2794 **cornea:** The transparent epithelial structure forming the anterior part of the external covering of the eye.

2795 **cortical cataract:** Opacification that appears in the lens cortex. Associated with induction by both solar and  
2796 ionizing radiation.

2797 **deoxyribonucleic acid (DNA):** Genetic material of cells; a complex molecule of high molecular weight consisting  
2798 of deoxyribose, phosphoric acid, and four bases which are arranged as two long chains that twist around each  
2799 other to form a double helix joined by hydrogen bonds between the complementary components.

2800 **deterministic effects:** Detrimental health effects for which the severity varies with the dose of radiation (or other  
2801 toxic substance), and for which a threshold usually exists (*i.e.*, causally determined by preceding events). ICRP  
2802 Publication 103 has restated this as: “Injury in a population of cells, characterized by a threshold dose and an  
2803 increase in the severity of the reaction as the dose is increased further. In some cases, deterministic effects are  
2804 modifiable by post-irradiation procedures including biological response modifiers.” It is common for  
2805 deterministic effects to be termed tissue reactions.

2806 **detriment:** Measure of stochastic effects from exposure to ionizing radiation that takes into account the  
2807 probability of fatal cancers, probability of severe hereditary effects in future generations, probability of  
2808 nonfatal cancers weighted by the lethality fraction, and relative years of life lost per fatal health effect.

2809 **dose:** General term denoting the mean energy imparted from ionizing radiation to a tissue or organ from either an  
2810 external source or from radionuclides in the body. When unspecified, dose refers to the quantity of absorbed  
2811 dose, measured in gray (1 Gy = 1 J. kg<sup>-1</sup>) or rad (1 rad  
2812 = 100 ergs g<sup>-1</sup>). Depending upon the context in which it is used, the generic term dose may also refer to  
2813 equivalent dose, effective dose or other dose-related quantities.

2814 **dose limit:** A limit on radiation dose that is applied for exposure to individuals in order to prevent the occurrence  
2815 of radiation-induced deterministic effects or to limit the probability of radiation-induced stochastic effects to an  
2816 acceptable level.

2817 **dose rate:** Dose per unit time; often expressed as an average over some time period (*e.g.*, a year).

2818 **dosimetry:** The science or technique of determining radiation dose.

2819 **effective dose:** The sum of the weighted equivalent doses for the radiosensitive tissues and organs of the body.  
2820 Each equivalent dose is modified by a tissue weighting factor that takes into account the relative radiation  
2821 detriment for the tissue or organ. The tissue weighting factor for a particular tissue or organ represents the  
2822 fraction of the total radiation detriment to the whole body attributed to that tissue when the whole body is  
2823 irradiated uniformly. The tissue weighting factors have been developed from a reference population of equal  
2824 numbers of both sexes and a wide range of ages. A similar quantity is effective dose equivalent (an earlier  
2825 formulation of effective dose) that is also the sum of weighted doses for the radiosensitive tissues and organs of  
2826 the body. These weighted doses (called dose equivalents) were also modified by tissue weighting factors (but  
2827 an earlier set of factors different than used for effective dose). The SI unit of effective dose (and effective dose  
2828 equivalent) is  $\text{J. kg}^{-1}$  with the special name sievert (Sv);  $1 \text{ Sv} = 1 \text{ J. kg}^{-1}$  (see equivalent dose).

2829 **electrons:** Subatomic charged particle. Negatively charged particles are parts of atoms. Both negatively and  
2830 positively charged electrons may be expelled from a radioactive atom when it disintegrates.

2831 **equivalent dose:** A quantity used for radiation protection purposes that takes into account the different  
2832 probabilities of stochastic effects that occur with the same absorbed dose delivered by radiations with different  
2833 radiation weighting factors (the factor by which the mean absorbed dose in a tissue or organ is modified to  
2834 account for the type and energy of radiation in determining the probability of stochastic effects). The SI unit of  
2835 equivalent dose is joule per kilogram ( $\text{J. kg}^{-1}$ ), with the special name sievert (Sv) (see effective dose).

2836 **excess absolute risk:** The excess risk attributed to exposure and usually expressed as the arithmetic difference  
2837 between the incidence or mortality rate of disease among those exposed and that obtained in the absence of  
2838 exposure. The resultant risk coefficient is usually normalized to a population base of 10,000 people and is  
2839 expressed as the number of excess cases per 10,000 persons per gray (organ dose) per year at risk [*i.e.*,  
2840  $(10^4 \text{ PY Gy})^{-1}$ ]. Absolute risk also may be presented on a lifetime (70 y) basis.

2841 **excess relative risk (ERR):** Proportional excess risk above and beyond the baseline risk, defined as the relative  
2842 risk (RR) minus one ( $\text{ERR} = \text{RR} - 1$ ). It is usually stated as ERR per unit dose, *e.g.*,  $\text{ERR Gy}^{-1}$ , and is derived  
2843 using a dose-response regression analysis.

2844 **exposure:** Most often used in a general sense meaning to be irradiated. When used as the specifically defined  
2845 radiation quantity, exposure is a measure of the ionization produced in air by x or gamma radiation. The unit  
2846 of exposure is coulomb per kilogram ( $\text{C kg}^{-1}$ ). The special unit for exposure is roentgen (R), where  $1 \text{ R} =$   
2847  $2.58 \times 10^{-4} \text{ C kg}^{-1}$ .

- 2848 **fluoroscopy (fluoro):** The process of producing a real-time image using x rays. The machine used for  
2849 visualization, in which the dynamic image appears in real time on a display screen is a fluoroscope.
- 2850 **fractionation:** The delivery of a given total dose of radiation as several smaller doses, separated by intervals of  
2851 time.
- 2852 **gamma radiation:** Electromagnetic radiation emitted in de-excitation of atomic nuclei, and frequently  
2853 occurring in decay of radionuclides. Also called gamma ray and sometimes shortened to gamma (*e.g.*,  
2854 gamma-emitting radionuclides) (see photon and x ray).
- 2855 **genetic effects:** Changes in reproductive cells that may result in detriment to offspring.
- 2856 **gray (Gy):** The SI special name for the unit of the quantities absorbed dose and air kerma.  $1 \text{ Gy} = 1 \text{ J} \cdot \text{kg}^{-1}$ .
- 2857 **heritable effects:** Changes in reproductive cells that may be passed on to offspring of persons or animals. Often  
2858 called genetic effects (see genetic effects).
- 2859 **incidence:** The rate of occurrence of a disease, usually expressed as number of cases per hundred-  
2860 thousand individuals per year [or per 100,000 person-years (PY)].
- 2861 **ionization:** The process by which a neutral atom or molecule acquires a positive or negative charge through  
2862 the loss or gain of an orbital electron.
- 2863 **ionizing radiation:** Any radiation capable of displacing electrons from atoms or molecules, thereby producing  
2864 ions. Examples include alpha radiation, beta radiation, gamma or x rays, and cosmic rays. Minimum energy  
2865 of ionizing radiation is a few electron volts (eV);  $1 \text{ eV} = 1.6 \times 10^{-19} \text{ J}$ .
- 2866 **irradiation:** Exposure to ionizing or nonionizing radiation (see also exposure).
- 2867 **lifetime risk:** The probability during one's lifetime of expressing a given health outcome.
- 2868 **LET:** Linear-energy transfer, the average amount of energy lost per unit of particle track length and expressed in  
2869  $\text{keV} \cdot \mu\text{m}^{-1}$ .
- 2870 **high-LET:** Radiation having a high linear-energy transfer (*e.g.*, protons, alpha particles, heavy ions, and  
2871 the interaction products of fast neutrons).
- 2872 **low-LET:** Radiation having a low linear-energy transfer (*e.g.*, electrons, x rays, and gamma rays).
- 2873 **meta-analysis:** In statistics evaluating epidemiologic studies, this comprises the use of statistical methods for  
2874 contrasting and combining results from different studies reported in the literature in the hope of identifying  
2875 patterns among study results, sources of disagreement among those results, or other interesting relationships  
2876 that may come to light in the context of multiple studies.
- 2877 **neutrons:** Particles with a mass similar to that of a proton, but with no electrical charge. Because they are  
2878 electrically neutral, they cannot be accelerated in an electrical field.
- 2879 **noncancer:** Health effects other than cancer (*e.g.*, cataracts, cardiovascular disease) that occur in the exposed  
2880 individual.

2881 **occupational dose:** The dose received by an individual in a restricted area, or in the course of employment in  
2882 which the individual's duties necessarily involve exposure to radiation (medical doses involving diagnosis  
2883 or treatment of the exposed individual that are not required as a condition of employment are excluded).

2884 **odds ratio (OR):** The odds of a disease for a group is the number of people with the disease divided by the number  
2885 in that group without the disease (call it  $p_i/q_i$  for group  $i$ ). Then for two groups,  $i$  and  $j$ , the OR is  $(p_i/q_i)/(p_j/q_j)$ ,  
2886 the ratio of the two sets of odds. The OR can be used in a case-control study where the relative risk (RR)  
2887 method normally cannot be applied. With an assumption that the disease is fairly rare (in, say, a  $y$ 's time), the  
2888 odds ratio is a good estimate of the RR. With a continuous dose variable, rather than just two groups, the OR is  
2889 usually modeled per unit dose using the logistic regression method.

2890 **photon:** Quantum of electromagnetic radiation, having no charge or mass, that exhibits both particle and  
2891 wave behavior, such as a gamma or x ray.

2892 **protons:** The nucleus of the hydrogen atom. Protons are positively charged.

2893 **radionuclide:** An unstable (*i.e.*, radioactive) nuclide. A species of atom characterized by the constitution of its  
2894 nucleus (*i.e.*, the number of protons and neutrons) and the excess energy available in the unstable nucleus.

2895 **relative biological effectiveness (RBE):** For a specific radiation ( $A$ ), the ratio of absorbed dose of a reference  
2896 radiation required to produce a specific level of response in a biological system to absorbed dose of radiation  
2897 ( $A$ ) required to produce an equal response. The reference radiation normally is x or gamma rays with an  
2898 average linear energy transfer of  $3.5 \text{ keV } \mu\text{m}^{-1}$  or less. Relative biological effectiveness generally depends on  
2899 dose, dose per fraction if the dose is fractionated, dose rate, and biological endpoint.

2900 **relative risk (RR):** The ratio of the risk of a given disease in those exposed to the risk of that disease in those not  
2901 exposed, usually expressed as a RR adjusted for ages and perhaps other factors.

2902 **risk:** probability of harm, sometimes combined with potential severity of that harm.

2903 **risk coefficient:** The increase in the annual incidence or mortality rate per unit dose: (1) absolute risk coefficient is  
2904 the observed minus the expected number of cases per person  $y$  at risk for a unit dose, and (2) the relative risk  
2905 coefficient is the fractional increase in the baseline incidence or mortality rate for a unit dose.

2906 **severe hazard:** A hazard that has the potential to cause death, severe injury, or occupational illness, significant  
2907 risk to the public, extensive environmental harm, or significant property damage.

2908 **severity:** In the context of this Report, the quality or power of afflicting, distressing, or paining an individual or  
2909 organ system from exposure to an environmental insult, such as ionizing radiation, that in the extreme would  
2910 cause pain or anguish and possible clinical sequelae in the individual.

2911 **sievert (Sv):** Special name for the SI unit of dose equivalent, equivalent dose, and effective dose.  $1 \text{ Sv} = 1 \text{ J. kg}^{-1}$ .

2912 **somatic effect:** Biological effects (of radiation or otherwise) that occur in the exposed individual, as opposed  
2913 to genetic (or heritable) effects which occur in the descendants of exposed individuals due to genetic  
2914 mutations in the germline.

2915 **stochastic:** Describes random events leading to effects whose probability of occurrence in an exposed  
2916 population (rather than severity in an affected individual) is a direct function of dose; these effects are  
2917 commonly regarded as having no threshold; cancer and hereditary effects are regarded as being stochastic.  
2918 **tissue reaction (deterministic effect):** Injury in populations of cells, characterized by a threshold dose and an  
2919 increase in the severity of the reaction as the dose is increased further. In some cases, tissue reactions are  
2920 modifiable by post-irradiation procedures including biological response modifiers (ICRP 2012).  
2921

2922 **Abbreviations, Acronyms and Symbols**

2923		
2924	ALARA	as low as reasonably achievable
2925	AECL	Atomic Energy of Canada Limited
2926	AHS	Adult Health Study (Radiation Effects Research Foundation)
2927	ANSI	American National Standards Institute
2928	AOP	adverse outcome pathway
2929	ARS	acute radiation syndrome
2930	BBDR	biologically-based dose response
2931	BNFL	British Nuclear Fuels Limited
2932	CVD	cardiovascular disease
2933	CED	committed effective dose
2934	CeVD	cerebrovascular disease
2935	CI	confidence interval
2936	CLL	chronic lymphocytic leukemia
2937	CNS	central nervous system
2938	CT	computed tomography
2939	CV	coefficient of variation
2940	CVD	cardiovascular disease
2941	DDREF	dose and dose-rate effectiveness factor
2942	DREF	dose rate effectiveness factor
2943	DSB	double-strand break
2944	EAR	excess absolute risk
2945	ERR	excess relative risk
2946	ESR	electron spin resonance
2947	FISH	fluorescence <i>in situ</i> hybridization
2948	HBRA	high background radiation area
2949	INWORKS	International Nuclear Workers Study
2950	LDEF	low dose effectiveness factor
2951	LET	linear energy transfer
2952	LNT	linear nonthreshold assumption or hypothesized model
2953	LSS	Life Span Study of atomic-bomb survivors (Radiation Effects Research Foundation)
2954	MWS	Million Workers Study
2955	NDR	National Dose Registry (Canada)

2956	NRRW	National Registry for Radiation Workers (United Kingdom) NTS Nevada
2957		Test Site
2958	OR	odds ratio
2959	RBE	relative biological effectiveness
2960	RBM	red bone marrow
2961	REL	recommended exposure limit
2962	RR	relative risk
2963	SIR	standardized incidence ratio
2964	SMR	standardized mortality ratio
2965	SNTS	Semipalatinsk Nuclear Test Site (Kazakhstan)
2966	TBI	total body irradiation
2967	TLD	thermoluminescence dosimeter/dosimetry
2968	TMI	Three Mile Island
2969		

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